

# Analyses for the paper: Measuring emotions during learning

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19 August 2019

## 1. PREPARATIONS

```
# Load packages
library(tidyverse)
library(rstan)
library(brms)
library(bayesplot)
library(ggmcmc) # for ggs posterior plot

# set rstan options
rstan::rstan_options(auto_write = T)
options(mc.cores = parallel::detectCores())
```

## Selection of statistical methods

- Bayesian statistics, as they are more intuitive to interpret.
- Mixed models, because events (level 1) are nested within participants (level 2) and texts (level 2). However, only random intercepts are included, as additional random effects would be hardly identifiable.
- Treat emotional self-reports as ordinal outcome variables. Cummulative family in brms, as we understand the Likert-scales as the categorization of a latent continuous construct (Buerkner & Vuorre, 2019: <https://journals.sagepub.com/doi/full/10.1177/2515245918823199> (<https://journals.sagepub.com/doi/full/10.1177/2515245918823199>)).
- Assumption of equal variances: “If unequal variances are theoretically possible – and they usually are – we also recommend incorporating them into the model” (Buerkner & Vuorre, 2019). However, models allowing for unequal variances did not converge and were therefore omitted.

## Choosing link function:

link-distributions available for cummulative models in brms (usually only minor impact on results): logit = logistic, probit = gaussian, cloglog = extreme value distribution <http://bayesium.com/which-link-function-logit-probit-or-cloglog/> (<http://bayesium.com/which-link-function-logit-probit-or-cloglog/>)

The choice should be made based on some combination of: - Knowledge of the response distribution, - Theoretical considerations, and - Empirical fit to the data.

<https://stats.stackexchange.com/questions/20523/difference-between-logit-and-probit-models> (<https://stats.stackexchange.com/questions/20523/difference-between-logit-and-probit-models>)

## interpretation of summary of fitted model

- Estimate is the mean of the posterior distribution, and corresponds to the frequentist point estimate
- Est.Error is the standard deviation of the posterior distribution

- thresholds in ordinal models are called “intercepts” in the output
- Visualisation of marginal effects for ordinal models: <https://github.com/paul-buerkner/brms/issues/190> (<https://github.com/paul-buerkner/brms/issues/190>)

## posterior predictive checks (not included in this document)

for ordinal models pp\_check not adequate: <https://github.com/stan-dev/bayesplot/issues/73> (<https://github.com/stan-dev/bayesplot/issues/73>) → use ppc: <https://mc-stan.org/bayesplot/articles/graphical-ppcs.html> (<https://mc-stan.org/bayesplot/articles/graphical-ppcs.html>)

## load data

```
# Load data
df <- read_csv("df_TEEM_final.csv")

# rename some variables
df <- df %>% rename("participant" = "subject_nr", "text" = "text_pic", "valence_post" = "SAM_LIKERT_POST")
```

## standardize predictors (aggregated from FaceReader)

(helps for model convergence and for the interpretation of the interaction effects)

```
df <- df %>%
  mutate(
    mean_interest = scale(mean_interest, center = T, scale = T),
    mean_boredom = scale(mean_boredom, center = T, scale = T),
    mean_valence = scale(mean_valence, center = T, scale = T)

df <- df %>%
  mutate(
    sd_interest = scale(sd_interest, center = T, scale = T),
    sd_boredom = scale(sd_boredom, center = T, scale = T),
    sd_valence = scale(sd_valence, center = T, scale = T)

df <- df %>%
  mutate(
    peak10_interest = scale(peak10_interest, center = T, scale = T),
    peak10_boredom = scale(peak10_boredom, center = T, scale = T),
    peak10_valence_pos = scale(peak10_valence_pos, center = T, scale = T),
    peak10_valence_neg = scale(peak10_valence_neg, center = T, scale = T)
```

## 2. INTEREST

### INTEREST restricted model

```
# complete cases only (drop NAs)
dfsub <- df %>% select(participant, interested_post, mean_interest, sd_interest, peak10_inter
est, text) %>% drop_na()

# restricted model
m0i_cloglog <- brm(
  interested_post ~ 1 + (1|participant) + (1|text),
  family = cumulative("cloglog"),
  prior = prior(cauchy(0, 10), class = sd),
  iter = 4000, warmup = 2000, chains = 4, cores = 4,
  control = list(adapt_delta = 0.99),
  inits = 0,
  data = dfsub,
  save_all_pars = T) # needed for bayes factor

summary(m0i_cloglog)
```

```
## Family: cumulative
## Links: mu = cloglog; disc = identity
## Formula: interested_post ~ 1 + (1 | participant) + (1 | text)
## Data: dfsub (Number of observations: 205)
## Samples: 4 chains, each with iter = 4000; warmup = 2000; thin = 1;
## total post-warmup samples = 8000
##
## Group-Level Effects:
## ~participant (Number of levels: 103)
## Estimate Est.Error l-95% CI u-95% CI Eff.Sample Rhat
## sd(Intercept) 1.52 0.28 1.01 2.10 1663 1.00
##
## ~text (Number of levels: 6)
## Estimate Est.Error l-95% CI u-95% CI Eff.Sample Rhat
## sd(Intercept) 0.73 0.43 0.24 1.83 2457 1.00
##
## Population-Level Effects:
## Estimate Est.Error l-95% CI u-95% CI Eff.Sample Rhat
## Intercept[1] -5.53 0.81 -7.26 -4.09 3740 1.00
## Intercept[2] -3.07 0.50 -4.11 -2.16 3798 1.00
## Intercept[3] -1.56 0.43 -2.41 -0.75 4062 1.00
## Intercept[4] 0.98 0.42 0.21 1.89 3773 1.00
##
## Samples were drawn using sampling(NUTS). For each parameter, Eff.Sample
## is a crude measure of effective sample size, and Rhat is the potential
## scale reduction factor on split chains (at convergence, Rhat = 1).
```

```

# other link functions
m0i_logit <- update(m0i_cloglog,
                    family = cumulative("logit"))

m0i_probit <- update(m0i_cloglog,
                    family = cumulative("probit"))
# compare different link functions using assimilation of leave-one-out-cross validation (looic)
m0i_logit <- add_criterion(m0i_logit, "loo", reloo = T) # "reloo = T" actually calculates MCMC
for problematic observations
m0i_probit <- add_criterion(m0i_probit, "loo", reloo = T)
m0i_cloglog <- add_criterion(m0i_cloglog, "loo", reloo = T)
print(loo_compare(m0i_logit, m0i_probit, m0i_cloglog, criterion="loo"), simplify = F) # clog
Log 1.5-3.5 SD better

```

```

##          elpd_diff se_diff elpd_loo se_elpd_loo p_loo se_p_loo looic
## m0i_cloglog    0.0      0.0 -239.6    11.8      74.9    5.8   479.1
## m0i_probit    -1.9      1.4 -241.5    11.5      71.5    6.0   483.0
## m0i_logit     -2.0      1.3 -241.5    11.6      76.5    5.7   483.0
##          se_looic
## m0i_cloglog    23.5
## m0i_probit     23.0
## m0i_logit      23.1

```

```

# chosen response distribution (link function) for final restricted model
m0i <- m0i_cloglog

