

March 24, 2021

The results below are generated from an R script.

```
library(rmarkdown)
library(ggplot2)

## Need help getting started? Try the R Graphics Cookbook: https://r-graphics.org

library(tidyverse)

## - Attaching packages ----- tidyverse 1.3.0 -
## v tibble 3.1.0    v dplyr 1.0.4
## v tidyr 1.1.2    v stringr 1.4.0
## v readr 1.4.0    v forcats 0.5.1
## v purrr 0.3.4
## - Conflicts ----- tidyverse_conflicts() -
## x dplyr::filter() masks stats::filter()
## x dplyr::lag()    masks stats::lag()

library(rsample)
library(caret)

## Loading required package: lattice
##
## Attaching package: 'caret'
## The following object is masked from 'package:purrr':
##
## lift

library(modelr)
library(parallel)
library(foreach)

##
## Attaching package: 'foreach'
## The following objects are masked from 'package:purrr':
##
## accumulate, when

library(FNN)
library(tseries)

## Registered S3 method overwritten by 'quantmod':
## method from
## as.zoo.data.frame zoo
##
## 'tseries' version: 0.10-48
##
## 'tseries' is a package for time series analysis and computational finance.
##
## See 'library(help="tseries")' for details.
```

```

library(stats)
library(urca)
library(AER)

## Loading required package: car
## Loading required package: carData
##
## Attaching package: 'car'
## The following object is masked from 'package:dplyr':
##
##   recode
## The following object is masked from 'package:purrr':
##
##   some
## Loading required package: lmtest
## Loading required package: zoo
##
## Attaching package: 'zoo'
## The following objects are masked from 'package:base':
##
##   as.Date, as.Date.numeric
## Loading required package: sandwich
## Loading required package: survival
##
## Attaching package: 'survival'
## The following object is masked from 'package:caret':
##
##   cluster

library(dynlm)
library(forecast)

## This is forecast 8.13
## Want to meet other forecasters? Join the International Institute of Forecasters:
## http://forecasters.org/

library(vrtest)
library(readxl)
library(stargazer)

##
## Please cite as:
## Hlavac, Marek (2018). stargazer: Well-Formatted Regression and Summary Statistics Tables.
## R package version 5.2.2. https://CRAN.R-project.org/package=stargazer

library(scales)

##
## Attaching package: 'scales'
## The following object is masked from 'package:purrr':
##
##   discard
## The following object is masked from 'package:readr':
##
##   col_factor

library(quantmod)

```

```

## Loading required package: xts
##
## Attaching package: 'xts'
## The following objects are masked from 'package:dplyr':
##
## first, last
## Loading required package: TTR

library(zoo)
library(fGarch)

## Loading required package: timeDate
## Loading required package: timeSeries
##
## Attaching package: 'timeSeries'
## The following object is masked from 'package:zoo':
##
## time<-
## Loading required package: fBasics
##
## Attaching package: 'fBasics'
## The following object is masked from 'package:TTR':
##
## volatility
## The following object is masked from 'package:car':
##
## densityPlot

library(LINselect)
library(aTSA)

##
## Attaching package: 'aTSA'
## The following object is masked from 'package:forecast':
##
## forecast
## The following objects are masked from 'package:tseries':
##
## adf.test, kpss.test, pp.test
## The following object is masked from 'package:graphics':
##
## identify

library(egcm)

#Time Series Question 1 Cointegration

# plot both interest series
plot(merge(as.zoo(irates$r1), as.zoo(irates$r3)),
      plot.type = "single",
      lty = c(2, 1),
      lwd = 0.5,
      xlab = "Date",
      ylab = "Percent per annum",

