Supplement for The Hawkes Edge Partition Model

A INFERENCE

Next we shall explain the Gibbs sampling algorithm to infer the parameters of the Hawkes-EPM.

A.1 GIBBS SAMPLING

The conditional intensity function of the Hawkes-EPM, for the directed events from u to v, is

$$\lambda_{u,v}(t) = \mu_{u,v} + \sum_{j:t_j \in \mathcal{H}_{v,k',k,u}(t)} \gamma_{k,k'}(t-t_j)$$

$$= \sum_{k,k'} \left\{ \mu_{u,k,k',v} + \sum_{j:t_j \in \mathcal{H}_{v,k',k,u}(t)} \alpha_{k,k'} \exp[-(t-t_j)/\delta] \right\}.$$
(1)

Sampling latent variables $\{z_i^s, z_i^d\}_{i=1}^N$: For each event (t_i, s_i, d_i) , we utilize an auxiliary binary variable b_i to denote whether i-th event is triggered by the base rate (exogenous) or by opposite past interactions (endogenous) as

$$(b_i \mid -) \sim \text{Bernoulli}(\mu_{s_i,d_i}/\lambda_{s_i,d_i}(t_i)).$$
 (2)

Then, we sample the latent patterns (z_i^s, z_i^d) for each event as

$$(z_i^s, z_i^d \mid -) \sim \begin{cases} \operatorname{Cat}\left(\frac{\{\mu_{s_i, k, k', d_i}\}_{k, k'=1}^K}{\lambda_{s_i, d_i}(t_i)}\right), & \text{if } b_i = 1\\ \operatorname{Cat}\left(\frac{\{\check{\lambda}_{s_i, k, k', d_i}(t_i)\}_{k, k'=1}^K}{\lambda_{s_i, d_i}(t_i)}\right), & \text{otherwise} \end{cases}$$
(3)

where $Cat(\cdot)$ denotes the categorical distribution, and we define

$$\check{\lambda}_{s_i,k,k',d_i}(t_i) \equiv \sum_{j:t_j \in \mathcal{H}_{d_i,k',k,s_i}(t)} \alpha_{kk'} \exp[-\delta(t_i - t_j)].$$
(4)

Given the sampled latent variables, we update the sufficient statistics as

$$\hat{m}_{u,k,k',v} \equiv \sum_{j} \mathbf{1}(b_j = 1, s_j = u, d_j = v, z_j^s = k, z_j^d = k'),$$

$$\check{m}_{u,k,k',v} \equiv \sum_{j} \mathbf{1}(b_j = 0, s_j = u, d_j = v, z_j^s = k, z_j^d = k').$$
(5)

The log-posterior of the observed temporal events $\mathcal{D} \equiv \{(t_i, s_i, d_i)\}_{i=1}^N$ is shown in Eq. 6

$$\mathcal{L}(\Theta) = \sum_{i} \log \left\{ \mu_{s_{i},d_{i}} + \sum_{k,k'} \sum_{j:t_{j} \in \mathcal{H}_{d_{i},k',k,s_{i}}(t_{i})} \alpha_{kk'} \exp\left[-(t_{i} - t_{j})/\delta\right] \right\}$$

$$- \sum_{i} \left\{ \mu_{s_{i},d_{i}} T + \sum_{k,k'} \sum_{j:t_{j} \in \mathcal{H}_{d_{i},k',k,s_{i}}(t_{i})} \alpha_{kk'} \delta(1 - \exp\left[-(t_{i} - t_{j})/\delta\right]) \right\}$$

$$+ \log \Pr(\Theta).$$
(6)

Sampling the kernel parameters $\{\alpha_{kk'}\}$: As we place gamma priors over $\alpha_{kk'}$ as $\alpha_{kk'} \sim \text{Gamma}(1,1)$, and thus we have

$$(\alpha_{kk'} \mid -) \sim \operatorname{Gamma}(1 + \check{m}_{\cdot k, k' \cdot},$$

$$1 / \left[1 + \sum_{i} \sum_{j: t_j \in \mathcal{H}_{d_i, k', k, s_i}(t_i)} \frac{1}{\delta} \left(1 - \exp\left[-\frac{(T - t_j)}{\delta} \right] \right) \right],$$

$$(7)$$

where $\check{m}_{\cdot k,k'} \equiv \sum_i \check{m}_{s_i,k,k',d_i}$, and $\check{m}_{\cdot k,k'}$ denotes the total number of endogenous events associated with the latent pattern (k,k').