```

## INTEREST mean

```

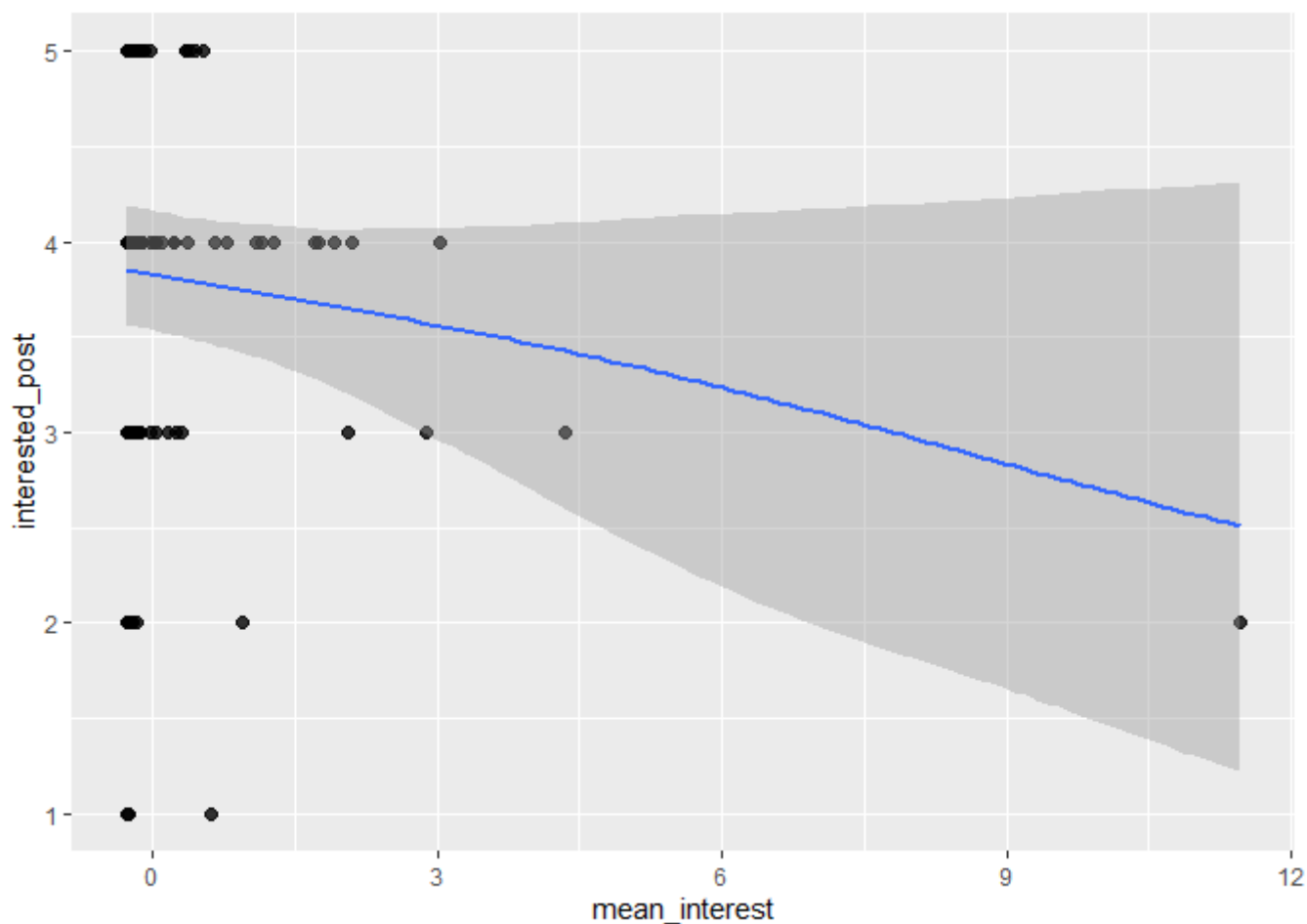
# full model including FaceReader's estimate as predictor
m1_imean <- update(m0i, formula. = ~ . + mean_interest,
                  prior = c(prior(normal(0, 10), class = b),
                             prior(cauchy(0, 10), class = sd)),
                  newdata = dfsub,
                  save_all_pars = T)

## model parameter
summary(m1_imean)

```

```
## Family: cumulative
## Links: mu = cloglog; disc = identity
## Formula: interested_post ~ (1 | participant) + (1 | text) + mean_interest
## Data: dfsub (Number of observations: 205)
## Samples: 4 chains, each with iter = 4000; warmup = 2000; thin = 1;
##           total post-warmup samples = 8000
##
## Group-Level Effects:
## ~participant (Number of levels: 103)
##           Estimate Est.Error 1-95% CI u-95% CI Eff.Sample Rhat
## sd(Intercept)      1.49      0.27   0.99   2.06      1475 1.00
##
## ~text (Number of levels: 6)
##           Estimate Est.Error 1-95% CI u-95% CI Eff.Sample Rhat
## sd(Intercept)      0.70      0.40   0.23   1.69      2715 1.00
##
## Population-Level Effects:
##           Estimate Est.Error 1-95% CI u-95% CI Eff.Sample Rhat
## Intercept[1]      -5.66      0.85  -7.53  -4.17      3859 1.00
## Intercept[2]      -3.08      0.49  -4.13  -2.18      3704 1.00
## Intercept[3]      -1.56      0.41  -2.43  -0.78      4271 1.00
## Intercept[4]       0.97      0.40   0.21   1.83      4362 1.00
## mean_interest     -0.23      0.16  -0.53   0.09      5384 1.00
##
## Samples were drawn using sampling(NUTS). For each parameter, Eff.Sample
## is a crude measure of effective sample size, and Rhat is the potential
## scale reduction factor on split chains (at convergence, Rhat = 1).
```

```
# plots
plot(marginal_effects(m1_imean, "mean_interest", categorical = F), points = T, point_args = c
(alpha = 0.8)) # shows the strong influence of two observations
```

