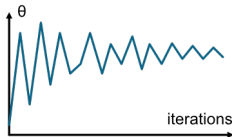




**POLITECNICO**  
MILANO 1863

# Coupled Markov chains with applications to Approximate Bayesian Computation for model based clustering

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15 February 2022



likelihood

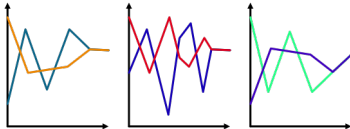
intractable



**Unbiased Markov chain  
Monte Carlo methods with  
couplings**



**Approximate Bayesian  
Computation**

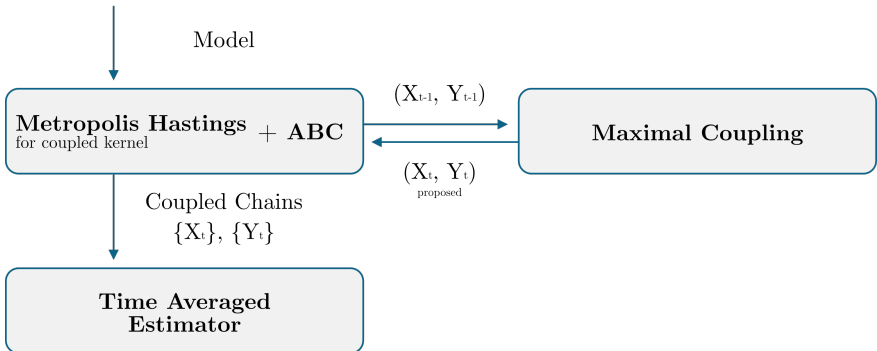


# The complete method: MCMC + Couplings + ABC

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## Implementation

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## Maximal coupling

Maximal coupling between two distributions  $p$  and  $q$

- sample  $X \sim p$ ;
- sample  $W|X \sim \mathcal{U}\{[0, p(X)]\}$ ;
- if  $W \leq q(X)$  then output  $(X, X)$ , otherwise:
  - ① sample  $Y \sim q$ ;
  - ② sample  $W^*|Y \sim \mathcal{U}\{[0, q(Y)]\}$  until  $W^* > p(Y)$
  - ③ output  $(X, Y)$ .

## Output

Distribution of a pair of random variables  $X, Y$  that maximizes  $\mathbb{P}(X = Y)$  subject to the marginal constraints  $X \sim p$  and  $Y \sim q$

# Metropolis Hastings with couplings and ABC

## Model

$$Y_i | \mu, \sigma^2 \stackrel{iid}{\sim} \mathcal{N}(\mu, \sigma^2)$$

$$\mu \sim \mathcal{N}(\mu_0, \sigma_0^2)$$

$$\sigma^2 \sim \text{InvGa}(a, b)$$

$$\mu_0 = 8, \quad \sigma_0^2 = 4$$

$$a = b = 1$$

### Initialization

set  $\theta = (\mu, \sigma^2)$

for  $k=1,2$ :

until  $K_h(||s_k^{(0)} - s_{obs}||) > 0$ :

- sample  $\theta_k^{(0)} \sim \pi(\theta)$
- simulate  $n_{obs}$  observations  $y_{ki} \sim \mathcal{N}(\mu^{(0)}, \sigma^{2(0)})$  from maximal coupling
- compute  $s_k^{(0)} = S(y_k)$  ;

## Iterations

for  $i = 1, \dots, M$ :

given  $\theta^{(i-1)} = (\mu^{(i-1)}, \sigma^{2(i-1)})$

- generate  $[\theta_1^{(i)}, \theta_2^{(i)}]$  from a maximal coupling given  $[\theta_1^{(i-1)}, \theta_2^{(i-1)}]$ ;
- for  $k=1,2$ :
  - ① simulate  $n_{obs}$  observations  $y_k^{(i)} \sim \mathcal{N}(\mu^{(i)}, \sigma^{2(i)})$  from maximal coupling;
  - ② compute  $s_k^{(i)} = S(y_k)$ ;
  - ③ accept  $\theta_k^{(i)}$  with probability

$$\frac{K_h(\|s_k^{(i)} - s_{obs}\|) \pi(\theta_k^{(i)})}{K_h(\|s_k^{(i-1)} - s_{obs}\|) \pi(\theta_k^{(i-1)})}$$