Sampling the base intensity $\{\mu_{u,k,k',v}\}$: As we have gamma prior over $\mu_{u,k,k',v}$ as $\mu_{u,k,k',v} \sim \text{Gamma}(\tilde{\mu}_{u,k,k',v}, 1/(\exp[-\mathbf{x}_{u,v}^T \boldsymbol{\beta}_{kk'}]))$, and thus we have

$$(\mu_{u,k,k',v} \mid -) \sim \operatorname{Gamma} \left(\tilde{\mu}_{u,k,k',v} + \hat{m}_{u,k,k',v}, 1/(T + \exp[-\mathbf{x}_{u,v}^{\mathrm{T}} \boldsymbol{\beta}_{kk'}]) \right), \tag{8}$$

Marginalizing out $\mu_{u,k,k',v}$ from the likelihood leads to

$$\Pr(\mathcal{D} \mid \mathbf{x}_{u,v}, \boldsymbol{\beta}_{kk'}) = \int \Pr(\mathcal{D} \mid \mu_{u,k,k',v}) \Pr(\mu_{u,k,k',v} \mid \mathbf{x}_{u,v}, \boldsymbol{\beta}_{kk'}) d\mu_{u,k,k',v}$$

$$\propto \text{NB}(\hat{m}_{u,k,k',v}; \tilde{\mu}_{u,k,k',v}, \sigma[\mathbf{x}_{u,v}^{T} \boldsymbol{\beta}_{kk'} + \log(T)]),$$

where $\sigma(x) = 1/(1 + \exp(-x))$ denotes the logistic function, and NB(·) denotes the Negative-Binomial distribution. Using the Pólya-Gamma data augmentation strategy (Zhou et al., 2012; Polson et al., 2013), we first sample

$$(\omega_{u,k,k',v} \mid -) \sim PG(\mu_{u,k,k',v} + \widehat{m}_{u,k,k',v}, \psi_{u,k,k',v}),$$

$$(\psi_{u,k,k',v} \mid -) \sim \mathcal{N}(\mu_{\psi}, \sigma_{\psi}),$$
(9)

where PG denotes a Pólya-Gamma draw, and where

$$\psi_{u,k,k',v} \equiv \mathbf{x}_{uv}^{\mathrm{T}} \boldsymbol{\beta}_{kk'} + \log(T\pi_{uv}),$$

$$\pi_{uv} \sim \log \mathcal{N}(0, \tau^{-1})$$

$$\sigma_{\psi} = [\omega_{u,k,k',v} + \tau]^{-1},$$

$$\mu_{\psi} = \sigma_{\psi} \left[(\hat{m}_{u,k,k',v} - \mu_{u,k,k',v})/2 + \tau(\mathbf{x}_{uv}^{\mathrm{T}} \boldsymbol{\beta}_{kk'} + \log(T)) \right],$$

where $\log \mathcal{N}(\cdot)$ denotes the lognormal distribution.

Sampling the regression coefficients $\{\beta_{kk'}\}$: Given $\{\psi_{kk'} \equiv (\psi_{1kk'1}, \dots, \psi_{Ukk'V})\}$, we sample $\{\boldsymbol{\beta}_{kk'}\}$ as

$$(\beta_{k \ k'} \mid -) \sim \mathcal{N}(\mu_{\beta}, \Sigma_{\beta}), \tag{10}$$

where
$$\Sigma_{\beta} = (\tau \mathbf{X}^{\mathrm{T}} \mathbf{X} + \mathbf{A})^{-1}$$
, $\mathbf{A} \equiv \mathrm{diag}[\nu_{1}^{-1}, \dots, \nu_{D}^{-1}]$, $\boldsymbol{\mu}_{\beta} = \tau \Sigma_{\beta} \mathbf{X}^{\mathrm{T}} (\boldsymbol{\psi}_{kk'} - \log(T))$, and $\mathbf{X} \equiv [\mathbf{x}_{11}, \dots, \mathbf{x}_{UV}]^{\mathrm{T}}$.