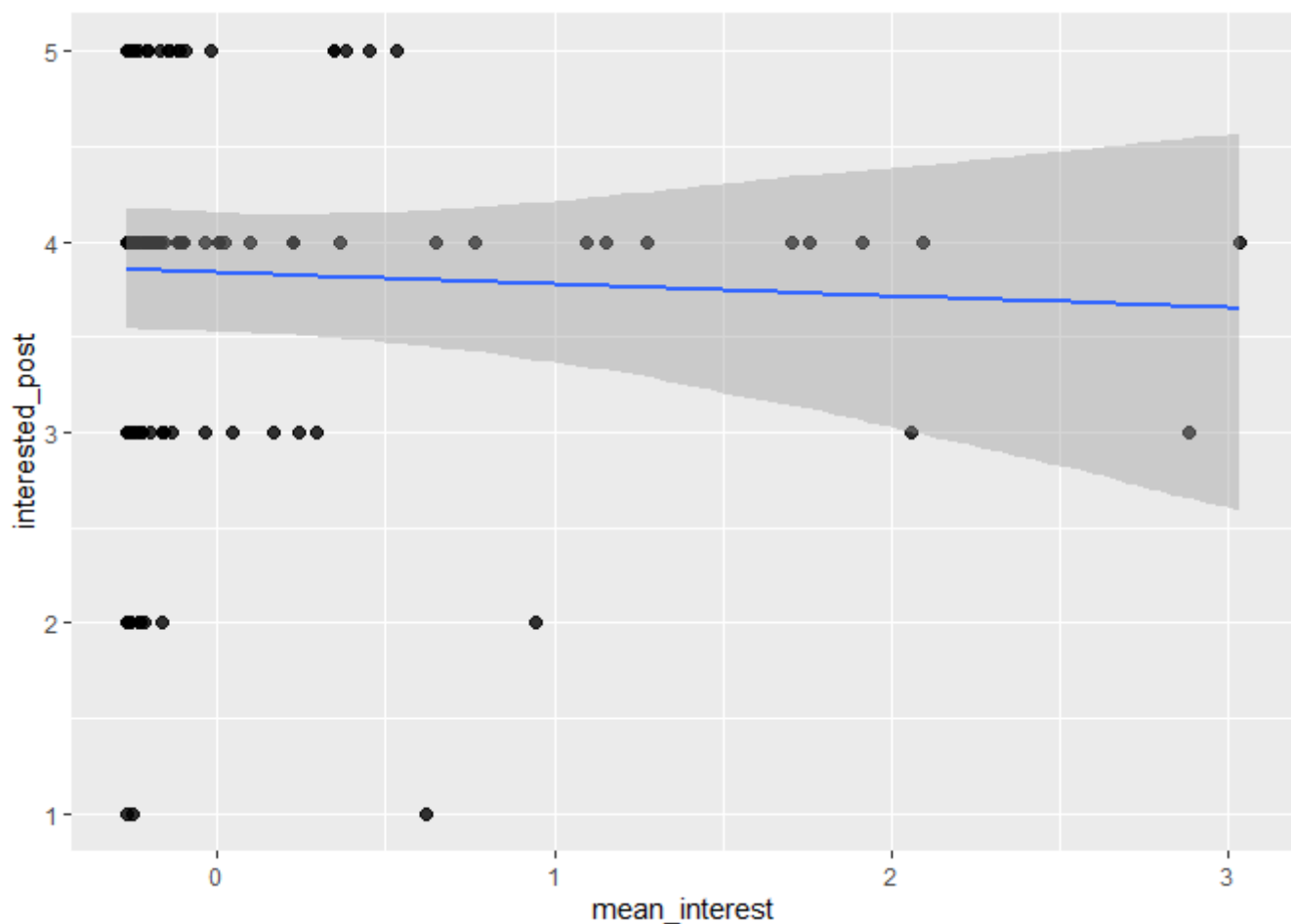


```
# Model without "outliers"
dfsub_out <- dfsub %>%
  filter(mean_interest < (mean(mean_interest) + 4*sd(mean_interest)) & mean_interest > (mean(
    mean_interest) - 4*sd(mean_interest)))
# resulting in two observations less
# dfsub %>% filter(mean_interest > (mean(mean_interest) + 4*sd(mean_interest)) | mean_interes
# t < (mean(mean_interest) - 4*sd(mean_interest))) # outliers are from one participant (highly e
# xpressive in video)

m1_imean_out <- update(m0i, formula. = ~ . + mean_interest,
  prior = c(prior(normal(0, 10), class = b),
    prior(cauchy(0, 10), class = sd)),
  newdata = dfsub_out,
  save_all_pars = T)
summary(m1_imean_out)
```

```
## Family: cumulative
## Links: mu = cloglog; disc = identity
## Formula: interested_post ~ (1 | participant) + (1 | text) + mean_interest
## Data: dfsub_out (Number of observations: 203)
## Samples: 4 chains, each with iter = 4000; warmup = 2000; thin = 1;
##           total post-warmup samples = 8000
##
## Group-Level Effects:
## ~participant (Number of levels: 102)
##           Estimate Est.Error 1-95% CI u-95% CI Eff.Sample Rhat
## sd(Intercept)      1.52      0.27   1.00   2.08      1748 1.00
##
## ~text (Number of levels: 6)
##           Estimate Est.Error 1-95% CI u-95% CI Eff.Sample Rhat
## sd(Intercept)      0.76      0.48   0.25   1.92      2629 1.00
##
## Population-Level Effects:
##           Estimate Est.Error 1-95% CI u-95% CI Eff.Sample Rhat
## Intercept[1]      -5.48      0.80  -7.22  -4.04      3600 1.00
## Intercept[2]      -3.09      0.50  -4.13  -2.18      3813 1.00
## Intercept[3]      -1.58      0.43  -2.45  -0.77      3964 1.00
## Intercept[4]       0.98      0.43   0.22   1.88      3946 1.00
## mean_interest     -0.15      0.32  -0.77   0.50      5419 1.00
##
## Samples were drawn using sampling(NUTS). For each parameter, Eff.Sample
## is a crude measure of effective sample size, and Rhat is the potential
## scale reduction factor on split chains (at convergence, Rhat = 1).
```

```
## plots
plot(marginal_effects(m1_imean_out, "mean_interest", categorical = F), points = T, point_args
= c(alpha = 0.8))
```



```
# choose final model
m1_imean <- m1_imean_out
```

## INTEREST mean\*SD

```
# the participant from before also outlier in SD? --> No.
#dfsub_out %>% filter(sd_interest > (mean(sd_interest) + 4*sd(sd_interest)) | sd_interest <
#  (mean(sd_interest) - 4*sd(sd_interest)))

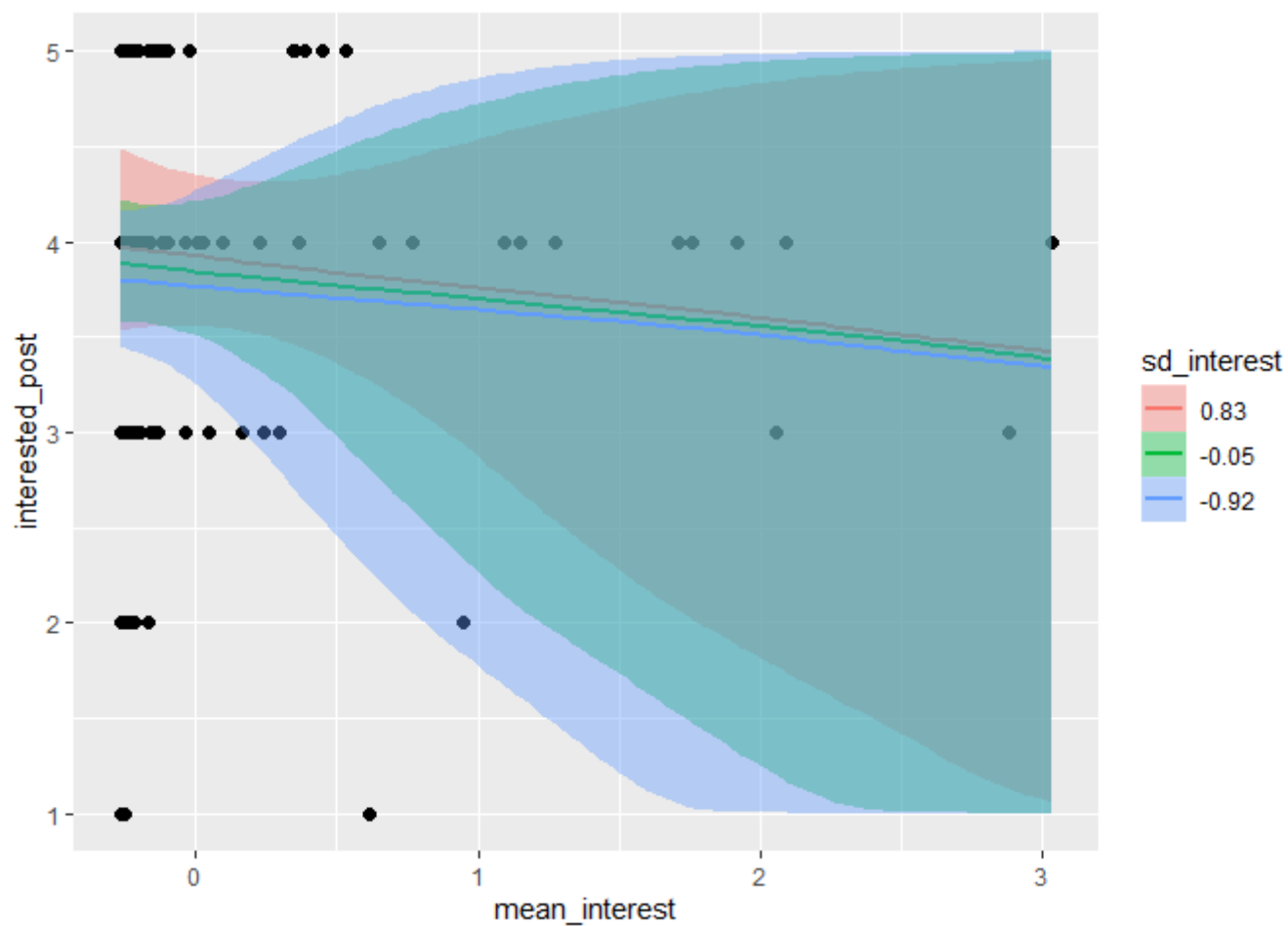
# full model including FaceReader's estimates as predictor
m1_imeanxsd <- update(m0i, formula. = ~ . + mean_interest*sd_interest,
  prior = c(prior(normal(0, 10), class = b),
    prior(cauchy(0, 10), class = sd)),
  newdata = dfsub_out, # without outliers of mean interest
  save_all_pars = T)

## model indicator
summary(m1_imeanxsd)
```

