

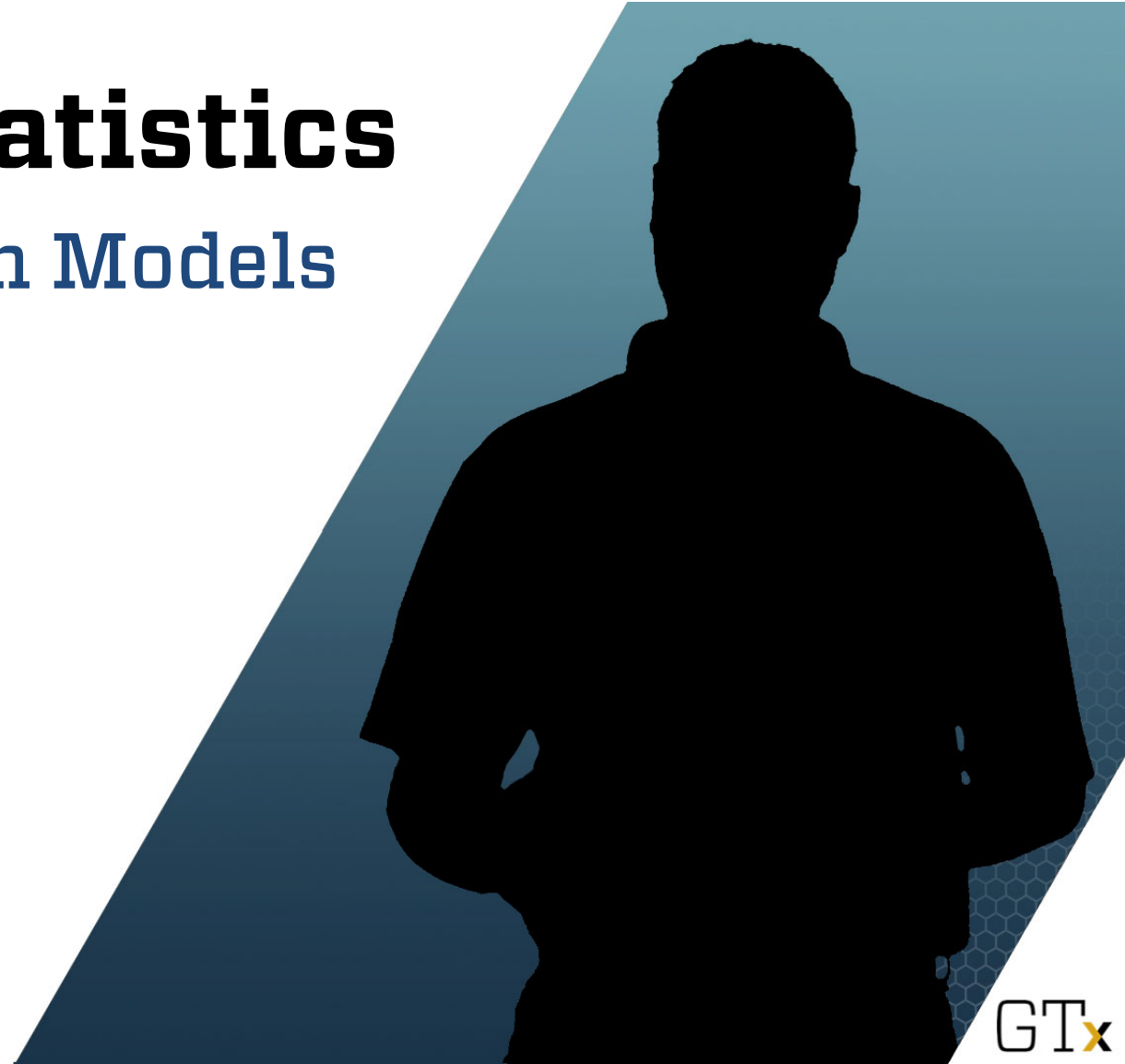
Bayesian Statistics

Various Bayesian Models

Brani Vidakovic

Professor

School of Industrial and Systems Engineering



GTx

Before We Begin...