```

```

ylim = c(-5, 30),
main = "Interest Rates")

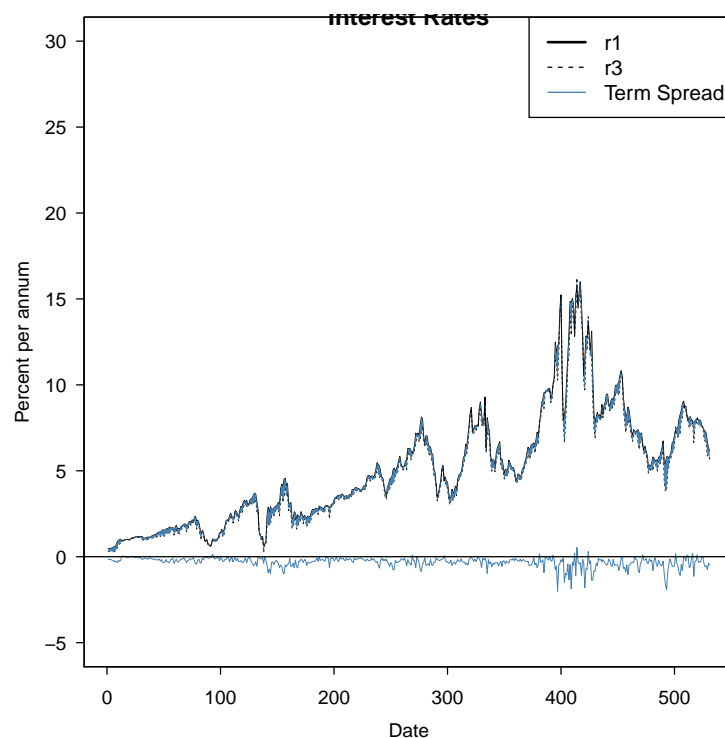
TSpread=irates$r1-irates$r3
# add the term spread series
lines(as.zoo(TSpread),
      col = "steelblue",
      lwd = 0.5,
      xlab = "Date",
      ylab = "Percent per annum",
      main = "Term Spread")

# shade the term spread
polygon(c(time(irates$r1), rev(time(irates$r3))),
       c(irates$r1, rev(irates$r3)),
       col = alpha("steelblue", alpha = 7),
       border = NA)

# add horizontal line add 0
abline(0, 0)

# add a legend
legend("topright",
      legend = c("r1", "r3", "Term Spread"),
      col = c("black", "black", "steelblue"),
      lwd = c(2, 1, 1),
      lty = c(1, 2, 1))

```



*#The plot suggests that r1 and r3 interest rates are cointegrated:
 #That is, both interest series seem to show the same behavior by sharing a
 #common stochastic trend. The term spread, which I get by taking the difference
 #between r1 and r3 interest rates, seems to be stationary. In fact, the
 #expectations theory of the term structure suggests the cointegrating coefficient
 #<U+03B8> to be 1. This is consistent with the visual result I obtain.*

#(a)

#The null hypothesis is that the data are non-stationary

```
r1_1=ur.df(irates$r1, type = c("none"), lags = 3)
summary(r1_1)
```

```
##
## #####
## # Augmented Dickey-Fuller Test Unit Root Test #
## #####
##
## Test regression none
##
##
## Call:
## lm(formula = z.diff ~ z.lag.1 - 1 + z.diff.lag)
##
## Residuals:
##      Min       1Q   Median       3Q      Max
## -4.4371 -0.1534  0.0398  0.2427  2.9574
##
## Coefficients:
##              Estimate Std. Error t value Pr(>|t|)
## z.lag.1      -0.003874   0.004579  -0.846  0.39797
## z.diff.lag1   0.022573   0.043495   0.519  0.60399
## z.diff.lag2  -0.014892   0.043496  -0.342  0.73220
## z.diff.lag3  -0.114004   0.043573  -2.616  0.00914 **
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
##
## Residual standard error: 0.6057 on 523 degrees of freedom
## Multiple R-squared:  0.01572, Adjusted R-squared:  0.008192
## F-statistic: 2.088 on 4 and 523 DF,  p-value: 0.08109
##
##
## Value of test-statistic is: -0.8459
##
## Critical values for test statistics:
##      1pct  5pct 10pct
## tau1 -2.58 -1.95 -1.62
```

#Ho: tau1: <U+03B3>=0

*#For type= "none," tau1 is the null hypothesis for gamma = 0.
 #Using the interest rates example,
 #I get a test-statistic of -0.8459. The Critical values for*

```

#test statistics are: tau1 -2.58 -1.95 -1.62.
#Given that the test statistic is within the all 3 regions
#(1%, 5%, 10%) where we fail to reject the null, we should presume
#the data is a random walk, i.e that a unit root is present.