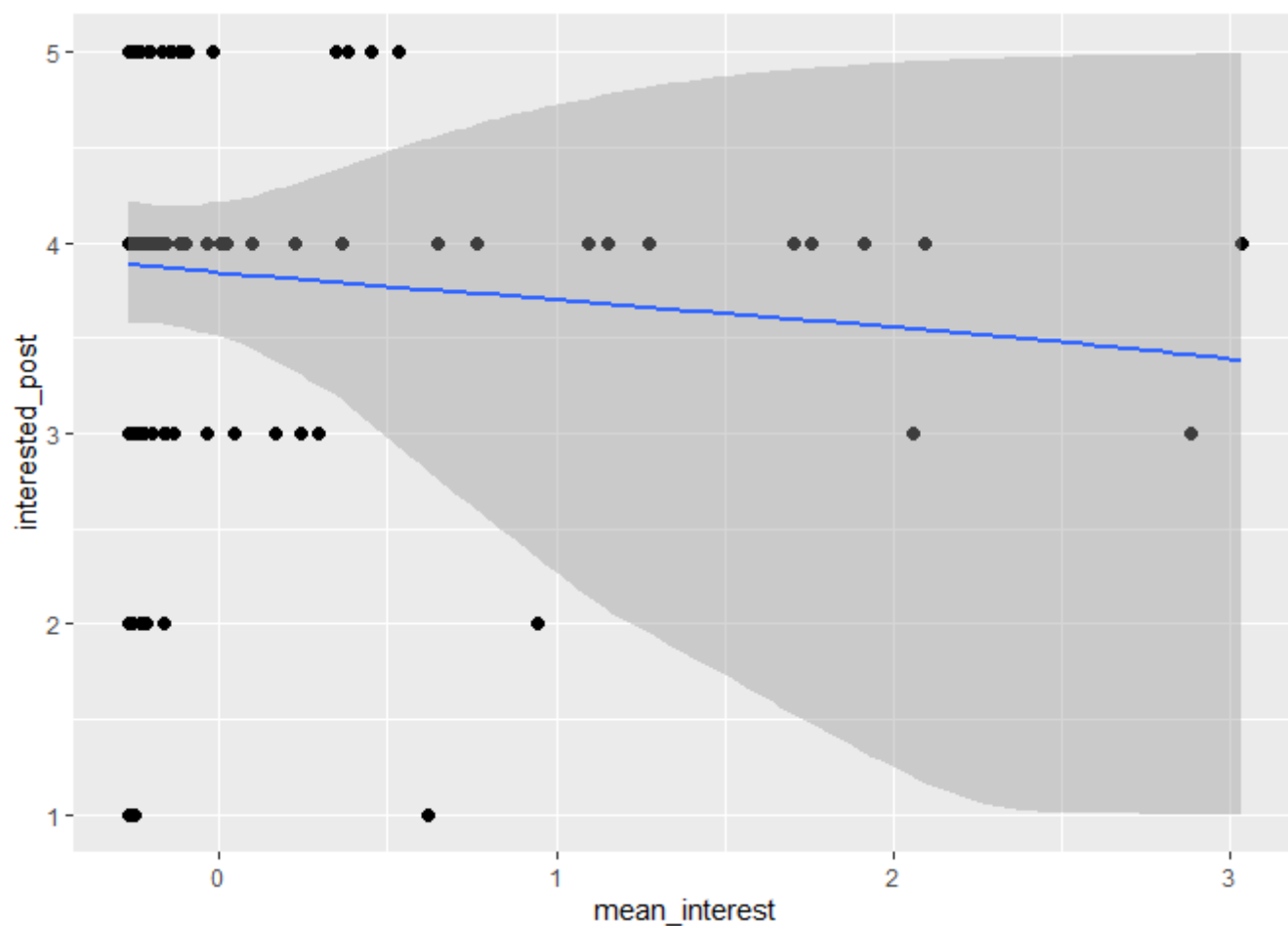


```
## Family: cumulative
## Links: mu = cloglog; disc = identity
## Formula: interested_post ~ (1 | participant) + (1 | text) + mean_interest + sd_interest +
mean_interest:sd_interest
## Data: dfsub_out (Number of observations: 203)
## Samples: 4 chains, each with iter = 4000; warmup = 2000; thin = 1;
## total post-warmup samples = 8000
##
## Group-Level Effects:
## ~participant (Number of levels: 102)
## Estimate Est.Error 1-95% CI u-95% CI Eff.Sample Rhat
## sd(Intercept) 1.54 0.28 1.02 2.14 1702 1.00
##
## ~text (Number of levels: 6)
## Estimate Est.Error 1-95% CI u-95% CI Eff.Sample Rhat
## sd(Intercept) 0.74 0.42 0.25 1.83 2771 1.00
##
## Population-Level Effects:
## Estimate Est.Error 1-95% CI u-95% CI Eff.Sample
## Intercept[1] -5.56 0.84 -7.35 -4.09 3422
## Intercept[2] -3.15 0.54 -4.29 -2.16 3661
## Intercept[3] -1.62 0.47 -2.58 -0.75 3718
## Intercept[4] 0.98 0.46 0.11 1.95 2842
## mean_interest -0.41 1.12 -2.70 1.72 3288
## sd_interest 0.24 0.38 -0.50 1.01 3714
## mean_interest:sd_interest -0.04 0.33 -0.65 0.66 3842
## Rhat
## Intercept[1] 1.00
## Intercept[2] 1.00
## Intercept[3] 1.00
## Intercept[4] 1.00
## mean_interest 1.00
## sd_interest 1.00
## mean_interest:sd_interest 1.00
##
## Samples were drawn using sampling(NUTS). For each parameter, Eff.Sample
## is a crude measure of effective sample size, and Rhat is the potential
## scale reduction factor on split chains (at convergence, Rhat = 1).
```

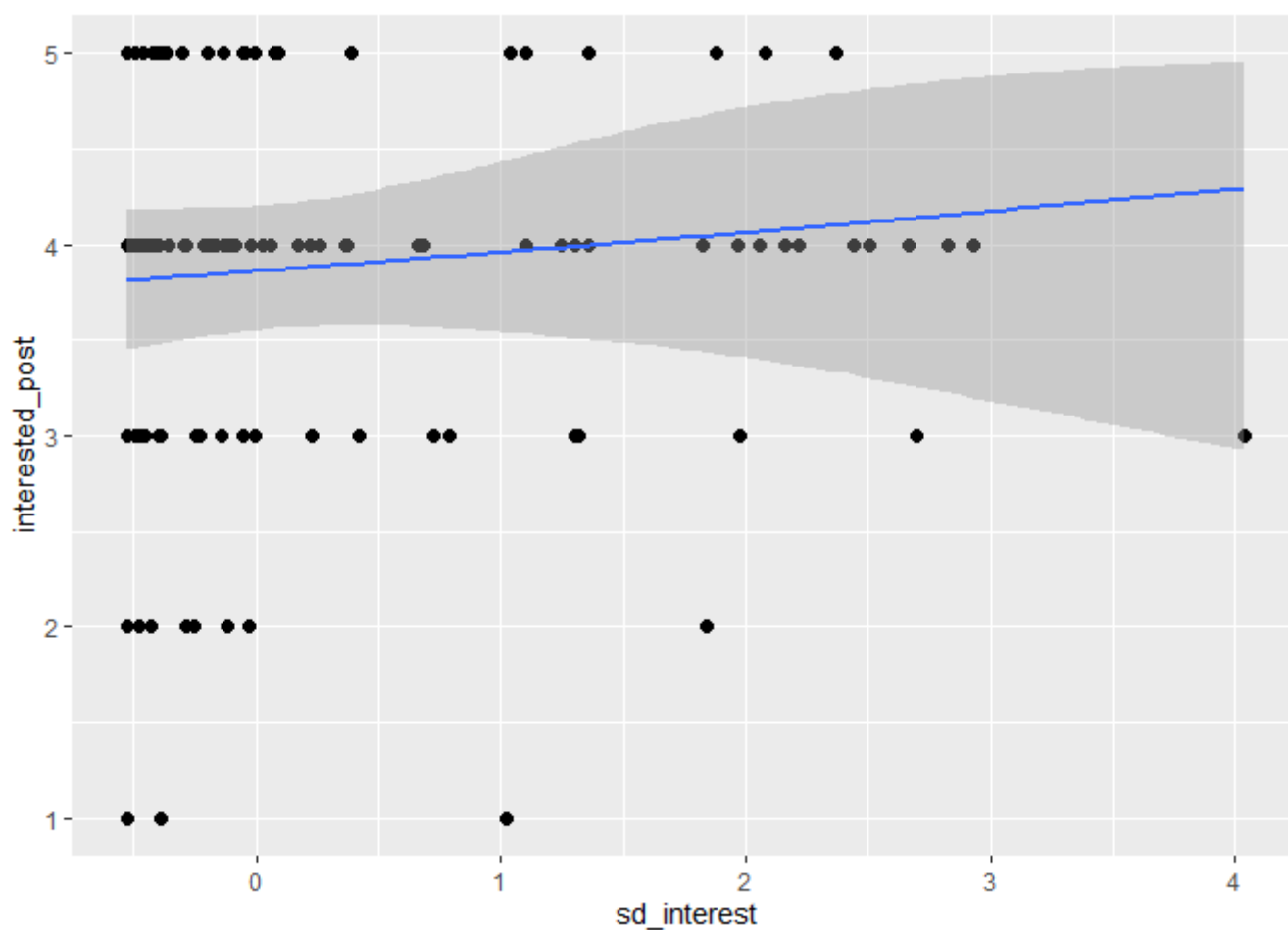
```
# plots
plot(marginal_effects(m1_imeanxsd,"mean_interest:sd_interest", categorical = F), points = T)
```



```
plot(marginal_effects(m1_imeanxsd,"mean_interest", categorical = F), points = T)
```



```
plot(marginal_effects(m1_imeanxsd,"sd_interest", categorical = F), points = T)
```



