#### HANU Probability and Statistics



#### HYPOTHESIS TESTING



#### **Lecture Contents**

Hypothesis Testing I



Hypothesis testing is to check if a statement about the population is right or wrong (in probabilistic sense).

A statistical hypothesis is an assertion or conjecture concerning one or more populations.

We could prove or reject the hypothesis from the evidence of data.



□What does this mean by "probabilistic sense"?

Since the conclusion is tested from finite sample -> probabilistic statement - i.e accept or reject the hypothesis with a probability (confident level) - p-value (probability of being wrong)



#### Hypothesis formulation for testing:

H<sub>0</sub> – Null Hypothesis

H₁ – Alternative Hypothesis

#### **Conclusions:**

reject  $H_0$  in favor of  $H_1$  because of sufficient evidence in the data or fail to reject  $H_0$  because of insufficient evidence in the data.

 $H_0$ : defendant is innocent,

 $H_1$ : defendant is guilty.



#### Hypothesis formulation for testing:

Example - certain type of cold vaccine is known to be only 25% effective after a period of 2 years. New vaccine has been developed and tested against 20 people, and shown to be effective on 8 of them. Does this mean that the new vaccine is more effective than the old?



#### Hypothesis formulation for testing:

Hypotheses for testing:

(Let p be the probability/proportion to be effective)

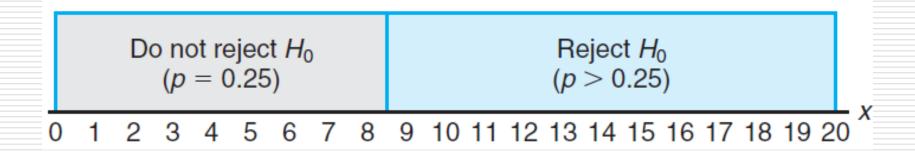
$$H_0$$
:  $p = 0.25$ ,

$$H_1$$
:  $p > 0.25$ .



#### □ Test Statistics:

We calculate the test statistics X an use X=8 as a boundary for our decision.





#### □ Probability of Errors:

#### Type I error:

Rejection of the null hypothesis when it is true is called a type I error.

#### Type II error:

Nonrejection of the null hypothesis when it is false is called a type II error.



#### □ Probability of Errors:

Table 10.1: Possible Situations for Testing a Statistical Hypothesis

	$H_0$ is true	$H_0$ is false
Do not reject $H_0$	Correct decision	Type II error
$\operatorname{Reject} H_0$	Type I error	Correct decision



#### □ Probability of Errors:

#### **Type I error:**

$$\alpha = P(\text{type I error}) = P\left(X > 8 \text{ when } p = \frac{1}{4}\right) = \sum_{x=9}^{20} b\left(x; 20, \frac{1}{4}\right)$$

$$=1-\sum_{x=0}^{8}b\left(x;20,\frac{1}{4}\right)=1-0.9591=0.0409.$$



#### □ Probability of Errors:

**Type II error**: Need an alternative for p.

$$\beta = P(\text{type II error}) = P\left(X \le 8 \text{ when } p = \frac{1}{2}\right)$$

$$= \sum_{x=0}^{8} b\left(x; 20, \frac{1}{2}\right) = 0.2517.$$



#### □ Probability of Errors:

**Type II error:** Lower type II error result in increase p.

$$\beta = P(\text{type II error}) = P(X \le 8 \text{ when } p = 0.7)$$
  
=  $\sum_{0.0051}^{8} b(x; 20, 0.7) = 0.0051.$ 



#### Continuous Random Variable:

Consider the null hypothesis that the average weight of male students in a certain college is 68 kilograms against the alternative hypothesis that it is unequal to 68. That is, we wish to test

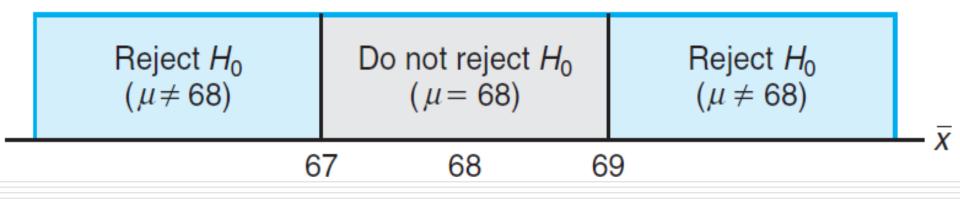
$$H_0$$
:  $\mu = 68$ ,

$$H_1: \ \mu \neq 68.$$



#### Continuous Random Variable:

We calculate the sample mean and define the critical region for making our decision on hypothesis testing.





