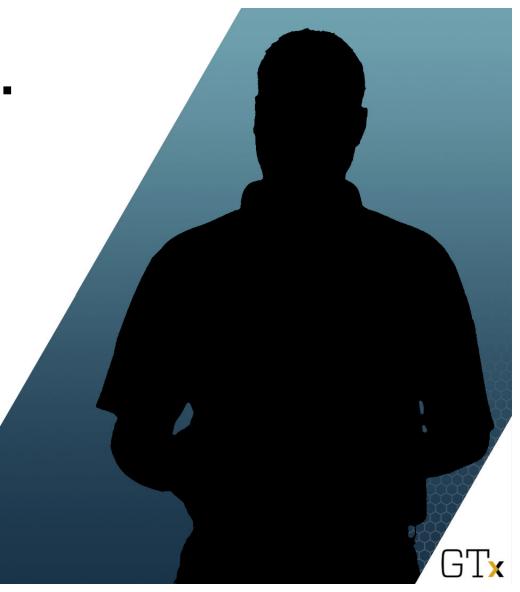




Parametric Approach

Non-Parametric Approach



Empirical Bayes

 Carl Morris (1983, JASA paper) divided Empirical Bayes to <u>parametric</u> and <u>non-parametric</u>.

Parametric Approach

$$X_i \mid \theta_i \stackrel{ind.}{\sim} f_i(x_i \mid \theta_i), \qquad i = 1, 2, ..., n$$

 $\theta_i \stackrel{i.i.d.}{\sim} \pi(\theta_i \mid \xi), \qquad \xi \text{ is common hyperparameter}$

Then
$$m_i(x_i|\xi) = \int f_i(x_i|\theta_i) \cdot \pi(\theta_i|\xi) d\theta_i$$
.
Also, $m(x_i|\xi) = \int \prod_{i=1}^n f_i(x_i|\theta_i) \cdot \prod_{i=1}^n \pi(\theta_i|\xi) d\theta_1 \cdots d\theta_n$

$$= \prod_{i=1}^n \int f_i(x_i|\theta_i) \cdot \pi(\theta_i|\xi) d\theta_i$$

$$= \prod_{i=1}^n m_i(x_i|\theta_i) \text{ independent}$$

From $m(\underline{x}|\xi) = \prod_{i=1}^n m_i(x_i|\xi)$ \longrightarrow X_i are marginally independent if $\theta_i \stackrel{i.i.d.}{\sim} \pi(\theta_i|\xi)$.

• If $f_i \equiv f$, then X_i are i.i.d. (marginally) Also, the posterior is

$$\pi(\theta_i|X_i,\xi) = \frac{f(x_i|\theta_i) \cdot \pi(\theta_i|\xi)}{m(x_i|\xi)}.$$

- ξ is unknown, can be estimated from X_1, X_2, \dots, X_n via
 - MLE (called MLE II approach)
 - MM (moment matching)

Jeremy in Empirical Bayes

Let X = (98, 107, 89, 88, 108) be Jeremy's scores on n = 5 independent IQ tests.

For this data:

$$X_i \mid \theta_i \stackrel{ind.}{\sim} N(\theta_i, \sigma^2), \ \sigma^2 \text{ known and } \sigma^2 = 80.$$

$$\theta_i \overset{i.i.d.}{\sim} N(\mu, \tau^2)$$
, Goal: estimate $\theta_i's$.

$$X_{i} \overset{i.i.d.}{\sim} N(\mu, \sigma^{2} + \tau^{2})$$

$$m\left(x \middle| \mu, \tau^{2}\right) = \prod_{i=1}^{n} m(x_{i} \middle| \mu, \tau^{2}) = \prod_{i=1}^{n} \frac{1}{\sqrt{2\pi(\sigma^{2} + \tau^{2})}} e^{-\frac{(x_{i} - \mu)^{2}}{2(\sigma^{2} + \tau^{2})}}$$

Then, the MLE of
$$\mu$$
 is $\hat{\mu} = \bar{X}$ and of τ^2 is $\hat{\tau}^2 = (s^2 - \sigma^2)_+ \equiv \max\{0, s^2 - \sigma^2\}, s^2 - \text{sample variance for } X.$

 With these estimators from the data X, the (estimated) posterior becomes:

$$\pi(\theta_i|X_i,\hat{\mu},\hat{\tau}^2) = N(\hat{B}\hat{\mu} + (1-\hat{B})x_i,(1-\hat{B})\cdot\sigma^2),$$
 where $\hat{\mu} = \bar{X},\hat{\tau}^2 = (s^2 - \sigma^2)_+$, and $\hat{B} = \frac{\sigma^2}{\sigma^2 + \hat{\tau}^2}$.

Thus, for Jeremy's data: $s^2 = 101$,

$$\hat{B} = \frac{\sigma^2}{\sigma^2 + \hat{\tau}^2} = \frac{80}{80 + (101 - 80)} = \frac{80}{101}$$

$$\hat{\theta}_1 = \frac{80}{101} \cdot 98 + \frac{21}{101} \cdot 98 = 98; \ \hat{\theta}_2 = \frac{80}{101} \cdot 98 + \frac{21}{101} \cdot 107 = 99.8713,$$
 etc.

Example: $X_i \sim \text{Pois}(\lambda_i), i = 1, 2, ..., n$ $\lambda_i \sim \text{Exp}(\mu), \ \pi(\lambda_i) = \mu e^{-\mu \lambda_i}$

Find EB estimators of λ_i .

