### Abstract

Gaussian Mixture Models (GMM), commonly used in pattern recognition and machine learning, providing a flexible probabilistic model for the data. Conventional algorithm for maximum llikelihood estimation of parameters is called Experctation-Minimization (EM) provides fixed point iteration procedure, not global optimal in general case and strongly depends on initialization. Stochastic search algorithms and in particular particle swarm optimization (PSO) are quite popular alternative for global optimization and have been already used for optizing GMM. Our contribution is that we suggest new covariance matrix parametrization that allows us to successfully apply PSO for optimizing parameters of GMM.

## Intoduction

## GMM parameters estimation

Provide statement for GMM: We consider a family of mixture of K multivariate Gaussian distributions in  $\mathbb{R}^d$  that are parametrized as follows:  $\theta = \{w_1, w_2, ..., w_K, \mu_1, \Sigma_1, ..., \mu_K, \Sigma_K\}$ . Where  $\{\mu_i, \Sigma_i\}$  is parametrization of ith multivariate Gaussian distribution,  $\mu_k \in \mathbb{R}^d$  is vector of means and covariance matrix  $\Sigma_k \in \mathbb{S}^d_{++}$  where  $\mathbb{S}^d_{++}$  is a set of symmetric positive definite  $d \times d$  matrices and  $w_k \in [0,1]$  is a weight of each gaussian distribution and  $\sum_{k=1}^K w_k = 1$ . Given a set of data  $X = \{x_1, ..., x_n\}, x_i \in \mathbb{R}^d$  which are i.i.d. Mixture probability desnity function is  $p(x|\theta) = \sum_{k=1}^K w_k p_k(x|\mu_k, \Sigma_k)$ . Our objective is to obtain maximum likelihood estimation of  $\theta$  through maximizing log-likelihood function:

$$log(X|\theta) = \sum_{i=1}^{N} log(\sum_{k=1}^{K} w_k p_k(x_i|\mu_k, \Sigma_k))$$

Since this functional is not convex we don't have covinient algorithm for obtaining global maxima.

# **Expectation Maximization**

Conventional algorithm for maximizing log-likelihood is Expectation-Minimization(EM). We do add latent variables to the model  $Z = \{z_1, ..., z_n\}, z_i \in \mathbb{Z}[0, k], z_i = k$  if  $x_i$  was generated from kth Gaussian component. Then we do add probability  $z_{ik} = q(z_i = k) = p(z_i = k | x_i, \mu_k, \Sigma_k)$  that  $x_i$  was generated from kth Gaussian component and reformulate out model and log-likelihood function considering that variables.

$$log p(X|\theta) = \mathbb{E}_{q(z)}[log \frac{p(X, Z|\theta)}{q(z)}] + KL(q||p)$$

Since  $KL(q||p) \ge 0$  we get that  $logp(X|\theta) \ge \mathbb{E}_{q(z)}[log\frac{p(X,Z|\theta)}{q(z)}]$ .  $\mathbb{E}_{q(z)}[log\frac{p(X,Z|\theta)}{q(z)}] = \mathcal{L}(q,\theta)$  is called Evidencee Lower Bound (ELBO) and we do optimize it instead of  $logp(X|\theta)$ . Then consider fixed point iteration: For fixed  $\theta^{(t)}$ , where t is the number of iteration:

$$q(Z)^{(t+1)} = argmax_q \mathcal{L}(q, \theta^{(t)}) = p(Z|X, \theta^{(t)})$$

$$z_{ik}^{(t+1)} = q(z_i = k)^{(t+1)} = \frac{w_k^{(t)} p_k(x_i, \theta_k^{(t)})}{\sum_{j=1}^K w_j^{(t)} p_j(x_i, \theta_j^{(t)})}$$

Then for fixed  $q(Z)^{(t+1)}$ 

$$\theta^{(t+1)} = argmax_{\theta} \mathcal{L}(q^{(t+1)}, \theta) = argmax_{\theta} \mathbb{E}_{q(z)} logp(X, Z | \theta)$$

That leads up to update of parameters:

$$w_k^{(t+1)} = \frac{1}{N} \sum_{i=1}^{N} z_{ik}^{(t+1)}$$

$$\mu_k^{(t+1)} = \frac{\sum_{i=1}^N z_{ik}^{(t+1)} x_i}{\sum_{i=1}^N z_{ik}^{(t+1)}}$$

$$\Sigma_k^{(t+1)} = \frac{\sum_{i+1}^N z_{ik} (x_i - \mu_k^{(t+1)}) (x_i - \mu_k^{(t+1)})^T}{\sum_{i=1}^N z_{ik}}$$

After a number of iteration EM converge and one the EM problems is that it strongly depend on initialization.