#### Continuous Random Variable:

Assume the standard deviation of the population of weights to be  $\sigma=3.6$ . For large samples, we may substitute s for  $\sigma$  if no other estimate of  $\sigma$  is available. Our decision statistic, based on a random sample of size n = 36, will be the sample mean , the most efficient estimator of  $\mu$ .



#### Continuous Random Variable:

The central limit theorem says that the distribution of  $\bar{x}$  is approximately normal with standard deviation:

$$\sigma_{\bar{X}} = \sigma / \sqrt{n} = 3.6/6 = 0.6.$$



#### Continuous Random Variable: Type I error:

$$\alpha = P(\bar{X} < 67 \text{ when } \mu = 68) + P(\bar{X} > 69 \text{ when } \mu = 68).$$

The z-values corresponding to  $\bar{x}_1 = 67$  and  $\bar{x}_2 = 69$  when  $H_0$  is true are

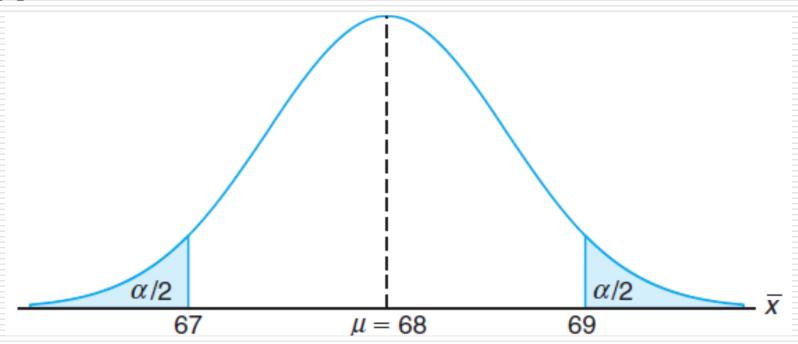
$$z_1 = \frac{67 - 68}{0.6} = -1.67$$
 and  $z_2 = \frac{69 - 68}{0.6} = 1.67$ .

Therefore,

$$\alpha = P(Z < -1.67) + P(Z > 1.67) = 2P(Z < -1.67) = 0.0950.$$



□ Continuous Random Variable: Type I error:





#### □ Continuous Random Variable:

Type II error: testing over alternatives

$$\beta = P(67 \le \bar{X} \le 69 \text{ when } \mu = 70).$$

The z-values corresponding to  $\bar{x}_1 = 67$  and  $\bar{x}_2 = 69$  when  $H_1$  is true are

$$z_1 = \frac{67 - 70}{0.45} = -6.67$$
 and  $z_2 = \frac{69 - 70}{0.45} = -2.22$ .

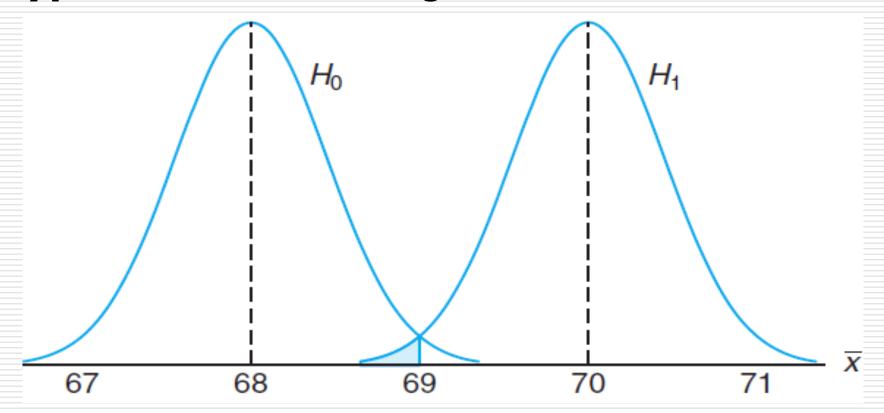
Therefore,

$$\beta = P(-6.67 < Z < -2.22) = P(Z < -2.22) - P(Z < -6.67)$$
$$= 0.0132 - 0.0000 = 0.0132.$$



Continuous Random Variable:

Type II error: testing over alternatives





#### □ Test Errors: properties

- 1. The type I error and type II error are related. A decrease in the probability of one generally results in an increase in the probability of the other.
- 2. The size of the critical region, and therefore the probability of committing a type I error, can always be reduced by adjusting the critical value(s).
- **3.** An increase in the sample size n will reduce  $\alpha$  and  $\beta$  simultaneously.
- 4. If the null hypothesis is false,  $\beta$  is a maximum when the true value of a parameter approaches the hypothesized value. The greater the distance between the true value and the hypothesized value, the smaller  $\beta$  will be.



