

# The Cramér-Rao Inequality on Singular Statistical Models

Hông Vân Lê<sup>1(✉)</sup>, Jürgen Jost<sup>2</sup>, and Lorenz Schwachhöfer<sup>3</sup>

<sup>1</sup> Mathematical Institute of ASCR, Žitná 25, 11567 Praha, Czech Republic  
hvle@math.cas.cz

<sup>2</sup> Max-Planck-Institut Für Mathematik in den Naturwissenschaften,  
Inselstrasse 22, 04103 Leipzig, Germany  
jost@mis.mpg.de

<sup>3</sup> Technische Universität Dortmund, Vogelpothsweg 87, 44221 Dortmund, Germany  
Lorenz.Schwachhoefer@math.uni-dortmund.de

**Abstract.** We introduce the notions of essential tangent space and reduced Fisher metric and extend the classical Cramér-Rao inequality to 2-integrable (possibly singular) statistical models for general  $\varphi$ -estimators, where  $\varphi$  is a  $V$ -valued feature function and  $V$  is a topological vector space. We show the existence of a  $\varphi$ -efficient estimator on strictly singular statistical models associated with a finite sample space and on a class of infinite dimensional exponential models that have been discovered by Fukumizu. We conclude that our general Cramér-Rao inequality is optimal.

## 1 $k$ -integrable Parametrized Measure Models and the Reduced Fisher Metric

In this section we recall the notion of a *k-integrable parametrized measure model* (Definitions 1, 3). Then we give a characterization of *k-integrability* (Theorem 1), which is important for later deriving the classical Cramér-Rao inequalities from our general Cramér-Rao inequality. Finally we introduce the notion of *essential tangent space* of a 2-integrable parametrized measure model (Definition 4) and the related notion of *reduced Fisher metric*.

**Notations.** For a measurable space  $\Omega$  and a finite measure  $\mu_0$  on  $\Omega$  we denote

$$\begin{aligned}\mathcal{P}(\Omega) &:= \{\mu : \mu \text{ a probability measure on } \Omega\}, \\ \mathcal{M}(\Omega) &:= \{\mu : \mu \text{ a finite measure on } \Omega\}, \\ \mathcal{S}(\Omega) &:= \{\mu : \mu \text{ a signed finite measure on } \Omega\}, \\ \mathcal{S}(\Omega, \mu_0) &= \{\mu = \phi \mu_0 : \phi \in L^1(\Omega, \mu_0)\}.\end{aligned}$$

---

Hông Vân Lê—Speaker, partially supported by RVO: 6798584.

**Definition 1** ([AJLS2016b, Definition 4.1]). Let  $\Omega$  be a measurable space.

1. A parametrized measure model is a triple  $(M, \Omega, \mathbf{p})$  where  $M$  is a (finite or infinite dimensional) Banach manifold and  $\mathbf{p} : M \rightarrow \mathcal{M}(\Omega) \subset \mathcal{S}(\Omega)$  is a Frechét- $C^1$ -map, which we shall call simply a  $C^1$ -map.
2. The triple  $(M, \Omega, \mathbf{p})$  is called a statistical model if it consists only of probability measures, i.e., such that the image of  $\mathbf{p}$  is contained in  $\mathcal{P}(\Omega)$ .
3. We call such a model dominated by  $\mu_0$  if the image of  $\mathbf{p}$  is contained in  $\mathcal{S}(\Omega, \mu_0)$ . In this case, we use the notation  $(M, \Omega, \mu_0, \mathbf{p})$  for this model.

Let  $(M, \Omega, \mathbf{p})$  be a parametrized measure model. It follows from [AJLS2016b, Proposition 2.1] that for all  $\xi \in M$  the differential  $d_\xi \mathbf{p}(V)$  is dominated by  $\mathbf{p}(\xi)$ . Hence the logarithmic derivative of  $\mathbf{p}$  at  $\xi$  in direction  $V$  [AJLS2016b, (4.2)]

$$\partial_V \log \mathbf{p}(\xi) := \frac{d\{d_\xi \mathbf{p}(V)\}}{d\mathbf{p}(\xi)} \quad (1)$$

is an element in  $L^1(\Omega, \mathbf{p}(\xi))$ . If measures  $\mathbf{p}(\xi)$ ,  $\xi \in M$ , are dominated by  $\mu_0$ , we also write

$$\mathbf{p}(\xi) = p(\xi) \cdot \mu_0 \text{ for some } p(\xi) \in L^1(\Omega, \mathbf{p}_0). \quad (2)$$