## Particle swarm optimization

Particle swarm optimization (PSO) is a population-based stochastic search algorithm that is inspired by the social interactions of swarm animals. In PSO each member of a polutaion is called particle. We have some function we want to optimize  $f(X_{\theta}), X_{\theta} \in \mathbb{R}^d$ . Each particle is represented as two vectors:  $X_i = \{X_{(\theta,i)}, x_v\}, X_{\theta}, x_v \in \mathbb{R}^d$  parameters and velocity respectively. Parameters part is the candidate solution that particle does represent and velocity is the part used for defining optimization step. We can evaluate criterion value for each particle  $f(X_{(\theta,i)})$ . As the optimization proodure we do iterative updates. At the each iteration we do remember the particle position that does have the best criterion among all the particle positions among all the iterations (global best) and is denoted as  $X^{(GB,t)}$ . Also through the search each particle remebers its personal best position  $f(X_i^{(PB,t)}) \geq f(X_{\theta,i}), i \in 1, ..., i_{last}$ 

We start from initializing particles by some way, choosing  $X^{(GB,1)}$  and  $X_i^{(PB,1)}$  and then we do run optimization steps as follows:

$$\begin{split} X_{v,i}^{(t+1)} &= \eta X_{v,i}^{(t)} + c_1 U_1^{(t)} (X_i^{(PB,t)} - X_{(\theta,i)}) + c_2 U_2^{(t)} (X^{(GB,t)} - X_{(\theta,i)}) \\ X_{(\theta,i)}^{(t+1)} &= X_{(\theta,i)}^{(t)} + X_{v,i}^{(t+1)} \\ X_i^{(PB,t+1)} &= X_i^{(PB,t^*)}, t* = argmax_t f(X_i^{(PB,t)}) \\ X^{(GB,t+1)} &= X_{\theta,i}^{(GB,t^*)}, t* = argmax_t f(X_i^{(GB,t)}) \end{split}$$

Where  $\eta$  is inertia coefficient,  $c_1$  and  $c_2$  are acceleration weights,  $U_1, U_2 \sim U[0, 1]$  random numbers and t is iteration number. We have some function we want to optimize.

## PSO for GMM

We do want apply PSO for optimizing GMM. Every particle parameters part will be representing full GMM candidate solution  $X_{(\theta,i)} = \{w_1, w_2, ..., w_K, \mu_1, \Sigma_1, ..., \mu_K, \Sigma_K\}_i$ . But we would need to maintain positive definite constraint on covariance matrix  $\Sigma \in \mathbb{S}_{++}^d$  and PSO update doesn't necessary do it.  $\Sigma_{new} = \Sigma_1 + (\Sigma_2 - \Sigma_3)$  may not be positive definite. Then to perform optimization on the manifold of positive definite matrices lets consider some parametrization of covariance matrix  $\Sigma = g(\vartheta)$  so that changing  $g(\vartheta) \in \mathbb{S}_{++}^d, \forall \vartheta \in \mathbb{R}^m$ . One possible parametrization is Cholesky decomposition  $\Sigma = LL^T$  but as a result we have  $\frac{d(d+1)}{2}$  parameters which are unbounded, can have very diverse values  $(-\infty, +\infty)$  and that doesn't allow us to properly optimze. Another parametrization through eigenvalues and givens rotation matrix angles was presented in [1].

## Eigenvalue and Givens rotation matrix angles parametrization

**Theorem 1** An arbitary covariance matrix with  $\frac{d(d+1)}{2}$  degrees of freedom can be parametrized by using d eigenvalues in a particular order and  $\frac{d(d-1)}{2}$  Givens rotation matrix angles  $\phi^{pq} \in [-\pi/4, 3\pi/4]$  for  $1 \leq p < q \leq d$  computed from the eigenvector matrix whose columns store the eigenvectors in the same order as the corresponding eigenvalues.

**Lemma 1** An eigenvactor matrix  $V \in \mathbb{R}^{d \times d}$  can be written as a product of  $\frac{d(d-1)}{2}$  Givens rotation matrices with angles  $\phi^{pq} \in [-\pi/4, 3\pi/4]$  and a diagonal matrix with  $\pm 1$  entries.