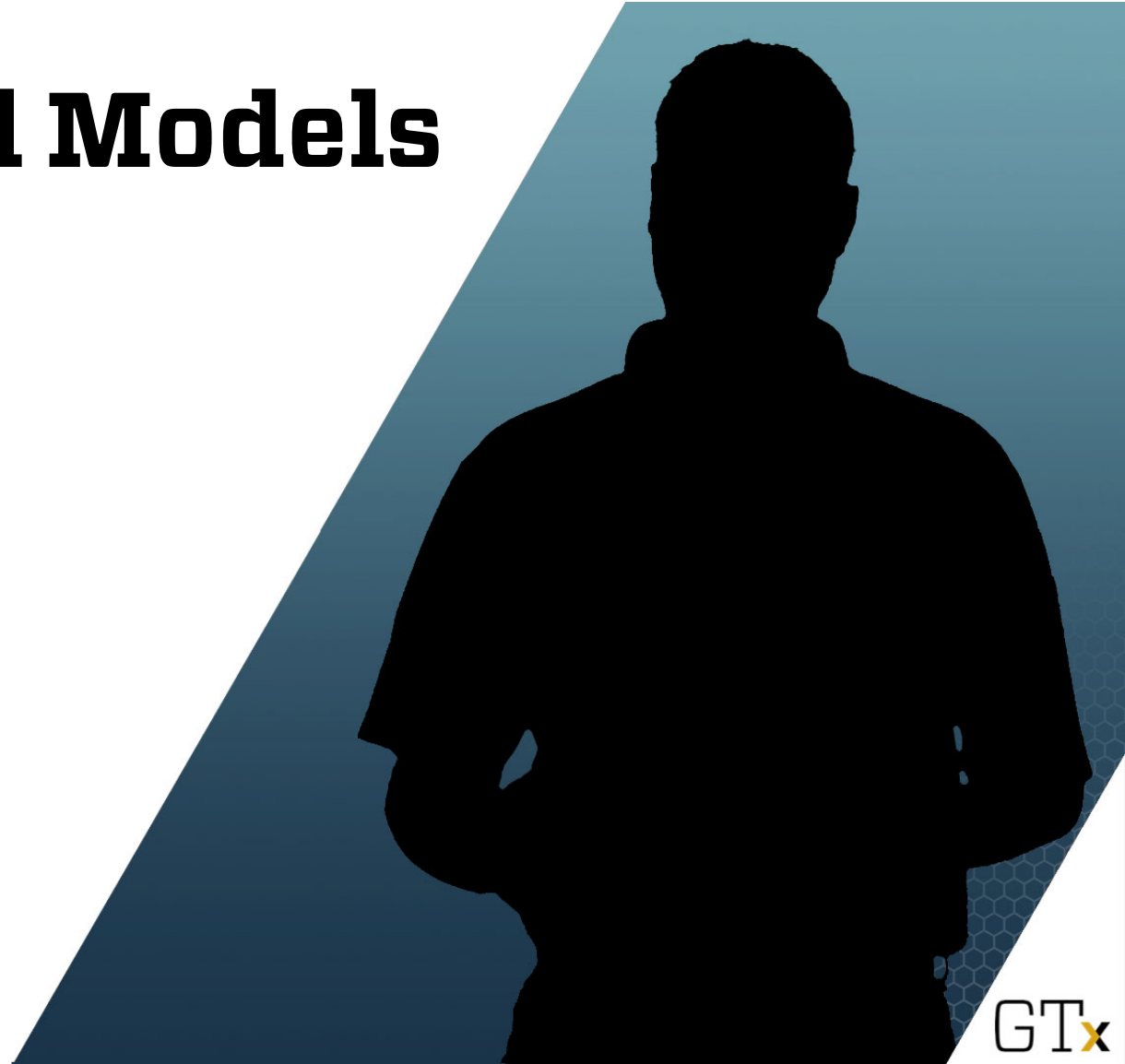
In this unit:

- Hierarchical Models
- Bayesian Linear Models
- Other Models



Hierarchical Models

- Why the hierarchy?
- Exchangeability
- Examples



Why the Hierarchy?

Prior $\pi(\theta) = \int \pi_1(\theta|\theta_1)\pi_2(\theta_1|\theta_2) \dots \pi_n(\theta_{n-1}|\theta_n)\pi_{n+1}(\theta_n)d\theta_1d\theta_2 \dots d\theta_n$

$$\left\{ \begin{array}{l} X|\theta \sim f(x|\theta) \\ \theta|\theta_1 \sim \pi_1(\theta|\theta_1) \\ \theta_1|\theta_2 \sim \pi_2(\theta_1|\theta_2) \\ \vdots \\ \theta_{n-1}|\theta_n \sim \pi_n(\theta_{n-1}|\theta_n) \\ \theta_n \sim \pi_{n+1}(\theta_n) \end{array} \right\} \Leftrightarrow \left\{ \begin{array}{l} X|\theta \sim f(x|\theta) \\ \theta \sim \pi(\theta) \end{array} \right\}$$

Why the Hierarchy?

$$X \leftarrow \theta \leftarrow \theta_1 \leftarrow \theta_2 \leftarrow \cdots \theta_{n-1} \leftarrow \theta_n$$

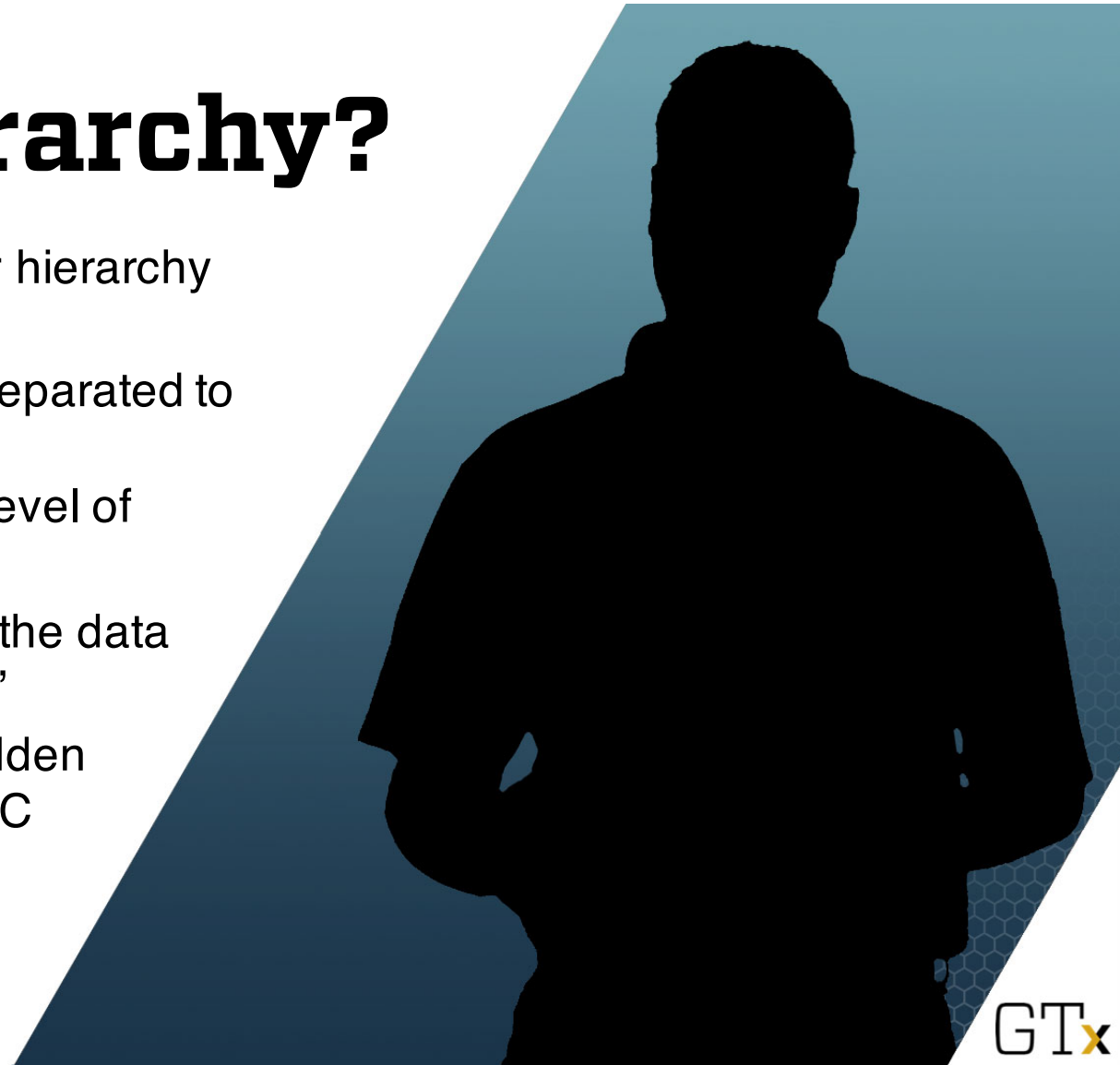
Joint distribution of $(X, \theta, \theta_1, \dots, \theta_n)$