## INTEREST mean of peaks

```
# remove outliers also in peak when from the same participant as in mean:
dfsub_outpeak <- dfsub %>% filter(peak10_interest < (mean(peak10_interest) + 4*sd(peak10_interest)) & peak10_interest > (mean(peak10_interest) - 4*sd(peak10_interest)))
# dfsub %>% filter(peak10_interest > (mean(peak10_interest) + 4*sd(peak10_interest)) | peak10_interest < (mean(peak10_interest) - 4*sd(peak10_interest))) # 2 outliers are from the same participant as before

# model including outliers
m1_ipeak <- update(m0i, formula. = ~ . + peak10_interest,
  prior = c(prior(normal(0, 10), class = b),
    prior(cauchy(0, 10), class = sd)),
  newdata = dfsub,
  save_all_pars = T)

summary(m1_ipeak)
```

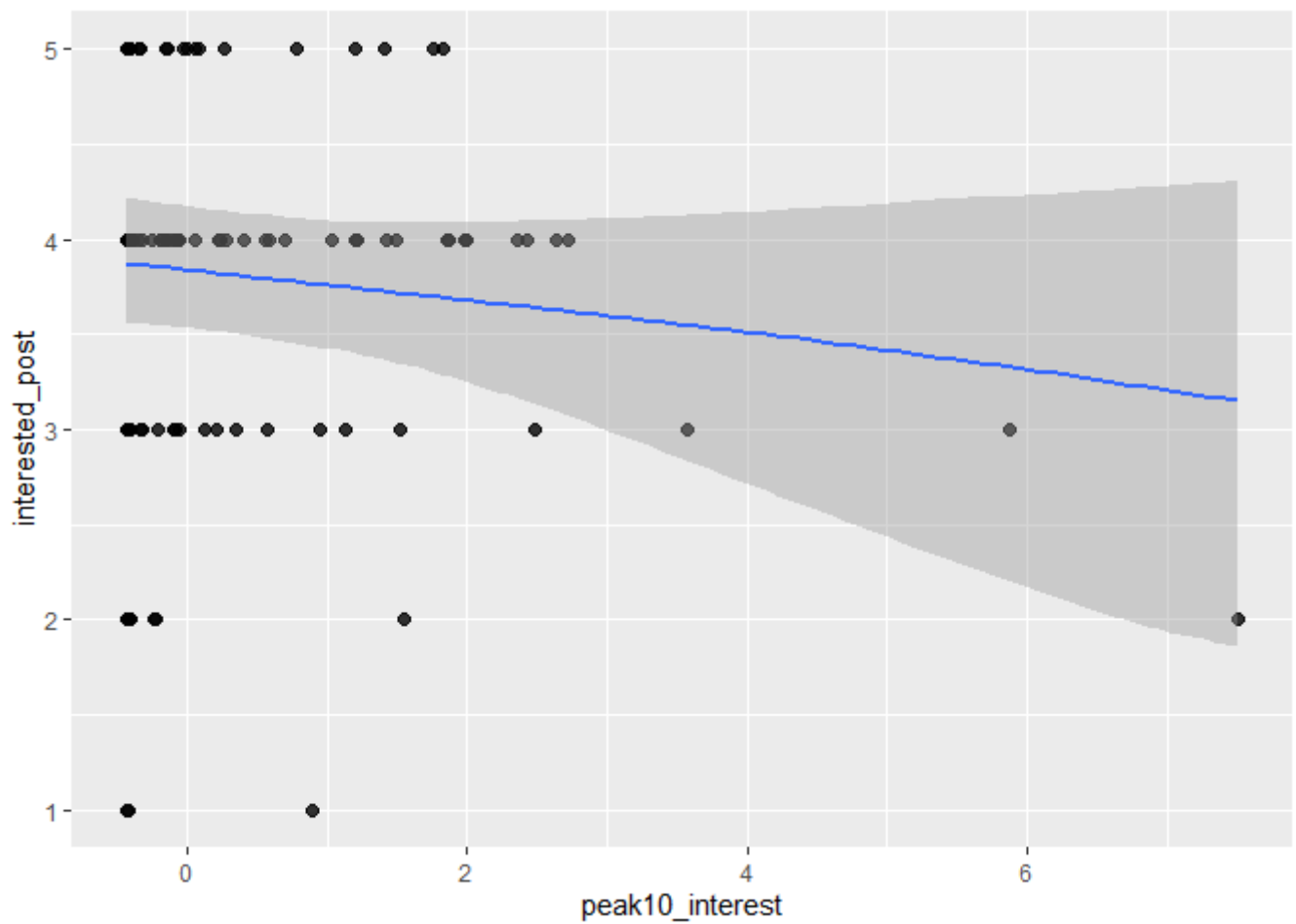
```
## Family: cumulative
## Links: mu = cloglog; disc = identity
## Formula: interested_post ~ (1 | participant) + (1 | text) + peak10_interest
## Data: dfsub (Number of observations: 205)
## Samples: 4 chains, each with iter = 4000; warmup = 2000; thin = 1;
##           total post-warmup samples = 8000
##
## Group-Level Effects:
## ~participant (Number of levels: 103)
##           Estimate Est.Error 1-95% CI u-95% CI Eff.Sample Rhat
## sd(Intercept)      1.54      0.27   1.03   2.12      1773 1.00
##
## ~text (Number of levels: 6)
##           Estimate Est.Error 1-95% CI u-95% CI Eff.Sample Rhat
## sd(Intercept)      0.74      0.45   0.25   1.83      2392 1.00
##
## Population-Level Effects:
##           Estimate Est.Error 1-95% CI u-95% CI Eff.Sample Rhat
## Intercept[1]       -5.60      0.83   -7.39   -4.13      4087 1.00
## Intercept[2]       -3.11      0.51   -4.19   -2.18      3378 1.00
## Intercept[3]       -1.58      0.43   -2.48   -0.77      3579 1.00
## Intercept[4]        0.98      0.42    0.21    1.85      3178 1.00
## peak10_interest    -0.21      0.17   -0.55    0.13      3982 1.00
##
## Samples were drawn using sampling(NUTS). For each parameter, Eff.Sample
## is a crude measure of effective sample size, and Rhat is the potential
## scale reduction factor on split chains (at convergence, Rhat = 1).
```

```
# model without outliers
m1_ipeak_out <- update(m0i, formula. = ~ . + peak10_interest,
                      prior = c(prior(normal(0, 10), class = b),
                                prior(cauchy(0, 10), class = sd)),
                      newdata = dfsub_outpeak,
                      save_all_pars = T)

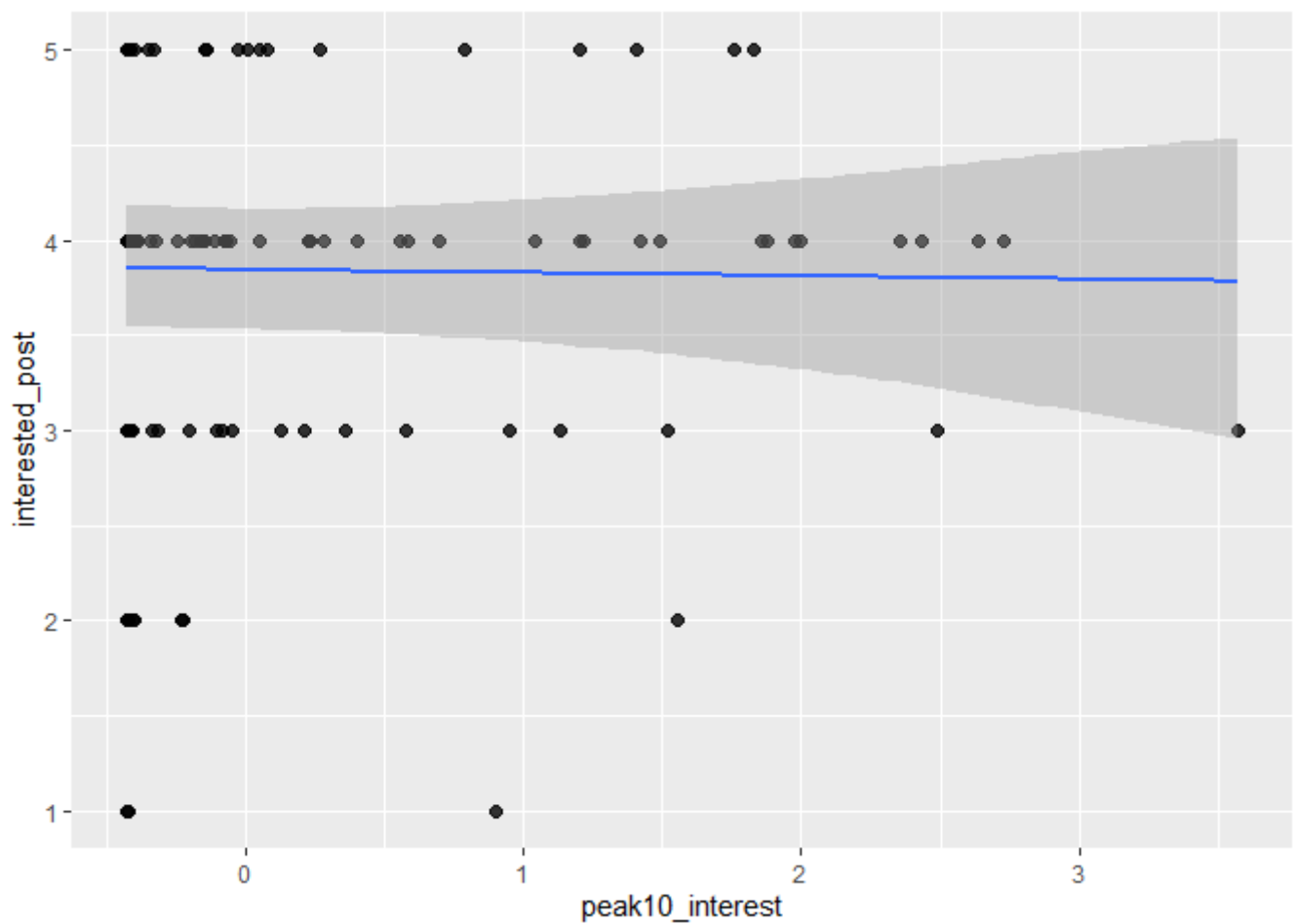
## model parameter
summary(m1_ipeak_out)
```

```
## Family: cumulative
## Links: mu = cloglog; disc = identity
## Formula: interested_post ~ (1 | participant) + (1 | text) + peak10_interest
## Data: dfsub_outpeak (Number of observations: 203)
## Samples: 4 chains, each with iter = 4000; warmup = 2000; thin = 1;
##           total post-warmup samples = 8000
##
## Group-Level Effects:
## ~participant (Number of levels: 102)
##           Estimate Est.Error 1-95% CI u-95% CI Eff.Sample Rhat
## sd(Intercept)      1.51      0.28    1.01    2.08      1615 1.00
##
## ~text (Number of levels: 6)
##           Estimate Est.Error 1-95% CI u-95% CI Eff.Sample Rhat
## sd(Intercept)      0.72      0.40    0.23    1.81      2618 1.00
##
## Population-Level Effects:
##           Estimate Est.Error 1-95% CI u-95% CI Eff.Sample Rhat
## Intercept[1]       -5.50      0.80   -7.20   -4.06      4020 1.00
## Intercept[2]       -3.10      0.50   -4.15   -2.19      3856 1.00
## Intercept[3]       -1.60      0.43   -2.45   -0.79      4200 1.00
## Intercept[4]        0.96      0.42    0.20    1.85      4070 1.00
## peak10_interest    -0.04      0.22   -0.48    0.41      4189 1.00
##
## Samples were drawn using sampling(NUTS). For each parameter, Eff.Sample
## is a crude measure of effective sample size, and Rhat is the potential
## scale reduction factor on split chains (at convergence, Rhat = 1).
