In a regression problem, we have a set of evidence variables X, , ,, XD and a response variable Y that we want to predict.

For instance, let's say we want to predict chalesterol level given age and weight:

often called
$$X$$
 (evidence vars) Y (response var) X_1 X_2 X_3 X_4 X_5 X_6 X_7 X_8 X_8 X_8 X_9 X_9

We have N training examples, each consisting of a vector x(n) and a scalar y(n), to learn from. Note that the entire dataset can be captured as an evidence matrix X and response vector y.

X (evid	lence n	natrix)	y (response vector)
[1	24	150	[182]
1 1	1	164	[182] 210 : 202

2) We assume the response variable is generated using the tollowing steps:

- we start with a vector of weights (one weight per evidence variable), e.g.

$$|w| = \begin{bmatrix} w_1 \\ w_2 \\ w_3 \end{bmatrix} = \begin{bmatrix} -50 \\ 2 \\ 1 \end{bmatrix}$$

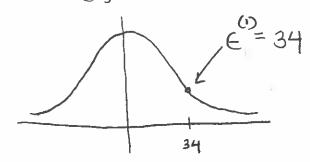
- we compute the weighted linear combination of the evidence variables (this will result in a single number):

$$\chi(i) = \left[\chi_1^{(i)} \chi_2^{(i)} \chi_3^{(i)} \right] \left[w_1 \right]$$
the offset allows us to shift the total p or down by a ...
$$w_3$$

to shift the total up or down by a constant factor

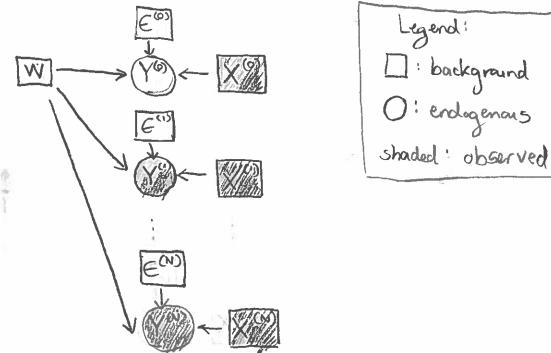
$$= -50.1 + 2.24 + 1.150$$
$$= 148$$

- we sample a random number & from a distribution Pe e.g. E' & Normal (0, 02)



- we compute the response variable as a function p of of x w and E, e.g

(3) Given this setup, we want to predict the value of an Unobserved response variable y given its observed evidence variables x and our training data. In its simplest formulation, the causal diagram looks as follows:



We assume a probability distribution P over the background variables such that all background variables are marginally independent, i.e.

 $P(w, \epsilon^{(n)}, x^{(n)}, x^{(n)}, x^{(n)}) = P_w(w)P_{\epsilon}(\epsilon^{(n)}) P_{\epsilon}(\epsilon^{(n)}) P_{\epsilon}(x^{(n)}) P_{\epsilon}(x^{(n)})$

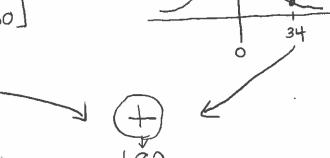
4) What's the deal with the E's?

We'll call flese stochastic terms. The idea is to allow some softness around the deterministic point x w. In other words, just because x w = 148 for age = 24 and weight = 150, that doesn't mean we want the model to claim that EVERY person whose age is 24 and whose weight is 150 must have a cholesterol level of EXACTLY 148.

Instead, we want their cholesterol levels to be disposed around 148.

5) Suppose we choose $P_{\epsilon} \sim Normal(0, \sigma^2)$ for some fixed variance σ^2 , and suppose we choose $P(z, \epsilon) = z + \epsilon$. This is what we did in (2), which gave us a cholesterol tenel of 148 + 34 = 182. For a 24-year-old, 1501b subject.