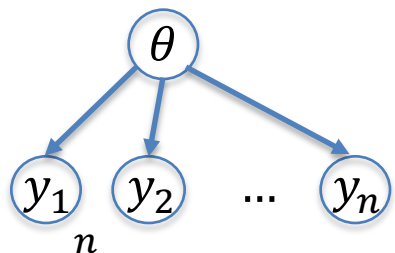
$$\begin{aligned} f(x, \theta, \theta_1, \dots, \theta_n) \propto & f(x|\theta, \theta_1, \dots, \theta_n) \times \\ & \pi_1(\theta|\theta_1, \theta_2, \dots, \theta_n) \times \\ & \pi_2(\theta_1|\theta_2, \dots, \theta_n) \times \\ & \vdots \\ & \pi_n(\theta_{n-1}|\theta_n) \times \\ & \pi_{n+1}(\theta_n). \end{aligned}$$

$$f(x, \theta, \theta_1, \dots, \theta_n) \propto f(x|\theta) \pi_1(\theta|\theta_1) \pi_2(\theta_1|\theta_2) \dots \pi_n(\theta_{n-1}|\theta_n) \pi_{n+1}(\theta_n)$$

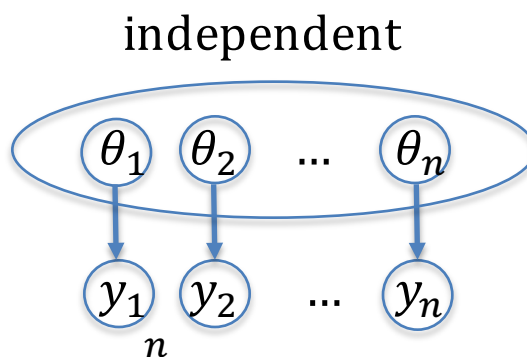
Why the Hierarchy?

- Modeling requirements ask for hierarchy (e.g. Bayesian Meta-Analysis)
- The prior information can be separated to structural part and subjective/noninformative part at higher level of hierarchy
- Robustness + objectivity: “Let the data “talk” about hyperparameters.”
- Computing issues (utilizing hidden mixtures, mixture priors, MCMC efficiency, etc.)

