

Unsupervised Learning

Introduction to clustering

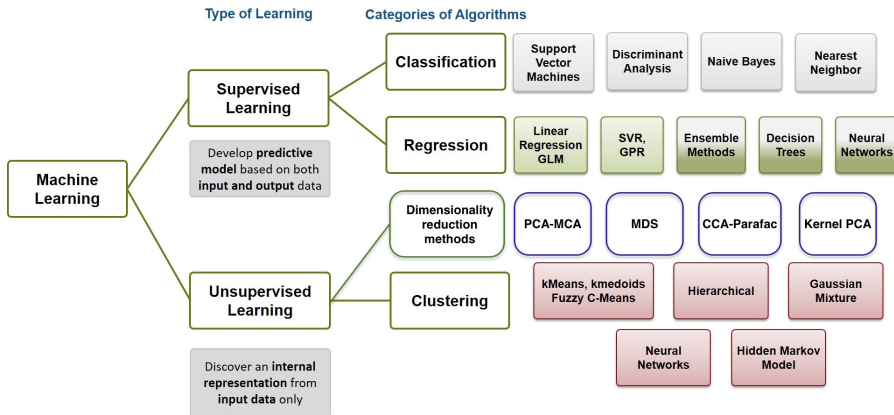
MAP 573, 2019 – Julien Chiquet

École Polytechnique, Autumn semester, 2019

<https://github.com/jchiquet/CourseUnsupervisedLearningX>



Machine Learning



Supervised vs Unsupervised Learning

Supervised Learning

- Training data $\mathcal{D}_n = \{(\mathbf{x}_1, y_1), \dots, (\mathbf{x}_n, y_n)\}$, $X_i \sim^{\text{i.i.d}} \mathbb{P}$
- Construct a predictor $\hat{f} : \mathcal{X} \rightarrow \mathcal{Y}$ using \mathcal{D}_n
- Loss $\ell(y, f(x))$ measures how well $f(x)$ predicts y
- Aim: minimize the generalization error
- Task: Regression, Classification

↪ The goal is clear: predict y based on x (regression, classification)

Unsupervised Learning

- Training data $\mathcal{D} = \{\mathbf{x}_1, \dots, \mathbf{x}_n\}$
- Loss? , Aim?
- Task: Dimension reduction, Clustering

↪ The goal is less well defined, and *validation* is questionable

Outline

① Clustering: introduction

- Motivating example

- Generalities

- Vocabulary

② Distance-based methods

- The K-means algorithm

- Hierarchical Agglomerative Clustering

③ Model-based approach

- Mixture models

- Expectation-Maximization algorithm

Outline

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Packages required for reproducing the slides

```
library(tidyverse)  # opinionated collection of packages for data manipulation
library(corrplot)   # fancy plots of matrices as images
library(GGally)     # extension to ggplot vizualization system
library(ggfortify)  # extension to ggplot vizualization system
library(mclust)     # Gaussian mixture models
library(aricode)    # fast computation of clustering measures
library(animation)  # kmeans animation slides
# color and plots themes
library(RColorBrewer)
pal <- brewer.pal(10, "Set3")
theme_set(theme_bw())
```

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Companion data set

Morphological Measurements on Leptograpsus Crabs

Description

The crabs data frame has 200 rows and 8 columns, describing 5 morphological measurements on 50 crabs each of two colour forms and both sexes, of the species *Leptograpsus variegatus* collected at Fremantle, W. Australia.

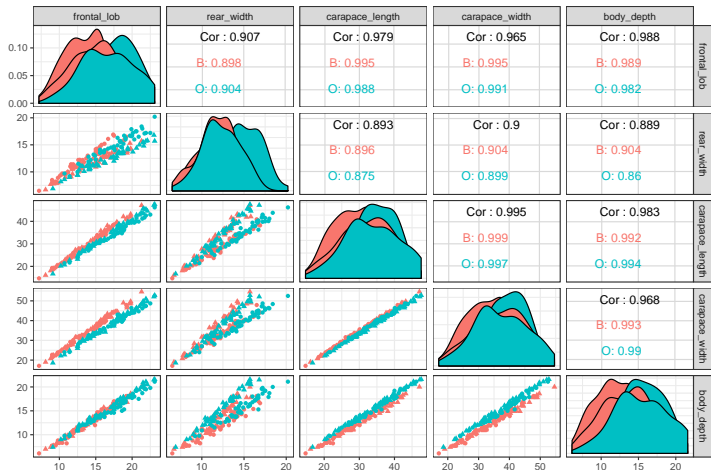
```
crabs <- MASS::crabs %>% select(-index) %>%  
  rename(sex = sex,  
         species = sp,  
         frontal_lob = FL,  
         rear_width = RW,  
         carapace_length = CL,  
         carapace_width = CW,  
         body_depth = BD)  
crabs %>% select(sex, species) %>% summary() %>% knitr::kable("latex")
```

	sex	species
	F:100	B:100
	M:100	O:100

Companion data set II

Pairs plot of attributes

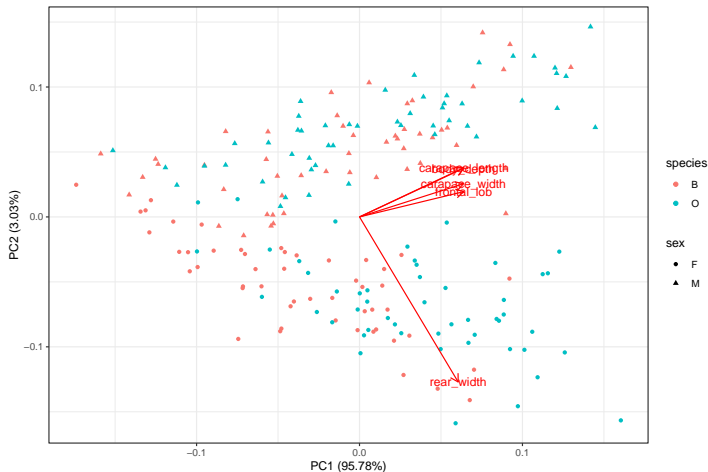
```
ggpairs(crabs, columns = 3:7, aes(colour = species, shape = sex))
```



Companion data set III

PCA on the attributes

```
prcomp(select(crabs, -species, -sex), scale. = TRUE) %>%  
  autoplot(loadings = TRUE, loadings.label = TRUE,  
    data = crabs, colour = 'species', shape = 'sex')
```



Remove size effect I

Carried by the 1st principal component

PCA is solved by SVD

$$\mathbf{X} = \mathbf{U}\mathbf{D}\mathbf{V}^{\top}.$$

We remove the best rank-1 approximation of \mathbf{X} to remove the *size effect*, carried by the first axis, that is,

$$\tilde{\mathbf{X}}^{(1)} = \mathbf{U}_{\bullet 1} d_{11} \mathbf{v}_{\bullet 1}^{\top}.$$

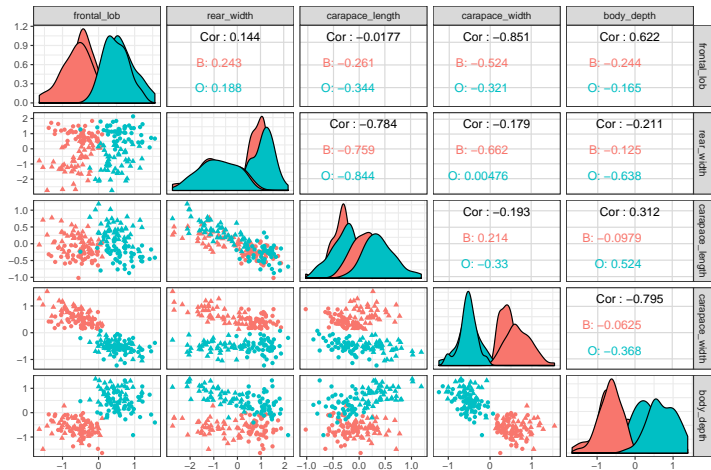
```
attributes <- select(crabs, -sex, -species)
SVD <- svd(attributes)
attributes_rank1 <- tcrossprod(SVD$u[, 1] * SVD$d[1], SVD$v[, 1])
crabs_corrected <- crabs
crabs_corrected[, 3:7] <- attributes - attributes_rank1
```

↪ Axis 1 explains a latent effect, here the size in the case at hand, common to all attributes.

