

# Non-negative Monte Carlo estimation of $f$ -divergences

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## Abstract

We show how to guarantee non-negative Monte Carlo estimations of  $f$ -divergences by considering the corresponding extended  $f$ -divergences.

## 1 Problem with naive Monte Carlo estimations of $f$ -divergences

Let  $(X, F, \mu)$  be a probability space [5] with  $X$  denoting the sample space,  $F$  the  $\sigma$ -algebra, and  $\mu$  a reference positive measure. The  $f$ -divergence [3, 6] between two probability measures  $P$  and  $Q$  both absolutely continuous with respect to  $\mu$  for a convex generator  $f : (0, \infty) \rightarrow \mathbb{R}$  strictly convex at 1 and satisfying  $f(1) = 0$  is

$$I_f(P : Q) = I_f(p : q) = \int p(x) f\left(\frac{q(x)}{p(x)}\right) d\mu(x),$$

where  $P = p d\mu$  and  $Q = q d\mu$  (i.e.,  $p$  and  $q$  are Radon-Nikodym derivatives with respect to  $\mu$ ). We use the following conventions:

$$0f\left(\frac{0}{0}\right) = 0, \quad f(0) = \lim_{u \rightarrow 0^+} f(u), \quad \forall a > 0, 0f\left(\frac{a}{0}\right) = \lim_{u \rightarrow 0^+} uf\left(\frac{a}{u}\right) = a \lim_{u \rightarrow \infty} \frac{f(u)}{u}.$$

When  $f(u) = -\log u$ , we retrieve the Kullback-Leibler divergence (KLD):

$$D_{\text{KL}}(p : q) = \int p(x) \log \frac{p(x)}{q(x)} d\mu(x).$$

The KLD is usually difficult to calculate in closed-form, say, for example, between statistical mixture models [7]. A common technique is to estimate the KLD using Monte Carlo sampling using a proposal distribution  $r$ :

$$\widehat{\text{KL}}_n(p : q) = \frac{1}{n} \sum_{i=1}^n \frac{p(x_i)}{r(x_i)} \log \frac{p(x_i)}{q(x_i)},$$

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where  $x_1, \dots, x_n \sim_{\text{iid}} r$ . When  $r$  is chosen as  $p$ , the KLD can be estimated as

$$\widehat{\text{KL}}_n(p : q) = \frac{1}{n} \sum_{i=1}^n \log \frac{p(x_i)}{q(x_i)}. \quad (1)$$

Monte Carlo estimators are consistent under mild conditions:  $\lim_{n \rightarrow \infty} \widehat{\text{KL}}_n(p : q) = \text{KL}(p : q)$ .

In practice, one problem when implementing Eq. 1, is that we may end up potentially with  $\widehat{\text{KL}}_n(p : q) < 0$ . This may have disastrous consequences as algorithms implemented by programs consider non-negative divergences to execute a correct workflow. The potential negative value problem of Eq. 1 comes from the fact that  $\sum_i p(x_i) \neq 1$  and  $\sum_i q(x_i) \neq 1$ .

## 2 Non-negative Monte Carlo estimates from extended $f$ -divergences

One way to circumvent this problem is to consider the extended  $f$ -divergences:

**Definition 1 (Extended  $f$ -divergence)** *The extended  $f$ -divergence for a convex generator  $f$ , strictly convex at 1 and satisfying  $f(1) = 0$  is defined by*

$$I_f^e(p : q) = \int p(x) \left( f\left(\frac{q(x)}{p(x)}\right) - f'(1) \left(\frac{q(x)}{p(x)} - 1\right) \right) d\mu(x).$$

Indeed, for a strictly convex generator  $f$ , let us consider the scalar Bregman divergence [2]:

$$B_f(a : b) = f(a) - f(b) - (a - b)f'(b) \geq 0. \quad (2)$$

Setting  $a = \frac{q(x)}{p(x)}$  and  $b = 1$  in Eq. 2, and using the fact that  $f(1) = 0$ , we get

$$f\left(\frac{q(x)}{p(x)}\right) - \left(\frac{q(x)}{p(x)} - 1\right) f'(1) \geq 0.$$

Therefore we define the *extended  $f$ -divergences* as

$$\boxed{I_f^e(p : q) = \int p(x) B_f\left(\frac{q(x)}{p(x)} : 1\right) d\mu(x) \geq 0.} \quad (3)$$

That is, the formula for the extended  $f$ -divergences is

$$I_f^e(p : q) = \int p(x) \left( f\left(\frac{q(x)}{p(x)}\right) - f'(1) \left(\frac{q(x)}{p(x)} - 1\right) \right) d\mu(x) \geq 0. \quad (4)$$

Then we estimate the extended  $f$ -divergence using importance sampling of the integral with respect to distribution  $r$ , using  $n$  variates  $x_1, \dots, x_n \sim_{\text{iid}} p$  as:

$$\boxed{\hat{I}_{f,n}(p : q) = \frac{1}{n} \sum_{i=1}^n f\left(\frac{q(x_i)}{p(x_i)}\right) - f'(1) \left(\frac{q(x_i)}{p(x_i)} - 1\right) \geq 0.}$$

For example, for the KLD, we obtain the following Monte Carlo estimator:

$$\widehat{\text{KL}}_n(p : q) = \frac{1}{n} \sum_{i=1}^n \left( \log \frac{p(x_i)}{q(x_i)} + \frac{q(x_i)}{p(x_i)} - 1 \right) \geq 0, \quad (5)$$

since the extended KLD is

$$D_{\text{KL}^e}(p : q) = \int \left( p(x) \log \frac{p(x)}{q(x)} + q(x) - p(x) \right) d\mu(x).$$

Eq. 5 can be interpreted as a sum of scalar Itakura-Saito divergences since the Itakura-Saito divergence is scale-invariant:  $\widehat{\text{KL}}_n(p : q) = \frac{1}{n} \sum_{i=1}^n D_{\text{IS}}(p(x_i) : q(x_i))$  with the scalar Itakura-Saito divergence

$$D_{\text{IS}}(a : b) = D_{\text{IS}}\left(\frac{a}{b} : 1\right) = \frac{a}{b} - \log \frac{a}{b} - 1 \geq 0,$$

a Bregman divergence obtained for the generator  $f(u) = -\log u$ .