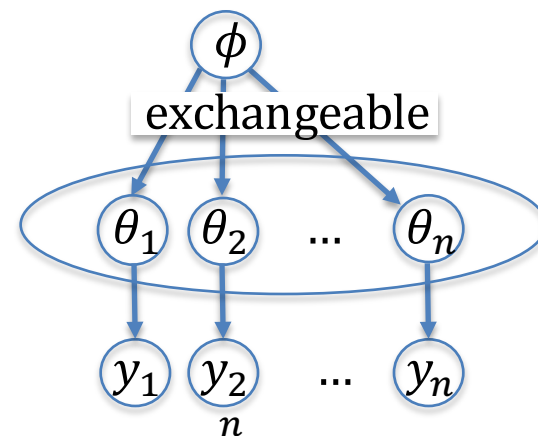




$$\prod_{i=1}^n f(y_i|\theta)\pi(\theta)$$



$$\prod_{i=1}^n f(y_i|\theta_i)\pi(\theta_i)$$



$$\pi(\phi) \prod_{i=1}^n f(y_i|\theta_i)\pi(\theta_i|\phi)$$

Y_1, Y_2, \dots, Y_n exchangeable

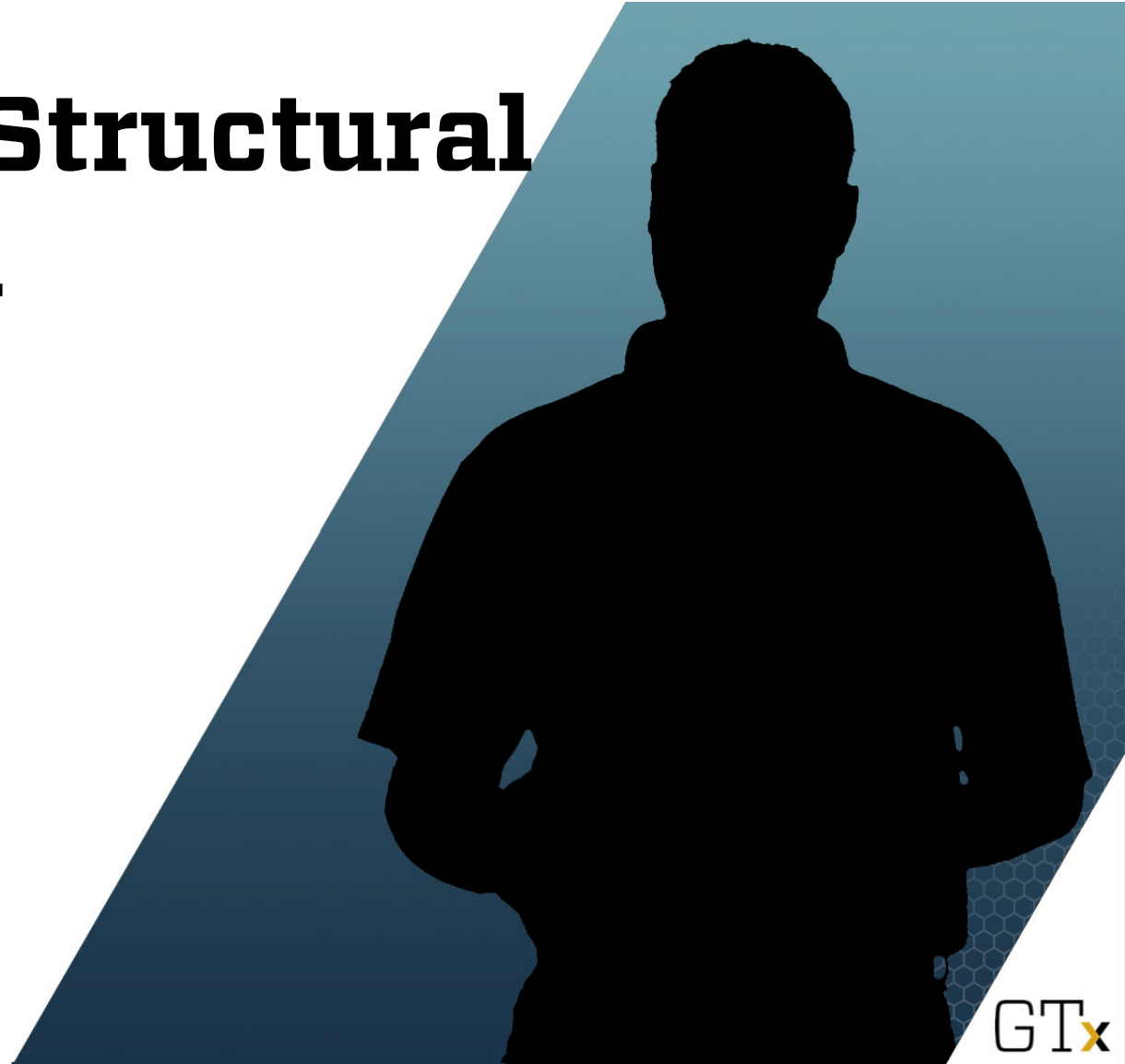
$$\Leftrightarrow (Y_1, \dots, Y_n) \stackrel{d}{=} (Y_{\pi_1}, Y_{\pi_2}, \dots, Y_{\pi_n})$$

where (π_1, \dots, π_n) is any permutation of $(1, 2, \dots, n)$

- $(X, Y) \sim \text{MVN}_2 \left(0, \begin{pmatrix} 1 & \rho \\ \rho & 1 \end{pmatrix} \right), \quad \rho \in (-1, 1)$

are exchangeable but not independent

Priors with Structural Information



- **Example:**

$$X|p \sim \text{Bin}(n, p)$$

$$p|k \sim \text{Beta}(k, k), \quad k \in \mathbb{N}$$

$$k|r \sim \text{Geom}(r), \quad p(k = i) = (1 - r)^{i-1}r, \quad i = 1, 2, \dots \quad 0 < r < 1$$

$$r \sim \text{Beta}(2, 2)$$

$$[p|k] \times [k|r] \times [r] \propto \frac{1}{B(k, k)} p^{k-1} (1 - p)^{k-1} (1 - r)^{k-1} r \frac{1}{B(2, 2)} r (1 - r),$$

$$[p|k] \times [k] \propto \frac{B(3, k + 1)}{B(k, k) B(2, 2)} p^{k-1} (1 - p)^{k-1}$$

$$[p] \propto \sum_{k=1}^{\infty} \frac{B(3, k + 1)}{B(k, k) B(2, 2)} p^{k-1} (1 - p)^{k-1} =$$

$$\frac{2p^4(4a - 15) - 4p^3(4a - 15) + 2p^2(11a - 25) - 2p(7a - 10) + (3a - 3)}{20p^4(1 - p)^4},$$

