Problem 1 . Gradient

$$g\left(A,\beta,\chi\right) = \sum_{i=1}^{p} \beta_{(i)} \left(\sum_{j=1}^{k} A_{(ij)},\chi_{(j)}\right)^{2}$$

BERP, AERPXK, XERK, compute 7,9, 7,9, VB9

$$g \in R$$
 $g(A, \beta, x) = \beta^{T}(Ax)^{3}$

$$\begin{cases} \nabla_X g = 3 A^T [\beta \circ (Ax)^2] \end{cases}$$

$$\nabla_{\beta} q = (Ax)^3$$

Proof:

$$\frac{\int}{\beta^{T} \alpha^{(2)}} \longrightarrow \frac{\partial g}{\partial \beta} = \alpha^{(2)} = (Ax)^{3}$$

$$\alpha^{(2)} : \frac{\partial g}{\partial \alpha^{(2)}} = \beta \qquad \frac{\partial g}{\partial \alpha^{(2)}_{(j)}} = \beta_{(j)}$$

$$\frac{\int}{(Z^{(2)})^{3}}$$

$$\frac{\partial}{\partial z^{(2)}} = \sum_{j} \frac{\partial g}{\partial \alpha^{(j)}_{(j)}} \frac{\partial \alpha^{(2)}_{(j)}}{\partial z^{(2)}} = \sum_{j} \beta_{(j)}$$

$$Z^{(i)} = \sum_{j} \frac{\partial g}{\partial z^{(i)}} = \sum_{j} \frac{\partial g}{\partial \alpha^{(j)}_{(j)}} \frac{\partial \alpha^{(i)}_{(j)}}{\partial z^{(i)}} = \sum_{j} \beta_{(j)} \frac{\partial e_{j} T(z^{(i)})^{3}}{\partial z^{(i)}} = \sum_{j} \beta_{(j)} 3(z^{(i)})^{2} e_{j}$$

$$= 3 \beta \circ (Ax)^{2}$$

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$$= \sum_{j} 3 \left[\beta \circ (Ax)^{2}\right]_{(j)} e_{j} x^{T} \qquad \chi : \frac{\partial 9}{\partial x} = \sum_{j} \frac{\partial 9}{\partial z^{(j)}} \frac{\partial z^{(j)}}{\partial x} = \sum_{j} 3 \left[\beta \circ (Ax)^{2}\right]_{(j)} \frac{\partial e_{j}^{T} Ax}{\partial x}$$

$$= 3 \left[\beta \circ (Ax)^{2}\right] x^{T} \qquad = 3 \sum_{j} \left[\beta \circ (Ax)^{2}\right]_{(j)} \left(e_{j}^{T} A\right)^{T}$$

=
$$3 \frac{5}{5} \left[\beta \circ (Ax)^{2}\right]_{(j)} A^{T} e_{j}$$

= $3 A^{T} \left[\beta \circ (Ax)^{2}\right]$

$$\frac{\partial Z_{(j)}^{(2)}}{\partial A} = \frac{\partial e_{j}^{T}Ax}{\partial A} = \frac{\partial tr(e_{j}^{T}Ax)}{\partial A}$$
$$= \frac{\partial tr(Axe_{j}^{T})}{\partial A} = \frac{\partial tr[A(e_{j}^{T}x)^{T}]}{\partial A}$$

$$= e_j \chi^T$$

Problem 2: Weighted Logistic Regression

suppose xi E Rd O E Rd

Since L(0) is a function of a vector $\in \mathbb{R}^d$ the gradient is also a vector $\in \mathbb{R}^d$