## Output

for  $k = 1, 2$

$$\mu_k^{(1)}, \dots, \mu_k^{(M)} \sim \pi_{ABC}(\mu | y_{obs});$$

$$\sigma_k^{2(1)}, \dots, \sigma_k^{2(M)} \sim \pi_{ABC}(\sigma | y_{obs}).$$

**Summary statistics:** Sample mean, Sample Variance

**Distance:**  $L^2$ -norm

**Kernel:**  $K(u) = \frac{1}{\sqrt{2\pi}} e^{-\frac{1}{2}u^2}$ ,  $K_h(u) = \frac{K(\frac{u}{h})}{h}$

**Proposal distribution:**  $\begin{bmatrix} \mu^* \\ \log(\sigma^{2*}) \end{bmatrix} \sim \mathcal{N} \left( \begin{bmatrix} \mu^{(i)} \\ \log(\sigma^{2(i)}) \end{bmatrix}, 0.1^2 \cdot \mathcal{I} \right)$

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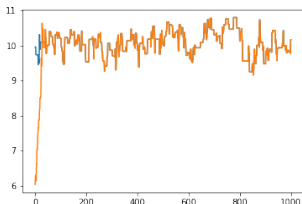
## Dataset

100 samples generated from a Gaussian distribution:

$$Y_{obs} \sim \mathcal{N}(\mu_{obs}, \sigma_{obs}^2)$$

$$\mu_{obs} = 10, \sigma_{obs}^2 = 3$$

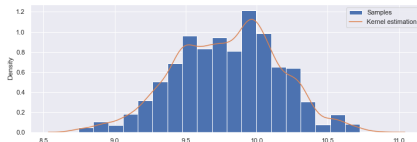
Single coupled chains



Samplings from all accepted chains

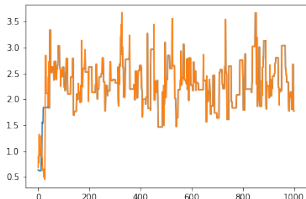


Sampling histogram with estimated distribution

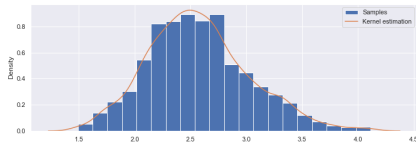


Time Averaged Estimators mean obtained:  $\mathbb{E}[H_{k:m}(X_\mu, Y_\mu)] \simeq 9.8$

Single coupled chains



Sampling histogram with estimated distribution



Samplings from all accepted chains



Time Averaged Estimators mean obtained:  $\mathbb{E}[H_{k:m}(X_{\sigma^2}, Y_{\sigma^2})] \simeq 2.56$

# Numerical experiment: g-and-k distribution

## Quantile function

$$r \in (0, 1) \mapsto a + b \left( 1 + 0.8 \left( \frac{1 - \exp(-g \cdot z(r))}{1 + \exp(-g \cdot z(r))} \right) \right) (1 + z(r)^2)^k \cdot z(r)$$

where  $z(r)$  is the  $r$ -th quantile of the standard Normal distribution.

## Prior distributions

$$a \sim \mathcal{U}([0, 10])$$

$$b \sim \mathcal{U}([0, 10])$$

$$g \sim \mathcal{U}([0, 10])$$

$$k \sim \mathcal{U}([0, 10])$$

**Summary statistics:** 10 quantiles

**Distance:**  $L^2$  – norm

**Kernel:**

$$K(u) = \frac{1}{\sqrt{2\pi}} e^{-\frac{1}{2}u^2}, \quad K_h(u) = \frac{K\left(\frac{u}{h}\right)}{h}$$

### Dataset

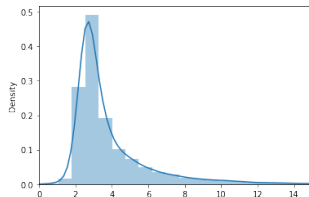
100 samples generated as follows:

$$a_{obs} = 3, b_{obs} = 1, g_{obs} = 2, k_{obs} = 0.5$$

$$y_{obs} \sim \text{quantilefunction}(z(r), \theta_{obs})$$

$$z(r) \sim \mathcal{N}(0, 1)$$

**Quantile distribution:**  $y_{obs}$





### Initialization

Given  $\theta = (a, b, g, k)$

for  $k=1,2$ :