$$p \in (0, 1) \quad a = \sqrt{(2p - 1)^2}$$

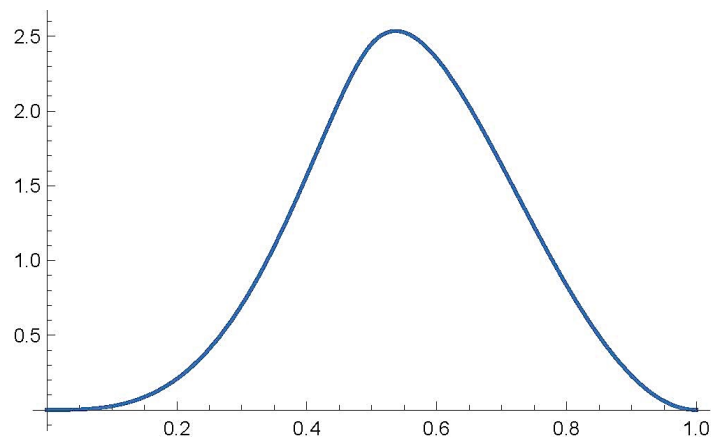
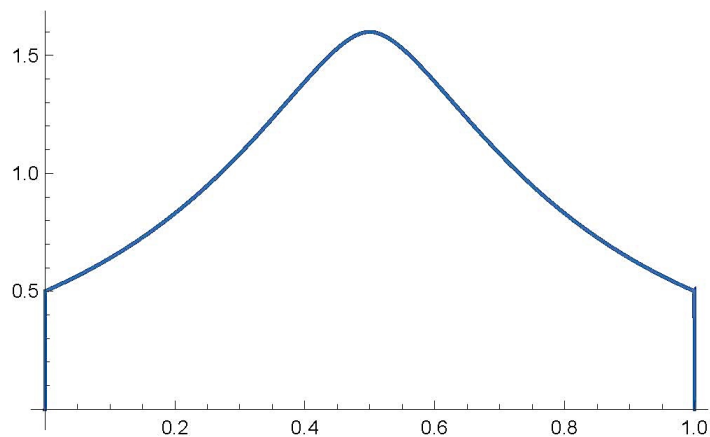
$$X|p \sim \text{Bin}(n, p): \quad f(x|p) = \binom{n}{x} p^x (1-p)^{n-x}$$

$$p \sim \pi(p); \quad \textbf{Assume } X = 3, n = 5$$

$$E^{\pi} p = \frac{1}{2} = 0.5$$

$$\hat{p} = \frac{X}{n} = \frac{3}{5} = 0.6$$

$$E^{\pi(\cdot|X)} p = \frac{80 \log 2 - 55}{56 \log 2 - 38} = 0.553481$$



OpenBUGS

File Edit Attributes Tools Info Model Inference Doodle Map PKBugs PKHelp Text Window Examples Manuals Help

basichierarhy

A SIMPLE HIERARCHICAL MODEL

Suppose that in a cohort of n people X of them have a particular condition. The likelihood is $[X|p] \sim \text{Bin}(n,p)$, and population proportion p is the parameter of interest. You believe that the proportion is close to $1/2$, but not quite sure. The appropriate prior on p would be Beta $\text{Be}(k,k)$ for k positive integer; it is symmetric about $k/(k+k)=1/2$. Thus, $[p|k] \sim \text{Be}(k,k)$. You are reluctant to specify k since this would specify the variance on p as $1/(8k+4)$. Instead you place a hyperprior on k , as geometric with parameter r , and hyperprior on r as $\text{Be}(2,2)$.

*** What is the Bayes estimator of p if $n=5$ and $X=3$?

*** What are the Bayes estimators of k and r .

```
model{
x ~ dbin(p, n)
p ~ dbeta(k,k)
k ~ dgeom(r)
# This works for OpenBUGS only. For WinBUGS
# instead of k ~ dgeom(r) one needs
# k1 ~ dnegbin(r,1)
# k <- k1+1
r ~ dbeta(2,2)
}
```

Delete

OpenBUGS

File Edit Attributes Tools Info Model Inference Doodle Map PKBugs PKHelp Text Window Examples Manuals Help

basichierarhy

*** What is the Bayes estimator of p if $n=5$ and $X=3$?
*** What are the Bayes estimators of k and r .

```
model{
x ~ dbin(p, n)
p ~ dbeta(k,k)
k ~ dgeom(r)
# This works for OpenBUGS only. For WinBUGS
# instead of k ~ dgeom(r) one needs
# k1 ~ dnegbin(r,1)
# k <- k1+1
r ~ dbeta(2,2)
}
```

DATA
list(n=5, x=3)

INITS
list(r=0.5, k=2, p=0.5)

Delete

Node statistics

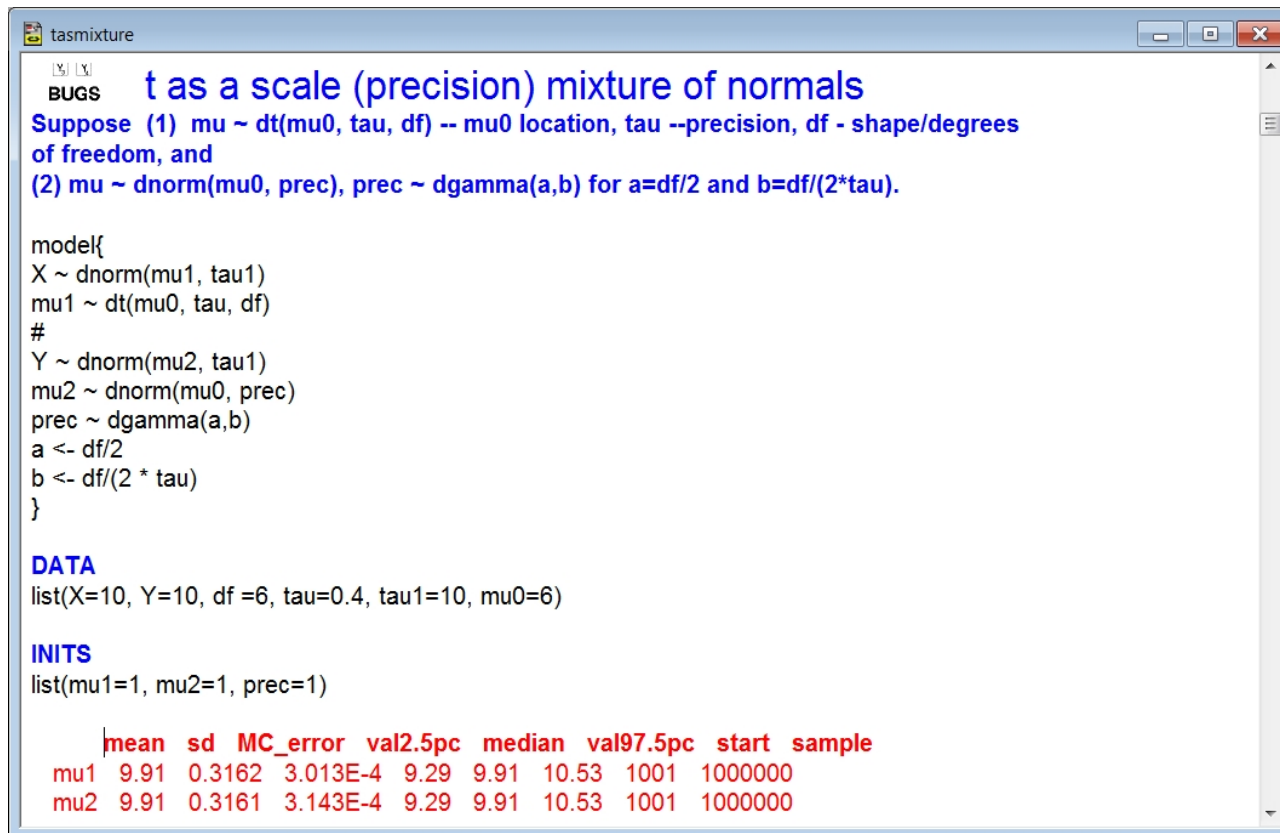
	mean	sd	MC_error	val2.5pc	median	val97.5pc	start	sample
k	3.713	9.717	0.02344	1.0	2.0	17.0	1001	1100000
p	0.5536	0.1546	1.527E-4	0.2488	0.5523	0.8514	1001	1100000
r	0.474	0.2259	4.038E-4	0.07948	0.4682	0.8965	1001	1100000