the jth entry of gradient is

$$\begin{bmatrix} \nabla L(\theta) \end{bmatrix}_{cj} = \frac{\partial L(\theta)}{\partial \theta_{cj}} \\
= \frac{\partial \sum_{i=1}^{n} d_{i} \left[\log \left(1 + \exp \left(x_{i}^{T} \theta \right) \right) - y_{i} x_{i}^{T} \theta \right]}{\partial \theta_{cj}} \\
= \sum_{i=1}^{n} d_{i} \frac{\partial \left[\log \left(1 + \exp \left(x_{i}^{T} \theta \right) \right) - y_{i} x_{i}^{T} \theta \right]}{\partial \theta_{cj}} \quad \text{linearity of clerivative} \\
= \sum_{i=1}^{n} d_{i} \frac{\partial \log \left(1 + \exp \left(x_{i}^{T} \theta \right) \right)}{\partial \theta_{cj}} \frac{\partial y_{i} x_{i}^{T} \theta}{\partial \theta_{cj}} \\
= \sum_{i=1}^{n} d_{i} \left[\frac{\exp \left(x_{i}^{T} \theta \right) \left[x_{i}^{T} \right]_{cj}}{1 + \exp \left(x_{i}^{T} \theta \right)} - y_{i}^{T} \left[x_{i}^{T} \right]_{cj}} \right]$$

$$= \sum_{i=1}^{h} \left[\operatorname{cl}_{i} \left(\frac{\exp(\alpha_{i} \tau_{\theta})}{1 + \exp(\alpha_{i} \tau_{\theta})} - y_{i} \right) \left[\chi_{i} \right]_{ij} \right]$$

Then
$$\nabla L(\theta) = \sum_{i=1}^{n} \left[di \left(\frac{e^{x} p(x_{i}^{T} \theta)}{1 + e^{x} p(x_{i}^{T} \theta)} - y_{i} \right) x_{i} \right]$$

Problem 3 Weighted Linear Regression

$$y_i = x_i^T \beta^* + w_i \qquad w_i \sim N(0, \sigma_i^2)$$

$$L(y; \chi, \beta^*, \sigma^2) = \prod_{i=1}^n \frac{1}{\sqrt{2\pi\sigma_i^2}} e^{xp} \left[-\frac{(y_i - \chi_i^T \beta^*)^2}{2\sigma_i^2} \right]$$

$$NLL = -\log L = -\log \left(\frac{1}{12\pi\sigma_{i}^{2}} \frac{1}{\sqrt{2\pi\sigma_{i}^{2}}} \exp \left[\frac{(y_{i} - \chi_{i}^{7}\beta^{*})^{2}}{2\sigma_{i}^{2}} \right] \right)$$

$$= -\sum_{i=1}^{n} \log \left(\frac{1}{\sqrt{2\pi\sigma_{i}^{2}}} \exp \left[-\frac{(y_{i} - \chi_{i}^{7}\beta^{*})^{2}}{2\sigma_{i}^{2}} \right] \right)$$

$$= -\sum_{i=1}^{n} \left[\log \frac{1}{\sqrt{2\pi\sigma_{i}^{2}}} - \frac{(y_{i} - \chi_{i}^{7}\beta^{*})^{2}}{2\sigma_{i}^{2}} \right]$$

$$= \frac{h}{2} \log_{2}\pi + \sum_{i=1}^{n} \log_{3}\sigma_{i} + \frac{1}{2\sigma_{i}^{2}} \sum_{i=1}^{n} (y_{i} - \chi_{i}^{7}\beta^{*})^{2}$$

The effect of adding a varible $w_i \sim N(0, \sigma_i^2)$ to a Gaussian compared to adding a constant $w \sim N(0, \sigma^2)$ to a Gaussian in standard linear regression is that When we try to find the optimal β by solving argmin $NLL(\beta)$ by setting $\nabla_{\beta} NLL(\beta) = 0$, every sample will contribute a $\frac{1}{2\sigma_i^2}$ ratio which means every sample is weighted

P4: Regularization
$$L(\theta) = f(\theta) + \frac{2}{2} \theta^{T} \theta$$

$$\theta_{k+1} = \theta_k - \eta_k (\nabla f(\theta_k) + \lambda \theta_k)$$

where g_k is the learning rate

1.41

11.