Remove size effect II

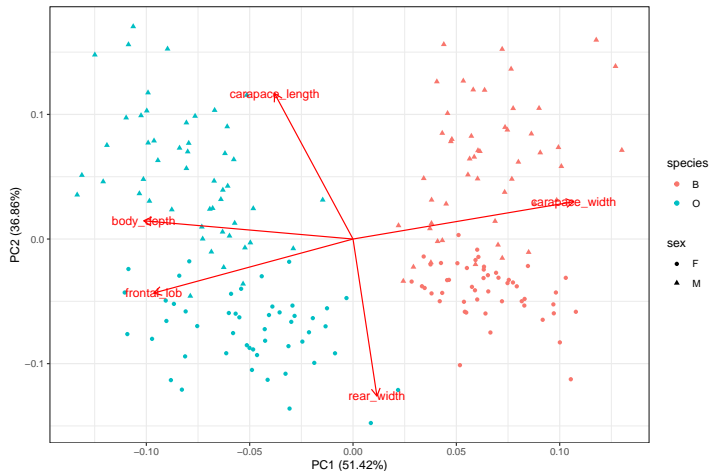
Carried by the 1st principal component

```
ggpairs(crabs_corrected, columns = 3:7, aes(colour = species, shape = sex))
```



PCA on corrected data

```
prcomp(select(crabs_corrected, -species, -sex), scale. = TRUE) %>%  
  autoplot(loadings = TRUE, loadings.label = TRUE,  
    data = crabs_corrected, colour = 'species', shape = 'sex')
```



Questions

- ① Could we automatically identify some grouping (**clustering**) between samples ?
- ② Would this clustering correspond to some known labels (sex, species)?
- ③ Does it matter?

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Clustering: general goals

Objective: construct a map f from \mathcal{D} to $\{1, \dots, K\}$ where K is a fixed number of clusters.

Careful! classification \neq clustering

- Classification presupposes the existence of classes
- Clustering labels only elements of the dataset
 - \rightsquigarrow no ground truth (no given labels)
 - \rightsquigarrow discovers a structure "natural" to the data
 - \rightsquigarrow not necessarily related to a known classification

Motivations

- describe large masses of data in a simplified way,
- structure a set of knowledge,
- reveal structures, hidden causes,
- use of the groups in further processing,
- ...

Clustering: challenges

Clustering quality

No obvious measure to define the **quality** of the clusters. Ideas:

- **Inner** homogeneity: samples in the same group should be similar
- **Outer** inhomogeneity: samples in different groups should be different

Number of clusters

Choice of the number of clusters K often complex

- No ground truth in unsupervised learning!
- Several solutions might be equally good

Two general approaches

- **distance-based**: require a distance/dissimilarity between $\{\mathbf{x}_i\}$
- **model-based**: require assumptions on the distribution \mathbb{P}

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Dissimilarity and Distance

Clustering requires a measure of resemblance between object

Definition ((dis)similarity)

Similarity (*resp.* Dissimilarity) measures the resemblance (*resp.* discrepancy) between objects based on several features.

For instance, two objects are similar if

- they share a certain feature
- their features are close according to a measure of proximity

Definition (distance/metric)

Dissimilarity can be measured by distances, *i.e.* a function d_{ij} between pairs in $\{\mathbf{x}_i\}$ s.t.

- $d_{ij} \geq 0$,
- $d_{ij} = 0 \Leftrightarrow \mathbf{x}_i = \mathbf{x}_j$,
- $d_{ij} = d_{ji}$,
- $d_{ik} \leq d_{ij} + d_{jk}$.

Dissimilarity and Distance

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Classification structures: Partition

Clustering leads to a grouping (or classification) of individuals into homogeneous classes

We consider two structures to describe this classification:

- partitions and
- hierarchies.

Definition (Partition)

A partition \mathcal{P} is a decomposition $\mathcal{P} = \{P_1, \dots, P_K\}$ of a finite ensemble Ω such that

- $P_k \cap P_{k'} = \emptyset$ for any $k \neq k'$
- $\bigcup_k P_k = \Omega$

In a set $\Omega = (\mathbf{x}_1, \dots, \mathbf{x}_n)$ partitioned into K classes, each element of the set belongs to a class and only one.

Classification structures: Hierarchy

Definition (Hierarchy)

A hierarchy \mathcal{H} is a non empty subset of a finite ensemble Ω such that

- $\Omega \in \mathcal{H}$,
- $\forall \mathbf{x} \in \Omega, \{\mathbf{x}\} \in \mathcal{H}$,
- $\forall H, H' \in \mathcal{H}$, then either $H \cap H' = \emptyset$, $H \subset H'$ or $H' \subset H$.

Definition (Index of a Hierarchy)

The index is a function $i: \mathcal{H} \rightarrow \mathbb{R}_+$ such that

- if $H \subset H'$ then $i(H) < i(H')$;
- if $\mathbf{x} \in \Omega$ then $i(\mathbf{x}) = 0$.

Properties (Partition and Hierarchy)

- *Each level of an indexed hierarchy is a partition;*
- *$\{\Omega, P_1, \dots, P_K, \mathbf{x}_1, \dots, \mathbf{x}_n\}$ is a hierarchy.*

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Clusterings Comparison: Contingency table

Definition

Consider two clusterings U and V of elements in Ω , into respectively $|U|$ and $|V|$ classes. The $|U| \times |V|$ contingency matrix stores at position (i, j) the number of elements that are simultaneously in cluster i of U and j of V .

$U \backslash V$	V_1	V_2	\dots	$V_{ V }$	Sums
U_1	n_{11}	n_{12}	\dots	$n_{1 V }$	$n_{1.}$
U_2	n_{21}	n_{22}	\dots	$n_{2 V }$	$n_{2.}$
\vdots	\vdots	\vdots	\ddots	\vdots	\vdots
$U_{ U }$	$n_{ U 1}$	$n_{ U 2}$	\dots	$n_{ U V }$	$n_{ U .}$
Sums	$n_{.1}$	$n_{.2}$	\dots	$n_{. V }$	$n_{..} = n$

Clusterings Comparison: Measures (I)

Definition (Rand index)

Given a set Ω of n elements and two partitions U and V to compare, define the following:

- a , the number of pairs in the same subset in U and in V
- b , the number of pairs in different subsets in U and in V

The Rand index, $RI \in [0, 1]$ is

$$RI = \frac{a + b}{\binom{n}{2}}$$

The Rand index can be viewed as a measure of the percentage of correct decisions:

$$RI = \frac{TP + TN}{\binom{n}{2}},$$

where TP, TN are true positive and true negative decisions.