Bayes estimator $X_i \sim \text{Pois}(\lambda_i)$, $\lambda_i \sim \text{Exp}(\mu)$

$$\lambda_i | X_i \sim \text{Ga}(x_i + 1, 1 + \mu)$$



 $\lambda_i | X_i \sim \operatorname{Ga}(x_i + 1, 1 + \mu)$ $\mathbb{E}(\lambda_i | X_i) = \frac{x_i + 1}{1 + \mu}, \text{ but } \mu \text{ may not be known}...$

Empirical Bayes

$$m(x_i) = \int_{0}^{+\infty} \frac{\lambda_i^{x_i}}{x_i!} e^{-\lambda_i} \cdot \mu e^{-\lambda_i \mu} d\lambda_i$$

$$m(x_{i}) = \int_{0}^{+\infty} \frac{\lambda_{i}^{x_{i}}}{x_{i}!} e^{-\lambda_{i}} \cdot \mu e^{-\lambda_{i}\mu} d\lambda_{i}$$

$$= \frac{1}{(1+\mu)^{x_{i}+1}} \cdot \mu \int_{0}^{+\infty} \frac{(1+\mu)^{x_{i}+1} \cdot \lambda_{i}^{x_{i}}}{P(x_{i}+1)} \cdot e^{-(1+\mu)\lambda_{i}} d\lambda_{i}$$
as integral of pdf of gamma $Ga(x_{i}+1,1+\mu)$

$$= \left(\frac{1}{1+\mu}\right)^{x_{i}} \frac{\mu}{1+\mu}, \ x_{i} = 0,1,...$$

This is geometric distribution!

Denote
$$\frac{\mu}{1+\mu} = p$$
 \longrightarrow $Ge(p)$: $P(X_i = x_i) = (1-p)^{x_i} \cdot p$

$$L = \prod_{i=1}^{n} m(x_i) = (1-p)^{\sum x_i} \cdot p^n$$

$$l = \log L = \sum x_i \cdot \log(1-p) + n \cdot \log p$$

$$l' = -\frac{\sum x_i}{1-p} + \frac{n}{p} = 0 \implies \hat{p} = \frac{n}{n+\sum x_i} = \frac{1}{1+\bar{x}}$$

Thus,
$$\frac{\mu}{1+\mu} = \frac{1}{1+\bar{x}}$$
 $\hat{\mu}$

Back to Bayes estimator with μ estimated from the data:

$$\frac{X_i+1}{1+\mu} \longrightarrow \frac{X_i+1}{1+\hat{\mu}} = \frac{X_i+1}{1+\frac{1}{\overline{X}}} = \frac{\overline{X}}{1+\overline{X}} \cdot (X_i+1).$$
Bayes EB

Thus,

$$\hat{\lambda}_i = \frac{\bar{X}}{1 + \bar{X}}(X_i + 1)$$



We assume only that parameters θ_i are i.i.d., no family of distribution is specified.

Use data to estimate marginal or the prior directly.

Pioneered by Herbert Robbins in 1950's.



Example:

Let
$$X_i | \lambda_i \sim \text{Pois}(\lambda_i), i = 1, ..., n$$

$$\lambda_i \stackrel{i.i.d.}{\sim} \pi(\lambda_i)$$

$$\hat{\lambda}_{i} = \frac{\int \lambda_{i} \cdot \frac{\lambda_{i}^{x_{i}}}{x_{i}!} e^{-\lambda_{i} \cdot \pi(\lambda_{i}) d\lambda_{i}}}{\int \frac{\lambda_{i}^{x_{i}}}{x_{i}!} e^{-\lambda_{i} \cdot \pi(\lambda_{i}) d\lambda_{i}}} = \frac{(x_{i}+1) \int \frac{\lambda_{i}^{x_{i}+1}}{(x_{i}+1)!} e^{-\lambda_{i} \cdot \pi(\lambda_{i}) d\lambda_{i}}}{\int \frac{\lambda_{i}^{x_{i}}}{x_{i}!} e^{-\lambda_{i} \cdot \pi(\lambda_{i}) d\lambda_{i}}} = (x_{i}+1) \frac{m_{\pi}(x_{i}+1)}{m_{\pi}(x_{i})}$$

Given $X_1, ..., X_n$, estimate m as \widehat{m} , and use \widehat{m} in

$$(\hat{\lambda})_{EB} = (x_i + 1) \cdot \frac{\widehat{m}(x_i+1)}{\widehat{m}(x_i)}$$

• Trivial: $\widehat{m}(x_i) = \frac{1}{n} \sum_{i=1}^{n} \mathbf{1}(X_i = x_i)$ = relative frequency of $x_i's$ in X_1, \dots, X_n

$$\left(\hat{\lambda}_i\right)_{EB} = (x_i + 1) \frac{\widehat{m}(x_i + 1)}{\frac{1}{n} + \widehat{m}(x_i)}$$

• Better estimators use smooth estimation of m(x).

Summary

In conclusion:

- Use of data to assess the prior.
 Bayesians consider prior information exogenous to observations
- In function estimation, EB is popular since it is difficult to formulate universal priors that will be efficient for any observed data
- NP EB has limited practical value
- Instead of EB, will use hierarchical Bayes' models