Until  $K_h(\|s_k^{(0)} - s_{obs}\|) > 0$ :

- Generate  $\theta_k^{(0)} \sim \mathcal{U}([0, 10]^4)$  from prior density.
- Generate  $z(r) \sim \mathcal{N}(0, 1)$
- Generate a sample of  $n_{obs}$  observations such that  $y_k \sim \text{quantilefunction}(z(r), \theta_k^0)$
- Compute  $s_k^{(0)} = S(y_k)$

## Iterations

for  $i$  in  $(1, \dots, M)$ :

- generate  $\theta_1^{(i)}, \theta_2^{(i)}$  from  $\text{maximalcoupling}(\theta_1^{(i-1)}, \theta_2^{(i-1)})$
- simulate  $n_{obs}$  observations  $y_1^{(i)}, y_2^{(i)}$  from maximal coupling
- Compute the summaries  $s_k^{(i)} = S(y_k^{(i)})$
- Accept  $\theta_k^{(i)}$  with probability

$$\frac{Kh(\|s_k^{(i)} - s_{obs}\|)\pi(\theta_k^{(i)})}{Kh(\|s_k^{(i-1)} - s_{obs}\|)\pi(\theta_k^{(i-1)})}$$

otherwise  $\theta_k^{(i)} = \theta_k^{(i-1)}$ ;

## Iterations

for  $i$  in  $(1, \dots, M)$ 

## Couplings

for  $j$  in  $(1, \dots, n_{obs})$ :

- ①  $z^{(i,j)}(r) \sim \mathcal{N}(0, 1)$

- ② generate:

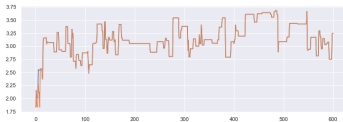
$$y_1^{(ij)} \sim \text{quantilefunction}(z^{(i,j)}(r), \theta_1^{(i-1)})$$

$$y_2^{(ij)} \sim \text{quantilefunction}(z^{(i,j)}(r), \theta_2^{(i-1)})$$

otherwise  $\theta_k^{(i)} = \theta_k^{(i-1)}$ ;

# Results for parameter a

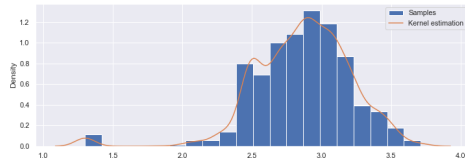
**Single coupled chains**



**Samplings from all accepted chains**

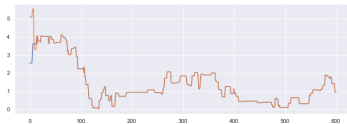


**Sampling histogram with estimated distribution**



# Results for parameter g

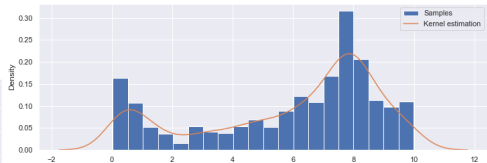
Single coupled chains



Samplings from all accepted chains



Sampling histogram with estimated distribution



Pierre Jacob, John O'Leary, and Yves Atchadé.

Unbiased markov chain monte carlo with couplings.

*Journal of the Royal Statistical Society: Series B (Statistical Methodology)*, 82, 08 2017.

Peter W. Glynn and Chang han Rhee.

Exact estimation for markov chain equilibrium expectations, 2014.

Jeffrey S. Rosenthal.

Faithful couplings of markov chains: Now equals forever.

*Advances in Applied Mathematics*, 18(3):372–381, 1997.

Dylan Cordaro.

Markov chain and coupling from the past.

2017.

Jinming Zhang.

Markov chains, mixing times and coupling methods with an application in social learning.

2020.

S. A. Sisson, Y. Fan, and M. A. Beaumont.

Overview of approximate bayesian computation, 2018.

Y. Fan and S. A. Sisson.

Abc samplers, 2018.

Dennis Prangle.

Summary statistics in approximate bayesian computation, 2015.