Clusterings Comparison: Measures (II)

The ARI (most popular) is a version of the RI adjusted for chance grouping of element (i.e., the expected similarity of all pair-wise comparisons).

Definition (Adjusted Rand-index)

$$ARI(U, V) = \frac{\sum_{i,j} \binom{n_{ij}}{2} - \left[\sum_i \binom{n_{i.}}{2} \sum_j \binom{n_{.j}}{2} \right] / \binom{n}{2}}{\frac{1}{2} \left[\sum_i \binom{n_{i.}}{2} + \sum_j \binom{n_{.j}}{2} \right] - \left[\sum_i \binom{n_{i.}}{2} \sum_j \binom{n_{.j}}{2} \right] / \binom{n}{2}}$$

Other popular measures:

- *NVI*, the normalized variation information
- *NID*, the normalized information distance
- *NMI*, the normalized mutual information

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- 2 Distance-based methods
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- 3 Model-based approach

References



The Elements of Statistical Learning,
T. Hastie, R. Tibshirani, J. Friedman

Chapter: 14 Unsupervised Learning, Section 3: Cluster Analysis

<https://web.stanford.edu/~hastie/ElemStatLearn/>



Classification non-supervisées,
É. Lebarbier, T. Mary-Huard

Chapitre 2 - méthode de partitionnement

<https://www.agroparistech.fr/IMG/pdf/ClassificationNonSupervisee-AgroParisTech.pdf>

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K-means heuristic

Idea

- ① Clustering is defined by a partition in K classes
- ② Minimize a criteria of clustering quality
- ③ Use Euclidean distances to measure dissimilarity

Criteria: intra-class variance/ Inertia "within"

Intra-class variance measures **inner** homogeneity

$$I_W = \sum_{k=1}^K \sum_{i=1}^n c_{ik} \|\mathbf{x}_i - \boldsymbol{\mu}_k\|_2^2,$$

where

- $\boldsymbol{\mu}_k$ are the centers (prototypes) of classes
- $c_{ik} = \mathbf{1}_{i \in \mathcal{P}_k}$ is a partition matrix

K-means algorithm

Ideally, one would solve

$$(\hat{\mathbf{c}}, \hat{\boldsymbol{\mu}}) = \arg \min_{(\mathbf{c}, \boldsymbol{\mu})} I_w((\mathbf{c}, \boldsymbol{\mu})), \quad \text{s.t.} \quad \mathbf{c} \text{ is a partition matrix.}$$

This problem is hard to solve but can be optimized locally as follows:

K-means algorithm (Loyds)

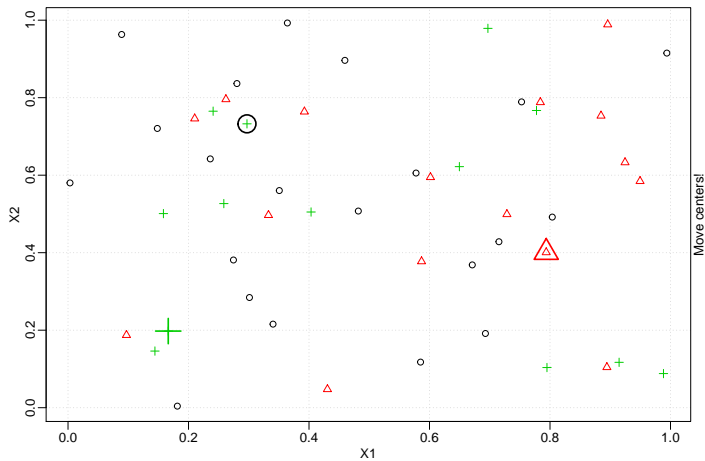
Initialization start by a (pseudo) random choice for the centers $\boldsymbol{\mu}_k$

Alternate until convergence

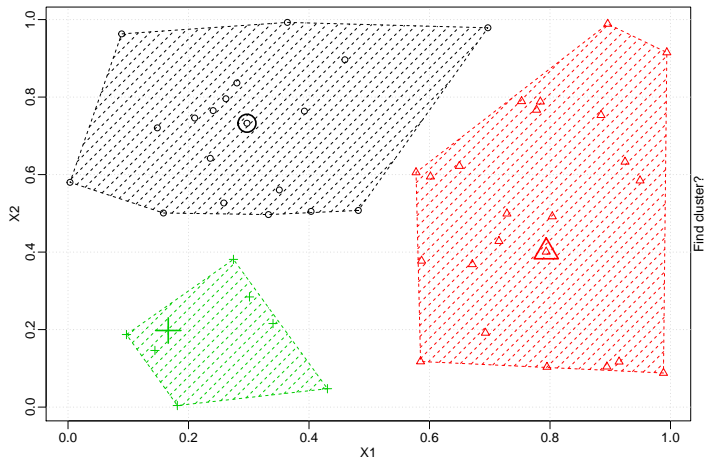
step 1 given $\boldsymbol{\mu}$, chose \mathbf{c} minimizing $I_w \equiv$ assign \mathbf{x}_i to the nearest prototype

step 2 given \mathbf{c} , chose $\boldsymbol{\mu}$ minimizing $I_w \equiv$ update $\boldsymbol{\mu}$ by the new means of classes