```

```

r1_2=ur.df(irates$r1, type = c("trend"), lags = 3)
summary(r1_2)

```

```

##
## #####
## # Augmented Dickey-Fuller Test Unit Root Test #
## #####
##
## Test regression trend
##
##
## Call:
## lm(formula = z.diff ~ z.lag.1 + 1 + tt + z.diff.lag)
##
## Residuals:
##      Min       1Q   Median       3Q      Max
## -4.2262 -0.1855  0.0007  0.1985  3.0382
##
## Coefficients:
##              Estimate Std. Error t value Pr(>|t|)
## (Intercept)  0.0480298  0.0531093   0.904  0.36622
## z.lag.1      -0.0466643  0.0145947  -3.197  0.00147 **
## tt           0.0007070  0.0003021   2.340  0.01965 *
## z.diff.lag1  0.0442049  0.0437827   1.010  0.31314
## z.diff.lag2  0.0068410  0.0437872   0.156  0.87591
## z.diff.lag3 -0.0928695  0.0438271  -2.119  0.03456 *
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
##
## Residual standard error: 0.6012 on 521 degrees of freedom
## Multiple R-squared:  0.03388, Adjusted R-squared:  0.02461
## F-statistic: 3.655 on 5 and 521 DF,  p-value: 0.002933
##
##
## Value of test-statistic is: -3.1973 3.5567 5.238
##
## Critical values for test statistics:
##      1pct  5pct 10pct
## tau3 -3.96 -3.41 -3.12
## phi2  6.09  4.68  4.03
## phi3  8.27  6.25  5.34

```

```

#For the type="trend":

```

```

#tau: <U+03B3>=0
#phi3: <U+03B3>=a2=0
#phi2: a0=<U+03B3>=a2=0

```

*#In this case, the test statistics are -3.1973 3.5567 5.238.
 #In each one of these cases, the t-statistics fall within the
 #"fail to reject the null" regions. Tau3 implies that we fail
 #to reject the null of unit root. Failing to reject phi3 implies
 #two things: 1) $\langle U+03B3 \rangle = 0$ (unit root) and 2) there is no time trend term, i.e., $a_2 = 0$.
 #Failing to reject phi2 implies 3 things: 1) $\langle U+03B3 \rangle = 0$ and 2) no time
 #trend term and 3) no drift term, i.e. that $\langle U+03B3 \rangle = 0$, that $a_0 = 0$, and that $a_2 = 0$.*

```
r1_3=ur.df(irates$r1, type = c( "drift"), lags = 3)
summary(r1_3)
```

```
##
## #####
## # Augmented Dickey-Fuller Test Unit Root Test #
## #####
##
## Test regression drift
##
##
## Call:
## lm(formula = z.diff ~ z.lag.1 + 1 + z.diff.lag)
##
## Residuals:
##      Min       1Q   Median       3Q      Max
## -4.3442 -0.1850 -0.0146  0.2083  3.0055
##
## Coefficients:
##              Estimate Std. Error t value Pr(>|t|)
## (Intercept)  0.101381   0.048173   2.105   0.0358 *
## z.lag.1      -0.018604   0.008356  -2.226   0.0264 *
## z.diff.lag1   0.028333   0.043439   0.652   0.5145
## z.diff.lag2  -0.009045   0.043443  -0.208   0.8352
## z.diff.lag3  -0.108248   0.043517  -2.487   0.0132 *
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
##
## Residual standard error: 0.6038 on 522 degrees of freedom
## Multiple R-squared:  0.02373, Adjusted R-squared:  0.01625
## F-statistic: 3.172 on 4 and 522 DF,  p-value: 0.01364
##
##
## Value of test-statistic is: -2.2264 2.5747
##
## Critical values for test statistics:
##      1pct  5pct 10pct
## tau2 -3.43 -2.86 -2.57
## phi1  6.43  4.59  3.78
```

#phi3: $\langle U+03B3 \rangle = a_2 = 0$

*#type = "drift" (where a_0 is "a sub-zero" and refers to the constant, or drift term)
 #"tau2" is still the $\gamma = 0$ null hypothesis. In this case, where the first test
 #statistic = -2.2264 is within the region of failing to reject the null, we should*