**Definition 2** ([AJLS2016b, Definition 4.2]). We say that a parametrized model  $(M, \Omega, \mu_0, \mathbf{p})$  has a regular density function if the density function  $p : \Omega \times M \rightarrow \mathbb{R}$  satisfying (2) can be chosen such that for all  $V \in T_\xi M$  the partial derivative  $\partial_V p(\cdot; \xi)$  exists and lies in  $L^1(\Omega, \mu_0)$  for some fixed  $\mu_0$ .

If the model has a positive regular density function, we have

$$\partial_V \log \mathbf{p}(\xi) = \partial_V \log p. \quad (3)$$

Next we recall the notion of  $k$ -integrability. On the set  $\mathcal{M}(\Omega)$  we define the preordering  $\mu_1 \leq \mu_2$  if  $\mu_2$  dominates  $\mu_1$ . Then  $(\mathcal{M}(\Omega), \leq)$  is a directed set, meaning that for any pair  $\mu_1, \mu_2 \in \mathcal{M}(\Omega)$  there is a  $\mu_0 \in \mathcal{M}(\Omega)$  dominating both of them (e.g.  $\mu_0 := \mu_1 + \mu_2$ ).

For fixed  $r \in (0, 1]$  and measures  $\mu_1 \leq \mu_2$  on  $\Omega$  we define the linear embedding

$$\iota_{\mu_2}^{\mu_1} : L^{1/r}(\Omega, \mu_1) \longrightarrow L^{1/r}(\Omega, \mu_2), \quad \phi \longmapsto \phi \left( \frac{d\mu_1}{d\mu_2} \right)^r.$$

Observe that

$$\begin{aligned} \|\iota_{\mu_2}^{\mu_1}(\phi)\|_{1/r} &= \left| \int_{\Omega} |\iota_{\mu_2}^{\mu_1}(\phi)|^{1/r} d\mu_2 \right|^r = \left| \int_{\Omega} |\phi|^{1/r} \frac{d\mu_1}{d\mu_2} d\mu_2 \right|^r \\ &= \left| \int_{\Omega} |\phi|^{1/r} d\mu_1 \right|^r = \|\phi\|_{1/r}^r. \end{aligned} \quad (4)$$

It has been proved that  $\iota_{\mu_2}^{\mu_1}$  is an isometry [AJLS2016b, (2.6)]. Moreover,  $\iota_{\mu_2}^{\mu_1} \iota_{\mu_3}^{\mu_2} = \iota_{\mu_3}^{\mu_1}$  whenever  $\mu_1 \leq \mu_2 \leq \mu_3$ . Then we define the *space of  $r$ -th roots of measures on  $\Omega$*  to be the directed limit over the directed set  $(\mathcal{M}(\Omega), \leq)$

$$\mathcal{S}^r(\Omega) := \varinjlim L^{1/r}(\Omega, \mu). \quad (5)$$

By [AJLS2016b, (2.9)] the space  $\mathcal{S}^r(\Omega)$  is a Banach space provided with the norm  $\|\phi\|_{1/r}$  defined in (4).

Denote the equivalence class of  $\phi \in L^{1/r}(\Omega, \mu)$  by  $\phi\mu^r$ , so that  $\mu^r \in \mathcal{S}^r(\Omega)$  is the equivalence class represented by  $1 \in L^{1/r}(\Omega, \mu)$ .

In [AJLS2016b, Proposition 2.2], for  $r \in (0, 1]$  and  $0 < k \leq 1/r$  we defined a map

$$\tilde{\pi}^k : \mathcal{S}^r(\Omega) \rightarrow \mathcal{S}^{rk}(\Omega), \phi \cdot \mu^r \mapsto \text{sign}(\phi)|\phi|^k \mu^{rk}.$$