Priors as (hidden) Mixtures



Priors as Mixtures

t_n – prior is a scale mixture of normals



```
tasmixture
BUGS t as a scale (precision) mixture of normals
Suppose (1) mu ~ dt(mu0, tau, df) -- mu0 location, tau --precision, df - shape/degrees
of freedom, and
(2) mu ~ dnorm(mu0, prec), prec ~ dgamma(a,b) for a=df/2 and b=df/(2*tau).

model{
X ~ dnorm(mu1, tau1)
mu1 ~ dt(mu0, tau, df)
#
Y ~ dnorm(mu2, tau1)
mu2 ~ dnorm(mu0, prec)
prec ~ dgamma(a,b)
a <- df/2
b <- df/(2 * tau)
}

DATA
list(X=10, Y=10, df =6, tau=0.4, tau1=10, mu0=6)

INITS
list(mu1=1, mu2=1, prec=1)

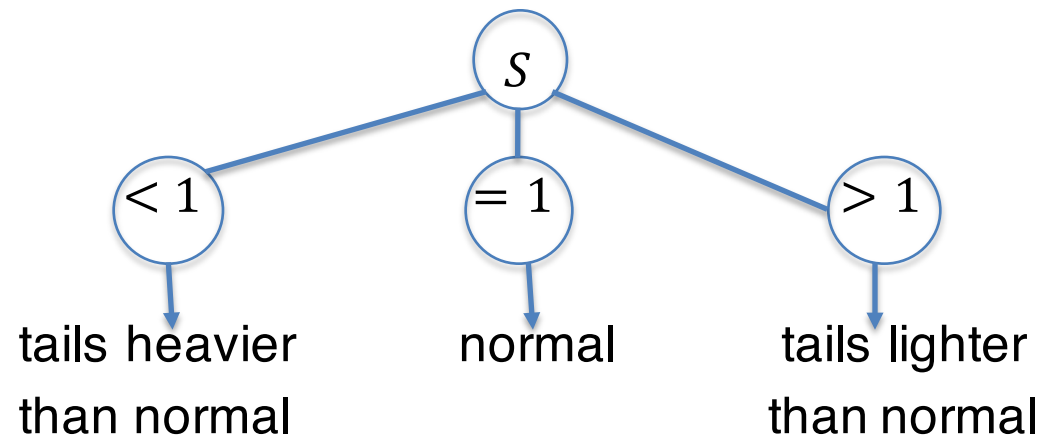
      mean sd MC_error val2.5pc median val97.5pc start sample
mu1  9.91 0.3162 3.013E-4  9.29  9.91  10.53  1001 1000000
mu2  9.91 0.3161 3.143E-4  9.29  9.91  10.53  1001 1000000
```

- **Example:**

Any symmetric unimodal distribution is a scale mixture of uniforms

$$y|\mu, \delta^2 \sim N(\mu, \delta^2) \Leftrightarrow \begin{cases} y|\mu, \delta^2, d \sim U(\mu - \sqrt{\delta^2 d}, \mu + \sqrt{\delta^2 d}) \\ d \sim \text{Ga}\left(\frac{3}{2}, \frac{1}{2}\right) \end{cases}$$

$$d \sim \text{Ga}\left(\frac{3}{2}, \frac{S}{2}\right)$$



Jeremy's IQ

Suppose, as before $X \sim N(\theta, \delta^2)$

$\theta \sim N(\mu, \tau^2), \delta^2, \tau^2, \mu$ **known**

$X = 98; \delta^2 = 80, \tau^2 = 120, \mu = 110$

$\Rightarrow \hat{\theta} = 102.8$

$\theta \sim U(\mu - \sqrt{\tau^2 d}, \mu + \sqrt{\tau^2 d})$

$d \sim \text{Ga}\left(\frac{3}{2}, \frac{s}{2}\right)$

$s = 1, \theta \sim N(\mu, \tau^2), \hat{\theta} = 102.8$


$s < 1, \theta \sim$ **heavy tailed**, $s = \frac{1}{2}, \hat{\theta} = 101.0$

$s > 1, \theta \sim$ **light tailed**, $s = 2, \hat{\theta} = 104.9$

OpenBUGS

File Edit Attributes Tools Info Model Inference Doodle Map PKBugs PKHelp Text Window Examples Manuals Help