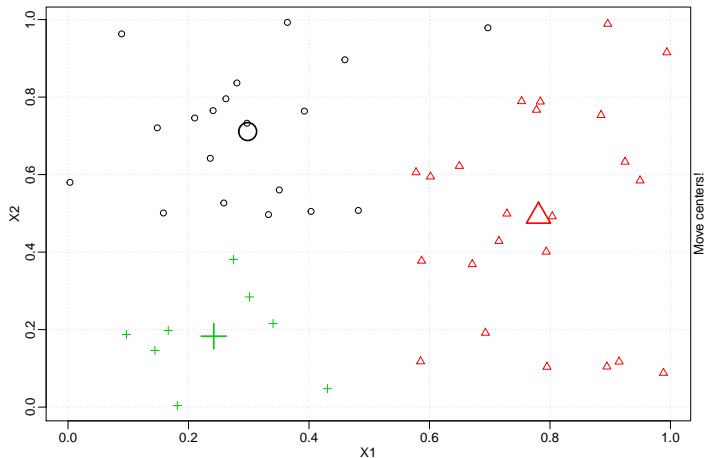
K-means in action I



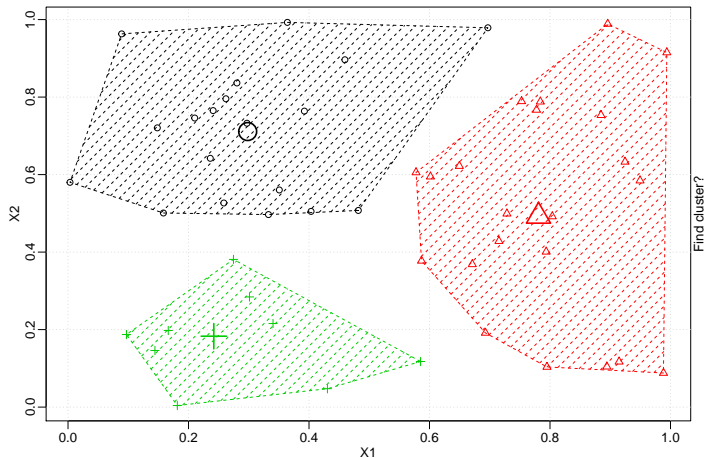
K-means in action II



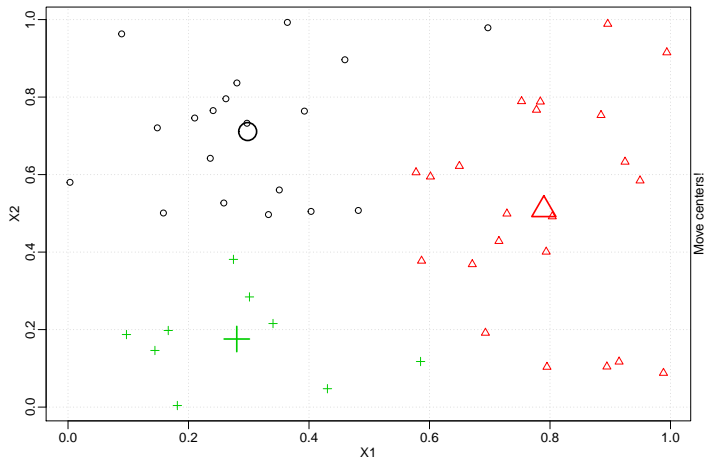
K-means in action III



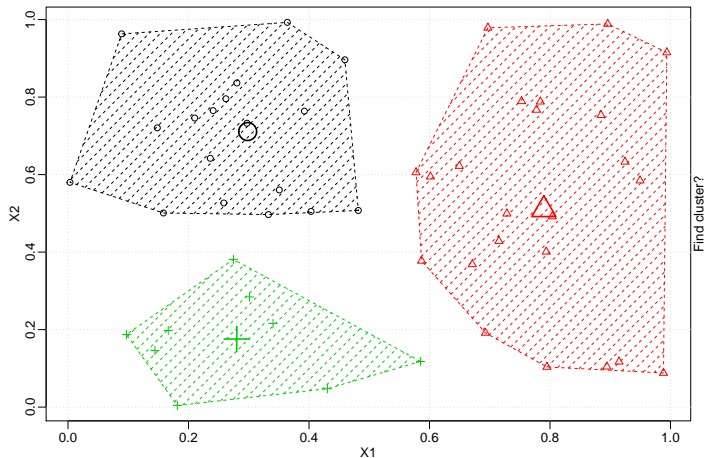
K-means in action IV



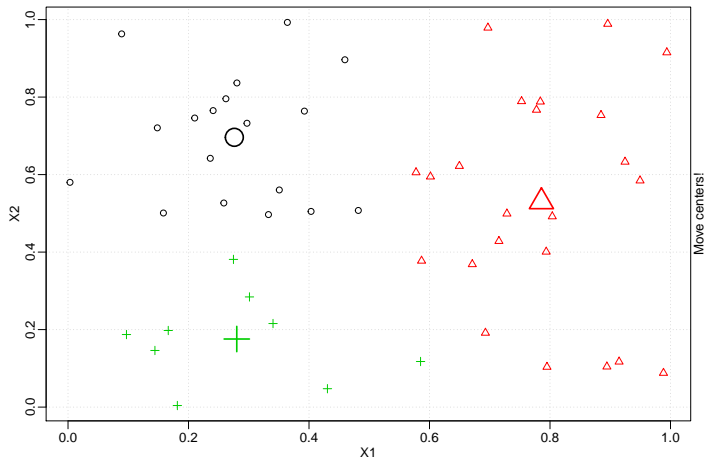
K-means in action V



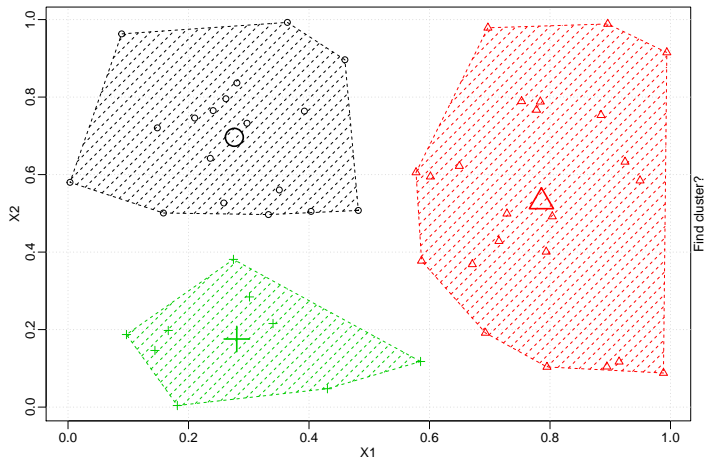
K-means in action VI



K-means in action VII



K-means in action VIII



K-means: properties

Other schemes

- **McQueen**: modify the mean each time a sample is assigned to a new cluster.
- **Hartigan**: modify the mean by removing the considered sample, assign it to the nearby center and recompute the new mean after assignment.

Initialization

No guarantee to converge to a global optimum

- Repeat and keep the best result
- k-Mean++: try to take them as separated as possible.

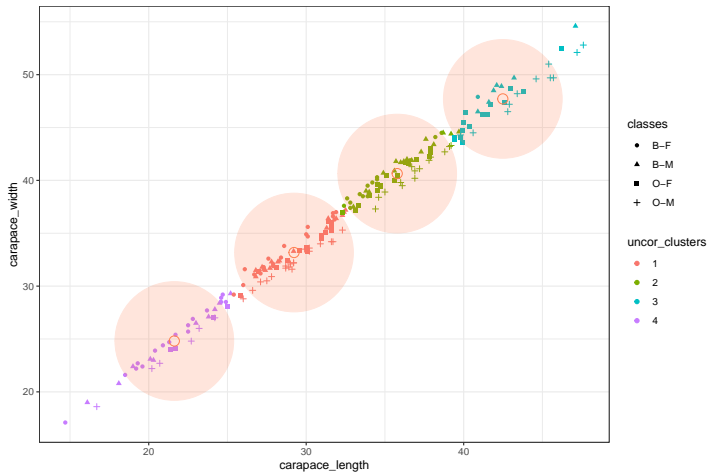
Complexity

$O(nKT)$ where T is the number of step in the algorithm.