The full procedure of our Gibbs sampler is summarized in Algorithm 1.

Algorithm 1 Gibbs Sampler for the Hawkes Edge Partition Model

Input: events data $\mathcal{D} = \{(t_i, s_i, d_i)\}_{i=1}^N, \{\Phi, \Omega\}$ inferred by the HGaP-EPM, maximum iterations

Output: $\{\mu_{u,k,k',v}\}, \{\alpha_{kk'}\}, \{(z_i^s, z_i^d)\}$

- 1: **for** $l = 1: \mathcal{J}$ **do**
- for n = 1:N do
- 3: Sample b_i (Eq. 2)
- 4:
- Sample the latent variables (z_i^s, z_i^d) (Eq. 3) Update the intensity function $\lambda_{u,v}(t_i)$ (Eq. 1) 5:
- 6: end for
- 7: Update $\hat{m}_{u,k,k',v}$ and $\check{m}_{u,k,k',v}$ (Eq. 5)
- Sample the base intensities $\{\mu_{u,k,k',v}\}$ (Eq. 8)
- Sample the parameters $\{\beta_{kk'}\}$, $\{\omega_{u,k,k',v}\}$, $\{\psi_{u,k,k',v}\}$ (Eqs. 10; 9)
- Sample the kernel parameters $\{\alpha_{k,k'}\}$ (Eq. 7) 10:
- 11: **end for**

B BASELINE MODELS

The Hawkes Edge Partition Model (Hawkes-EPM) For each pair of nodes (u, v), $u, v \in \mathcal{V}$, and $u \neq v$,

$$\begin{split} & \mu_{u,k,k',v} \sim \operatorname{Gamma}(\tilde{\mu}_{u,k,k',v}, 1/(\exp[-\mathbf{x}_{u,v}^{\mathrm{T}}\boldsymbol{\beta}_{kk'}])), \\ & \tilde{\mu}_{u,k,k',v} \equiv \phi_{u,k}\Omega_{k,k'}\phi_{v,k'}, \\ & \boldsymbol{\beta}_{k,k'} \sim \mathcal{N}(\mathbf{0}, \mathbf{A}), \\ & \alpha_{kk'} \sim \operatorname{Gamma}(e_0, 1/f_0), \\ & \lambda_{u,v}(t) = \sum_{k,k'} \left\{ \mu_{u,k,k',v} + \sum_{j:t_j \in \mathcal{H}_{v,k',k,u}(t)} \alpha_{kk'} \exp[-(t-t_j)/\delta] \right\}, \\ & N_{uv}(t) \sim \operatorname{Hawkes Process}(\lambda_{uv}(t)), \end{split}$$

where $\mathbf{A} \equiv \operatorname{diag}[\nu_1^{-1}, \dots, \nu_D^{-1}].$

The Hawkes Dual Latent Space (DLS) (Yang et al., 2017) For each pair of nodes (u, v), $u, v \in \mathcal{V}$, and $u \neq v$,

$$\begin{aligned} \mathbf{z}_{v} &\sim \mathcal{N}(\mathbf{0}, \sigma^{2} \mathbf{I}_{d \times d}), \\ \boldsymbol{\mu}_{v} &\sim \mathcal{N}(\mathbf{0}, \sigma^{2}_{\mu} \mathbf{I}_{d \times d}), \\ \boldsymbol{\epsilon}_{v}^{(b)} &\sim \mathcal{N}(\mathbf{0}, \sigma^{2}_{\epsilon} \mathbf{I}_{d \times d}), \\ \mathbf{x}_{v}^{(b)} &\sim \boldsymbol{\mu}_{v} + \boldsymbol{\epsilon}_{v}^{(b)}, \\ \lambda_{uv}(t) &= \phi \ e^{-\|\mathbf{z}_{u} - \mathbf{z}_{v}\|_{2}^{2}} + \sum_{j: t_{j} \in \mathcal{H}_{v,u}(t)} \sum_{b=1}^{B} \beta \ e^{-\|\mathbf{x}_{u}^{(b)} - \mathbf{x}_{v}^{(b)}\|_{2}^{2}} \ \gamma_{b}(t - t_{j}), \\ N_{uv}(t) &\sim \text{Hawkes Process}(\lambda_{uv}(t)). \end{aligned}$$