```

```
# plots
plot(marginal_effects(m1_ipeak,"peak10_interest", categorical = F), points = T, point_args =
c(alpha = 0.8))
```



```
plot(marginal_effects(m1_ipeak_out,"peak10_interest", categorical = F), points = T, point_args = c(alpha = 0.8))
```



```
# define final model
m1_peak <- m1_ipeak_out
```

### 3. BOREDOM

## BOREDOM restricted model

```
# complete cases only (drop NAs)
dfsubb <- df %>% select(participant, bored_post, mean_boredom, sd_boredom, peak10_boredom, te
xt) %>% drop_na()

# restricted model
m0b_cloglog <- brm(
  bored_post ~ 1 + (1|participant) + (1|text),
  family = cumulative("cloglog"),
  prior = prior(cauchy(0, 10), class = sd),
  iter = 4000, warmup = 2000, chains = 4, cores = 4,
  control = list(adapt_delta = 0.99),
  inits = 0,
  data = dfsubb,
  save_all_pars = T)

summary(m0b_cloglog)
```

```
## Family: cumulative
## Links: mu = cloglog; disc = identity
## Formula: bored_post ~ 1 + (1 | participant) + (1 | text)
## Data: dfsubb (Number of observations: 204)
## Samples: 4 chains, each with iter = 4000; warmup = 2000; thin = 1;
## total post-warmup samples = 8000
##
## Group-Level Effects:
## ~participant (Number of levels: 102)
## Estimate Est.Error l-95% CI u-95% CI Eff.Sample Rhat
## sd(Intercept) 1.49 0.36 0.86 2.29 1554 1.00
##
## ~text (Number of levels: 6)
## Estimate Est.Error l-95% CI u-95% CI Eff.Sample Rhat
## sd(Intercept) 0.41 0.31 0.03 1.17 2204 1.00
##
## Population-Level Effects:
## Estimate Est.Error l-95% CI u-95% CI Eff.Sample Rhat
## Intercept[1] 0.45 0.31 -0.12 1.13 2614 1.00
## Intercept[2] 1.76 0.41 1.05 2.67 2019 1.00
## Intercept[3] 2.95 0.59 1.93 4.24 2034 1.00
## Intercept[4] 3.53 0.73 2.30 5.16 2163 1.00
##
## Samples were drawn using sampling(NUTS). For each parameter, Eff.Sample
## is a crude measure of effective sample size, and Rhat is the potential
## scale reduction factor on split chains (at convergence, Rhat = 1).
```

```
# other link functions
m0b_logit <- update(m0b_cloglog,
                    family = cumulative("logit"))

m0b_probit <- update(m0b_cloglog,
                    family = cumulative("probit"))

# compare different link functions
m0b_logit <- add_criterion(m0b_logit, "loo")
m0b_probit <- add_criterion(m0b_probit, "loo")
m0b_cloglog <- add_criterion(m0b_cloglog, "loo")
print(loo_compare(m0b_logit, m0b_probit, m0b_cloglog, criterion="loo", simplify = F))
```

```
##           elpd_diff se_diff elpd_loo se_elpd_loo p_loo se_p_loo looic
## m0b_cloglog    0.0      0.0  -157.5    11.7      52.1    5.2   315.0
## m0b_probit   -8.6      1.8  -166.1    13.1      56.6    5.6   332.2
## m0b_logit   -12.4      2.6  -170.0    13.6      61.3    5.9   339.9
##           se_looic
## m0b_cloglog   23.4
## m0b_probit    26.2
## m0b_logit     27.2
```

```
# chosen response distribution (link function)
m0b <- m0b_cloglog
```