K-means in R on uncorrected data set I

```
uncor_kmeans_res <- crabs %>%  
  select(-species, -sex) %>%  
  kmeans(4, nstart = 10)  
uncor_clusters <- as.factor(uncor_kmeans_res$cluster)  
uncor_centers  <- as_tibble(uncor_kmeans_res$centers)  
classes <- paste(crabs_corrected$species, crabs_corrected$sex, sep = "-")  
  
crabs %>%  
  ggplot(aes(x = carapace_length, y = carapace_width, color = uncor_clusters)) +  
  geom_point(aes(shape = classes)) +  
  geom_point(data = uncor_centers, color = 'coral', size = 4, pch = 21) +  
  geom_point(data = uncor_centers, color = 'coral', size = 50, alpha = 0.2)
```

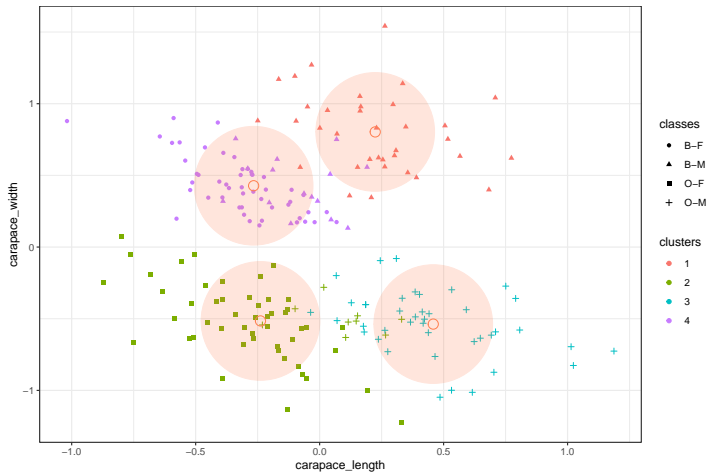
K-means in R on uncorrected data set II



K-means in R on corrected crabs data set I

```
kmeans_res <- crabs_corrected %>%  
  select(-species, -sex) %>%  
  kmeans(4, nstart = 10)  
clusters <- as.factor(kmeans_res$cluster)  
centers <- as.tibble(kmeans_res$centers)  
classes <- paste(crabs_corrected$species, crabs_corrected$sex, sep = "-")  
  
crabs_corrected %>%  
  ggplot(aes(x = carapace_length, y = carapace_width, color = clusters)) +  
  geom_point(aes(shape = classes)) +  
  geom_point(data = centers, color = 'coral', size = 4, pch = 21) +  
  geom_point(data = centers, color = 'coral', size = 50, alpha = 0.2)
```

K-means in R on corrected crabs data set II



Clustering comparison

```
aricode::ARI(clusters, classes)

## [1] 0.7223637

aricode::ARI(uncor_clusters, classes)

## [1] 0.01573617
```

```
knitr::kable(table(clusters, classes),
caption = "Estimating structure with k-means")
```

Table: Estimating structure with k-means

B-F	B-M	O-F	O-M
0	35	0	0
0	0	50	9
0	0	0	41
50	15	0	0

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Agglomerative Clustering: Heuristic

Idea

- ① Start with small clusters (e.g. one cluster \equiv one individual)
- ② Merge the most similar clusters sequentially (and greedily)
- ③ Stops when all individuals are in the same groups

Ingredients

- ① a dissimilarity measure (distance between individuals)
- ② a merging criterion Δ (dissimilarity between clusters)

- + Generates a hierarchy of clustering instead of a single partition
- Need to select the number of cluster afterwards

Agglomerative Clustering: general algorithm

Algorithm

- 1 Start with $(\mathcal{C}_k^{(0)}) = (\{\mathbf{x}_i\})$ the collection of all singletons.
- 2 At step s , we have $n - s$ clusters $(\mathcal{C}_k^{(s)})$:
 - Find the two most similar clusters according to a criterion Δ :

$$(k, \ell) = \arg \min_{(k', \ell')} \Delta(\mathcal{C}_{k'}^{(s)}, \mathcal{C}_{\ell'}^{(s)})$$

- Merge $\mathcal{C}_k^{(s)}$ and $\mathcal{C}_\ell^{(s)}$ into $\mathcal{C}_k^{(s+1)}$
 - Update the distances between $\mathcal{C}_k^{(s+1)}$ and the remaining clusters
- 3 Repeat until there is only one cluster.

Complexity

- In general $O(n^3)$
- Can be reduced to $O(n^2)$ if bounding the number of merges

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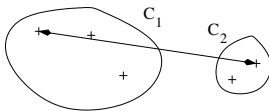
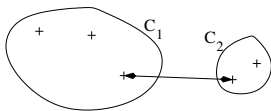
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Merging criterion based on the distance between points

- Single linkage (or minimum linkage):

$$\Delta(\mathcal{C}_k, \mathcal{C}_\ell) = \min_{\mathbf{x}_i \in \mathcal{C}_k, \mathbf{x}_j \in \mathcal{C}_\ell} d(\mathbf{x}_i, \mathbf{x}_j)$$

- Complete linkage (or maximum linkage):

$$\Delta(\mathcal{C}_k, \mathcal{C}_\ell) = \max_{\mathbf{x}_i \in \mathcal{C}_k} \max_{\mathbf{x}_j \in \mathcal{C}_\ell} d(\mathbf{x}_i, \mathbf{x}_j)$$

- Average linkage (or group linkage):

$$\Delta(\mathcal{C}_k, \mathcal{C}_\ell) = \frac{1}{|\mathcal{C}_k||\mathcal{C}_\ell|} \sum_{\mathbf{x}_i \in \mathcal{C}_k} \sum_{\mathbf{x}_j \in \mathcal{C}_\ell} d(\mathbf{x}_i, \mathbf{x}_j)$$

Ward's criteria

Merging criterion based on distance to the mean

Ward's criterion:

$$\begin{aligned}\Delta(\mathcal{C}_k, \mathcal{C}_\ell) = & \sum_{\mathbf{x}_i \in \mathcal{C}_k} (d^2(\mathbf{x}_i, \boldsymbol{\mu}_{\mathcal{C}_k \cup \mathcal{C}_\ell}) - d^2(\mathbf{x}_i, \boldsymbol{\mu}_{\mathcal{C}_k})) \\ & + \sum_{\mathbf{x}_j \in \mathcal{C}_\ell} (d^2(\mathbf{x}_j, \boldsymbol{\mu}_{\mathcal{C}_k \cup \mathcal{C}_\ell}) - d^2(\mathbf{x}_j, \boldsymbol{\mu}_{\mathcal{C}_\ell}))\end{aligned}$$

Euclidean case

If d is the Euclidean distance, then

$$\Delta(\mathcal{C}_k, \mathcal{C}_\ell) = \frac{2|\mathcal{C}_k||\mathcal{C}_\ell|}{|\mathcal{C}_k| + |\mathcal{C}_\ell|} d^2(\boldsymbol{\mu}_{\mathcal{C}_k}, \boldsymbol{\mu}_{\mathcal{C}_\ell})$$

Ward's criteria: details

Recall that the inertia measures the homogeneity of th size-K clustering

$$I_W = \sum_{k=1}^K \sum_{\mathbf{x}_i \in \mathcal{C}_k} \|\mathbf{x}_i - \boldsymbol{\mu}_{\mathcal{C}_k}\|_2^2, \quad I_B = \sum_{k=1}^K n_k \|\boldsymbol{\mu}_k - \boldsymbol{\mu}\|_2^2$$

Consider the following two partitions

- $\mathcal{P} = (\mathcal{C}_1, \dots, \mathcal{C}_K)$ at one level of the hierarchy Ω
- \mathcal{P}' is \mathcal{P} once $\mathcal{C}_k, \mathcal{C}_\ell$ merged