Notice that the extended  $f$ -divergence is a  $f$ -divergence for the generator

$$f_e(u) = f(u) - f'(1)(u - 1).$$

We check that the generator  $f_e$  satisfies both  $f(1) = 0$  and  $f'(1) = 0$ , and we have  $I_f^e(p : q) = I_{f_e}(p : q)$ . Thus  $D_{\text{KL}^e}(p : q) = I_{f_{\text{KL}}}^e(p : q)$  with  $f_{\text{KL}}^e(u) = -\log u + u - 1$ .

Let us remark that we only need to have the scalar function strictly convex at 1 to ensure that  $B_f\left(\frac{a}{b} : 1\right) \geq 0$ . Indeed, we may use the definition of Bregman divergences extended to strictly convex functions but not necessarily smooth functions [4, 8]:

$$B_f(x : y) = \max_{g(y) \in \partial f(y)} \{f(x) - f(y) - (x - y)g(y)\},$$

where  $\partial f(y)$  denotes the subderivative of  $f$  at  $y$ .

### 3 Conclusion

The  $f$ -divergence  $I_f(p : q) = \int p(x) f\left(\frac{q(x)}{p(x)}\right) d\mu(x)$  is defined for a convex generator satisfying  $f(1) = 0$  since it follows from Jensen inequality that  $I_f(p : q) \geq f\left(\int p(x) \frac{q(x)}{p(x)} d\mu(x)\right) = f(1) = 0$ . For densities, the generator  $f$  is equivalent to the family of generators  $f_\lambda(u) = f(u) + \lambda(u - 1)$  where  $\lambda \in \mathbb{R}$ :  $I_f(p : q) = I_{f_\lambda}(p : q)$ . We showed that we can express the  $f$ -divergence as a scaled integral of a scalar Bregman divergence:  $I_f(p : q) = \int p(x) B_f\left(\frac{q(x)}{p(x)} : 1\right) d\mu(x)$  provided that  $f'(1) = 0$ . This can always be done by choosing the equivalent generator  $f_\lambda$  such that  $f'_\lambda(1) = f'(1) + \lambda = 0$ , i.e.  $\lambda = -f'(1)$ . It follows that in order to have the  $f$ -divergences satisfying the law of the indiscernibles, we need to have strict convexity of  $f$  at 1. Expressing the  $f$ -divergence using a Bregman divergence allows one to

1. calculate non-negative Monte Carlo estimates  $\hat{I}_f(p : q) = \frac{1}{s} \sum_{i=1}^s \frac{p(x_i)}{r(x_i)} B_f\left(\frac{q(x_i)}{p(x_i)} : 1\right) \geq 0$  where  $x_1, \dots, x_s \sim_{\text{id}} r$ , a proposal distribution, and
2. extend the  $f$ -divergences to positive densities.

Furthermore, noticing that  $I_{\lambda f}(p : q) = \lambda I_f(p : q)$  for  $\lambda > 0$ , we may enforce that  $f''(1) = 1$ , and obtain a *standard  $f$ -divergence* [1] which enjoys the property that  $I_f(p_\theta(x) : p_{\theta+d\theta}(x)) = d\theta^\top I(\theta) d\theta$ , where  $I(\theta)$  denotes the Fisher information matrix of the parametric family  $\{p_\theta\}_\theta$  of densities.

## References

- [1] S. Amari. *Information Geometry and Its Applications*. Applied Mathematical Sciences. Springer Japan, 2016.
- [2] Lev M Bregman. The relaxation method of finding the common point of convex sets and its application to the solution of problems in convex programming. *USSR computational mathematics and mathematical physics*, 7(3):200–217, 1967.
- [3] Imre Csiszár. Information-type measures of difference of probability distributions and indirect observation. *studia scientiarum Mathematicarum Hungarica*, 2:229–318, 1967.
- [4] G. J. Gordon. *Approximate solutions to Markov decision processes*. PhD thesis, Department of Computer Science, Carnegie Mellon University, 1999.
- [5] Robert W Keener. *Theoretical statistics: Topics for a core course*. Springer, 2011.
- [6] Frank Nielsen and Richard Nock. On the chi square and higher-order chi distances for approximating  $f$ -divergences. *IEEE Signal Processing Letters*, 21(1):10–13, 2013.
- [7] Frank Nielsen and Ke Sun. Guaranteed bounds on the Kullback–Leibler divergence of univariate mixtures. *IEEE Signal Processing Letters*, 23(11):1543–1546, 2016.
- [8] Matus Telgarsky and Sanjoy Dasgupta. Agglomerative Bregman clustering. In *Proceedings of the 29th International Conference on Machine Learning*, pages 1011–1018. Omnipress, 2012.