#presume a unit root, that $\langle U+03B3 \rangle = 0$ at either one of the significance levels of 1%, 5% or that of 10%. The phi1 term refers to the second hypothesis, which is a combined null hypothesis of $a_0 = \gamma = 0$. This means that BOTH of the values are tested to be 0 at the same time. If $p < 0.05$, we reject the null, and presume that AT LEAST one of these is false--i.e. one or both of the terms a_0 or γ are not 0. Failing to reject this null implies that BOTH a_0 AND $\gamma = 0$, implying 1) that $\langle U+03B3 \rangle = 0$ therefore a unit root is present, AND 2) $a_0 = 0$, so there is no drift term. Since a t-statistic of 2.5747 is smaller than any of the critical values of phi1, we fail to reject the null that one or both terms are not zero. Unit root is present.

```
r3_1=ur.df(irates$r3, type = c("none"), lags = 3)
summary(r3_1)

##
## #####
## # Augmented Dickey-Fuller Test Unit Root Test #
## #####
##
## Test regression none
##
##
## Call:
## lm(formula = z.diff ~ z.lag.1 - 1 + z.diff.lag)
##
## Residuals:
##      Min       1Q   Median       3Q      Max
## -4.4734 -0.0967  0.0372  0.2073  2.6232
##
## Coefficients:
##              Estimate Std. Error t value Pr(>|t|)
## z.lag.1      -0.003059   0.003873  -0.790   0.4300
## z.diff.lag1   0.120264   0.043762   2.748   0.0062 **
## z.diff.lag2  -0.078709   0.043933  -1.792   0.0738 .
## z.diff.lag3  -0.012009   0.043831  -0.274   0.7842
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
##
## Residual standard error: 0.5398 on 523 degrees of freedom
## Multiple R-squared:  0.02025, Adjusted R-squared:  0.01276
## F-statistic: 2.702 on 4 and 523 DF, p-value: 0.02991
##
##
## Value of test-statistic is: -0.7899
##
## Critical values for test statistics:
##      1pct  5pct 10pct
## tau1 -2.58 -1.95 -1.62
```

#tau: $\langle U+03B3 \rangle = 0$

#For type= "none," tau1 in R output is the null hypothesis for $\gamma = 0$.
#Using the interest rates example,


```
#I get a value of test-statistic is -0.7899. The critical values for test
#statistics are: tau1 -2.58 -1.95 -1.62. The test statistic falls within
#the all 3 regions (1%, 5%, 10%). As such, we fail to reject the null hypothesis.
#We presume that a unit root is present.
```

```
r3_2=ur.df(irates$r3, type = c("trend"), lags = 3)
summary(r3_2)

##
## #####
## # Augmented Dickey-Fuller Test Unit Root Test #
## #####
##
## Test regression trend
##
##
## Call:
## lm(formula = z.diff ~ z.lag.1 + 1 + tt + z.diff.lag)
##
## Residuals:
##      Min       1Q   Median       3Q      Max
## -4.2608 -0.1380  0.0019  0.1626  2.8282
##
## Coefficients:
##              Estimate Std. Error t value Pr(>|t|)
## (Intercept)  0.0481030  0.0475005   1.013  0.31168
## z.lag.1      -0.0405648  0.0127394  -3.184  0.00154 **
## tt           0.0006391  0.0002720   2.350  0.01914 *
## z.diff.lag1  0.1386243  0.0438889   3.159  0.00168 **
## z.diff.lag2 -0.0598721  0.0440638  -1.359  0.17481
## z.diff.lag3  0.0070132  0.0439794   0.159  0.87336
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
##
## Residual standard error: 0.5358 on 521 degrees of freedom
## Multiple R-squared:  0.03824, Adjusted R-squared:  0.02901
## F-statistic: 4.143 on 5 and 521 DF, p-value: 0.001064
##
##
## Value of test-statistic is: -3.1842 3.5294 5.1915
##
## Critical values for test statistics:
##      1pct  5pct 10pct
## tau3 -3.96 -3.41 -3.12
## phi2  6.09  4.68  4.03
## phi3  8.27  6.25  5.34

#tau: <U+03B3>=0
#phi3: <U+03B3>=a2=0
#phi2: a0=<U+03B3>=a2=0

#The test statistics are -3.1842 3.5294 5.1915.
```

```

#In this case, although we fail to reject the null of unit root at
#the 1% and 5% significance levels, we are able to reject it at the
#10% significance level. On the other hand, we fail to reject the null
#hypothesis that there is no time trend and unit root in the data.
#Failing to reject phi2 implies 3 things: 1)  $\langle U+03B3 \rangle = 0$  AND 2) no time trend
#term AND 3) no drift term, i.e. that  $\langle U+03B3 \rangle = 0$ , that  $a_0 = 0$ , and that  $a_2 = 0$ .

r3_3=ur.df(irates$r3, type = c( "drift"), lags = 3)
summary(r3_3)

##
## #####
## # Augmented Dickey-Fuller Test Unit Root Test #
## #####
##
## Test regression drift
##
##
## Call:
## lm(formula = z.diff ~ z.lag.1 + 1 + z.diff.lag)
##
## Residuals:
##      Min       1Q   Median       3Q      Max
## -4.3846 -0.1297 -0.0150  0.1849  2.7038
##
## Coefficients:
##              Estimate Std. Error t value Pr(>|t|)
## (Intercept)   0.091960   0.043870   2.096  0.03654 *
## z.lag.1       -0.015850   0.007220  -2.195  0.02859 *
## z.diff.lag1    0.124623   0.043671   2.854  0.00449 **
## z.diff.lag2   -0.073745   0.043855  -1.682  0.09325 .
## z.diff.lag3   -0.007202   0.043750  -0.165  0.86930
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
##
## Residual standard error: 0.5381 on 522 degrees of freedom
## Multiple R-squared:  0.02805, Adjusted R-squared:  0.0206
## F-statistic: 3.766 on 4 and 522 DF, p-value: 0.004962
##
##
## Value of test-statistic is: -2.1951 2.511
##
## Critical values for test statistics:
##      1pct  5pct 10pct
## tau2 -3.43 -2.86 -2.57
## phi1  6.43  4.59  3.78

#phi3:  $\langle U+03B3 \rangle = a_2 = 0$ 

#In this case, where the first test statistic = -2.1951 is within the region
#of failing to reject the null, we should presume a unit root, that  $\langle U+03B3 \rangle = 0$  at
#either the significance level of 1%, 5% or that of 10%.
#The phi1 term refers to the second hypothesis, which is a combined null hypothesis