For  $1 \leq k \leq 1/r$  the map  $\tilde{\pi}^k$  is a  $C^1$ -map between Banach spaces [AJLS2016b, (2.13)]. Using the same analogy, we set [AJLS2016b, (4.3)]

$$\mathbf{p}^{1/k} := \tilde{\pi}^{1/k} \circ \mathbf{p} : M \rightarrow \mathcal{S}^{1/k}(\Omega) \quad (6)$$

and

$$d_\xi \mathbf{p}^{1/k}(V) := \frac{1}{k} \partial_V \log \mathbf{p}(\xi) \mathbf{p}^{1/k}(\xi) \in \mathcal{S}^{1/k}(\Omega, \mathbf{p}(\xi)). \quad (7)$$

**Definition 3** ([JLS2017a, Definition 2.6]). *A parametrized measure model  $(M, \Omega, \mathbf{p})$  is called  $k$ -integrable, if the map  $\mathbf{p}^{1/k}$  from (6) is a Fréchet- $C^1$ -map.*

The  $k$ -integrability of parametrized measure models can be characterized in different ways.

**Theorem 1** ([JLS2017a, Theorem 2.7]). *Let  $(M, \Omega, \mathbf{p})$  be a parametrized measure model. Then the model is  $k$ -integrable if and only if the map*

$$V \mapsto \|d\mathbf{p}^{1/k}(V)\|_k < \infty \quad (8)$$

*defined on  $TM$  is continuous.*

Thus,  $(M, \Omega, \mathbf{p})$  is  $k$ -integrable if and only if the map  $d\mathbf{p}^{1/k} : M \rightarrow \mathcal{S}^{1/k}(\Omega)$  from (7) is well defined (i.e.,  $\partial_V \log \mathbf{p}(\xi) \in L^k(\Omega, \mathbf{p}(\xi))$ ) and continuous. In particular, the definition of  $k$ -integrability in Definition 3 above is equivalent to that in [AJLS2016b, Definition 4.4] and [AJLS2015, Definition 2.4].

**Remark 1.** 1. The Fisher metric  $\mathbf{g}$  on a 2-integrable parametrized measure model  $(M, \Omega, \mathbf{p})$  is defined as follows for  $v, w \in T_\xi M$

$$\mathbf{g}_\xi(v, w) := \langle \partial_v \log \mathbf{p}; \partial_w \log \mathbf{p} \rangle_{L^2(\Omega, \mathbf{p}(\xi))} = \langle d\mathbf{p}^{1/2}(v); d\mathbf{p}^{1/2}(w) \rangle_{\mathcal{S}^{1/2}(\Omega)} \quad (9)$$

2. The standard notion of a statistical model always assumes that it is dominated by some measure and has a positive regular density function (e.g. [Borovkov1998, p. 140, 147], [BKRW1998, p. 23], [AN2000, Sect. 2.1], [AJLS2015, Definition 2.4]). In fact, the definition of a parametrized measure model or statistical model in [AJLS2015, Definition 2.4] is equivalent to a parametrized measure model or statistical model with a positive regular density function in the sense of Definition 2, see also [AJLS2016] for detailed discussion.

Let  $(M, \Omega, \mathbf{p})$  be a 2-integrable parametrized measure model. Formula (9) shows that the kernel of the Fisher metric  $\mathbf{g}$  at  $\xi \in M$  coincides with the kernel of the map  $\Lambda_\xi : T_\xi M \rightarrow L^2(\Omega, \mathbf{p}(\xi))$ ,  $V \mapsto \partial_V(\log \mathbf{p})$ . In other words, the degeneracy of the Fisher metric  $\mathbf{g}$  is caused by the non-effectiveness of the parametrisation of the family  $\mathbf{p}(\xi)$  by the map  $\mathbf{p}$ . The tangent cone  $T_{\mathbf{p}(\xi)}\mathbf{p}(M)$  of the image  $\mathbf{p}(M) \subset \mathcal{S}(\Omega)$  is isomorphic to the quotient  $T_\xi M / \ker \Lambda_\xi$ . This motivates the following

**Definition 4.** ([JLS2017a, Definition 2.9]). *The quotient  $\hat{T}_\xi M := T_\xi M / \ker \Lambda_\xi$  will be called the essential tangent space of  $M$  at  $\xi$ .*

Clearly the Fisher metric  $\mathbf{g}$  descends to a non-degenerated metric  $\hat{\mathbf{g}}$  on  $\hat{T}M$ , which we shall call the *reduced Fisher metric*.