Then

$$I_B(\mathcal{P}) - I_B(\mathcal{P}') = \frac{|\mathcal{C}_k||\mathcal{C}_\ell|}{|\mathcal{C}_k| + |\mathcal{C}_\ell|} d^2(\boldsymbol{\mu}_{\mathcal{C}_k}, \boldsymbol{\mu}_{\mathcal{C}_\ell}) = \frac{1}{2} \Delta(\mathcal{C}_k, \mathcal{C}_\ell).$$

- ↪ At each step, Ward limits the loss (increase) of the intra (inter) class variance
- ↪ Defines an indexed hierarchy (height of the dendrogram)
- ↪ Same criteria as in the K-means algorithm

Ward's criteria: details

Recall that the inertia measures the homogeneity of the size-K clustering

$$I_W = \sum_{k=1}^K \sum_{\mathbf{x}_i \in \mathcal{C}_k} \|\mathbf{x}_i - \boldsymbol{\mu}_{\mathcal{C}_k}\|_2^2, \quad I_B = \sum_{k=1}^K n_k \|\boldsymbol{\mu}_k - \boldsymbol{\mu}\|_2^2$$

Consider the following two partitions

- $\mathcal{P} = (\mathcal{C}_1, \dots, \mathcal{C}_K)$ at one level of the hierarchy Ω
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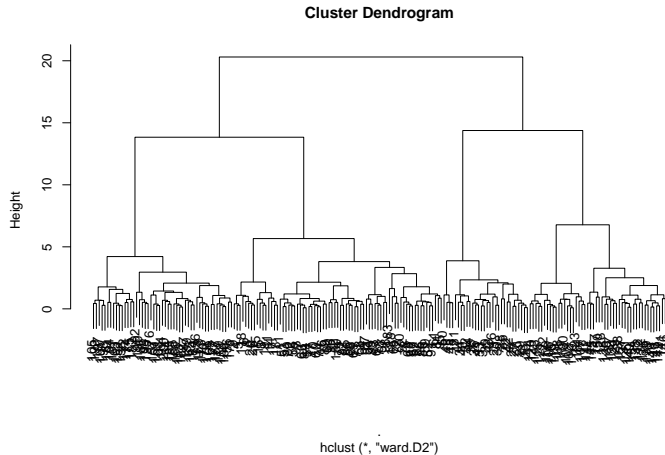
Then

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Ward agglomerative clustering in R

```
Ward <- crabs_corrected %>%  
  select(-sex, -species) %>%  
  dist(method = "euclidean") %>%  
  hclust(method = "ward.D2")  
plot(Ward)
```



Ward agglomerative clustering in R: comparison I

Compare with out reference classification and k-means

```
aricode::ARI(cutree(Ward, 4), classes)
```

```
## [1] 0.6829729
```

```
aricode::ARI(cutree(Ward, 4), clusters)
```

```
## [1] 0.7999974
```

```
knitr::kable(table(clusters, cutree(Ward,4)),  
caption = "k-means vs Ward")
```

Table: k-means vs Ward

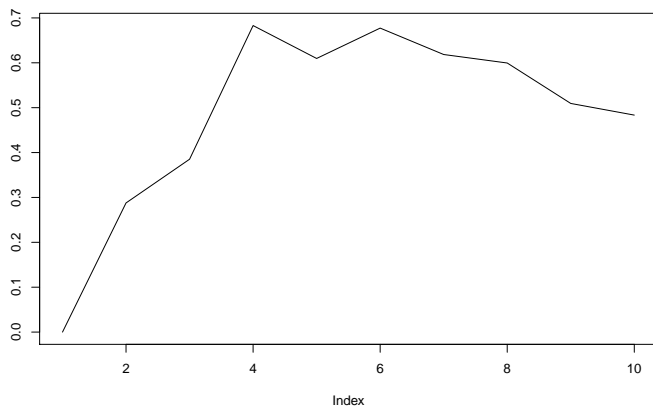
1	2	3	4
5	30	0	0
2	0	49	8
1	0	0	40
65	0	0	0

Ward agglomerative clustering in R: comparison II

Optimize over a range of values

```
Ward %>% cutree(k = 1:10) %>% as.data.frame() %>% as.list() %>%  
  sapply(aricode::ARI, classes) %>% plot(type = "l")
```

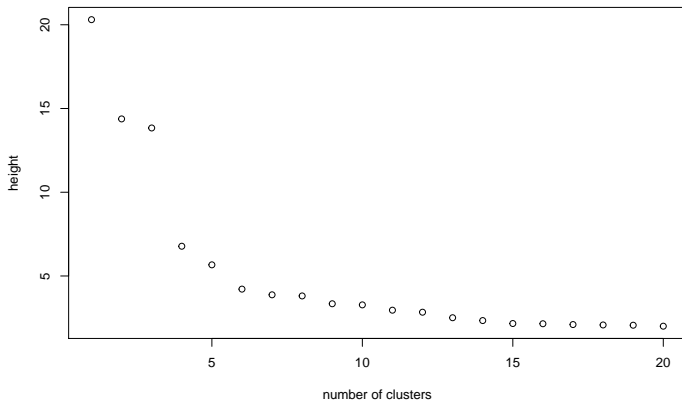
Ward agglomerative clustering in R: comparison III



Look at Ward intra-class variance

Ward agglomerative clustering in R: comparison IV

```
plot(rev(Wards$height)[1:20], xlab = "number of clusters", ylab = "height")
```

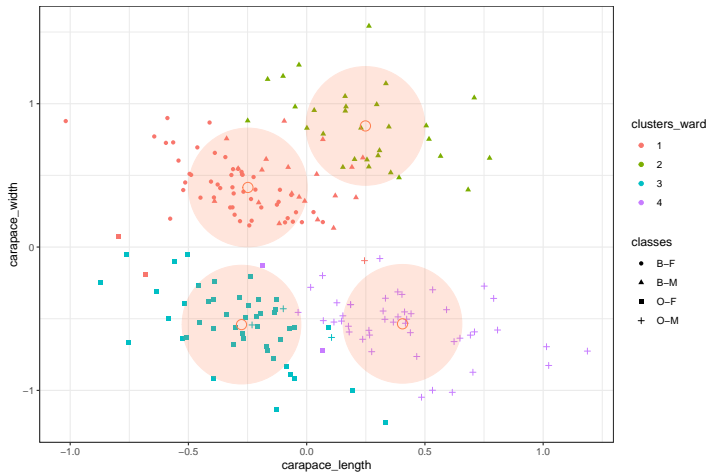


Ward agglomerative clustering in R: projection I

```
clusters_ward <- as.factor(cutree(Ward, 4))
centers_ward <- select(crabs_corrected, -sex, -species) %>%
  aggregate(list(cutree(Ward, 4)), mean) %>% as_tibble() %>% select(-Group.1)

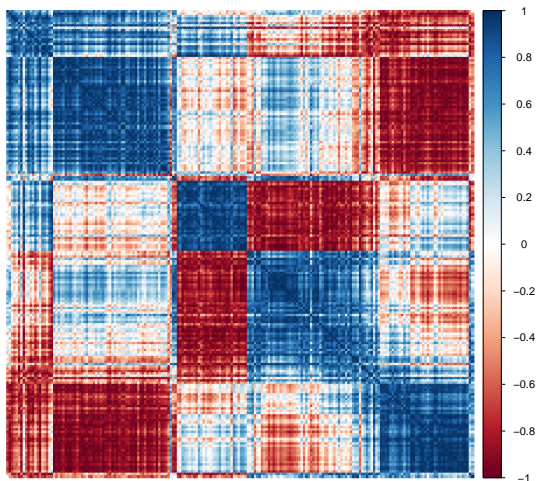
crabs_corrected %>%
  ggplot(aes(x = carapace_length, y = carapace_width, color = clusters_ward)) +
  geom_point(aes(shape = classes)) +
  geom_point(data = centers_ward, color = 'coral', size = 4, pch = 21) +
  geom_point(data = centers_ward, color = 'coral', size = 50, alpha = 0.2)
```