```

```

#of  $\alpha_0 = \gamma = 0$ .
#This means that BOTH of the values are tested to be 0 at the same time.
#Since a t-statistic of 2.511 is smaller than any of the critical values of  $\phi_{11}$ ,
#we fail to reject the null that one or both terms
#are not zero.

#In part (a) I conclude that it is plausible to model both interest rate series as  $I(1)$ .

#(b) Perform a regression by OLS explaining r1 from r3. Test for cointegration.

lm_r1=lm(irates$r1~irates$r3, data=irates)

ur.df(window(irates$r1) - window(irates$r3),
      lags = 3,
      selectlags = "AIC",
      type = "drift")

##
## #####
## # Augmented Dickey-Fuller Test Unit Root / Cointegration Test #
## #####
##
## The value of the test statistic is: -8.0431 32.3517

#(c) Perform a regression by OLS explaining r3 from r1. Test for cointegration.

lm_r3=lm(irates$r3~irates$r1, data=irates)

ur.df(window(irates$r3) - window(irates$r1),
      lags = 3,
      selectlags = "AIC",
      type = "drift")

##
## #####
## # Augmented Dickey-Fuller Test Unit Root / Cointegration Test #
## #####
##
## The value of the test statistic is: -8.0431 32.3517

#Here I test if term spread is stationary (cointegration of interest rates) using ADF.
#The test leads me to reject the null hypothesis of non-stationarity of the term spread
#series at any of the standard significance levels. This is strong evidence in favor
#of the alternative hypothesis that the term spread is stationary, implying cointegration
#of r1 and r3 interest rates.

#(d) What is cointegration? Interpret your findings above intuitively.

#When the dependent( $Y_t$ ) and the independent( $X_t$ ) variables are  $I(1)$  and if
#there is a  $c$  such that  $Y_t - cX_t$  is  $I(0)$ ,  $X_t$  and  $Y_t$  are said to be cointegrated.
#In other words, cointegration of  $X_t$  and  $Y_t$  means that  $X_t$  and  $Y_t$  have the same
#or common stochastic trend and that this trend can be eliminated by taking a

```

#difference of the series such that the resulting series is stationary.

#Time Series Question 2

#I) Cholesky Decomposition

#(c)

#The Cholesky decomposition of a Hermitian positive-definite matrix A , is a decomposition of the form $A=LL^*$, where L is a lower triangular matrix with real and positive diagonal entries, and L^* is the conjugate transpose of L . Every Hermitian positive-definite matrix (and thus also every real-valued symmetric positive-definite matrix) has a unique Cholesky decomposition. In other words, Cholesky decomposition reduces a symmetric matrix into a lower-triangular matrix. When multiplied by its transpose this matrix produces the original symmetric matrix. In our example we use it for making certain assumptions on the variables we are interested in, which allow us to assume that the contemporaneous effect (short-run effect) is equal to zero on our variables of interest. This is an assumption we make and an assumption we need to defend. However, without the Cholesky decomposition, representing this assumption into an equation would be more difficult. We know that the Cholesky decomposition will give us the parameters we want because we are able to represent our parameters in a lower-triangular form where we define as to whether or not certain variables are allowed, by assumption, to affect one another.