#### One vs. Two-tailed Tests:

#### One-sided tests:

$$H_0$$
:  $\theta = \theta_0$ ,  $H_0$ :  $\theta = \theta_0$ ,

$$H_1$$
:  $\theta > \theta_0$   $H_1$ :  $\theta < \theta_0$ ,

$$H_0$$
:  $\theta = \theta_0$ ,

$$H_1: \theta < \theta_0,$$

#### Two-sided tests:

$$H_0$$
:  $\theta = \theta_0$ ,

$$H_1: \theta \neq \theta_0,$$



## How to choose Null and Alternative Hypotheses?

- Choose the Null hypothesis with equality (often to reject it).
- $\square$  Choose the Alternative ( $H_1$ ) (one-tailed, two-tailed) depending on what do we need when  $H_0$  is rejected.



#### Significant Level

- □ To reject  $H_0$  we need to compute the  $P(H_0)$  (P-value) if this value is smaller than a threshold (significant level)  $\alpha$ , then we could reject  $H_0$
- □ It is customary to choose  $\alpha$ =0.05 or 0.01 ( $z_{\alpha/2}$  –values for 0.05 and 0.01 are 1.96 and 2.57 respectively)



#### **Hypothesis Testing Steps:**

- 1. State the null and alternative hypotheses.
- 2. Choose a fixed significance level  $\alpha$ .
- 3. Choose an appropriate test statistic and establish the critical region based on  $\alpha$ .
- 4. Reject  $H_0$  if the computed test statistic is in the critical region. Otherwise, do not reject.
- 5. Draw scientific or engineering conclusions.



## Tests concerning mean: Single sample When variance is known:

$$H_0: \mu = \mu_0,$$

$$H_1: \mu \neq \mu_0.$$

$$Z = \frac{\bar{X} - \mu}{\sigma / \sqrt{n}}.$$

$$P\left(-z_{\alpha/2} < \frac{\bar{X} - \mu_0}{\sigma/\sqrt{n}} < z_{\alpha/2}\right) = 1 - \alpha$$



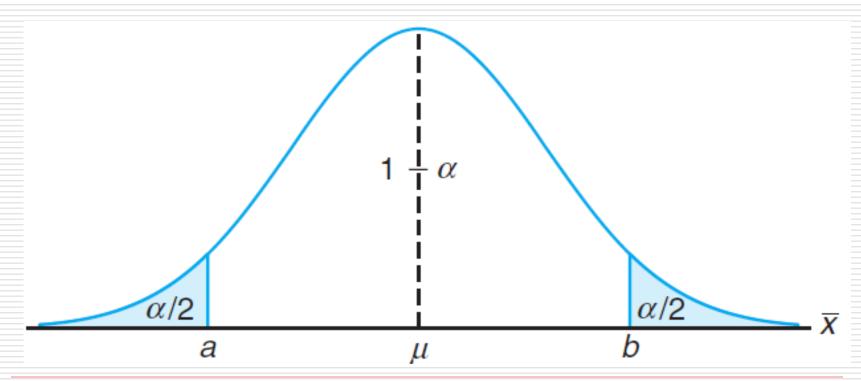
## Tests concerning mean: Single sample When variance is known:

$$z = \frac{\bar{x} - \mu_0}{\sigma / \sqrt{n}} > z_{\alpha/2}$$
 or  $z = \frac{\bar{x} - \mu_0}{\sigma / \sqrt{n}} < -z_{\alpha/2}$ 

If  $-z_{\alpha/2} < z < z_{\alpha/2}$ , do not reject  $H_0$ . Rejection of  $H_0$ , of course, implies acceptance of the alternative hypothesis  $\mu \neq \mu_0$ . With this definition of the critical region, it should be clear that there will be probability  $\alpha$  of rejecting  $H_0$  (falling into the critical region) when, indeed,  $\mu = \mu_0$ .