Denote by  $\hat{T}^{\hat{\mathbf{g}}}M$  the fiberwise completion of  $\hat{T}M$  wrt the reduced Fisher metric  $\hat{\mathbf{g}}$ . Its inverse  $\hat{\mathbf{g}}^{-1}$  is a well-defined quadratic form on the fibers of the dual bundle  $\hat{T}^{*, \hat{\mathbf{g}}^{-1}}M$ , which we can therefore identify with  $\hat{T}^{\hat{\mathbf{g}}}M$ .

## 2 The General Cramér-Rao Inequality

In this section we assume that  $(M, \Omega, \mathbf{p})$  is a 2-integrable measure model. We introduce the notion of a regular function on a measure space  $\Omega$  (Definition 5), state a rule of differentiation under integral sign (Proposition 1) and derive a general Cramér-Rao inequality (Theorem 2).

We set for  $k \in \mathbb{N}^+$

$$L_M^k(\Omega) := \{f \in L^k(\Omega, \mathbf{p}(\xi)) \text{ for all } \xi \in M\}.$$

**Definition 5.** *Let  $(M, \Omega, \mathbf{p})$  be a parametrized measure model. We call an element  $f \in L_M^k(\Omega)$  regular if the function  $\xi \mapsto \|f\|_{L^k(\Omega, \mathbf{p}(\xi))}$  is locally bounded, i.e. if for all  $\xi_0 \in M$*

$$\limsup_{\xi \rightarrow \xi_0} \|f\|_{L^k(\Omega, \mathbf{p}(\xi))} < \infty.$$

The regularity of a function  $f$  is important for the validity of differentiation under the integral sign.

**Proposition 1.** *Let  $k, k' > 1$  be dual indices, i.e.  $k^{-1} + k'^{-1} = 1$ , and let  $(M, \Omega, \mathbf{p})$  be a  $k'$ -integrable parametrized measure model. If  $f \in L_M^k(\Omega)$  is regular, then the map*

$$M \longrightarrow \mathbb{R}, \quad \xi \longmapsto \mathbb{E}_{\mathbf{p}(\xi)}(f) = \int_{\Omega} f \, d\mathbf{p}(\xi) \quad (10)$$

*is Gâteaux-differentiable, and for  $X \in TM$  the Gâteaux-derivative is*

$$\partial_X \mathbb{E}_{\mathbf{p}(\xi)}(f) = \mathbb{E}_{\mathbf{p}(\xi)}(f \, \partial_X \log \mathbf{p}(\xi)) = \int_{\Omega} f \, \partial_X \log \mathbf{p}(\xi) \, d\mathbf{p}(\xi). \quad (11)$$

Let  $V$  be a topological vector space over the real field  $\mathbb{R}$ , possibly infinite dimensional. We denote by  $V^M$  the vector space of all  $V$ -valued functions on  $M$ . A  $V$ -valued function  $\varphi$  will stand for the coordinate functions on  $M$ , or in general, a feature of  $M$  (cf. [BKRW1998]). Let  $V^*$  denote the dual space of  $V$ . Later, for  $l \in V^*$  we denote the composition  $l \circ \varphi$  by  $\varphi^l$ . This should be considered as the  $l$ -th coordinate of  $\varphi$ .

Assume that  $(M, \Omega, \mathbf{p})$  is a 2-integrable parametrized measure model. A Gateaux-differentiable function  $f$  on  $M$  whose differential  $df$  vanishes on  $\ker d\mathbf{p} \subset TP$  will be called a *visible function*.

Recall that an *estimator* is a map  $\hat{\sigma} : \Omega \rightarrow M$ . If  $k, k' > 1$  are dual indices, i.e.,  $k^{-1} + k'^{-1} = 1$ , and given a  $k'$ -integrable parametrized measure model  $(M, \Omega, \mathbf{p})$  and a function  $\varphi \in V^M$ , we define

$$L_{\varphi}^k(M, \Omega) := \{\hat{\sigma} : \Omega \rightarrow M \mid \varphi^l \circ \hat{\sigma} \in L_M^k(\Omega) \text{ for all } l \in V^*\}.$$

We call an estimator  $\hat{\sigma} \in L_{\varphi}^k(M, \Omega)$   $\varphi$ -regular if  $\varphi^l \circ \hat{\sigma} \in L_M^k(\Omega)$  is regular for all  $l \in V^*$ .