For proof see [1].  $\Sigma \in \mathbb{S}_{++}^d$  is symmetrid hence it is diagonizable in orthogonal basis and we can represent it as  $\Sigma = V\Lambda V^T = \sum_{i=1}^d \lambda_i v_i v_i^T$ . Regarding Lemma 1 that we don't need to store the diagonal matrix  $\pm 1$  entries because  $v_i v_i^T = (-v_i)(-v_i)^T$ . Then we can construct our covariance matrix as  $\Sigma = V\Lambda V^T = \sum_{i=1}^d \lambda_i v_i v_i^T = \sum_{i=1}^d \lambda_i \prod_{p=1}^d \prod_{q=p+1}^d G(p,q,\phi^{pq})$ , where  $G(p,q,\phi^{pq})$  is Givens rotation matrix with for dimensions p and q with  $\phi^{pq}$  rotation angle. Also notice since our eigendecomposition is unique up to eigenvalues/eigenvectors permutation so we need to maintain order of eigenvalues and corresponding Givens rotation angles after making the decomposition. To tackle this issue authors proposed algorithm of eigenvectors ordering w.r.t. reference eigenvectors matrix  $V_{ref}$  which happens to be the personal best for each particle.

Algorithm with such parametrization stated as follows:

### Algorithm 1

Input: d-dimensional dataset with N samples, number of iterations  $T_1$ , number of iteration for EM local convergence  $T_2$ , number of components K, number of particles M, PSO hyperparameters  $\eta, c_1, c_2$ 

```
Output: GMM solution \{w_1, w_2, ..., w_K, \mu_1, \Sigma_1, ..., \mu_K, \Sigma_K\}
for t = 1 to T_1 do
   for m = 1 to M do
       Construct K eigenvalue matrices
       Construct K eigenvector matrices by multiplying Givens rotation angles
       Run EM for local convergence for T_2 iterations T_2: number of EM iterations for each PSO iteration
       Compute K eigenvalue and eigenvector matrices via singular value decomposition of new covariance
matrices
       Reorder eigenvalues and eigenvectors of each covariance matrix according to personal best
       Extract Givens rotation angles using QR factorization
       Replace particle's means, eigenvalues, and angles
       Calculate log-likelihood
       Update personal best
   end for
   Update global best
   for m = 1 to M do
       Update particle's means, eigenvalues, and angles
   end for
end for
```

But this parametrization still have a problem since it has  $\frac{d(d+1)}{2}$  parameters and we need to compute SVD and QR factorization for each Gaussian covariance matrix for each particle for each iteration.

#### Low rank delta parametrization

We propose novel method of parametrization not the covariance matrix but the addition to it.  $\Sigma_{new} = \Sigma + \Delta(\vartheta)$ . Using property that for  $\Sigma_1, \Sigma_2 \in \mathbb{S}^d_{++}$  holds true  $\Sigma_1 + \Sigma_2 \in \mathbb{S}^d_{++}$ . We choose such  $\Delta$  that  $\Delta(\vartheta) \in \mathbb{S}^d_{++}$ ,  $\forall \vartheta \in R^m$ . Than we get that that  $\Sigma + \Delta(\vartheta) \in \mathbb{S}^d_{++}$ .

Let's try to avoid  $\frac{d(d+1)}{2}$  parameters for covariance matrix. We can do that by making low rank addition, in other words making  $\Delta$  such that  $rank(\Delta(\vartheta)) < d$  and hence we can lower number of dimensions for  $\vartheta$  so that  $\vartheta \in \mathbb{R}^m, m < \frac{d(d+1)}{2}$ .