## Tests concerning mean: Single sample When variance is known:





Tests concerning mean: Single sample
When variance is known: One-tailed tests

$$H_0$$
:  $\mu = \mu_0$ ,  $H_1$ :  $\mu > \mu_0$ .

The critical region is when  $Z>Z_{\alpha}$  ( $Z<-Z_{\alpha}$ )



Tests concerning mean: Single sample When variance is known: Example A random sample of 100 recorded deaths in the United States during the past year showed an average life span of 71.8 years. Assuming a population standard deviation of 8.9 years, does this seem to indicate that the mean life span today is greater than 70 years? Use a 0.05 level of significance.



# Tests concerning mean: Single sample When variance is known: Example

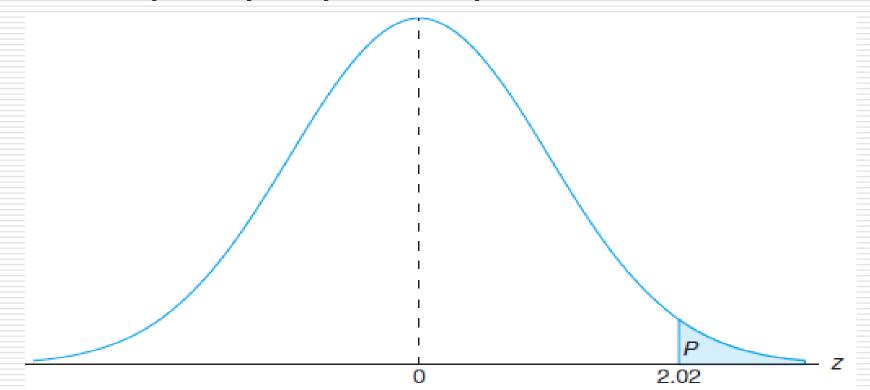
- 1.  $H_0$ :  $\mu = 70$  years.
- 2.  $H_1$ :  $\mu > 70$  years.
- 3.  $\alpha = 0.05$ .
- 4. Critical region: z > 1.645, where  $z = \frac{\bar{x} \mu_0}{\sigma/\sqrt{n}}$ .
- 5. Computations:  $\bar{x} = 71.8$  years,  $\sigma = 8.9$  years, and hence  $z = \frac{71.8 70}{8.9 / \sqrt{100}} = 2.02$ .
- 6. Decision: Reject  $H_0$  and conclude that the mean life span today is greater than 70 years.



Tests concerning mean: Single sample

When variance is known: Example

P-value(2.02)=P(Z>2.02)=0.0217





# Tests concerning mean: Single sample When variance is known: Example

A manufacturer of sports equipment has developed a new synthetic fishing line that the company claims has a mean breaking strength of 8 kilograms with a standard deviation of 0.5 kilogram. Test the hypothesis that  $\mu = 8$  kilograms against the alternative that  $\mu \neq 8$  kilograms if a random sample of 50 lines is tested and found to have a mean breaking strength of 7.8 kilograms. Use a 0.01 level of significance.



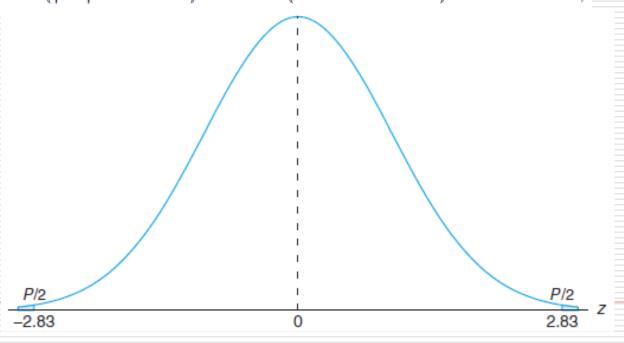
# Tests concerning mean: Single sample When variance is known: Example

- 1.  $H_0$ :  $\mu = 8$  kilograms.
- 2.  $H_1$ :  $\mu \neq 8$  kilograms.
- 3.  $\alpha = 0.01$ .
- 4. Critical region: z < -2.575 and z > 2.575, where  $z = \frac{\bar{x} \mu_0}{\sigma/\sqrt{n}}$ .
- 5. Computations:  $\bar{x} = 7.8$  kilograms, n = 50, and hence  $z = \frac{7.8 8}{0.5/\sqrt{50}} = -2.83$ .
- 6. Decision: Reject  $H_0$  and conclude that the average breaking strength is not equal to 8 but is, in fact, less than 8 kilograms.