Jeremymus



Jeremy with Normal Prior as a Scale Mixture of Uniforms

Any Symmetric Unimodal distribution is a Scale Mixture of Uniforms

$\theta | \mu, \tau^2 \sim N(\mu, \tau^2)$
 is the same as
 $\theta | \mu, \tau^2, d \sim U(\mu - \tau \cdot d^{0.5}, \mu + \tau \cdot d^{0.5})$
 $d \sim \text{Ga}(3/2, 1/2)$
 If
 $d \sim \text{Ga}(3/2, s/2)$, $s < 1$ prior is heavy tails,
 for $s > 1$ light tails. Heavy/light compared to normal.

Jeremy: $y=98$, $\sigma^2=80$, $\tau^2=120$, $\mu=110$

```

model{
  y ~ dnorm(theta, precy) #precy = 1/sigma^2 = 1/80
  lb <- mu - sqrt(d/precth) #precth = 1/tau^2 = 1/120
  ub <- mu + sqrt(d/precth)
  theta ~ dunif(lb, ub)
  #s <- 1      #normal
  #s <- 1/2    #tails heavier than normal
  s <- 2       #tails lighter than normal

  beta <- s/2
  d ~ dgamma(1.5, beta)
}
  
```

OpenBUGS version 3.2.3 rev 1012

Jeremymus

DATA

```
list(y = 98, precy = 0.0125, precth = 0.0083333, mu = 110)
```

INITS

```
list(theta=100, d=2)
```

s=1

	mean	sd	MC_error	val2.5pc	median	val97.5pc	start	sample
d	2.828	2.233	0.008332	0.2259	2.276	8.576	1001	300000
theta	102.8	6.928	0.01805	89.27	102.8	116.4	1001	300000

s = 1/2 (heavy tail)

	mean	sd	MC_error	val2.5pc	median	val97.5pc	start	sample
d	5.163	4.196	0.01406	0.4123	4.074	16.15	1001	300000
theta	101.0	7.757	0.01756	85.81	101.0	116.2	1001	300000

s=2 (light tail)

	mean	sd	MC_error	val2.5pc	median	val97.5pc	start	sample
d	1.509	1.191	0.004587	0.117	1.217	4.571	1001	300000
theta	104.9	5.876	0.01513	93.37	104.8	116.4	1001	300000