Any  $\hat{\sigma} \in L_{\varphi}^k(M, \Omega)$  induces a  $V^{**}$ -valued function  $\varphi_{\hat{\sigma}}$  on  $M$  by computing the expectation of the composition  $\varphi \circ \hat{\sigma}$  as follows

$$\langle \varphi_{\hat{\sigma}}(\xi), l \rangle := \mathbb{E}_{\mathbf{p}(\xi)}(\varphi^l \circ \hat{\sigma}) = \int_{\Omega} \varphi^l \circ \hat{\sigma} \, d\mathbf{p}(\xi) \quad (12)$$

for any  $l \in V^*$ . If  $\hat{\sigma} \in L_{\varphi}^k(M, \Omega)$  is  $\varphi$ -regular, then Proposition 1 immediately implies that  $\varphi_{\hat{\sigma}} : M \rightarrow V^{**}$  is visible with Gâteaux-derivative

$$\langle \partial_X \varphi_{\hat{\sigma}}(\xi), l \rangle = \int_{\Omega} \varphi^l \circ \hat{\sigma} \cdot \partial_X \log \mathbf{p}(\xi) \, d\mathbf{p}(\xi). \quad (13)$$

Let  $pr : TM \rightarrow \hat{T}M$  denote the natural projection.

**Definition 6.** ([JLS2017a, Definition 3.8]). *A section  $\xi \mapsto \nabla_{\hat{\mathbf{g}}} f(\xi) \in \hat{T}_{\xi}^{\hat{\mathbf{g}}} M$  will be called the generalized Fisher gradient of a visible function  $f$ , if for all  $X \in T_{\xi} M$  we have*

$$df(X) = \hat{\mathbf{g}}(pr(X), \nabla_{\hat{\mathbf{g}}} f).$$

*If the generalized gradient belongs to  $\hat{T}M$  we will call it the Fisher gradient.*

We set (cf. [Le2016])

$$\mathcal{L}_1^k(\Omega) := \{(f, \mu) \mid \mu \in \mathcal{M}(\Omega) \text{ and } f \in L^k(\Omega, \mu)\}.$$

For a map  $\mathbf{p} : P \rightarrow \mathcal{M}(\Omega)$  we denote by  $\mathbf{p}^*(\mathcal{L}_1^k(\Omega))$  the pull-back “fibration” (also called the fiber product)  $P \times_{\mathcal{M}(\Omega)} \mathcal{L}_1^k(\Omega)$ .

**Definition 7.** ([JLS2017a, Definition 3.10]). *Let  $h$  be a visible function on  $M$ . A section*

$$M \rightarrow \mathbf{p}^*(\mathcal{L}_1^2(\Omega)), \xi \mapsto \nabla h_\xi \in L^2(\Omega, \mathbf{p}(\xi)),$$

*is called a pre-gradient of  $h$ , if for all  $\xi \in M$  and  $X \in T_\xi M$  we have*

$$dh(X) = \mathbb{E}_{\mathbf{p}(\xi)}((\partial_X \log \mathbf{p}) \cdot \nabla h_\xi).$$

**Proposition 2.** ([JLS2017a, Proposition 3.12]).

1. *Let  $(M, \Omega, \mathbf{p})$  be a 2-integrable measure model and  $f \in L_M^2(\Omega, V)$  is a regular function. Then the section of the pullback fibration  $\mathbf{p}^*(\mathcal{L}_1^2(\Omega))$  defined by  $\xi \mapsto f \in L^2(\Omega, \mathbf{p}(\xi))$  is a pre-gradient of the visible function  $E_{\mathbf{p}(\xi)}(f)$ .*
2. *Let  $(P, \Omega, \mathbf{p})$  be a 2-integrable statistical model and  $f \in L_P^2(\Omega, V)$ . Then the section of the pullback fibration  $\mathbf{p}^*(\mathcal{L}_1^2(\Omega))$  defined by  $\xi \mapsto f - \mathbb{E}_{\mathbf{p}(\xi)}(f) \in L^2(\Omega, \mathbf{p}(\xi))$  is a pre-gradient of the visible function  $E_{\mathbf{p}(\xi)}(f)$ .*