## BOREDOM mean

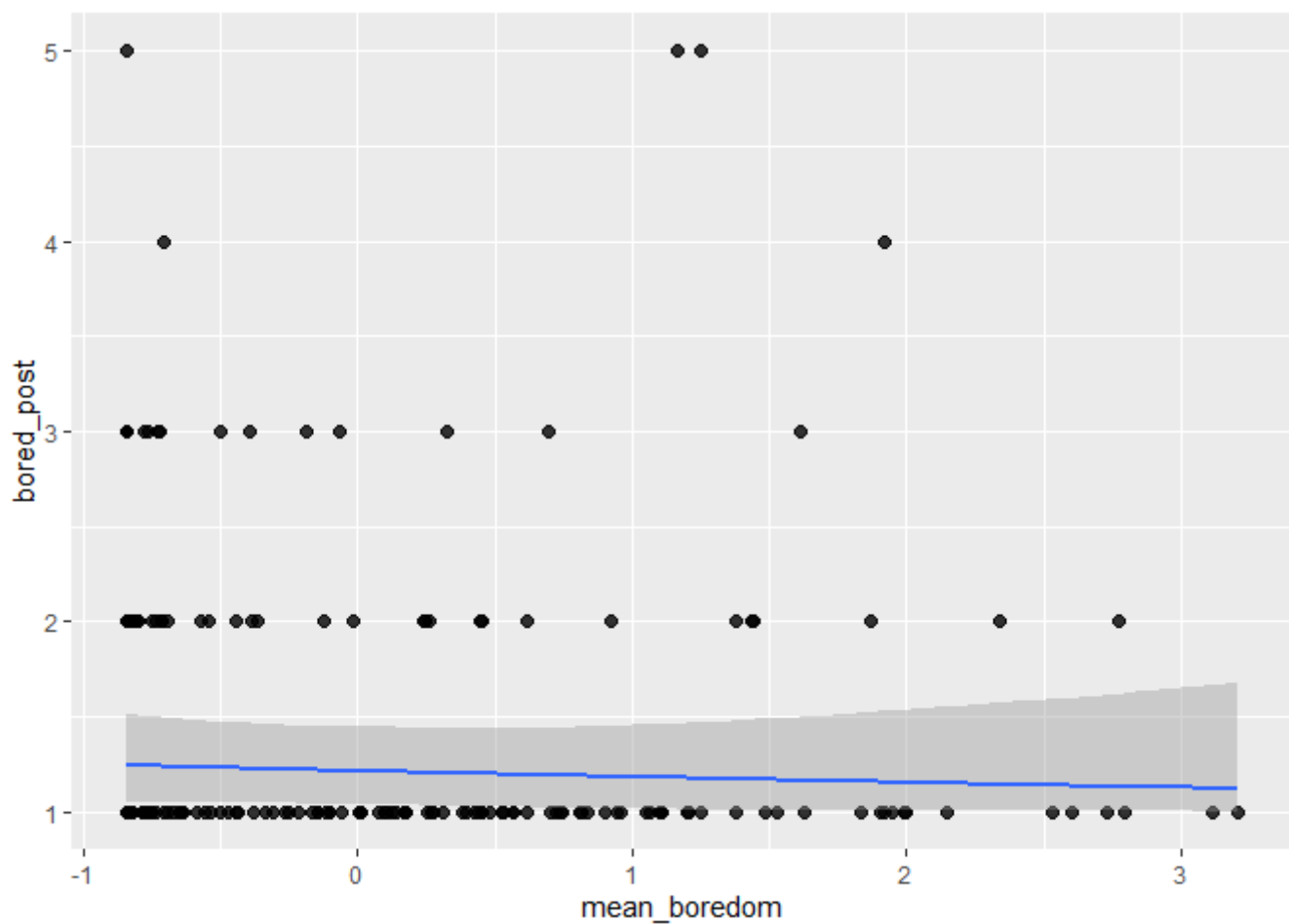
```
# full model including FaceReader's estimate as predictor
m1_bmean <- update(m0b, formula. = ~ . + mean_boredom,
                  prior = c(prior(normal(0, 10), class = b),
                            prior(cauchy(0, 10), class = sd)),
                  newdata = dfsubb,
                  save_all_pars = T)

## model parameter
summary(m1_bmean)
```



```
## Family: cumulative
## Links: mu = cloglog; disc = identity
## Formula: bored_post ~ (1 | participant) + (1 | text) + mean_boredom
## Data: dfsubb (Number of observations: 204)
## Samples: 4 chains, each with iter = 4000; warmup = 2000; thin = 1;
##           total post-warmup samples = 8000
##
## Group-Level Effects:
## ~participant (Number of levels: 102)
##           Estimate Est.Error 1-95% CI u-95% CI Eff.Sample Rhat
## sd(Intercept)      1.56      0.38   0.90   2.41      1451 1.00
##
## ~text (Number of levels: 6)
##           Estimate Est.Error 1-95% CI u-95% CI Eff.Sample Rhat
## sd(Intercept)      0.44      0.33   0.03   1.27      2363 1.00
##
## Population-Level Effects:
##           Estimate Est.Error 1-95% CI u-95% CI Eff.Sample Rhat
## Intercept[1]       0.48      0.33  -0.11   1.20      2655 1.00
## Intercept[2]       1.83      0.43   1.11   2.81      2025 1.00
## Intercept[3]       3.03      0.61   2.00   4.42      1852 1.00
## Intercept[4]       3.64      0.76   2.36   5.41      1995 1.00
## mean_boredom      -0.10      0.19  -0.50   0.26      3101 1.00
##
## Samples were drawn using sampling(NUTS). For each parameter, Eff.Sample
## is a crude measure of effective sample size, and Rhat is the potential
## scale reduction factor on split chains (at convergence, Rhat = 1).
```

```
# plots
plot(marginal_effects(m1_bmean, "mean_boredom", categorical = F), points = T, point_args = c
(alpha = 0.8))
```