We propose low rank delta parametrization satisfying all the above requirements as follows:

$$\Sigma_{new} = \Sigma + diag(a_1^2, ..., a_d^2) + \sum_{i=1}^{R} b_i b_i^T$$

Notice that  $diag(a_1^2,..,a_d^2) \in \mathbb{S}_{++}^d$ ,  $\forall a_i \in R$  and  $b_ib_i^T \in \mathbb{S}_{++}^d$ ,  $\forall b_i \in \mathbb{R}^d$ .  $rank(\sum_{i=1}^R b_ib_i^T) \leq R$  for approximating parameters of covariance matrix and  $diag(a_1^2,..,a_d^2)$  because we need special attention to diagonal elements. Total number of parameters is (R+1)d. For the initialization we run EM alogorithm and get  $\{w_1,w_2,..,w_K,\mu_1,\Sigma_1,..,\mu_K,\Sigma_K\}$  parameters, we freeze  $\Sigma_{(i,base)}$ , and set  $w_1,w_2,...,w_K,\mu_1,...,\mu_K$  as particle initial parameters. Then we randomly initialize low rank addition  $a_i \sim N(0,\sigma_1)$  and  $b_i \sim N(0,\sigma_2I_d)$  and compute covariance matrices for each Gaussian for mth particle as  $\Sigma_{(i,m)} = \Sigma_{(i,base)} + diag(a_{(1,m)}^2,...,a_{(d,m)}^2) + \sum_{i=1}^R b_{(i,m)} b_{(i,m)}^T$ .

Then PSO particle will contain parameters as follows:  $\{w_1, w_2, ..., w_K, \mu_1, ..., \mu_K, a_1, ..., a_d, b_1, ..., b_r\}$ Then PSO algorithm with such a parametrization will look as follows:

## Algorithm 2

Input: d-dimensional dataset with N samples, number of iterations T, number of initial EM iterations  $T_{EM}$ , number of components K, number of particles M, rank of addition R, PSO hyperparameters  $\eta, c_1, c_2$ 

```
Output: GMM solution \{w_1, w_2, ..., w_K, \mu_1, \Sigma_1, ..., \mu_K, \Sigma_K\}
run EM for T_{EM} iterations
for m = 1 to M do
    a^{(0)} \sim N(0, \sigma_1 I_d)
                                                                                      ▶ Initialize parameters for particles
    for r = 1 to R do b_r^{(0)} \sim N(0, \sigma_2 I_d)
    end for
end for
for t = 1 to T do
    for m = 1 to M do
        Calculate covariance matrices using low rank addition parameters
        Calculate log-likelihood
        Update personal best
    end for
    Update global best
    for m = 1 to M do
        Update particle's \mu_m^{(t)}, w_m^{(t)}, a_m^{(t)}, b_m^{(t)}
    end for
end for
return Global best
```

### Low rank delta parametrization for PSO with EM reinit

EM is still very effictive algo to search for a local minima so let's try to utilize it inside our PSO framework. Unfortunately we are not able to extract our parametrization from covariance matrix in a unique way so we will have an opportunity to init EM using constructed from PSO particles parameters weights, means and covariance matrices, then run EM and but we won't have an opportunity to construct PSO particle back using resulting EM weights, means and cavariances. So let's try to use PSO optimization as a proxy metric. Then we'd have criterion corresponding to each PSO particle parametrization itself,  $\theta_{pso} = \{w_1, w_2, ..., w_K, \mu_1, ..., \mu_K, a_1, ..., a_d, b_1, ..., b_r\}$ , from  $\theta_{pso}$  we'll construct GMM parameters  $\theta$  in a deterministic way  $\theta_{pso} = \{w_1, w_2, ..., w_K, \mu_1, ..., \mu_K, a_1, ..., a_d, b_1, ..., b_r\} \rightarrow \{w_1, w_2, ..., w_K, \mu_1, \sum_1, ..., \mu_K, \sum_K\} = \theta$ .  $logp(x|\theta)$  will be criterion corresponding to each PSO particle parametrization itself. And we are able to launch EM starting from  $\theta$  and obtain new set of GMM parameters  $\theta_{EM}$ .  $EM(init = \theta) = \theta_{EM}$  and corresponding criterion will be  $logp(x|\theta_{EM})$ . So we'll optimize using PSO w.r.t. first criterion but we'll evaluate second criterion and output as an alogrithm result parameters  $\theta_{EM}$ .

Also for economy of computations we can run EM reinitialization for particles each N iterations (for example 10).