For an estimator  $\hat{\sigma} \in L_\varphi^2(P, \Omega)$  we define the *variance of  $\hat{\sigma}$  w.r.t.  $\varphi$*  to be the quadratic form  $V_{\mathbf{p}(\xi)}^\varphi[\hat{\sigma}]$  on  $V^*$  such that for all  $l, k \in V^*$  we have [JLS2017a, (4.3)]

$$V_{\mathbf{p}(\xi)}^\varphi[\hat{\sigma}](l, k) := E_{\mathbf{p}(\xi)}[(\varphi^l \circ \hat{\sigma} - E_{\mathbf{p}(\xi)}(\varphi^l \circ \hat{\sigma})) \cdot (\varphi^k \circ \hat{\sigma} - E_{\mathbf{p}(\xi)}(\varphi^k \circ \hat{\sigma}))]. \quad (14)$$

We regard  $\|d\varphi_\sigma^l\|_{\hat{\mathbf{g}}^{-1}}^2(\xi)$  as a quadratic form on  $V^*$  and denote the latter one by  $(\hat{\mathbf{g}}_\sigma^\varphi)^{-1}(\xi)$ , i.e.

$$(\hat{\mathbf{g}}_\sigma^\varphi)^{-1}(\xi)(l, k) := \langle d\varphi_\sigma^l, d\varphi_\sigma^k \rangle_{\hat{\mathbf{g}}^{-1}}(\xi).$$

**Theorem 2 (General Cramér-Rao inequality)** ([JLS2017a, Theorem 4.4]). *Let  $(P, \Omega, \mathbf{p})$  be a 2-integrable statistical model,  $\varphi$  a  $V$ -valued function on  $P$  and  $\hat{\sigma} \in L_\varphi^2(P, \Omega)$  a  $\varphi$ -regular estimator. Then the difference  $V_{\mathbf{p}(\xi)}^\varphi[\hat{\sigma}] - (\hat{\mathbf{g}}_\sigma^\varphi)^{-1}(\xi)$  is a positive semi-definite quadratic form on  $V^*$  for any  $\xi \in P$ .*

*Remark 2.* Assume that  $V$  is finite dimensional and  $\varphi$  is a coordinate mapping. Then  $\mathbf{g} = \hat{\mathbf{g}}$ ,  $d\varphi^l = d\xi^l$ , and abbreviating  $b_\sigma^\varphi$  as  $b$ , we write

$$(\mathbf{g}_\sigma^\varphi)^{-1}(\xi)(l, k) = \left\langle \sum_i \left( \frac{\partial \xi^l}{\partial \xi^i} + \frac{\partial b^l}{\partial \xi^i} \right) d\xi^i, \sum_j \left( \frac{\partial \xi^k}{\partial \xi^j} + \frac{\partial b^k}{\partial \xi^j} \right) d\xi^j \right\rangle_{\mathbf{g}^{-1}}(\xi). \quad (15)$$

Let  $D(\xi)$  be the linear transformation of  $V$  whose matrix coordinates are

$$D(\xi)_k^l := \frac{\partial b^l}{\partial \xi^k}.$$

Using (15) we rewrite the Cramér-Rao inequality in Theorem 2 as follows

$$V_{\xi}[\hat{\sigma}] \geq (E + D(\xi))\mathfrak{g}^{-1}(\xi)(E + D(\xi))^T. \quad (16)$$

The inequality (16) coincides with the Cramér-Rao inequality in [Borovkov1998, Theorem 1.A, p. 147]. By Theorem 1, the condition (R) in [Borovkov1998, p. 140, 147] for the validity of the Cramér-Rao inequality is essentially equivalent to the 2-integrability of the (finite dimensional) statistical model with positive density function under consideration, more precisely Borokov ignores/excludes the points  $x \in \Omega$  where the density function vanishes for computing the Fisher metric. Borovkov also uses the  $\varphi$ -regularity assumption, written as  $\mathbb{E}_{\theta}((\theta^*)^2) < c < \infty$  for  $\theta \in \Theta$ , see also [Borovkov1998, Lemma 1, p. 141] for a more precise formulation. Classical versions of Cramér-Rao inequalities, as in e.g. [CT2006], [AN2000], are special cases of the Cramér-Rao inequality in [Borovkov1998]. We refer the reader to [JLS2017a] for comparison of our Cramér-Rao inequality with more recent Cramér-Rao inequalities in parametric statistics.

