

# Posterior Consistency in Bayesian Nonparametric Models

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# Framework

# Modeling the dataset

- **Dataset:**  $(x_i)_{i=1}^n$ , where  $x_i \in \mathcal{X}$  for all  $i$
- $\mathcal{X}$  is a **Polish space**, endowed such space with its Borel sigma algebra,  $\mathcal{B}(\mathcal{X})$
- We consider a **random sequence** defined on an underlying experiment:

$$(\Omega, \mathcal{F}, \mathbb{P}) \xrightarrow{X} (\mathcal{X}^\infty, \mathcal{B}(\mathcal{X}^\infty)) \xrightarrow{X^{(n)}} (\mathcal{X}^n, \mathcal{B}(\mathcal{X}^n))$$

where  $X := (X_i)_{i \geq 1}$  and  $X^{(n)} := (X_i)_{i=1}^n$

- We introduce a **random parameter**:

$$\theta : (\Omega, \mathcal{F}, \mathbb{P}) \rightarrow (\Theta, \mathcal{B}(\Theta))$$

where  $\Theta$  is Polish as well, and  $\theta$  is distributed according to  $\Pi$ , the **prior**

- For all  $n \in \mathbb{N}$  we define a **regular conditional probability** on  $\mathcal{X}^n$  through the family:

$$\mathcal{M}_n = \{P_\theta^n : \theta \in \Theta\}$$

- We use a product structure for  $P_\theta^n$ , so that  $X_i | \theta \stackrel{iid}{\sim} P_\theta$ .

- Under regularity conditions, there exists a **Kolmogorov extension**  $P_\theta^\infty$  of  $\{P_\theta^n\}$  on  $\mathcal{X}^\infty$ .
- In turn, **there exists a probability measure  $\mathbf{P}$**  on  $\mathcal{X}^\infty \times \Theta$ , such that for all  $A \in \mathcal{B}(\mathcal{X}^\infty)$ , for all  $B \in \mathcal{B}(\Theta)$

$$\mathbb{P}(X \in A, \theta \in B) = \mathbf{P}(A \times B) = \int_B P_\theta^\infty(A) d\Pi(\theta)$$

- In turn, there exists a regular conditional probability  $\bar{\Pi}_n(\cdot | X^{(n)})$  on  $\mathcal{X}^\infty \times \Theta$  from which we can define both a **posterior**  $\Pi_n$  on  $\Theta$  and a **predictive**  $P_{\Pi_n}$  on  $\mathcal{X}^\infty$ , integrating out  $\mathcal{X}^\infty$  or  $\Theta$ , respectively.

# Exchangeability

## Definition

A random sequence  $X = (X_i)_{i \geq 1}$  is exchangeable if and only if for all  $n \in \mathbb{N}$ , for all bijective functions  $\sigma : \{1, \dots, n\} \rightarrow \{1, \dots, n\}$  the distribution of  $(X_1, \dots, X_n)$  is the same as the distribution of  $(X_{\sigma(1)}, \dots, X_{\sigma(n)})$ .

# De Finetti Theorem

## Theorem

Let  $\mathcal{P}_{\mathcal{X}}$  be the space of probability measures on  $\mathcal{X}$ , and let  $\mathcal{B}(\mathcal{P}_{\mathcal{X}})$  be its Borel sigma-algebra.  $X$  is exchangeable if and only if there exists a **random probability measure**  $\theta$  on  $(\mathcal{P}_{\mathcal{X}}, \mathcal{B}(\mathcal{P}_{\mathcal{X}}))$  distributed according to  $\Pi$  such that, for all  $B \in \mathcal{B}(\mathcal{X}^{\infty})$ :

$$\mathbb{P}(X \in B) = \int_{\Theta} \theta^{\infty}(B) d\Pi(\theta)$$

where  $\theta^{\infty}$  is a probability measure on  $(\mathcal{X}^{\infty}, \mathcal{B}(\mathcal{X}^{\infty}))$  such that for all rectangles  $C = \times_{i=1}^{\infty} C_i$ , with  $C_i \in \mathcal{B}(\mathcal{X})$  for all  $i$ :

$$\theta^{\infty}(C) = \prod_{i=1}^{\infty} \theta(C_i)$$



- In light of de Finetti's theorem, it is very natural to set  $\Theta = \mathcal{P}_{\mathcal{X}}$ .
- However, we could also use a **different parametric space**:  $\hat{\Theta}$ , endowed with its Borel sigma-algebra  $\mathcal{B}(\hat{\Theta})$ .
- Let  $\tilde{g} : \hat{\Theta} \rightarrow \mathcal{P}_{\mathcal{X}}$  be defined by  $\tilde{g}(\hat{\theta}) = P_{\hat{\theta}}$ . Once we restrict the codomain to  $\Theta = \tilde{g}(\hat{\Theta})$ ,  $g : \hat{\theta} \mapsto P_{\hat{\theta}}$  is **bijective**. We will assume that  $g$  is also **bimeasurable**, i.e. measurable and with a measurable inverse.