#(d)

#After imposing the short-run restrictions (i.e., after we specify if any entry has any relationship with any other entry, and/or if any entry is zero or 1), the A_0 matrix will be a 3 by 3 matrix. The first row will have a value of 'a' in the first column and values of 0 in the other two columns. This is to represent the fact that the first entry cannot depend contemporaneously on the other two entries. The second row of the matrix has a value of 'b' in the first column, a value 'c' in the second column and a value of '0' in the third column. This is to show that while the second entry can depend contemporaneously on the first entry, it cannot depend contemporaneously on the third entry. The third row of the matrix takes a value of 'd' in the first column, a value of 'e' in the second column, and a value of 'f' in the third column. In this row we do not make any restriction assumptions as to what the relationship among entries is.

#II) Short Run Restrictions

#(a)

#The intuition behind the restrictions imposed is to move from the parameters of a reduced form VAR to the structural VAR parameters of interest. Otherwise, the model cannot be identified. That is, because the coefficients in the identification of structural parameters is unknown, we need to impose theoretical restrictions to reduce the number of unknown structural parameters to be less than or equal to the number of estimated parameters of the variance-covariance matrix of the VAR residuals. Finding dynamic patterns consistent with the structural model used for the identification would provide evidence in support of the theoretical model. Otherwise, the theory

```
#will be invalid or the empirical model somewhat misspecified.  
#It is also important to note that we cannot estimate the structural form with OLS because  
#one of the main assumptions in time series is violated. The assumption violated is that  
#the regressors are correlated with the error term. The Ao matrix is problematic because it  
#includes all the contemporaneous relations between endogenous variables.
```

```
 #(c)
```

```
#The effects that are contemporaneously zero as a result of short run restrictions  
#are the aggregate demand shock on oil production, the oil-specific demand shock on  
#oil production and the oil-specific demand shock on real activity. This is because  
#we are assuming that the shock on the real price does not affect contemporaneous  
#production and the shock on real economic activity does not affect contemporaneous  
#production. And then we are also assuming that the real oil price doesn't have a  
#contemporaneous effect on real economic activity. This is because we assume that  
#all these have lagged effects in the economy, and hence the zero starting value.
```

```
 #(d)
```

```
#Intuitively, the first line of graphs explains the effect of an oil supply shock on  
#oil production. Here we have not made any identifying assumptions. Instead, what we  
#observe in the first graph is a negative oil supply shock that begins to improve with a lag.  
#On the other hand, the oil supply shock seems to have little to no effect on the real  
#economic activity. Similarly, it appears that the oil supply shock has no effect on the price  
#of oil. The second line of graphs explains the aggregate demand shock on oil production,  
#real economic activity and real price of oil. Aggregate demand shock does not seem to  
#significantly affect the oil production. On the other hand, we observe a positive aggregate  
#demand shock in the real economic activity, which remains fairly stable in the very shock run,  
#but starts slightly decreasing the longer run. The aggregate demand shock on the real price  
#of oil starts by having no effect, but gradually shows an increase as time passes.  
#The third line of graphs shows the oil-specific demand shock on oil production, real  
#economic activity and real price of oil. The oil specific demand shock does not seem  
#to have a huge impact on oil production and the latter remains fairly stable across time.  
#The oil specific demand shock on real economic activity starts from zero, by assumption,  
#then increases slightly, then goes back to zero again in the longer run. It is the oil  
#specific shock that seems to drive the real price of oil. It begins as a positive shock and  
#then gradually declines.
```