### 3 Optimality of the General Cramér-Rao Inequality

To investigate the optimality of our general Cramér-Rao inequality we introduce the following

**Definition 8** ([JLS2017b]). *Assume that  $\varphi$  is a  $V$ -valued function on  $P$ , where  $(P, \Omega, \mathbf{p})$  is a 2-integrable statistical model. A  $\varphi$ -regular estimator  $\hat{\sigma} \in L_{\varphi}^2(P, \Omega)$  will be called  $\varphi$ -efficient, if  $V_{\mathbf{p}(\xi)}^{\varphi} = (\hat{\mathfrak{g}}_{\hat{\sigma}}^{\varphi})^{-1}(\xi)$  for all  $\xi \in P$ .*

If a statistical model  $(P, \Omega, \mathbf{p})$  admits a  $\varphi$ -efficient estimator, the Cramér-Rao inequality is optimal on  $(P, \Omega, \mathbf{p})$ .

*Example 1.* Assume that  $(P \subset \mathbb{R}^n, \Omega \subset \mathbb{R}^n, \mathbf{p})$  is a minimal full regular exponential family,  $\varphi : P \rightarrow \mathbb{R}^n$  - the canonical embedding  $P \rightarrow \mathbb{R}^n$ , and  $\hat{\sigma} : \Omega \rightarrow P$  - the mean value parametrization. Then it is well known that  $\hat{\sigma}$  is an unbiased  $\varphi$ -efficient estimator, see e.g. [Brown1986, Theorem 3.6, p. 74]. Let  $S$  be a submanifold in  $P$  and  $f : P' \rightarrow P$  is a blowing-up of  $P$  along  $S$ , i.e.  $f$  is a smooth surjective map such that  $\ker df$  is non-trivial exactly at  $f^{-1}(S)$ . Then  $(P', \Omega, \mathbf{p} \circ f)$  is a strictly singular statistical model which admits an unbiased  $\varphi$ -efficient estimator, since  $(P, \Omega, \mathbf{p})$  admits unbiased  $\varphi$ -efficient estimator.

*Example 2.* Let  $\Omega_n$  be a finite set of  $n$  elements. Let  $A : \Omega_n \rightarrow \mathbb{R}_+^d$  be a map, where  $d \leq m - 1$ . We define an exponential family  $P^A(\cdot|\theta) \subset \mathcal{M}(\Omega_m)$  with parameter  $\theta$  in  $\mathbb{R}^d$  as follows.

$$P^A(x|\theta) = Z_A(\theta) \cdot \exp(\theta, A(x)), \text{ for } \theta \in \mathbb{R}^d, \text{ and } x \in \Omega_m. \quad (17)$$

Here  $Z_A(\theta)$  is the normalizing factor such that  $P^A(\cdot|\theta) \cdot \mu_0$  is a probability measure, where  $\mu_0$  is the counting measure on  $\Omega_m$ :  $\mu_0(x_i) = 1$  for  $x_i \in \Omega_m$ .

Denote  $A^l(x) := \langle l, A(x) \rangle$  for  $l \in (\mathbb{R}^d)^*$ . We set

$$\begin{aligned}\hat{\sigma} : \Omega_n \rightarrow \mathbb{R}^d, \quad x \mapsto \log A(x) &:= (\log A^1(x), \dots, \log A^d(x)), \\ \varphi : \mathbb{R}^d \rightarrow \mathbb{R}_+^d \subset \mathbb{R}^d, \quad \theta \mapsto \exp \theta.\end{aligned}$$

Then  $\hat{\sigma}$  is a (possibly biased)  $\varphi$ -efficient estimator [JLS2017b]. Using blowing-up, we obtain strictly singular statistical models admitting (possibly biased)  $\varphi$ -efficient estimators.

