

$$1) \tilde{\beta} = (H+D)y = Hy + Dy = \beta^* + Dy$$

$$E(\tilde{\beta}) = E(\beta^*) + E(Dy) = \beta + E(Dy)$$

$$\begin{aligned} \text{Var}(\tilde{\beta}) &= \text{Var}(\beta^* + Dy) = \text{Var}(\beta^*) + D \text{Var}(y) D^T \\ &= \text{Var}(\beta^*) + D(y y^T) D^T \geq \text{Var}(\beta^*) \end{aligned}$$

$$\text{Var}(\tilde{\beta}) > \text{Var}(\beta^*) \quad \forall D \neq 0$$

The assumption made is that $y = X\beta + \varepsilon$, $\varepsilon \sim N(0, \sigma^2)$

$$2.1) \beta^*_{\text{ridge}} = (X^T X + \lambda I)^{-1} X^T y \Rightarrow$$

$$E(\beta^*_{\text{ridge}}) = E((X^T X + \lambda I)^{-1} X^T y) = (X^T X + \lambda I)^{-1} X^T E(y)$$

$$y = X\beta + \varepsilon \Rightarrow E(\beta^*_{\text{ridge}}) = (X^T X + \lambda I)^{-1} X^T (E(X\beta) + \underbrace{E(\varepsilon)}_{0})$$

$$E(\beta^*_{\text{ridge}}) = (X^T X + \lambda I)^{-1} X^T X \beta \neq \beta \quad (\text{unless } \lambda = 0)$$

$$2.2) X = U D V^T \text{ and } X^T = V D U^T$$

$$\begin{aligned} \beta^*_{\text{ridge}} &= (V D U^T U D V^T + \lambda I)^{-1} V D U^T y = (V D D V^T + \lambda I)^{-1} V D U^T y \\ &= (V D D V^T + V \lambda I V^T)^{-1} V D U^T y = V (D D + \lambda I)^{-1} V^T V D U^T y \\ &= V (D D + \lambda I)^{-1} D U^T y \end{aligned}$$

It is usually much easier to find $(D D + \lambda I)^{-1}$ than $(X^T X + \lambda I)^{-1}$.

Therefore, it is worth to do this decomposition when the computational cost of this operation is lower than the one of finding $(X^T X + \lambda I)^{-1}$. An example of this is when we are testing multiple λ .

$$2.3) \text{Var}((X^T X + \lambda I)^{-1} X^T Y)$$

$$= (X^T X + \lambda I)^{-1} X^T \text{Var}(Y) X (X^T X + \lambda I)^{-1} \quad \text{but } Y = X\beta + \varepsilon \Rightarrow \text{Var}(Y) = \text{Var}(\varepsilon) = \sigma^2$$

$$= \sigma^2 (X^T X + \lambda I)^{-1} X^T X (X^T X + \lambda I)^{-1}$$

$$\text{Var}(\beta_{OLS}^*) = \sigma^2 (X^T X)^{-1} \cdot X^T X \cdot (X^T X)^{-1}$$

$$\lambda \geq 0 \Rightarrow (X^T X + \lambda I)^{-1} \leq (X^T X)^{-1}$$

↓

$$\sigma^2 (X^T X + \lambda I)^{-1} X^T X (X^T X + \lambda I)^{-1} \leq \sigma^2 (X^T X)^{-1} X^T X (X^T X)^{-1}$$

$$\text{Var}(\beta_{ridge}^*) \leq \text{Var}(\beta_{OLS}^*)$$

$$2.4) (X^T X + \lambda_1 I)^{-1} X^T X (X^T X + \lambda_1 I)^{-1} < (X^T X + \lambda_2 I)^{-1} X^T X (X^T X + \lambda_2 I)^{-1}$$

for $\lambda_1 > \lambda_2 \Rightarrow$ when λ increases the variance decreases.

$$b(\beta_{ridge}^*) = E(\beta_{ridge}^*) - \beta = (X^T X + \lambda I)^{-1} X^T X \beta - \beta$$

$$= ((X^T X + \lambda I)^{-1} X^T X - I) \beta \quad \lambda = 0 \Rightarrow b(\beta_{ridge}^*) = 0$$

as λ increases, the bias increase in absolute value.

$$2.5) \beta_{ridge}^* = (X^T X + \lambda I)^{-1} X^T Y = (1 + \lambda)^{-1} I X^T Y = \frac{X^T Y}{(1 + \lambda)}$$

$$\beta_{OLS}^* = (X^T X)^{-1} X^T Y = X^T Y$$

$$\beta_{ridge}^* = \frac{\beta_{OLS}^*}{(1 + \lambda)}$$

$$3) f(\beta) = (Y - X\beta)^T (Y - X\beta) + \lambda_2 \|\beta\|_2^2 + \lambda_1 \|\beta\|_1$$

$$= Y^T Y - 2Y^T X\beta + \beta^T \cancel{X^T X}^I \beta + \lambda_2 \beta^T \beta + \lambda_1 \|\beta\|_1$$

$$= Y^T Y - 2Y^T X\beta + (1 + \lambda_2) \beta^T \beta + \lambda_1 \|\beta\|_1$$

$$2f(\beta) = -2(X^T X)^T + 2(1 + \lambda_2)\beta + \lambda_1 2\|\beta\|_1$$

$$\arg \min_{\beta} f(\beta) \Rightarrow 0 = -2X^T Y + 2(1 + \lambda_2)\beta + \lambda_1 2\|\beta\|_1$$

$$2\|\beta_j\|_1 \begin{cases} [-1, 1], & \beta_j = 0 \\ 1, & \beta_j > 0 \\ -1, & \beta_j < 0 \end{cases} \quad \beta^* \in \frac{2X^T Y - 2\|\beta\|_1}{2(1 + \lambda_2)}$$

for each β_j in β

$$\beta_j^* \in \begin{cases} \left\{ \frac{X^T Y - \frac{\lambda_1}{2}}{1 + \lambda_2} \right\}, & \beta_j^* > 0 \\ \left\{ \frac{X^T Y + \frac{\lambda_1}{2}}{1 + \lambda_2} \right\}, & \beta_j^* < 0 \\ \left[\frac{X^T Y - \frac{\lambda_1}{2}}{1 + \lambda_2}, \frac{X^T Y + \frac{\lambda_1}{2}}{1 + \lambda_2} \right], & \beta_j^* = 0 \end{cases}$$

$$\beta_{OLS}^* = (\cancel{X^T X}^I)^{-1} X^T Y = X^T Y$$

$$\Rightarrow \beta_j^* = \begin{cases} \frac{\beta_{OLSj}^* + \frac{\lambda_1}{2}}{1 + \lambda_2}, & \beta_{OLSj}^* \neq \frac{\lambda_1}{2} \\ \left[0, \frac{\lambda_1}{1 + \lambda_2} \right], & \beta_{OLSj}^* = \frac{\lambda_1}{2} \end{cases}$$