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Estimating a Probability Distribution: Setting

- Let p(y) represent a probability distribution on \mathcal{Y} .
- p(y) is **unknown** and we want to **estimate** it.
- Assume that p(y) is either a
 - ullet probability density function on a continuous space \mathcal{Y} , or a
 - probability mass function on a discrete space \mathcal{Y} .
- Typical y's:
 - $y = \mathbf{R}$; $y = \mathbf{R}^d$ [typical continuous distributions]
 - $\mathcal{Y} = \{-1, 1\}$ [e.g. binary classification]
 - $\mathcal{Y} = \{0, 1, 2, \dots, K\}$ [e.g. multiclass problem]
 - $y = \{0, 1, 2, 3, 4...\}$ [unbounded counts]

Evaluating a Probability Distribution Estimate

- Before we talk about estimation, let's talk about evaluation.
- Somebody gives us an estimate of the probability distribution

$$\hat{p}(y)$$
.

- How can we evaluate how good it is?
- We want $\hat{p}(y)$ to be descriptive of **future** data.

Likelihood of a Predicted Distribution

Suppose we have

$$\mathfrak{D} = (y_1, \dots, y_n)$$
 sampled i.i.d. from true distribution $p(y)$.

• Then the **likelihood** of \hat{p} for the data \mathcal{D} is defined to be

$$\hat{\rho}(\mathcal{D}) = \prod_{i=1}^{n} \hat{\rho}(y_i).$$

• If \hat{p} is a probability mass function, then likelihood is probability.



Parametric Models

Definition

A parametric model is a set of probability distributions indexed by a parameter $\theta \in \Theta$. We denote this as

$$\{p(y;\theta) \mid \theta \in \Theta\},\$$

where θ is the parameter and Θ is the parameter space.

- Below we'll give some examples of common parametric models.
 - But it's worth doing research to find a parametric model most appropriate for your data.
- We'll sometimes say family of distributions for a probability model.

Poisson Family

- Support $\mathcal{Y} = \{0, 1, 2, 3, \ldots\}.$
- Parameter space: $\{\lambda \in \mathbb{R} \mid \lambda > 0\}$
- Probability mass function on $k \in \mathcal{Y}$:

$$p(k;\lambda) = \lambda^k e^{-\lambda}/(k!)$$

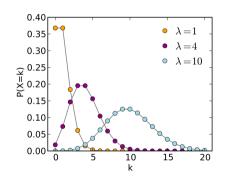
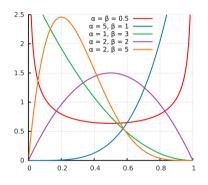


Figure is "Poisson pmf" by Skbkekas - Own work. Licensed under CC BY 3.0 via Wikimedia Commons - http://commons.wikimedia.org/wiki/File:Poisson_pmf.svg#/media/File:Poisson_pmf.svg.

Beta Family

- Support y = (0,1). [The unit interval.]
- Parameter space: $\{\theta = (\alpha, \beta) \mid \alpha, \beta > 0\}$
- Probability density function on $y \in \mathcal{Y}$:

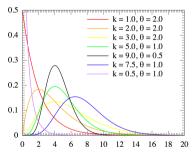
$$p(y; a, b) = \frac{y^{\alpha - 1} (1 - y)^{\beta - 1}}{B(\alpha, \beta)}$$



Gamma Family

- Support $\mathcal{Y} = (0, \infty)$. [Positive real numbers]
- Parameter space: $\{\theta = (k, \theta) \mid k > 0, \theta > 0\}$
- Probability density function on $y \in \mathcal{Y}$:

$$p(y; k, \theta) = \frac{1}{\Gamma(k)\theta^k} x^{k-1} e^{-y/\theta}.$$



• Special cases: exponential distribution, chi-squared distribution, Erlang distribution

Figure from Wikipedia https://commons.wikimedia.org/wiki/File:Gamma_distribution_pdf.svg.