P5 Concepts

a) $X \in \mathbb{R}^{100 \times 4}$, $\chi_i \in \mathbb{R}^4$ is the ith flower sample in the training set.

$$y_i = \begin{cases} 1 & \text{species } A \\ 0 & \text{species } B \end{cases}$$

For a new test case $v \in R^4$, we calculate its distance between each flower sample in the set, so we have 100 distances, then we select the top 5 flower sample data points with shortest distances

Then we look at the species of these 5 data points, we choose the most common species to be the label of the new test case.

- b) k=1 will lead to an overfitted model, label of a new test case will be the species of the 1st nearest data point, without capturing comprehensive. features of the data set. Test error will be very large.
- C) Because the data set only has 100 data points. Set k=100 then the label of a new test case will always be the species of majority of the data set.

$$\nabla_{\theta} l(x_{i}^{\mathsf{T}}\theta, y_{i}) = l'(x_{i}^{\mathsf{T}}\theta, y_{i}) \nabla_{\theta}(x_{i}^{\mathsf{T}}\theta)
= \left[(y - y') \middle|_{(x_{i}^{\mathsf{T}}\theta, y_{i})} \right] \chi_{i}
= (y_{i} - x_{i}^{\mathsf{T}}\theta) \chi_{i}$$

$$\nabla_{\theta} l (\chi; \tau_{\theta}, y_{i}) = l'(\chi; \tau_{\theta}, y_{i}) \nabla_{\theta} (\chi; \tau_{\theta})$$

$$= \left[sign (y - y') | (\chi; \tau_{\theta}, y_{i}) \right] \cdot \chi_{i}$$

$$= sign (y_{i} - \chi; \tau_{\theta}) \chi_{i}$$

Likelihood
$$L(\theta^*) = \prod_{i=1}^{n} f_{\theta}(y_i) = \left\{\prod_{i=1}^{n} \exp\left[-\left(y_i - x_i^{\mathsf{T}} \theta^*\right)\right] \quad y_i - x_i^{\mathsf{T}} \theta^* \ge 0\right\}$$

$$y_i - x_i^{\mathsf{T}} \theta^* < 0$$

$$Log L(0^*) = log L(0^*) = log \stackrel{\text{T}}{\underset{i=1}{\square}} exp \left[-(y_i - x_i^{\intercal} 0^*) \right]$$
$$= \stackrel{\text{T}}{\underset{i=1}{\square}} \left(\chi_i^{\intercal} 0^* - y_i \right)$$

=)
$$\log L(0^*) = \sum_{i=1}^{n} (x_i \circ x_i - y_i)$$
 for $y_i - x_i \circ x_i = 0$.

Problem 8 MLE

$$\chi_{i} \sim N(0, \sigma^{2}) \quad i.i.d$$

$$p(\chi_{i}|\sigma^{2}) = \frac{1}{\sqrt{2\pi\sigma^{2}}} \exp\left(-\frac{\chi_{i}^{2}}{2\sigma^{2}}\right)$$

$$L(\chi_{i}|\sigma^{2}) = \prod_{i=1}^{n} \left(\frac{1}{\sqrt{2\pi\sigma^{2}}} \exp\left(-\frac{\chi_{i}^{2}}{2\sigma^{2}}\right)\right)$$

$$NLL(\sigma^{2}) = -\log L(\chi_{i}|\sigma^{2}) = -\frac{n}{2} \log\left[\frac{1}{\sqrt{2\pi\sigma^{2}}} \exp\left(-\frac{\chi_{i}^{2}}{2\sigma^{2}}\right)\right]$$

$$= \frac{n}{2} \log_{2} 2\pi + \frac{n}{2} \log_{2} \sigma^{2} + \frac{1}{2\sigma^{2}} \sum_{j=1}^{n} \chi_{i}^{2}$$

$$\frac{\partial NLL(\sigma^{2})}{\partial \sigma^{2}} = \frac{\partial\left(\frac{n}{2}\log_{2} \sigma^{2} + \frac{1}{2\sigma^{2}}\sum_{j=1}^{n} \chi_{i}^{2}\right)}{\partial \sigma^{2}}$$

