Easily calculating and programming the skew Bhattacharyya coefficients and related divergences

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The α -skew Bhattacharyya coefficient [?] (for $\alpha \in (0,1)$ is a similarity measure defined by

$$\rho_{\alpha}[p:q] = \int p^{\alpha}(x)q^{1-\alpha}(x)d\mu(x) = \rho_{-\alpha}[q:p].$$

We have $0 < \rho_{\alpha}[p:q] \leq 1$.

The Bhattacharyya coefficient can be used to define dissimilarities like the α -skew Bhattacharyya distances:

$$D_{\text{Bhat}}[p:q] = -\log \rho_{\alpha}[p:q],$$

or the α -divergences for $\alpha \in \mathbb{R}$:

$$D_{\alpha}[p:q] = \begin{cases} \frac{4}{1-\alpha^2} \left(1 - \rho_{\frac{1-\alpha}{2}}(p:q)\right), & \alpha \notin \{-1,1\} \\ D_{\mathrm{KL}}[p:q], & \alpha = -1 \\ D_{\mathrm{KL}}[q:p], & \alpha = 1. \end{cases},$$

where D_{KL} denotes the Kullback-Leibler divergence:

$$D_{\mathrm{KL}}(p:q) = \int p(x) \log \left(\frac{p(x)}{q(x)}\right) \mathrm{d}\mu(x).$$

We shall consider the α -Bhattacharyya coefficient between multivariate Gaussian distributions with the same covariance matrix, where the probability density of a multivariate Gaussian distribution $p_{\mu,\Sigma}(x)$ with mean $\mu \in \mathbb{R}^d$ and covariance matrix Σ is:

$$p_{\mu,\Sigma}(x) = \frac{1}{\sqrt{\det(2\pi\Sigma)}} \exp\left(-\frac{1}{2}(x-\mu)^{\top}\Sigma^{-1}(x-\mu)\right).$$

It is reported in [?] (page 46):

$$\rho_{\alpha}[p_{\mu,\Sigma_{1}}:p_{\mu,\Sigma_{2}}] = \frac{|\Sigma_{1}|^{\frac{1-\alpha}{2}}|\Sigma_{2}|^{\frac{\alpha}{2}}}{|(1-\alpha)\Sigma_{1}+\alpha\Sigma_{2}|^{\frac{1}{2}}}.$$

We give the following calculation recipe in [?]:

• Since $\rho_{\alpha}[p_{\theta_1}:p_{\theta_2}] = \exp(-J_{F,\alpha}(\theta_1:\theta_2))$ for densities $p_{\theta}(x) = \exp(\theta^{\top}t(x) - F(\theta) + k(x))$ belonging to an exponential family (where $J_{F,\alpha}$ is a skew Jensen divergence) and since $J_{F,\alpha} = J_{G,\alpha}$ for $G(\theta) = F(\theta) + a^{\top}\theta + b$, let us choose $G(\theta) = -\log p_{\theta}(x)$. Furthermore, the densities are parameterized by their usual parameters λ which may differ from their natural parameters θ . Thus we have

$$\rho_{\alpha}[p_{\theta_1}:p_{\theta_2}] = \frac{p_{\lambda_{\alpha}}(\omega)}{p_{\lambda_1}(\omega)^{\alpha} p_{\lambda_2}(\omega)^{1-\alpha}}, \quad \forall \omega \in \mathbb{R}^d,$$

where $\lambda_{\alpha} = \lambda((1-\alpha)\theta_1 + \alpha\theta_2)$.

• To get λ_{α} , let us write $p_{\lambda_1}^{\alpha}(\omega)p_{\lambda_2}^{1-\alpha}(\omega) \propto \exp(\lambda_{\alpha}^{\top}t(\omega))$. Thus we do not need to explicitly calculate the log-normalizer $F(\theta)$ to get λ_{α} .

For the α -Bhattacharyya coefficient between same-mean Gaussian distributions, let us choose $\omega = \mu$ so that

$$p_{\mu,\Sigma}(\mu) = \frac{1}{\sqrt{\det(2\pi\Sigma)}}.$$

Therefore

$$p_{\mu,\Sigma_1}(\mu)^{\alpha} p_{\mu,\Sigma_2}(\mu)^{1-\alpha} = \frac{1}{\sqrt{(2\pi I)}|\Sigma_1|^{\alpha}|\Sigma_2|^{1-\alpha}}.$$

Hence

$$\Sigma_{\alpha} =$$

Overall, we have

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References