DETERMINISTIC TERM $W^T \times = \begin{bmatrix} -50 & 2 & 1 \end{bmatrix} \begin{bmatrix} 1 \\ 24 \\ 150 \end{bmatrix}$ = 148



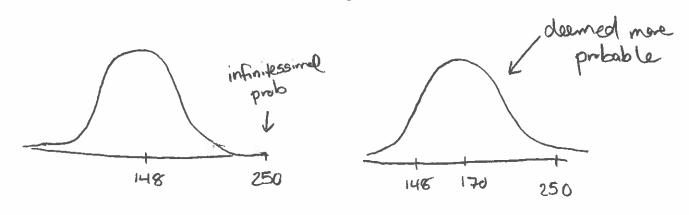
1 These choices:

$$\begin{bmatrix}
P_{\epsilon} \sim \text{Normal}(0, \sigma^{2}) \\
P(Z, \epsilon) = Z + \epsilon
\end{bmatrix}$$

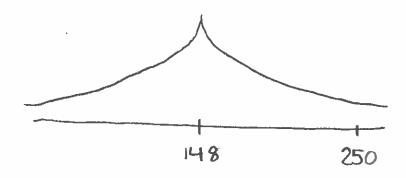
give us the ordinary linear regression model

But it may not be the best choice. One possible downside of the normal distribution is that it has rapidly diminishing tails, so a normal distribution centered at 148 may give a nearly infinitessimal probability to 250.

In practice, this means that using the normal distribution makes regression very sensitive to arthus and norse, because even you have a thousand 24-year-olds with cholesterol 148, just one 24-year-old with cholesterol 250 can cause the learned mean to shift significantly upwards, because this will be deemed more probable than even one person with cholesterol 250:



(8) If this is a problem for you, you can use a "heavy-tailed" distribution (i.e. they diminish in probability much more slowly). One example is the Laplace distribution.



Often, however, such distributions are not as computationally convenient.

This chaice: $\begin{bmatrix}
P_{\epsilon} \sim Laplace(0, b) \\
P(z, \epsilon) = z + \epsilon
\end{bmatrix}$ gives us the "robust" linear regression model

10) Another common choice for Pc and p comes into play when the response variable is Boolean-valued, e.g.:

X (evidence vars)

X1 X2 X3

(offset) (age) (weight)

$$\chi^{(1)} = \begin{bmatrix} 1 & 24 & 150 \end{bmatrix}$$
 $\chi^{(2)} = \begin{bmatrix} 1 & 50 & 164 \end{bmatrix}$

Y (response var)

$$=y^{(N)}$$

1) We could use ordinary linear regression, but then we end up predicting values in the range (-00,00), rather than restricting ausselves to the set 20, 13.

Another chaice

$$P_{\epsilon} \sim \text{Uniform}(0,1)$$

$$P(z,\epsilon) = 1$$

$$E < (1+e^{-z})^{-1}(z) = \begin{cases} 1 & \text{if } E < \frac{1}{1+e^{-z}} \\ 0 & \text{o.w.} \end{cases}$$

gives us the famous logistic regression model

REGRESSION PROBLEMS
(i) Let's unpack this: we generate a random number & between O and 1 from Constant (0,1)
1. (ii) i. we compute a threshold
$ \gamma = \frac{1}{1 + e^{-\chi^{c} n} \omega} $ (note: $0 < \gamma < 1$)
(iii) $y^{(n)} = \begin{cases} 1 & \text{if } \epsilon < \tau \\ 0 & \text{otherwise} \end{cases}$
In other words, the probability that $y''=1$ is equal to $\frac{1}{1+e^{-x^m}w}$
(13) This function, $f(z) = (1 + e^{-z})^{-1}$, looks like this:
when z is negative, $f(z) < \frac{1}{2}$ when z is positive, $f(z) > \frac{1}{2}$

is called the logistic function (or the logit, or the sigmoid).

(H) No matter what regression model we're using, we typically want to predict the value of the unobserved response variable y'(0) given its observed evidence variables x'(0) (recall x'(0) is a vector) and our other observations (i.e. x'(1), y'(1) for n ≥ 1).