$$= \frac{n}{2} \frac{1}{\sigma^{2}} - \frac{n}{2} \frac{\chi_{i}^{2}}{2\sigma^{2}} - \frac{1}{(\sigma^{2})^{2}}$$

$$\operatorname{Set} \frac{\partial NLL(\sigma^{2})}{\partial \sigma^{2}} = 0$$

$$\Rightarrow \frac{1}{n^2 = \sum_{i=1}^{n} \chi_i^2}$$

The MLE estimate for σ^2 is $\sum_{i=1}^{n} \chi_i^2$

Problem 9 MLE $\widehat{\theta}_{j} = \log \frac{N_{j}}{n} \quad N_{j} \text{ is the number of } [y_{i}]_{(j)} = 1 \quad \text{in n samples}$ $y_{i} \in \mathbb{R}^{k}$ $y_{i} = \begin{bmatrix} 0 \\ 1 \\ 0 \end{bmatrix} \Rightarrow 2^{th} \underset{now}{[y_{i}]_{(2)}} = 1 \quad \Rightarrow \quad Z_{i} = 2$ $y_{i} = \begin{bmatrix} 0 \\ \vdots \\ 0 \end{bmatrix} \Rightarrow j^{th} \underset{now}{\text{row}}$ $y_{i} = \begin{bmatrix} 0 \\ \vdots \\ 0 \end{bmatrix} \Rightarrow j^{th} \underset{now}{\text{row}}$ $y_{i} = \begin{bmatrix} 0 \\ \vdots \\ 0 \end{bmatrix} \Rightarrow j^{th} \underset{now}{\text{row}}$

For convenience, denote vector $Z \in \mathbb{R}^n$ $Z_{(i)} = k$ if $[y_i]_{(k)} = 1$ clefine $N_{(i)} = n$ number of $Z_{(i)} = j$ $N \in \mathbb{R}^k$ $\sum_{j=1}^k N_{(j)} = n$

 $L(y;\theta) = L(z;\theta) = \prod_{j=1}^{n} P(y) = \prod_{j=1}^{n} (P_{j})^{N_{ij}},$ $= -\frac{1}{2} \log (P_{j})^{N_{ij}},$ $= -\frac{1}{2} N_{ij}, \log (P_{j})^{N_{ij}},$ $= -[\frac{1}{2} N_{ij}, \log (P_{j}) + \lambda (1 - \frac{1}{2} P_{j})]$ $= \frac{1}{2} P_{j} = 1$

 $\begin{cases}
\frac{\partial NLL(p_{j}, \lambda)}{\partial p_{j}} = \frac{N(j) - \lambda = 0}{p_{j}} \\
\frac{\partial NLL(p_{j}, \lambda)}{\partial \lambda} = 1 - \sum_{j=1}^{k} p_{j} = 0
\end{cases}
\Rightarrow \begin{cases}
\hat{p}_{j} = \frac{N(j)}{\lambda} = \frac{N(j)}{n} = \exp(\hat{\theta}_{j}) \\
\lambda = n
\end{cases}$ $\hat{p}_{j} = \log \frac{N(j)}{n}$

Also we could verify: $\sum_{j=1}^{k} \exp(\hat{\theta}_{j}) = \sum_{j=1}^{k} \exp(\log \frac{N(j)}{n}) = \sum_{j=1}^{k} \frac{N(j)}{n} = \frac{1}{n} \cdot n = 1$

Problem 10 Concept

First I will standardize the new data point by the same way training set has clone $\chi' = \frac{\chi - m}{s}$ (where m is mean, s is standard deviation of training set). Then apply function h to χ' to predict its label. Because the nearness of NN is based on distance, without standardize columns, if one variable is on the scale of billions while another is on tens. then scale of billions will contributes more to error. Thus the prediction will bias towards samples that are close in billions scale feature and ignore the other. Since the training set was standardized, the test set should also be standardized to the same scale.