The R session information (including the OS info, R version and all packages used):

```
sessionInfo()  
  
## R version 4.0.4 (2021-02-15)  
## Platform: x86_64-w64-mingw32/x64 (64-bit)  
## Running under: Windows 10 x64 (build 19041)  
##  
## Matrix products: default  
##  
## locale:  
## [1] LC_COLLATE=English_United Kingdom.1252 LC_CTYPE=English_United Kingdom.1252  
## [3] LC_MONETARY=English_United Kingdom.1252 LC_NUMERIC=C  
## [5] LC_TIME=English_United Kingdom.1252  
##
```

```
## attached base packages:
## [1] parallel stats graphics grDevices utils datasets methods base
##
## other attached packages:
## [1] egcm_1.0.12 aTSA_3.1.2 LINselect_1.1.3 fGarch_3042.83.2
## [5] fBasics_3042.89.1 timeSeries_3062.100 timeDate_3043.102 quantmod_0.4.18
## [9] TTR_0.24.2 xts_0.12.1 scales_1.1.1 stargazer_5.2.2
## [13] readxl_1.3.1 vrtest_0.97 forecast_8.13 dynlm_0.3-6
## [17] AER_1.2-9 survival_3.2-7 sandwich_3.0-0 lmtest_0.9-38
## [21] zoo_1.8-8 car_3.0-10 carData_3.0-4 urca_1.3-0
## [25] tseries_0.10-48 FNN_1.1.3 foreach_1.5.1 modelr_0.1.8
## [29] caret_6.0-86 lattice_0.20-41 rsample_0.0.9 forcats_0.5.1
## [33] stringr_1.4.0 dplyr_1.0.4 purrr_0.3.4 readr_1.4.0
## [37] tidyr_1.1.2 tibble_3.1.0 tidyverse_1.3.0 ggplot2_3.3.3
## [41] rmarkdown_2.7
##
## loaded via a namespace (and not attached):
## [1] colorspace_2.0-0 ellipsis_0.3.1 class_7.3-18 rio_0.5.26
## [5] pls_2.7-3 fs_1.5.0 rstudioapi_0.13 farver_2.1.0
## [9] listenv_0.8.0 furrr_0.2.2 mvtnorm_1.1-1 prodlim_2019.11.13
## [13] fansi_0.4.2 lubridate_1.7.9.2 xml2_1.3.2 codetools_0.2-18
## [17] splines_4.0.4 knitr_1.31 Formula_1.2-4 jsonlite_1.7.2
## [21] pROC_1.17.0.1 broom_0.7.5 dbplyr_2.1.0 compiler_4.0.4
## [25] httr_1.4.2 backports_1.2.1 assertthat_0.2.1 Matrix_1.3-2
## [29] cli_2.3.1 lars_1.2 htmltools_0.5.1.1 tools_4.0.4
## [33] gtable_0.3.0 glue_1.4.2 reshape2_1.4.4 Rcpp_1.0.6
## [37] fracdiff_1.5-1 cellranger_1.1.0 vctr_0.3.6 nlme_3.1-152
## [41] iterators_1.0.13 gower_0.2.2 xfun_0.22 globals_0.14.0
## [45] spatial_7.3-13 openxlsx_4.2.3 rvest_1.0.0 lifecycle_1.0.0
## [49] gtools_3.8.2 future_1.21.0 MASS_7.3-53 ipred_0.9-9
## [53] hms_1.0.0 elasticnet_1.3 yaml_2.2.1 curl_4.3
## [57] rpart_4.1-15 stringi_1.5.3 highr_0.8 randomForest_4.6-14
## [61] zip_2.1.1 lava_1.6.9 rlang_0.4.10 pkgconfig_2.0.3
## [65] pracma_2.3.3 evaluate_0.14 recipes_0.1.15 tidyselect_1.1.0
## [69] parallelly_1.23.0 plyr_1.8.6 magrittr_2.0.1 R6_2.5.0
## [73] generics_0.1.0 DBI_1.1.1 pillar_1.5.1 haven_2.3.1
## [77] foreign_0.8-81 withr_2.4.1 abind_1.4-5 nnet_7.3-15
## [81] crayon_1.4.1 utf8_1.1.4 grid_4.0.4 data.table_1.14.0
## [85] ModelMetrics_1.2.2.2 reprex_1.0.0 digest_0.6.27 stats4_4.0.4
## [89] munsell_0.5.0 quadprog_1.5-8

Sys.time()

## [1] "2021-03-24 16:50:58 CDT"
```

QUESTION 2.

(a) I) CHOLESKY DECOMPOSITION

VECTOR OF STRUCTURAL SHOCKS:

$$\begin{bmatrix} \varepsilon_{1t} \\ \varepsilon_{2t} \\ \varepsilon_{3t} \end{bmatrix}$$

By assumption, the structural shocks aren't correlated. As such, the covariance of the diagonals in the structural shocks are zero. The other assumption we make is that the structural shocks have a mean zero & a variance-covariance matrix Ω of the following form:

$$\begin{bmatrix} \varepsilon_{1t} \\ \varepsilon_{2t} \\ \varepsilon_{3t} \end{bmatrix} \sim iid \left[\begin{pmatrix} 0 \\ 0 \\ 0 \end{pmatrix}, \underbrace{\begin{pmatrix} \sigma_1^2 & 0 & 0 \\ 0 & \sigma_2^2 & 0 \\ 0 & 0 & \sigma_3^2 \end{pmatrix}}_{\text{variance-covariance matrix with exogenous shocks to each variable.}} \right]$$

variance-covariance matrix with exogenous shocks to each variable.