In other words, we want argmax P(y'(1)|x'(1), x'(1), y'(1), y'(

(5) The exact computation would be:

= argmax
$$\int P(y^{(0)}, w|_{x^{(0)}}, x^{(0)}, ..., x^{(N)}, y^{(1)}, ..., y^{(N)}) dw$$
[Law of Total Probability]

= argmax
$$\int P(y^{(0)}|w,x^{(0)}) P(w|x^{(1)},x^{(N)},y^{(1)},y^{(N)}) dw$$

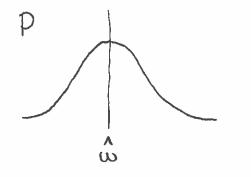
$$\left[d-separation\right]$$

(16) This integral is going to be messy, so let's make a simplifying assumption. Instead of using the actual distribution over weights:

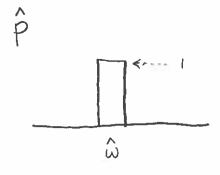
let's use a much simpler distribution:

$$\hat{P}(w|x^{(i)}, ..., x^{(N)}, y^{(i)}, ..., y^{(N)}) = \begin{cases} 1 & \text{if } w = \operatorname{argmax} P(w|x^{(i)}, ..., y^{(N)}) \\ 0 & \text{o.w.} \end{cases}$$

This is called the point estimate approach, because it concentrates all probability mass anto be most likely value:







17) Let's make this approximation:

$$\approx \underset{y(0)}{\operatorname{argmax}} \operatorname{SP}(y^{(0)}|_{w,x^{(0)}}) \hat{P}(w|_{x^{(1)},...,x^{(N)}},y^{(1)},...,y^{(N)}) dw$$

where
$$\hat{w} = \operatorname{argmax} P(w|x^{(1)}, ..., x^{(N)}, y^{(1)}, ..., y^{(N)})$$

(8) So the paint estimate approach to predicting an unknown response y (0) has two steps:

(a) compute the most probable weight vector û given the observations!

(6) compute the most probable response $\hat{y}^{(0)}$ given \hat{w} and evidence $\hat{x}^{(0)}$:

$$\hat{y}^{(0)} = \operatorname{argmax} P(y^{(0)} | \hat{w}, x^{(0)})$$

$$\hat{\omega} = \underset{\omega}{\operatorname{argmax}} P(\omega | \chi^{(1)}, ..., \chi^{(N)}, \chi^{(N)}, \chi^{(N)})$$

= argmax
$$P(x^{(i)}, ..., x^{(i)}, y^{(i)}, ..., y^{(i)}|w)P(w)$$

$$P(x^{(i)}, ..., y^{(i)})$$
[Bayes lule]

= argmax
$$P(x^{(i)}|w)P(x^{(i)}|x^{(i)}w).....P(y^{(N)}|x^{(i)},...,y^{(N-i)}w)P(w)$$

[Chain Rule]

= argmax
$$P(x^{(i)}) P(x^{(i)}) \cdots P(x^{(w)}) P(y^{(i)} | w, x^{(i)}) \cdots P(y^{(w)} | w, x^{(w)}) \cdots P(y^{(w)} | w, x^{(w)})$$

= argmax
$$P(w) \prod_{n=1}^{N} P(y^{(n)} | w, x^{(n)})$$

20 In short:

(a) compute
$$\hat{w} = \operatorname{argmax} P(w) \prod_{n=1}^{N} P(y^{(n)}|w, x^{(n)})$$

(b) compute
$$\hat{y}^{(6)} = \underset{y(6)}{\operatorname{argmax}} P(y^{(6)} | \hat{\omega}, x^{(6)})$$

This is called the MAP (maximum a posteriori) estimate.

(a) A special case of the MAP estimate assumes that P(w) is the same for every possible w. Since it then becomes a constant factor, we can drop how comit from the argmax:

(a) compute $\hat{w} = \operatorname{argmax} \prod_{n=1}^{N} P(y^{(n)}|w,x^{(n)})$

(b) compute $\hat{y}^{(0)} = \underset{y(0)}{\operatorname{argmax}} P(\hat{y}^{(0)}|\hat{\omega}_{x}^{(0)})$

This is called the MLE (maximum likelihood estimate).

P how can this be, if there's an infinite space of w-vectors? don't worry about it