Likelihood in a Parametric Model

Suppose we have a parametric model $\{p(y;\theta) \mid \theta \in \Theta\}$ and a sample $\mathcal{D} = (y_1, \dots, y_n)$.

• The likelihood of parameter estimate $\hat{\theta} \in \Theta$ for sample \mathcal{D} is

$$p(\mathcal{D}; \hat{\theta}) = \prod_{i=1}^{n} p(y_i; \hat{\theta}).$$

• In practice, we prefer to work with the log-likelihood. Same maximizer, but

$$\log p(\mathcal{D}; \hat{\theta}) = \sum_{i=1}^{n} \log p(y_i; \hat{\theta}),$$

and sums are easier to work with than products.

• Suppose $\mathcal{D} = (y_1, \dots, y_n)$ is an i.i.d. sample from some distribution.

Definition

A maximum likelihood estimator (MLE) for θ in the model $\{p(y;\theta) \mid \theta \in \Theta\}$ is

$$\begin{split} \hat{\theta} &\in & \underset{\theta \in \Theta}{\arg\max} \log p(\mathcal{D}, \hat{\theta}) \\ &= & \underset{\theta \in \Theta}{\arg\max} \sum_{i=1}^{n} \log p(y_i; \theta). \end{split}$$

- Finding the MLE is an optimization problem.
- For some model families, calculus gives a closed form for the MLE.
- Can also use numerical methods we know (e.g. SGD).

MLE Existence

- In certain situations, the MLE may not exist.
- But there is usually a good reason for this.
- e.g. Gaussian family $\left\{\mathcal{N}(\mu, \sigma^2) \mid \mu \in \mathbf{R}, \sigma^2 > 0\right\}$
- We have a single observation y.
- Is there an MLE?
- Taking $\mu = y$ and $\sigma^2 \to 0$ drives likelihood to infinity.
- MLE doesn't exist.

Example: MLE for Poisson

- Observed counts $\mathcal{D} = (k_1, \dots, k_n)$ for taxi cab pickups over n weeks.
 - k_i is number of pickups at Penn Station Mon, 7-8pm, for week i.
- We want to fit a Poisson distribution to this data.
- The Poisson log-likelihood for a single count is

$$\log[p(k;\lambda)] = \log\left[\frac{\lambda^k e^{-\lambda}}{k!}\right]$$
$$= k \log \lambda - \lambda - \log(k!)$$

• The full log-likelihood is

$$\log p(\mathcal{D}, \lambda) = \sum_{i=1}^{n} [k_i \log \lambda - \lambda - \log(k_i!)].$$

Example: MLE for Poisson

The full log-likelihood is

$$\log p(\mathcal{D}, \lambda) = \sum_{i=1}^{n} [k_i \log \lambda - \lambda - \log (k_i!)]$$

• First order condition gives

$$0 = \frac{\partial}{\partial \lambda} [\log p(\mathcal{D}, \lambda)] = \sum_{i=1}^{n} \left[\frac{k_i}{\lambda} - 1 \right]$$

$$\Longrightarrow \lambda = \frac{1}{n} \sum_{i=1}^{n} k_i$$

• So MLE $\hat{\lambda}$ is just the mean of the counts.

Test Set Log Likelihood for Penn Station, Mon-Fri 7-8pm

Method	Test Log-Likelihood
Poisson	-392.16
Negative Binomial	-188.67
Histogram (Bin width $= 7$)	$-\infty$
.95 Histogram +.05 NegBin	-203.89

Estimating Distributions, Overfitting, and Hypothesis Spaces

- Just as in classification and regression, MLE can overfit!
- Example Probability Models:
 - $\mathcal{F} = \{ \text{Poisson distributions} \}.$
 - $\mathcal{F} = \{ \text{Negative binomial distributions} \}.$
 - $\mathcal{F} = \{\text{Histogram with 10 bins}\}\$
 - $\mathcal{F} = \{\text{Histogram with bin for every } y \in \mathcal{Y}\}\ [\text{will likely overfit for continuous data}]$
- How to judge which model works the best?
- Choose the model with the highest likelihood on validation set.