1 General definitions

1.1 Basic

Sample variance

$$S_X^2 = \frac{1}{n-1} \sum_{i=1}^n (X_i - \bar{X})^2$$
 (1.1)

• Sample correlation coefficient

$$r_{X,Y} = \frac{\sum_{i=1}^{n} (X_i - \bar{X})(Y_i - \bar{Y})}{(n-1)\sqrt{(S_X^2 S_Y^2)}}$$
(1.2)

- QQ-plot for cumulative distribution function F is the set of points $\left(q_F\left(\frac{i}{n+1}\right), x_{(i)}\right)$, where $q_F(\cdot)$ is the quantile function for the distribution.
- Mean Squared Error (MSE)

$$MSE(\theta; T(X), g(\theta)) = \mathbb{E}_{\theta} (T(X) - g(\theta))^{2}$$
(1.3)

• Bias-variance decomposition

$$MSE(\theta; T(X)) = var_{\theta}T + (\mathbb{E}_{\theta}T(X) - g(\theta))^{2}$$
(1.4)

• Empirical distribution function

$$\hat{F}(x) = \sum_{i=1}^{n} \mathbb{I}(X_i \le x) \tag{1.5}$$

1.2 k-th order statistic $X_{(k)}$

 $X_{(k)} - k - th$ order statistic distribution for n i.i.d. variables from continuous distribution F.

$$f_{(k)}(x) = \frac{n!}{(k-1)!(n-k)!} F(x)^{k-1} (1 - F(x))^{n-k} f(x)$$
 (1.6)

$$F_{(k)}(x) = \sum_{j=k}^{n} \binom{n}{j} F(x)^{j} (1 - F(x))^{n-j}$$
(1.7)

$$\mathbb{E}F(X_{(k)}) = \frac{k}{n+1} \tag{1.8}$$

2 Important distributions

- Poisson distribution Poisson(λ), $\lambda > 0$
 - $-\lambda$ is the average number of events per interval
 - pdf

$$p_{\lambda}(k) = e^{-\lambda} \frac{\lambda^k}{k!} \tag{2.1}$$

• Geometric distribution $G(\theta)$, $0 \le \theta \le 1$

$$f_{\theta}(k) = (1 - \theta)^{1 - k} \theta \tag{2.2}$$

- cdf

$$F_{\theta}(k) = 1 - (1 - \theta)^k \tag{2.3}$$

• Exponential distribution $F(x; \lambda)$

$$f_{\lambda}(x) = \lambda e^{-\lambda x} \tag{2.4}$$

- cdf

$$F_{\lambda}(x) = 1 - e^{-\lambda x} \tag{2.5}$$

$$-\mathbb{E}_{\lambda}X = 1/\lambda$$

• Beta distribution $B(\alpha, \beta), \ \alpha, \beta > 0$

- pdf

$$f_{\alpha,\beta}(x) = \frac{x^{\alpha-1}(1-x)^{\beta-1}}{B(\alpha,\beta)}, \ B(\alpha,\beta) \equiv \frac{\Gamma(\alpha)\Gamma(\beta)}{\Gamma(\alpha+\beta)}$$
(2.6)

- Weibull distribution
 - α and λ are the "shape" and 'inverse scale" parameters.
 - pdf

$$f_{\lambda,\alpha}(x) = \lambda^{\alpha} \alpha x^{\alpha - 1} e^{-(\lambda x)^{\alpha}}$$
(2.7)

- cdf

$$F_{\lambda,\alpha}(x) = 1 - e^{-(\lambda x)^{\alpha}} \tag{2.8}$$

- Gamma distribution $\Gamma(\alpha, \lambda), \ \alpha > 0, \lambda > 0$
 - $-\alpha$ and λ are known as "shape" and "inverse scale" parameters.
 - pdf

$$f_{\alpha,\lambda}(x) = \frac{x^{\alpha-1}\lambda^{\alpha}e^{-\lambda x}}{\Gamma(\alpha)}$$
 (2.9)

– cdf (where $\gamma(s,x)=\int_0^x t^{s-1}e^{-t}dt$ — is the "incomplete gamma function")

$$F_{\alpha,\lambda}(x) = \frac{\gamma(\alpha, x\beta)}{\Gamma(\alpha)}$$
 (2.10)

3 Fundamental results

Theorem 3.1. Let $X_1, \ldots X_n$ be an i.i.d. ramdom variables from the $N(\mu, \sigma^2)$ distribution, then

- 1. \bar{X} is $N(\mu, \sigma^2/n)$ distributed;
- 2. $(n-1)S_X^2/\sigma^2$ is χ_{n-1}^2 -distributed (see 1.1);
- 3. \bar{X} and S_X^2 are independent;
- 4. $\sqrt{n}(\bar{X} \mu)/\sqrt{S_X^2}$ has the t_{n-1} distribution.

Proof.
$$||X||^2 - n\bar{X}^2 = (n-1)S_X^2$$

4 Estimators

4.1 Maximum of *n* uniformally distributed statistics

Set up: $X_1, X_2, ... X_n$ i.i.d. drown from $U[0, \theta]$, where θ is the parameter of interest.

- $\bullet \ \hat{\theta} = 2\bar{X}_n$
 - method of moments estimator
 - $-\ unbiased$
 - $MSE(\theta, \hat{\theta}) = \frac{\theta^2}{3n}$, see (1.6)

 $\bullet~X_{(n)}$ — n-th order statistic, i.e. maximum.

$$-\mathbb{E}_{\theta}X_{(n)} = \frac{n}{n+1}\theta$$
, see (1.6)

$$- \text{MSE}(\theta, X_{(n)}) = \frac{2\theta^2}{(n+2)(n+1)}$$

- $\bullet \ \frac{n+2}{n+1}X_{(n)}$
 - best estimator of the form $cX_{(n)}$
 - $MSE(\theta, \frac{n+2}{n+1}X_{(n)}) = \frac{\theta^2}{(n+1)^2}$

4.2 Univariate normal distribution

•
$$(\hat{\mu}, \hat{\sigma}^2) = \left(\bar{X}_n, \frac{1}{n} \sum_{i=1}^n (X_i - \bar{X}_n)\right) = \left(\bar{X}_n, \frac{n-1}{n} S_X^2\right)$$

- maximum likelihood estimator
- method of moments estimator
- $-\hat{\mu}$ is unbiased
- $-\mathbb{E}_{(\mu,\sigma^2)}\hat{\sigma}^2 = \frac{n-1}{n}\sigma^2$

4.3 Empirical distribution function

Let $X_1, \ldots X_n$ be an i.i.d. sample drawn from the distribution F.

- The empirical distribution function (ecdf) $\hat{F}(x) = \sum_{i=1}^{n} \mathbb{I}(X_i \leq x)$ (see 1.1)
 - unbiased
 - $\mathbf{cov}_F\left(\hat{F}(u), \hat{F}(v)\right) = n^{-1}(F(\min(u, v)) F(u)F(v))$ positively correlated