

# Dissimilarities, divergences, and distances

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This is a working document which will be frequently updated with materials concerning the discrepancy between two distributions.

## 1 Statistical distances between densities with computationally intractable normalizers

Consider a density  $p(x) = \frac{\tilde{p}(x)}{Z_p}$  where  $\tilde{p}(x)$  is an unnormalized *computable* density and  $Z_p = \int p(x)d\mu(x)$  the *computationally intractable* normalizer (also called in statistical physics the partition function or free energy). A statistical distance  $D[p_1 : p_2]$  between two densities  $p_1(x) = \frac{\tilde{p}_1(x)}{Z_{p_1}}$  and  $p_2(x) = \frac{\tilde{p}_2(x)}{Z_{p_2}}$  with computationally intractable normalizers  $Z_{p_1}$  and  $Z_{p_2}$  is said *projective* (or two-sided *homogeneous*) if and only if

$$\forall \lambda_1 > 0, \lambda_2 > 0, \quad D[p_1 : p_2] = D[\lambda_1 p_1 : \lambda_2 p_2].$$

In particular, letting  $\lambda_1 = Z_{p_1}$  and  $\lambda_2 = Z_{p_2}$ , we have

$$D[p_1 : p_2] = D[\tilde{p}_1 : \tilde{p}_2].$$

Notice that the rhs. does not rely on the computationally intractable normalizers. These projective distances are useful in statistical inference based on minimum distance estimators [2] (see next Section).

Here are a few statistical projective distances:

- **$\gamma$ -divergences** ( $\gamma > 0$ ) [5, 3]:

$$D_\gamma[p : q] := \log \left( \int_{\mathbb{R}} q^{\alpha+1} \right) - \left( 1 + \frac{1}{\alpha} \right) \log \left( \int_{\mathbb{R}} q^\alpha p \right) + \frac{1}{\alpha} \log \left( \int_{\mathbb{R}} p^{\alpha+1} \right), \quad \gamma \geq 0$$

When  $\gamma \rightarrow 0$ , we have [3]  $D_\gamma[p : q] = D_{\text{KL}}[p : q]$ , the Kullback-Leibler divergence (KLD). For example, we can estimate the KLD between two densities of an exponential-polynomial family by Monte Carlo stochastic integration of the  $\gamma$ -divergence for a small value of  $\gamma$  [9].

The  $\gamma$ -divergences (projective, Bregman-type) and the density power divergence [1] (non-projective, Bregman-type divergence):

$$D_{\alpha}^{\text{dpd}}[p : q] := \int_{\mathbb{R}} q^{\alpha+1} - \left(1 + \frac{1}{\alpha}\right) \int_{\mathbb{R}} q^{\alpha} p + \frac{1}{\alpha} \int_{\mathbb{R}} p^{\alpha+1}, \quad \alpha \geq 0,$$

can be encapsulated into the family of  $\Phi$ -power divergences [12] (functional density power divergence class):

$$D_{\phi,\alpha}[p : q] := \phi \left( \int_{\mathbb{R}} q^{\alpha+1} \right) - \left(1 + \frac{1}{\alpha}\right) \phi \left( \int_{\mathbb{R}} q^{\alpha} p \right) + \frac{1}{\alpha} \phi \left( \int_{\mathbb{R}} p^{\alpha+1} \right), \quad \alpha \geq 0,$$

where  $\phi(e^x)$  convex and strictly increasing,  $\phi$  continuous and twice continuously differentiable with finite second order derivatives. We have  $D_{\phi,0}[p : q] = \phi'(1) \int_{\mathbb{R}} p(x) \log \frac{p(x)}{q(x)} d\mu(x) = \phi'(1) D_{\text{KL}}[p : q]$ .