Meta-analysis via Hierarchical Models

Rats Example



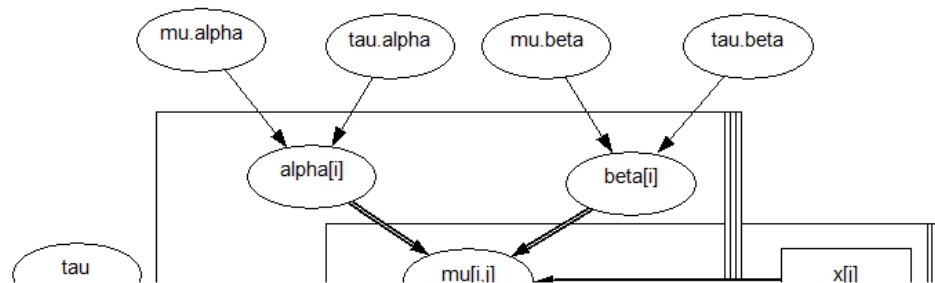
GTx

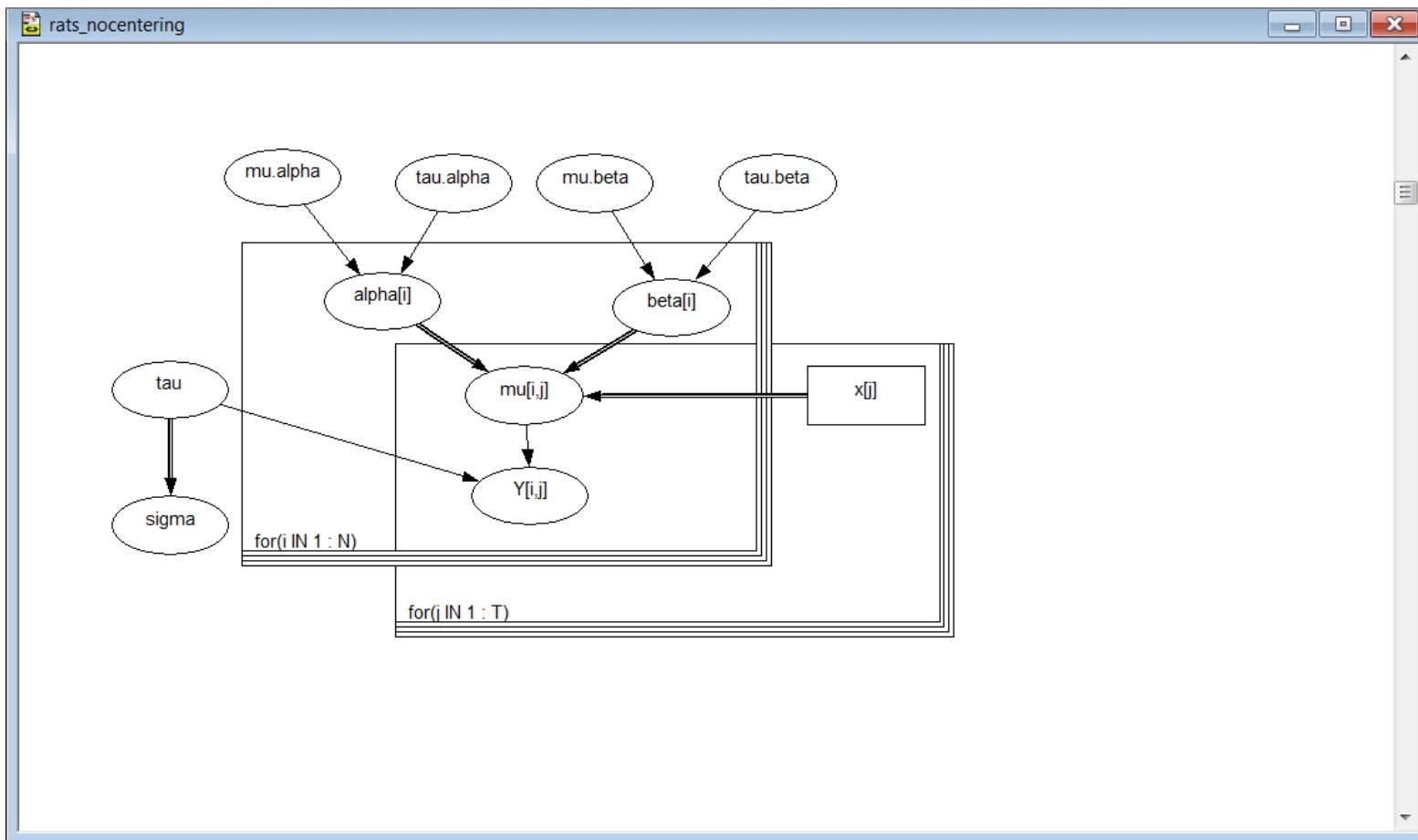
rats_nocentering

RATS

This example is taken from Gelfand *et al* (1990), and concerns 30 young rats whose weights were measured weekly for five weeks. Part of the data is shown below, where Y_{ij} is the weight of the i th rat measured at age x_j .

	Weights Y_{ij} of rat i on day x_j				
	$x_j = 8$	15	22	29	36
Rat 1	151	199	246	283	320
Rat 2	145	199	249	293	354
.....					
Rat 30	153	200	244	286	324







```
list(alpha = c(100, 100, 100, 100, 100, 100, 100, 100, 100, 100, 100, 100, 100, 100, 100,  
              100, 100, 100, 100, 100, 100, 100, 100, 100, 100, 100, 100, 100, 100),  
      beta = c(6, 6, 6, 6, 6, 6, 6, 6, 6, 6, 6, 6, 6, 6, 6, 6, 6, 6, 6, 6, 6, 6, 6, 6, 6, 6),  
      mu.alpha = 100, mu.beta = 6, tau = 1, tau.alpha = 1, tau.beta = 1)
```

```
rats_nocentering  
⇒  
list(x = c(8.0, 15.0, 22.0, 29.0, 36.0), N = 30, T = 5,  
      Y = structure(  
        .Data = c(151, 199, 246, 283, 320,  
                  145, 199, 249, 293, 354,  
                  147, 214, 263, 312, 328,  
                  155, 200, 237, 272, 297,  
                  135, 188, 230, 280, 323,  
                  159, 210, 252, 298, 331,  
                  141, 189, 231, 275, 305,  
                  159, 201, 248, 297, 338,  
                  177, 236, 285, 350, 376,  
                  134, 182, 220, 260, 296,  
                  160, 208, 261, 313, 352,  
                  143, 188, 220, 273, 314,  
                  154, 200, 244, 289, 325,  
                  171, 221, 270, 326, 358,  
                  163, 216, 242, 281, 312,  
                  160, 207, 248, 288, 324,  
                  142, 187, 234, 280, 316,  
                  156, 203, 243, 283, 317,  
                  157, 212, 259, 307, 336,  
                  152, 203, 246, 286, 321,  
                  154, 205, 253, 298, 334,  
                  139, 190, 225, 267, 302,  
                  146, 191, 229, 272, 302,  
                  157, 211, 250, 285, 323,  
                  132, 185, 237, 286, 331,  
                  160, 207, 257, 303, 345,  
                  169 216 261 295 333
```

```

139, 190, 225, 267, 302,
146, 191, 229, 272, 302,
157, 211, 250, 285, 323,
132, 185, 237, 286, 331,
160, 207, 257, 303, 345,
169, 216, 261, 295, 333,
157, 205, 248, 289, 316,
137, 180, 219, 258, 291,
153, 200, 244, 286, 324),
.Dim = c(30,5)))

```

↔

INITS

```

list(alpha = c(100, 100, 100, 100, 100, 100, 100, 100, 100, 100, 100, 100, 100, 100, 100,
100, 100, 100, 100, 100, 100, 100, 100, 100, 100, 100, 100, 100, 100, 100, 100),
beta = c(6, 6, 6, 6, 6, 6, 6, 6, 6, 6, 6, 6, 6, 6, 6, 6, 6, 6, 6, 6, 6, 6, 6, 6, 6, 6, 6, 6, 6, 6),
mu.alpha = 100, mu.beta = 6, tau = 1, tau.alpha = 1, tau.beta = 1)

```

	mean	sd	MC_error	val2.5pc	median	val97.5pc	start	sample
mu.alpha	106.6	2.308	0.01523	102.0	106.6	111.1	1001	100000
mu.beta	6.185	0.1056	6.625E-4	5.977	6.185	6.392	1001	100000
sigma	6.15	0.4677	0.003131	5.314	6.121	7.148	1001	100000
tau	0.0269	0.004047	2.674E-5	0.01957	0.02669	0.03541	1001	100000
tau.alpha	0.01013	0.00406	3.91E-5	0.004623	0.009401	0.01995	1001	100000
tau.beta	4.328	1.495	0.01078	2.104	4.101	7.882	1001	100000