In [Fukumizu2009] Fukumizu constructed a large class of infinite dimensional exponential families using reproducing kernel Hilbert spaces (RKHS). Assume that  $\Omega$  is a topological space and  $\mu$  is a Borel probability measure such that  $\text{sppt}(\mu) = \Omega$ . Let  $k : \Omega \times \Omega \rightarrow \mathbb{R}$  be a continuous positive definite kernel on  $\Omega$ . It is known that for a positive definite kernel  $k$  on  $\Omega$  there exists a unique RKHS  $\mathcal{H}_k$  such that

1.  $\mathcal{H}_k$  consists of functions on  $\Omega$ ,
2. Functions of the form  $\sum_{i=1}^m a_i k(\cdot, x_i)$  are dense in  $\mathcal{H}_k$ ,
3. For all  $f \in \mathcal{H}_k$  we have  $\langle f, k(\cdot, x) \rangle = f(x)$  for all  $x \in \Omega$ ,
4.  $\mathcal{H}_k$  contains the constant functions  $c|_\Omega$ ,  $c \in \mathbb{R}$ .

For a given positive definite kernel  $k$  on  $\Omega$  we set

$$\hat{k} : \Omega \rightarrow \mathcal{H}_k, \quad \hat{k}(x) := k(\cdot, x).$$

**Theorem 3** ([JLS2017b]). *Assume that  $\Omega$  is a complete topological space and  $\mu$  is a Borel probability measure with  $\text{sppt}(\mu) = \Omega$ . Suppose that a kernel  $k$  on  $\Omega$  is bounded and satisfies the following relation whenever  $x, y \in \Omega$*

$$\hat{k}(x) - \hat{k}(y) = c|_\Omega \in \mathcal{H}_k \implies c|_\Omega = 0 \in \mathcal{H}_k. \quad (18)$$

Let

$$\mathcal{P}_\mu := \{f \in L^1(\Omega, \mu) \cap C^0(\Omega) \mid f > 0 \text{ and } \int_\Omega f d\mu = 1\}.$$

Set

$$\mathbf{p} : \mathcal{P}_\mu \rightarrow \mathcal{M}(\Omega), \quad f \mapsto f \cdot \mu_0.$$

Then there exists a map  $\varphi : \mathcal{P}_\mu \rightarrow \mathcal{H}_k$  such that  $(\mathcal{P}_\mu, \Omega, \mathbf{p})$  admits a  $\varphi$ -efficient estimator.

## References

- [AJLS2015] Ay, N., Jost, J., Lê, H.V., Schwachhöfer, L.: Information geometry and sufficient statistics. *Probab. Theory Relat. Fields* **162**, 327–364 (2015)
- [AJLS2016] Ay, N., Jost, J., Lê, H.V., Schwachhöfer, L.: *Information Geometry*. Springer (2017). *Ergebnisse der Mathematik und ihrer Grenzgebiete*



- [AJLS2016b] Ay, N., Jost, J., Lê, H.V., Schwachhöfer, L.: Parametrized measure models. (accepted for Bernoulli Journal). [arXiv:1510.07305](#)
- [AN2000] Amari, S., Nagaoka, H.: *Methods of Information Geometry*. Translations of Mathematical Monographs, vol. 191. American Mathematical Society (2000)
- [BKRW1998] Bickel, P., Klaassen, C.A.J., Ritov, Y., Wellner, J.A.: *Efficient and Adaptive Estimation for Semiparametric Models*. Springer, New York (1998)
- [Borovkov1998] Borovkov, A.A.: *Mathematical Statistics*. Gordon and Breach Science Publishers (1998)
- [Brown1986] Brown, L.D.: *Fundamentals of Statistical Families with Applications in Statistical Decision Theory*, Lecture Notes-Monograph Series, vol. 9. IMS (1986)
- [CT2006] Cover, T.M., Thomas, J.A.: *Elements of Information Theory*, 2nd edn. Wiley, Hoboken (2006)
- [Fukumizu2009] Fukumizu, K.: Exponential manifold by reproducing kernel Hilbert spaces. In: Gibilisco, P., Riccomagno, E., Rogantin, M.-P., Winn, H. (eds.) *Algebraic and Geometric methods in Statistics*, pp. 291–306. Cambridge University Press (2009)
- [JLS2017a] Jost, J., Lê, H.V., Schwachhöfer, L.: The Cramér-Rao inequality on singular statistical models I. [arXiv:1703.09403](#)
- [JLS2017b] Jost, J., Lê, H.V., Schwachhöfer, L.: The Cramér-Rao inequality on singular statistical models II (2017, preprint)
- [Le2016] Lê, H.V.: The uniqueness of the Fisher metric as information metric. *AIMS* **69**, 879–896 (2017)