Resulting algorithm will be as follows:

## Experiments

## Experiments setup

### Synthetic Data

We'll use model of generatin synthtic data from [1]. Data sets will be generated from some Gaussian Mixture Model itself with various dimensions  $d \in \{5, 10, 15, 25, 35, 50, 70\}$ , number of components will be  $M \in \{5, 10, 15, 20\}$ , with sample size  $N \in \{500, 1000, 2000\}$ . Procedure of generating data sets will be as follows:

#### Real world data

We'll be testing on the following real world datasets:

• Cloud Dataset: 10 dimensions, 1024 data points

#### Algorithm 3

Input: d-dimensional dataset with N samples, number of iterations T, number of initial EM iterations  $T_{EM}$ , number of components K, number of particles M, rank of addition R, PSO hyperparameters  $\eta, c_1, c_2$ 

```
Output: GMM solution \{w_1, w_2, ..., w_K, \mu_1, \Sigma_1, ..., \mu_K, \Sigma_K\}
run EM for T_{EM} iterations
for m=1 to M do
    a^{(0)} \sim N(0, \sigma_1 I_d)
                                                                                                ▶ Initialize parameters for particles
    for r = 1 to R do
b_r^{(0)} \sim N(0, \sigma_2 I_d)
    end for
end for
for t = 1 to T do
    for m = 1 to M do
         Calculate covariance matrices using low rank addition parameters
         Calculate log-likelihood
         Update personal best
    end for
    Update global best
    for m = 1 to M do
         Update particle's \mu_m^{(t)}, w_m^{(t)}, a_m^{(t)}, b_m^{(t)}
    end for
end for
for m=1 to M do
                                                                                                              \triangleright Each N_{reinit} iterations
    Calculate covariance matrices using low rank addition parameters for each particle's personal best
    \theta_{EM} = \text{EM}(\{w_{1,m}^{(PB)}, w_{2,m}^{(PB)}, ..., w_{K,m}^{(PB)}, \mu_{1,m}^{(PB)}, \Sigma_{1,m}^{(PB)}, ..., \mu_{K,m}^{(PB)}, \Sigma_{K,m}^{(PB)}\}) Calculate log-likelihood (logp(\theta_{EM}))
end for
Update global best EM reinit
return Global best EM reinit
```

#### Algorithm 4

```
Input: dimensionality d, number of components M, number of samples N
  Output: \{x_1, x_2, ..., x_N\}, where x_i \in \mathbb{R}^d
w \sim U[0,1]^d
w := w / \sum_{i=1}^{d} w_i for m = 1 to M do
    \mu_m \sim U[0, 100]^d
    for i = 1 to d do
        \lambda_j \sim U[1, 16]

⊳ Sample eigenvalues

        for j = i + 1 to d do
            \phi_{(p,q)} \sim U[-\pi/4, 3\pi/4]

    ▷ Sample Givens rotation angles

        \Sigma_m = f(\lambda_1, ..., \lambda_d, \phi_{(1,1)}, \phi_{(d,d)})
                                                      ▷ Construtt covariance matrix from Givens rotation angles and
eigenvalues
    end for
    Sample x_1, ..., x_N from M Gaussian Mixtures w.r.t. w
end for
return x_1,..,x_N
```

• Breast Cancer Dataset: 22 dimensions, 569 data points

• Landsat Satelite Dataset: 36 dimensions, 6435 data points

Experimenting on rank of addition							
Dataset	Rank	Amplidute	Variation of init	LogLikelihood			
Cloud	3	0	0	0			
Cloud	5	0	0	0			
Cloud	7	0	0	0			
Breast Cancer	5	0	0	0			
Breast Cancer	10	0	0	0			
Breast Cancer	15	0	0	0			

Experimenting on variation of initialization							
Dataset	Variation of init	Amplidute	Rank	LogLikelihood			
Cloud	0.03	0	0	0			
Cloud	0.001	0	0	0			
Cloud	0.003	0	0	0			
Cloud	0.0001	0	0	0			
Breast Cancer	0.03	0	0	0			
Breast Cancer	0.001	0	0	0			
Breast Cancer	0.003	0	0	0			
Breast Cancer	0.0001	0	0	0			

Experimenting on Amplitude (learning rate)							
Dataset	Amplidute	Variation of init	Rank	LogLikelihood			
Cloud	0.3	0	0	0			
Cloud	0.1	0	0	0			
Cloud	0.03	0	0	0			
Cloud	0.01	0	0	0			
Breast Cancer	0.3	0	0	0			
Breast Cancer	0.1	0	0	0			
Breast Cancer	0.03	0	0	0			
Breast Cancer	0.01	0	0	0			

# References

[1] Çağlar Arı, Selim Aksoy, and Orhan Arıkan. Maximum likelihood estimation of gaussian mixture models using stochastic search. *Pattern Recognition*, 45(7):2804–2816, 2012.