Ward agglomerative clustering in R: projection II



Reordered correlation matrix between individuals

```
C <- cor(t(select(crabs_corrected, -sex, -species)))  
C <- C[order(clusters_ward), order(clusters_ward)]  
corrplot(C, method = "color", tl.pos = "n")
```



Outline

- ① Clustering: introduction
- ② Distance-based methods
- ③ Model-based approach
 - Mixture models
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References



Pattern recognition and machine learning,
Christopher Bishop

Chapter 9: Mixture Models and EM

<http://users.isr.ist.utl.pt/~wurmd/Livros/school/>



Models with Hidden Structure with Applications in Biology and
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Master MathSV Course

[https:](https://www6.inra.fr/mia-paris/content/download/4587/42934/version/1/file/ModelsHiddenStruct-Biology.pdf)

[//www6.inra.fr/mia-paris/content/download/4587/42934/version/1/file/ModelsHiddenStruct-Biology.pdf](https://www6.inra.fr/mia-paris/content/download/4587/42934/version/1/file/ModelsHiddenStruct-Biology.pdf)



Classification non-supervisées,

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Chapitre 3 - méthode probabiliste: le modèle de mélange

<https://www.agroparistech.fr/IMG/pdf/ClassificationNonSupervisee-AgroParisTech.pdf>

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Latent variables models

Definition

A **latent variable model** is a statistical model that relates, for $i = 1, \dots, n$ individuals,

- a set of **manifest** (observed) variables $\mathbf{X} = (X_i, i = 1, \dots, n)$ to
- a set of **latent** (unobserved) variables $\mathbf{Z} = (Z_i, i = 1, \dots, n)$.

Common assumption: conditional independence

$$\mathbb{P}((X_1, \dots, X_n) | (Z_1, \dots, Z_n)) = \prod_{i=1}^n \mathbb{P}(X_i | Z_i).$$

Famous examples

- $(Z_i, i \geq 1)$ is Markov chain: **Markov models**
- Z_i categorical and independent: **mixture models**

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Mixture models: the latent variables

When (Z_1, \dots, Z_n) are independent categorical variables, they give a **natural (latent) classification of the observations** (X_1, \dots, X_n) – or **labels**.

Notations

Let (Z_1, \dots, Z_n) be *iid* categorical variables with distribution

$$\mathbb{P}(i \in q) = \mathbb{P}(Z_i = q) = \alpha_q, \quad \text{s.t.} \quad \sum_{q=1}^Q \alpha_q = 1.$$

Alternative (equivalent) notation

Let $Z_i = (Z_{i1}, \dots, Z_{iQ})$ be an indicator vector of label for i :

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By definition, $Z_i \sim \mathcal{M}(1, \alpha)$, with $\alpha = (\alpha_1, \dots, \alpha_Q)$.

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Mixture models: the manifest variables

A mixture model represents the **presence of subpopulations** within an overall population as follows:

$$\mathbb{P}(X_i) = \sum_{z_i \in \mathcal{Z}_i} \mathbb{P}(X_i, Z_i) = \sum_{Z_i \in \mathcal{Z}_i} \mathbb{P}(X_i | Z_i) \mathbb{P}(Z_i).$$

Conditional distribution of the manifest variables

We assume a **parametric distribution** of X in each subpopulation

$$X_i | \{Z_i = q\} \sim \mathbb{P}_{\theta_q} \quad \left(\Leftrightarrow X_i | \{Z_{iq}\} = 1 \sim \mathbb{P}_{\theta_q} \right)$$

The specificity of each class is handled by $\{\theta_q\}_{q=1}^Q$.

Mixture models: likelihoods

The complete-data likelihood

It is the joint distribution of (X_i, Z_i) :

$$\mathbb{P}(X_i, Z_i) = \alpha_{Z_i} \mathbb{P}_{\theta_{Z_i}}(X_i)$$

The incomplete-data likelihood

It is the marginal distribution of X_i once Z_i integrated:

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↪ A mixture model is a sum of distributions weighed by the proportion of each subpopulation.

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Intractability of the Likelihood

Maximum Likelihood Estimator

The MLE aims to maximize the (marginal) likelihood of the observations:

$$L(\boldsymbol{\theta}; \mathbf{X}) = \mathbb{P}_{\boldsymbol{\theta}}((X_1, \dots, X_n)) = \int_{\mathbf{Z} \in \mathcal{Z}} \mathbb{P}_{\boldsymbol{\theta}}(\mathbf{X}, \mathbf{Z}) d\mathbf{Z}$$

Integrations are summation over $\{1, \dots, Q\}$: we have Q^n terms !

Intractable summation

With mixture models, for $\boldsymbol{\theta} = (\boldsymbol{\theta}_1, \dots, \boldsymbol{\theta}_Q)$ we have

$$\log L(\boldsymbol{\theta}; \mathbf{X}) = \sum_{i=1}^n \log \left\{ \sum_{q=1}^Q \alpha_q \mathbb{P}_{\boldsymbol{\theta}_q}(X_i) \right\}.$$

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↪ Direct maximization of the likelihood is impossible in practice

Bayes decision rule / Maximum *a posteriori*

Principle

Affect an individual i to the subpopulation which is the most likely according to the data:

$$\tau_{iq} = \mathbb{P}(Z_{iq} = 1 | X_i = x_i)$$

This is the **posterior probability** for $i \in q$.

Application of the Bayes Theorem

It is straightforward to show that

$$\tau_{iq} = \frac{\alpha_q \mathbb{P}_{\theta_q}(x_i)}{\sum_{q=1}^Q \alpha_q \mathbb{P}_{\theta_q}(x_i)}$$

Principle of the EM algorithm

If θ were known

...estimating the **posterior probability** $\mathbb{P}(Z_i|\mathbf{X})$ of \mathbf{Z} should be easy

By means of the Bayes decision rule

If \mathbf{Z} were known...

...estimating the **best set of parameter** θ should be easy

This is close to usual maximum likelihood estimation

EM principle

Maximize the marginal likelihood iteratively:

- ① Initialize θ
- ② Compute the probability of \mathbf{Z} given θ
- ③ Get a better θ with the new \mathbf{Z}
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Formal algorithm

Initialization: start from a good guess either of \mathbf{Z} or $\boldsymbol{\theta}$, then iterate 1-2

1. Expectation step

Calculate the expected value of the loglikelihood under the current $\boldsymbol{\theta}$

$$Q\left(\boldsymbol{\theta}|\boldsymbol{\theta}^{(t)}\right)=\mathbb{E}_{\mathbf{Z}|\mathbf{X};\boldsymbol{\theta}^{(t)}}\left[\log L(\boldsymbol{\theta};\mathbf{X},\mathbf{Z})\right] \quad (\text{needs } \mathbb{P}_{\boldsymbol{\theta}^{(t)}}(\mathbf{Z}|\mathbf{X}))$$

2. Maximization step

Find the parameters that maximize this quantity

$$\boldsymbol{\theta}^{(t+1)}=\arg \max _{\boldsymbol{\theta}} Q\left(\boldsymbol{\theta}|\boldsymbol{\theta}^{(t)}\right)$$

Stop when $\left\|\boldsymbol{\theta}^{(t+1)}-\boldsymbol{\theta}^{(t)}\right\|<\varepsilon$ or $\left\|Q^{(t+1)}-Q^{(t)}\right\|<\varepsilon$

(Basic) Convergence analysis

Theorem

At each step of the EM algorithm, the loglikelihood increases. EM thus reaches a local optimum.