## Consistency: definition and justification

# Consistency: definition

## Definition

Let  $\hat{\Theta}$  be Polish. The posterior  $\hat{\Pi}_n(\cdot|X^{(n)})$  is said to be consistent at  $\hat{\theta}_0$  if

$$\hat{\Pi}_n(\cdot|X^{(n)}) \implies \delta_{\hat{\theta}_0}$$

either  $\theta_0^\infty$ -almost surely or in  $\theta_0^n$ -probability.

## Why is consistency relevant for a subjective Bayesian?

Let  $(\alpha_n)_{n \geq 1}$  and  $(\beta_n)_{n \geq 1}$  be sequences of probability measures defined on  $\hat{\Theta}$ .

$\alpha_n$  and  $\beta_n$  **merge weakly** if  $R(\alpha_n, \beta_n) \rightarrow 0$  for all  $R : \mathcal{P}_{\hat{\Theta}} \times \mathcal{P}_{\hat{\Theta}} \rightarrow \mathbb{R}$  such that

- $R(\alpha, \alpha) = 0$  for all  $\alpha \in \mathcal{P}_{\hat{\Theta}}$
- $R$  is continuous.

Let  $G_{\hat{\Pi}} = \{(\hat{\theta}_0, x) \in \hat{\Theta} \times \mathcal{X}^\infty : \hat{\Pi}_n(\cdot | X^{(n)} = x^{(n)}) \implies \delta_{\hat{\theta}_0}\} \subset \hat{\Theta} \times \mathcal{X}^\infty$   
 be the good set where consistency holds.

## Theorem

Let  $g : \hat{\theta} \mapsto \theta = P_{\hat{\theta}}$  be an omeomorphism. Let  $\hat{\Pi}$  be a prior on  $\hat{\Theta}$ , and let us pick a specific version of the posterior  $\hat{\Pi}_n$ .

The following are equivalent.

- 1  $\hat{\Pi}_n$  is consistent for all  $\hat{\theta}_0 \in \hat{\Theta}$
- 2 for all  $\hat{\nu} \in \mathcal{P}_{\hat{\Theta}}$ ,  $\mathbf{P}_{\hat{\nu}}(G_{\hat{\Pi}}) = 1$
- 3 for all  $\hat{\nu} \in \mathcal{P}_{\hat{\Theta}}$ ,  $\hat{\Pi}_n$  and  $\hat{\nu}_n$  merge weakly  $\mathbf{P}_{\hat{\nu}}$ -almost surely.
- 4 for all  $\hat{\nu} \in \mathcal{P}_{\hat{\Theta}}$ ,  $P_{\hat{\Pi}_n}$  and  $P_{\hat{\nu}_n}$  merge  $\mathbf{P}_{\hat{\nu}}$ -almost surely.

## Consistency: Schwartz

# Background

- Let  $\Theta = \{\theta \in \mathcal{P}_{\mathcal{X}} : \theta \ll \nu\} \subset \mathcal{P}_{\mathcal{X}}$
- Let  $\hat{\Theta} = \{[f_{\theta}] : \theta \in \Theta\}$ , where  $[f_{\theta}]$  is the **class** of Radon-Nikodym derivatives of  $\theta$  w.r.t.  $\nu$ .
- Let  $\tau$  be the restriction of the weak topology on  $\Theta$ . Let  $h : \theta \mapsto [f_{\theta}]$  and  $g := h^{-1}$ . We **induce** a topology on  $\hat{\Theta}$  through  $g$  and  $\tau$ :

$$\hat{\tau} := \{A \subset \hat{\Theta} : A = g^{-1}(B), B \in \tau\}$$

By construction,  $g$  is an isomorphism and so **bimeasurable**. Thus, priors defined on  $(\Theta, \mathcal{B}(\Theta))$  induce priors on  $(\hat{\Theta}, \mathcal{B}(\hat{\Theta}))$  and viceversa.

- For all  $\hat{\theta}_0 \in \hat{\Theta}$ , for all  $\epsilon > 0$  we define a **Kullback-Leibler neighborhood** of  $\hat{\theta}_0$  to be:

$$KL(\hat{\theta}_0, \epsilon) := \{\hat{\theta} \in \hat{\Theta} : KL(\hat{\theta}_0, \hat{\theta}) < \epsilon\}$$

- Then, the **Kullback-Leibler support** of  $\hat{\Pi}$  is:

$$KL(\hat{\Pi}) \stackrel{d}{=} \{\hat{\theta} \in \hat{\Theta} : \forall \epsilon > 0, \hat{\Pi}(KL(\hat{\theta}, \epsilon)) > 0\}$$

- Does  $KL(\hat{\theta}_0, \epsilon) \in \mathcal{B}(\hat{\Theta})$  for all  $\epsilon$  and for all  $\theta_0$ ? We will **assume** so.



## Theorem

Let  $\hat{\Pi}$  be a prior on  $\hat{\Theta}$ , let  $\hat{\theta}_0 \in \hat{\Theta}$ . If  $\hat{\theta}_0 \in KL(\hat{\Pi})$  then  $\hat{\Pi}_n(\cdot | X^{(n)}) \implies \delta_{\hat{\theta}_0}, \theta_0^\infty$ -almost surely.