- **Cauchy-Schwarz divergence** [4] (CSD, projective)

$$D_{\text{CS}}[p : q] = -\log \left( \frac{\int p(x)q(x)d\mu(x)}{\sqrt{\int p(x)^2 d\mu(x) \int q(x)^2 d\mu(x)}} \right) = D_{\text{CS}}[\lambda_1 p : \lambda_2 q], \forall \lambda_1 > 0, \lambda_2 > 0,$$

and **Hölder divergences** [11] (HD, projective, which generalizes the CSD):

$$D_{\alpha,\gamma}^{\text{Hölder}}[p : q] = -\log \left( \frac{\int_{\mathcal{X}} p(x)^{\gamma/\alpha} q(x)^{\gamma/\beta} dx}{\left(\int_{\mathcal{X}} p(x)^{\gamma} dx\right)^{1/\alpha} \left(\int_{\mathcal{X}} q(x)^{\gamma} dx\right)^{1/\beta}} \right), \quad \frac{1}{\alpha} + \frac{1}{\beta} = 1.$$

We have

$$\forall \lambda_1 > 0, \lambda_2 > 0, D_{\alpha,\gamma}^{\text{Hölder}}[\lambda_1 p : \lambda_2 q] = D_{\alpha,\gamma}^{\text{Hölder}}[p : q],$$

and

$$D_{2,2}^{\text{Hölder}}[p : q] = D_{\text{CS}}[p : q].$$

Hölder divergences between two densities  $p_{\theta_p}$  and  $p_{\theta_q}$  of an exponential family with cumulant function  $F(\theta)$  is available in closed-form [11]:

$$D_{\alpha,\gamma}^{\text{Hölder}}[p : q] = \frac{1}{\alpha} F(\gamma\theta_p) + \frac{1}{\beta} F(\gamma\theta_q) - F\left(\frac{\gamma}{\alpha}\theta_p + \frac{\gamma}{\beta}\theta_q\right)$$

The CSD is available in closed-form between mixtures of an exponential family with a conic natural parameter [7]: This includes the case of Gaussian mixture models [6].

- **Hilbert distance** [10] (projective): Consider two probability mass functions  $p = (p_1, \dots, p_d)$  and  $q = (q_1, \dots, q_d)$  of the  $d$ -dimensional probability simplex. Then the Hilbert distance is

$$D^{\text{Hilbert}}[p : q] = \log \left( \frac{\max_{i \in \{1, \dots, d\}} \frac{p_i}{q_i}}{\min_{j \in \{1, \dots, d\}} \frac{p_j}{q_j}} \right).$$

We have

$$\forall \lambda_1 > 0, \lambda_2 > 0, D^{\text{Hilbert}}[\lambda_1 p : \lambda_2 q] = D^{\text{Hilbert}}[p : q].$$

## 2 Statistical distances between empirical distributions and densities with computationally intractable normalizers

When estimating the parameter  $\hat{\theta}$  for a parametric family of distributions  $\{p_\theta\}$  from i.i.d. observations  $\mathcal{S} = \{x_1, \dots, x_n\}$ , we can define a minimum distance estimator:

$$\hat{\theta} = \arg \min_{\theta} D[p_{\mathcal{S}} : p_{\theta}],$$

where  $p_{\mathcal{S}} = \frac{1}{n} \sum_{i=1}^n \delta_{x_i}$  is the empirical distribution (normalized). Thus we need only a right-sided projective divergence to estimate models with computationally intractable normalizers.

- **Hyvärinen divergence** [] (also called **Fisher divergence**):

$$D^{\text{Hyvärinen}}[p : p_{\theta}] := \frac{1}{2} \int \|\nabla_x \log p(x) - \nabla_x \log p_{\theta}(x)\|^2 p(x) dx.$$

The Hyvärinen divergence has been extended for order- $\alpha$  Hyvärinen divergences [8] (for  $\alpha > 0$ ):

$$D_{\alpha}^{\text{Hyvärinen}}[p : q] := \frac{1}{2} \int p(x)^{\alpha} (\nabla_x \log p(x) - \nabla_x \log q(x))^2 dx, \quad \alpha > 0.$$

This column is also available in pdf: filename `Distance.pdf`

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