Proof.

On board.



Choosing the number of component

Reminder: Bayesian Information Criterion

The BIC is a model selection criterion which penalizes the adjustment to the data by the number of parameter in model \mathcal{M} as follows:

$$\text{BIC}(\mathcal{M}) = \log L(\hat{\boldsymbol{\theta}}; \mathbf{X}) - \frac{1}{2} \log(n) \text{df}(\mathcal{M}).$$

Integrated Classification Criterion

It is an adaptation working with the complete-data likelihood:

$$\begin{aligned} \text{ICL}(\mathcal{M}) &= \log L(\hat{\boldsymbol{\theta}}; \mathbf{X}, \hat{\mathbf{Z}}) + \frac{1}{2} \log(n) \text{df}(\mathcal{M}) \\ &= \text{BIC} - \mathcal{H}(\mathbb{P}(\hat{\mathbf{Z}}|\mathbf{X})), \end{aligned}$$

where the entropy \mathcal{H} measures the separability of the subpopulations.

⇒ We choose $\mathcal{M}(Q)$ that maximizes either BIC or ICL

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Popular model: Gaussian Multivariate mixture models

The distribution of X_i conditional on the label of i is assumed to be a multivariate Gaussian distribution with unknown parameters:

$$X_i | i \in q \sim \mathcal{N}(\boldsymbol{\mu}_q, \boldsymbol{\Sigma}_q)$$

Complete Likelihood (\mathbf{X}, \mathbf{Z})

The model complete loglikelihood is

$$\log L(\boldsymbol{\mu}, \boldsymbol{\Sigma}; \mathbf{X}, \mathbf{Z}) = \sum_{i=1}^n \sum_{q=1}^Q Z_{iq} \left(\log \alpha_q - \frac{1}{2} \log \det(\boldsymbol{\Sigma}_q) - \frac{1}{2} \|\mathbf{x}_i - \boldsymbol{\mu}_q\|_{\boldsymbol{\Sigma}_q^{-1}}^2 \right) + c$$

↪ Implementation of the univariate case during the labs.

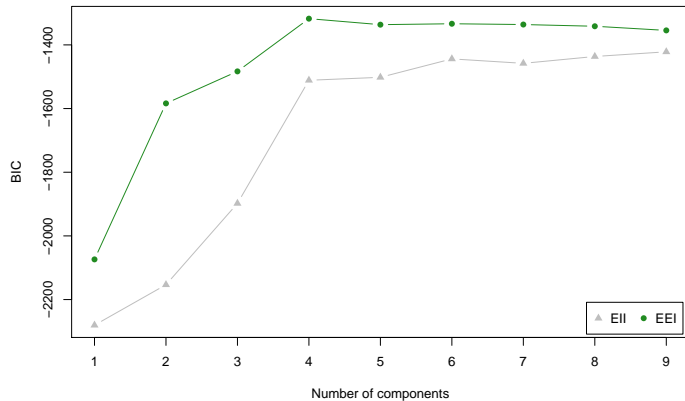
Gaussian mixture model in R I

The package Mclust is a great reference

See <https://cran.r-project.org/web/packages/mclust/vignettes/mclust.html>

```
GMM <- crabs_corrected %>%  
  select(-sex, -species) %>%  
  Mclust(modelNames = c("EII", "EEI"))  
plot(GMM, 'BIC')
```

Gaussian mixture model in R II



Gaussian mixture model in R III

```
aricode::ARI(GMM$classification, classes)

## [1] 0.6451662

aricode::ARI(GMM$classification, clusters)

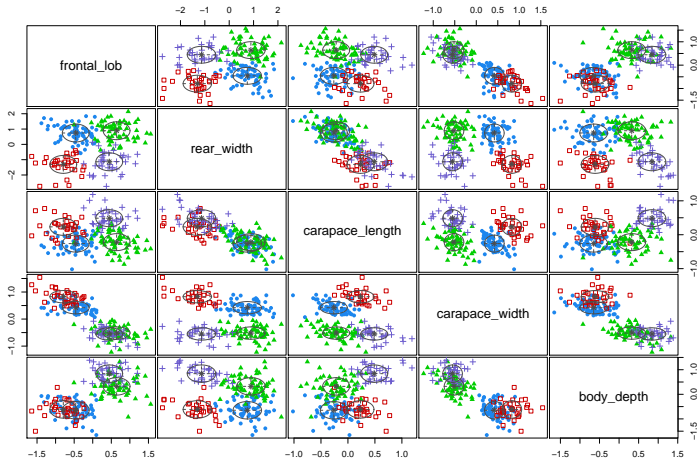
## [1] 0.8746812

aricode::ARI(GMM$classification, clusters_ward)

## [1] 0.8209783

plot(GMM, 'classification')
```

Gaussian mixture model in R IV



Example: data generation

We first generate data with 4 components:

```
mu1 <- 5    ; sigma1 <- 1; n1 <- 100
mu2 <- 10   ; sigma2 <- 1; n2 <- 200
mu3 <- 15   ; sigma3 <- 2; n3 <- 50
mu4 <- 20   ; sigma4 <- 3; n4 <- 100
cl <- rep(1:4,c(n1,n2,n3,n4))
x <- c(rnorm(n1,mu1,sigma1),rnorm(n2,mu2,sigma2),
      rnorm(n3,mu3,sigma3),rnorm(n4,mu4,sigma4))
n <- length(x)

## we randomize the class ordering
rnd <- sample(1:n)
cl <- cl[rnd]
x <- x[rnd]

alpha <- c(n1,n2,n3,n4)/n
```

Example: data generation - plot I

Let us plot the data and the theoretical mixture.

```
curve(alpha[1]*dnorm(x,mu1,sigma1) +  
      alpha[2]*dnorm(x,mu2,sigma2) +  
      alpha[3]*dnorm(x,mu3,sigma3) +  
      alpha[4]*dnorm(x,mu4,sigma3),  
      col="blue", lty=1, from=0,to=30, n=1000,  
      main="Theoretical Gaussian mixture and its components",  
      xlab="x", ylab="density")  
curve(alpha[1]*dnorm(x,mu1,sigma1), col="red", add=TRUE, lty=2)  
curve(alpha[2]*dnorm(x,mu2,sigma2), col="red", add=TRUE, lty=2)  
curve(alpha[3]*dnorm(x,mu3,sigma3), col="red", add=TRUE, lty=2)  
curve(alpha[4]*dnorm(x,mu4,sigma4), col="red", add=TRUE, lty=2)  
rug(x)
```

Example: data generation - plot II