The Community Hawkes Independent (CHIP) model

$$c_{u} \sim \operatorname{Categorical}(\pi_{1}, \dots, \pi_{k}), \qquad \forall u \in \mathcal{V}$$

$$\lambda_{uv}(t) = \phi_{c_{u}, c_{v}} + \sum_{j: t_{j} \in \mathcal{H}_{v, u}(t)} \alpha_{c_{u}, c_{v}} \exp\{-(t - t_{j})/\beta_{c_{u}, c_{v}}\},$$

$$N_{uv}(t) \sim \operatorname{Hawkes Process}(\lambda_{uv}(t)).$$

The Hawkes Stochastic Block (Hawkes-SBM) model

$$\begin{split} c_u &\sim \text{Categorical}(\pi_1, \dots, \pi_k), & \forall u \in \mathcal{V} \\ \lambda_{k,k'}(t) &= \phi_{k,k'} + \sum_{j: t_j \in \mathcal{H}_{k',k}(t)} \alpha_{k,k'} \exp\{-(t-t_j)/\beta_{k,k'}\}, \\ N_{k,k'}(t) &\sim \text{Hawkes Process}(\lambda_{k,k'}(t)). \end{split}$$

The Mutually Exciting Hawkes processes (MHPs) model

$$\lambda_{uv}(t) = \phi + \sum_{j:t_j \in \mathcal{H}_{v,u}(t)} \sum_{b=1}^{B} \beta_b \, \gamma_b(t - t_j),$$

$$N_{uv}(t) \sim \text{Hawkes Process}(\lambda_{uv}(t)).$$

Poisson process (PPs) model

$$\lambda_{uv}(t) = \phi_{uv},$$

 $N_{uv}(t) \sim \text{Poisson Process}(\lambda_{uv}(t)).$

C NUMERICAL SIMULATIONS

In this experiment we use synthetic data to evaluate the performance of the Hawkes-EPM in estimating the kernel parameters. We consider a collection of nodes $|\mathcal{V}|=100$, and K=4 latent communities. We generated the base rate $\mu_k \sim \mathrm{Uniform}[0,1]$, and set the kernel parameters $[\alpha_1,\alpha_2,\alpha_3,\alpha_4]=[0.5,0.88,1.38,1.96]$, and $\delta=0.45$. Via the derived Gibbs sampler, the Hawkes-EPM infers the number of latent communities. As shown in Figure (1), the posterior distributions of the estimated $\{\alpha_k\}$ concentrate toward the true values as the number of observed events is increasing.

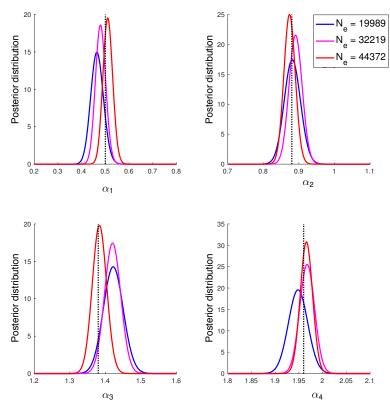


Figure 1: The posterior distribution of the estimated parameters $\{\alpha_k\}$ for the four simulations with the number of events N_e . The dashed line indicates the true values of $\{\alpha_k\}$.

D ADDITIONAL RESULTS

Figures 2 to 4 present the additional plots of the intensities of the interaction events between the nations: Iran (IRN)-USA, Israel (ISR)-Leban (LEB), Israel (ISR)-Palestin(PAL),Iraq (IRQ)-Israel (ISR), Iraq (IRQ)-Kuwait (KUW), Iraq (IRQ)-Saudi Arabi (SAU), USA-Kuwait (KUW), Iraq (IRQ)-Turkey (TUR), United Kingdom (UNK)-Iraq (IRQ).