# Tests concerning mean: Single sample When variance is known: Example

$$P = P(|Z| > 2.83) = 2P(Z < -2.83) = 0.0046,$$





# Tests concerning mean: Single sample When variance is unknown:

For the two-sided hypothesis

$$H_0$$
:  $\mu = \mu_0$ ,

$$H_1: \ \mu \neq \mu_0,$$

we reject  $H_0$  at significance level  $\alpha$  when the computed t-statistic

$$t = \frac{\bar{x} - \mu_0}{s / \sqrt{n}}$$

exceeds  $t_{\alpha/2,n-1}$  or is less than  $-t_{\alpha/2,n-1}$ .



# Tests concerning mean: Single samples When variance is unknown: Example

The Edison Electric Institute has published figures on the number of kilowatt hours used annually by various home appliances. It is claimed that a vacuum cleaner uses an average of 46 kilowatt hours per year. If a random sample of 12 homes included in a planned study indicates that vacuum cleaners use an average of 42 kilowatt hours per year with a standard deviation of 11.9 kilowatt hours, does this suggest at the 0.05 level of significance that vacuum cleaners use, on average, less than 46 kilowatt hours annually? Assume the population of kilowatt hours to be normal.



# Tests concerning mean: Single samples When variance is unknown: Example

- 1.  $H_0$ :  $\mu = 46$  kilowatt hours.
- 2.  $H_1$ :  $\mu < 46$  kilowatt hours.
- 3.  $\alpha = 0.05$ .
- 4. Critical region: t < -1.796, where  $t = \frac{\bar{x} \mu_0}{s/\sqrt{n}}$  with 11 degrees of freedom.
- 5. Computations:  $\bar{x} = 42$  kilowatt hours, s = 11.9 kilowatt hours, and n = 12. Hence,

$$t = \frac{42 - 46}{11.9/\sqrt{12}} = -1.16, \qquad P = P(T < -1.16) \approx 0.135.$$

6. Decision: Do not reject  $H_0$  and conclude that the average number of kilowatt hours used annually by home vacuum cleaners is not significantly less than 46.



# Tests concerning mean: Two samples When the variances are known:

Compute statistics. The test is similar to single sample

$$Z = \frac{(\bar{X}_1 - \bar{X}_2) - (\mu_1 - \mu_2)}{\sqrt{\sigma_1^2/n_1 + \sigma_2^2/n_2}}$$



## Tests concerning mean: Two samples When variances are unknown but equal:

For the two-sided hypothesis

$$H_0$$
:  $\mu_1 = \mu_2$ ,  $H_1$ :  $\mu_1 \neq \mu_2$ ,

we reject  $H_0$  at significance level  $\alpha$  when the computed t-statistic

$$t = \frac{(\bar{x}_1 - \bar{x}_2) - d_0}{s_p \sqrt{1/n_1 + 1/n_2}},$$

where

$$s_p^2 = \frac{s_1^2(n_1 - 1) + s_2^2(n_2 - 1)}{n_1 + n_2 - 2}$$

exceeds  $t_{\alpha/2,n_1+n_2-2}$  or is less than  $-t_{\alpha/2,n_1+n_2-2}$ .



#### Tests concerning mean: Two samples Example -

An experiment was performed to compare the abrasive wear of two different laminated materials. Twelve pieces of material 1 were tested by exposing each piece to a machine measuring wear. Ten pieces of material 2 were similarly tested. In each case, the depth of wear was observed. The samples of material 1 gave an average (coded) wear of 85 units with a sample standard deviation of 4, while the samples of material 2 gave an average of 81 with a sample standard deviation of 5. Can we conclude at the 0.05 level of significance that the abrasive wear of material 1 exceeds that of material 2 by more than 2 units? Assume the populations to be approximately normal with equal variances.



## Tests concerning mean: Two samples

#### **Example -**

Let  $\mu_1$  and  $\mu_2$  represent the population means of the abrasive wear for material 1 and material 2, respectively.