(b) The reduced form model :

$$y_t = B^{-1} \Gamma_0 + B^{-1} \Gamma_1 y_{t-1} + B^{-1} \varepsilon_t$$

$$y_t = C_{3 \times 1} + \Phi_{3 \times 3} y_{t-1} + e_t$$

Reduced form shocks :

$$e_{1t} = B^{-1} \varepsilon_{1t}$$

$$e_{2t} = \varepsilon_{2t} + \beta_{21} e_{1t}$$

$$e_{3t} = \varepsilon_{3t} + \beta_{31} e_{1t} + \beta_{32} \varepsilon_{2t}$$

Continuation of I) CHOLESKY DECOMPOSITION (b) :

VARIANCE - COVARIANCE MATRIX :

$$E[e_t] = 0$$

$$E[e_t e_t'] = E[B^{-1} \varepsilon_t \varepsilon_t' (B^{-1})']$$

$$= B^{-1} E[\varepsilon_t \varepsilon_t'] (B^{-1})'$$

$$= B^{-1} D (B^{-1})'$$

$$= \Omega = \begin{bmatrix} w_{11} & w_{12} & w_{13} \\ w_{21} & w_{22} & w_{23} \\ w_{31} & w_{23} & w_{33} \end{bmatrix}$$

The variance - covariance matrix of the reduced - form shocks is not diagonal. This is because the forecast errors are affected by the shocks.

Furthermore, the variance - covariance matrix is a positive definite symmetric matrix.

It depends on B and a diagonal matrix with positive elements.

II) SHORT RUN RESTRICTIONS

(b) Reduced-form VAR:

$$Z_t = A_0^{-1} \alpha + \sum_{i=1}^{24} A_0^{-1} A_i Z_{t-i} + A_0^{-1} \varepsilon_t$$

Reduced-form shocks:

$$A_0^{-1} \varepsilon_t = e_t = \begin{bmatrix} e_t^{\Delta \text{prod}} \\ e_t^{\text{rea}} \\ e_t^{\text{rpo}} \end{bmatrix} = \begin{bmatrix} a_{11} & 0 & 0 \\ a_{21} & a_{22} & 0 \\ a_{31} & a_{32} & a_{33} \end{bmatrix} \begin{bmatrix} \varepsilon_t^{\Delta \text{prod}} \\ \varepsilon_t^{\text{rea}} \\ \varepsilon_t^{\text{rpo}} \end{bmatrix}$$

From this we can identify ε_t from e_t .

- $e_t^{\Delta \text{prod}} = a_{11} \varepsilon_t^{\Delta \text{prod}}$
- $e_t^{\text{rea}} = a_{21} \varepsilon_t^{\Delta \text{prod}} + a_{22} \varepsilon_t^{\text{rea}}$
- $e_t^{\text{rpo}} = a_{31} \varepsilon_t^{\Delta \text{prod}} + a_{32} \varepsilon_t^{\text{rea}} + a_{33} \varepsilon_t^{\text{rpo}}$

For $a_{11} = a_{22} = a_{33} = 1$ due to the Cholesky Decomposition:

- $\rightarrow e_t^{\Delta \text{prod}} = \varepsilon_t^{\Delta \text{prod}}$
- $\rightarrow e_t^{\text{rea}} = a_{21} \varepsilon_t^{\Delta \text{prod}} + \varepsilon_t^{\text{rea}}$
- $\rightarrow e_t^{\text{rpo}} = a_{31} \varepsilon_t^{\Delta \text{prod}} + a_{32} \varepsilon_t^{\text{rea}} + \varepsilon_t^{\text{rpo}}$

We obtain the structural shocks:

$$\Rightarrow \varepsilon_t^{\Delta \text{prod}} = e_t^{\Delta \text{prod}}$$

$$\Rightarrow \varepsilon_t^{\text{rea}} = e_t^{\text{rea}} - a_{21} \varepsilon_t^{\Delta \text{prod}}$$

$$\Rightarrow \varepsilon_t^{\text{rpo}} = e_t^{\text{rpo}} - a_{31} \varepsilon_t^{\Delta \text{prod}} - a_{32} \varepsilon_t^{\text{rea}}$$

SHORT-RUN RESTRICTIONS:

1. Solve for the identification problem
 2. Use the recursive identification method.
 3. Construct a set of uncorrelated structural shocks directly from the reduced-form shocks.
 4. Assume certain shocks have effect on only some variables at time t .
- B is the lower triangular in VAR(1).