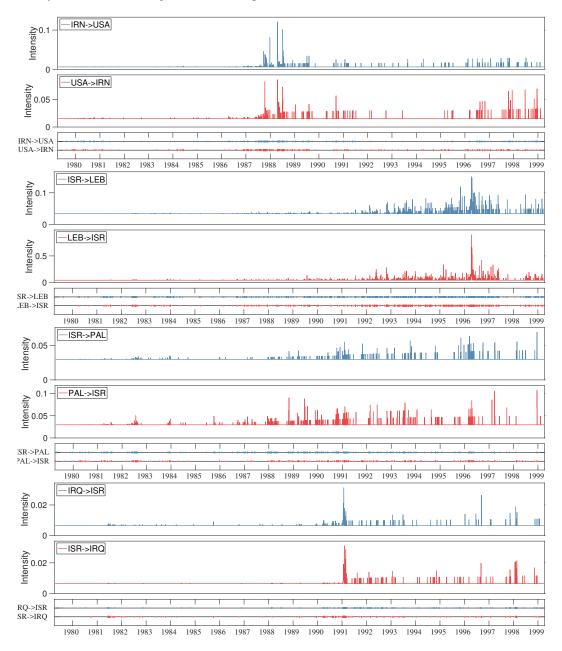


Figure 2: The plots show the intensity of interaction events among nations inferred by the Hawkes-EPM in the Gulf dataset.

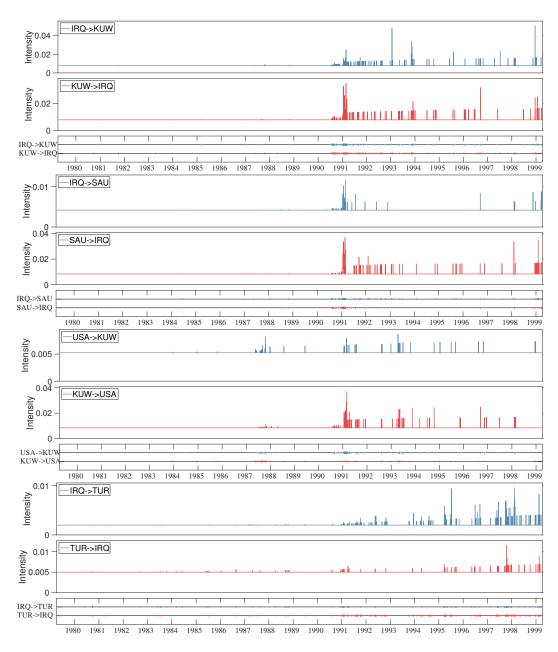


Figure 3: The plots show the intensity of interaction events among nations inferred by the Hawkes-EPM in the Gulf dataset.

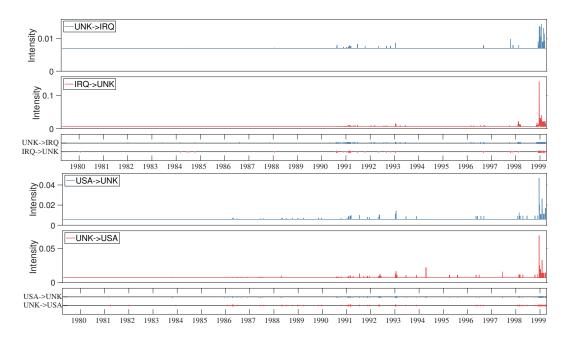


Figure 4: The plots show the intensity of interaction events among nations inferred by the Hawkes-EPM in the Gulf dataset.

References

- Polson, N. G. et al. (2013). Bayesian inference for logistic models using Pólya–Gamma latent variables. *Journal of the American Statistical Association*, 108(504):1339–1349.
- Yang, J. et al. (2017). Decoupling homophily and reciprocity with latent space network models. In *Proceedings* of the Thirty-Third Conference on Uncertainty in Artificial Intelligence.
- Zhou, M. et al. (2012). Lognormal and gamma mixed negative binomial regression. In *Proceedings of the 31st International Conference on Machine Learning*, pages 1343–1350.