- 1.  $H_0$ :  $\mu_1 \mu_2 = 2$ .
- 2.  $H_1$ :  $\mu_1 \mu_2 > 2$ .
- 3.  $\alpha = 0.05$ .
- 4. Critical region: t > 1.725, where  $t = \frac{(\bar{x}_1 \bar{x}_2) d_0}{s_p \sqrt{1/n_1 + 1/n_2}}$  with v = 20 degrees of freedom.
- 5. Computations:

$$\bar{x}_1 = 85,$$
  $s_1 = 4,$   $n_1 = 12,$   $\bar{x}_2 = 81,$   $s_2 = 5,$   $n_2 = 10.$ 



#### Tests concerning mean: Two samples Example -

Hence

$$s_p = \sqrt{\frac{(11)(16) + (9)(25)}{12 + 10 - 2}} = 4.478,$$
  
 $t = \frac{(85 - 81) - 2}{4.478\sqrt{1/12 + 1/10}} = 1.04,$   
 $P = P(T > 1.04) \approx 0.16.$  (See Table A.4.)

6. Decision: Do not reject  $H_0$ . We are unable to conclude that the abrasive wear of material 1 exceeds that of material 2 by more than 2 units.



#### Tests concerning mean: Two samples Unknown and unequal variances

$$T' = \frac{(\bar{X}_1 - \bar{X}_2) - d_0}{\sqrt{s_1^2/n_1 + s_2^2/n_2}}$$

has an approximate t-distribution with approximate degrees of freedom

$$v = \frac{(s_1^2/n_1 + s_2^2/n_2)^2}{(s_1^2/n_1)^2/(n_1 - 1) + (s_2^2/n_2)^2/(n_2 - 1)}.$$

As a result, the test procedure is to not reject  $H_0$  when

$$-t_{\alpha/2,v} < t' < t_{\alpha/2,v},$$



**Tests concerning mean: Two samples** 

Pairwise t-tests: For "before" vs "after"

Blood Sample Data: In a study conducted in the Forestry and Wildlife Department at Virginia Tech, J. A. Wesson examined the influence of the drug succinylcholine on the circulation levels of androgens in the blood. Blood samples were taken from wild, free-ranging deer immediately after they had received an intramuscular injection of succinylcholine administered using darts and a capture gun. A second blood sample was obtained from each deer 30 minutes after the



**Tests concerning mean: Two samples** 

Pairwise t-tests: For "before" vs "after"

first sample, after which the deer was released. The levels of androgens at time of capture and 30 minutes later, measured in nanograms per milliliter (ng/mL), for 15 deer are given in Table. Assuming that the populations of androgen levels at time of injection and 30 minutes later are normally distributed, test at the 0.05 level of significance whether the androgen concentrations are altered after 30 minutes.



#### **Tests concerning mean: Two samples**

Table 10.2: Data for Case Study 10.1

Androgen (ng/mL)			
$\mathbf{Deer}$	At Time of Injection	30 Minutes after Injection	$\boldsymbol{d_i}$
1	2.76	7.02	4.26
2	5.18	3.10	-2.08
3	2.68	5.44	2.76
4	3.05	3.99	0.94
5	4.10	5.21	1.11
6	7.05	10.26	3.21
7	6.60	13.91	7.31
8	4.79	18.53	13.74
9	7.39	7.91	0.52
10	7.30	4.85	-2.45
11	11.78	11.10	-0.68
12	3.90	3.74	-0.16
13	26.00	94.03	68.03
14	67.48	94.03	26.55
15	17.04	41.70	24.66



#### Tests concerning mean: Two samples

Let  $\mu_1$  and  $\mu_2$  be the average androgen concentration at the time of injection and 30 minutes later, respectively. We proceed as follows:

- 1.  $H_0$ :  $\mu_1 = \mu_2$  or  $\mu_D = \mu_1 \mu_2 = 0$ .
- 2.  $H_1$ :  $\mu_1 \neq \mu_2$  or  $\mu_D = \mu_1 \mu_2 \neq 0$ .
- 3.  $\alpha = 0.05$ .
- 4. Critical region: t < -2.145 and t > 2.145, where  $t = \frac{\overline{d} d_0}{s_D/\sqrt{n}}$  with v = 14 degrees of freedom.
- 5. Computations: The sample mean and standard deviation for the  $d_i$  are

$$\overline{d} = 9.848$$
 and  $s_d = 18.474$ .



#### Tests concerning mean: Two samples

Therefore,

$$t = \frac{9.848 - 0}{18.474/\sqrt{15}} = 2.06.$$

6. Though the t-statistic is not significant at the 0.05 level, from Table A.4,

$$P = P(|T| > 2.06) \approx 0.06.$$

As a result, there is some evidence that there is a difference in mean circulating levels of androgen.



#### **Further Readings**

- ☐ Textbook 2 Chapter 10.
- □ Textbook 1- Chapter 8-9.