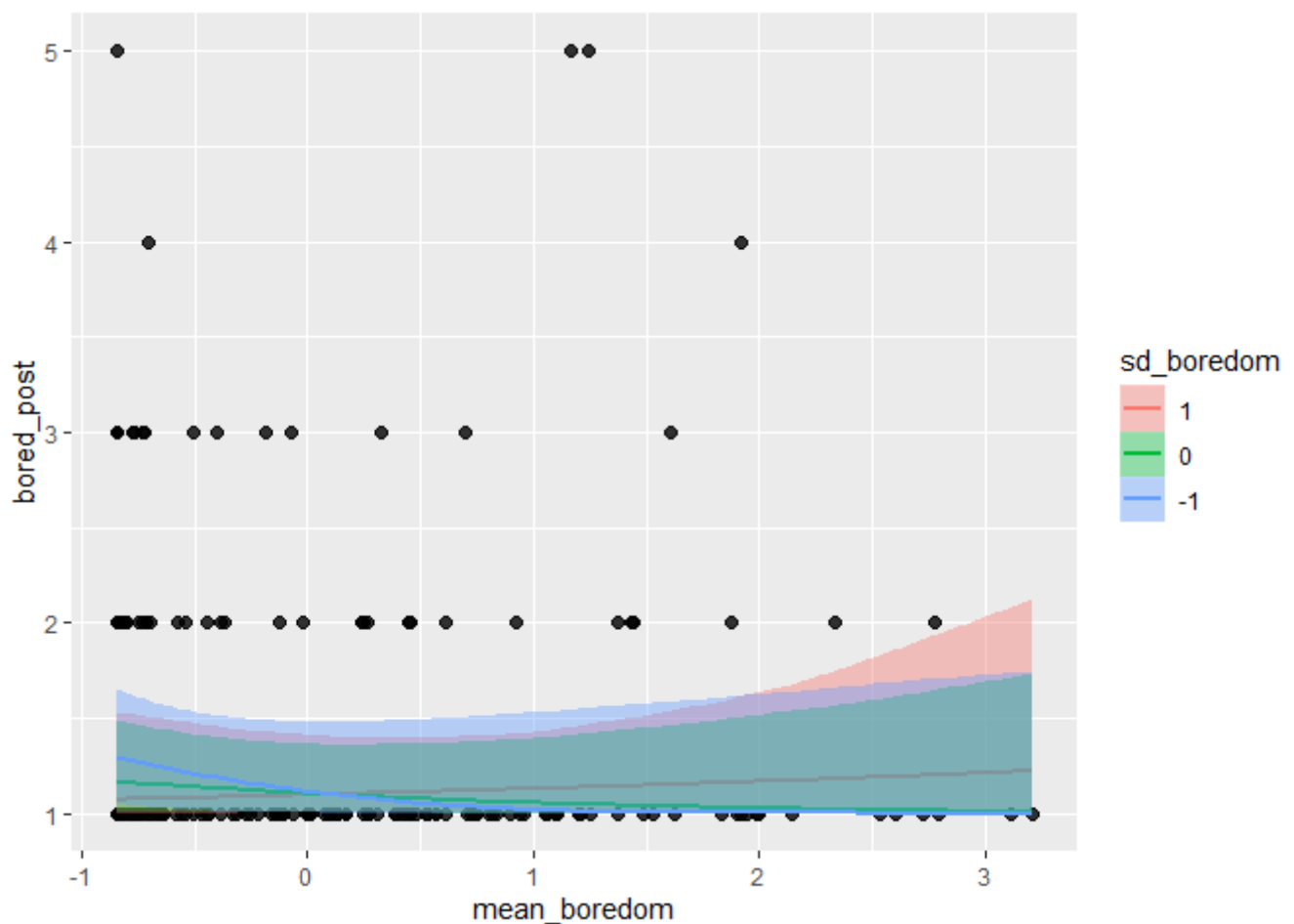
## BOREDOM mean\*SD

```
# full model including FaceReader's estimates as predictor
m1_bmeanxsd <- update(m0b, formula. = ~ . + mean_boredom*sd_boredom,
  prior = c(prior(normal(0, 10), class = b),
    prior(cauchy(0, 10), class = sd)),
  newdata = dfsubb,
  save_all_pars = T)

## model indicators
summary(m1_bmeanxsd)
```

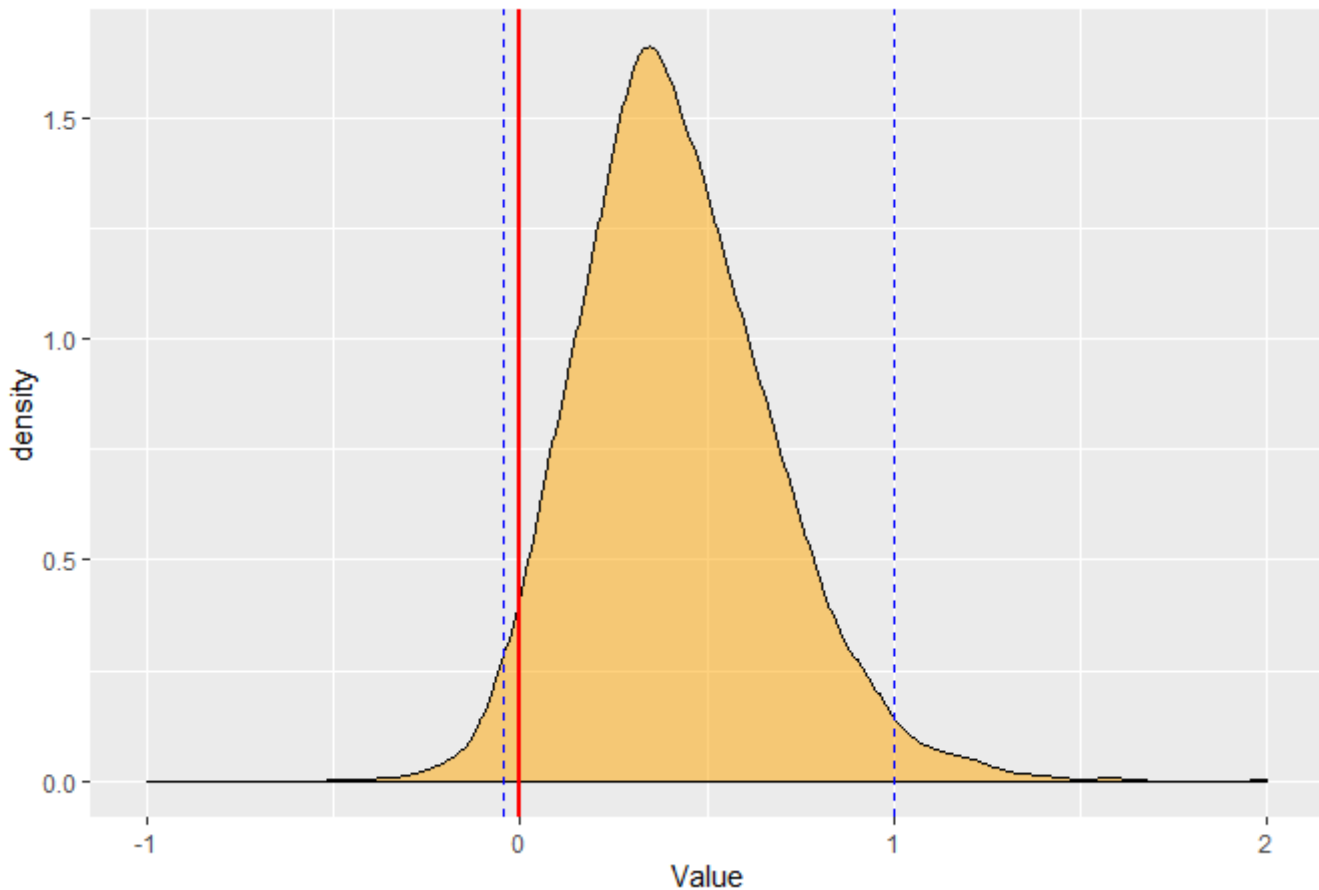
```
## Family: cumulative
## Links: mu = cloglog; disc = identity
## Formula: bored_post ~ (1 | participant) + (1 | text) + mean_boredom + sd_boredom + mean_bo
redom:sd_boredom
## Data: dfsubb (Number of observations: 204)
## Samples: 4 chains, each with iter = 4000; warmup = 2000; thin = 1;
## total post-warmup samples = 8000
##
## Group-Level Effects:
## ~participant (Number of levels: 102)
## Estimate Est.Error 1-95% CI u-95% CI Eff.Sample Rhat
## sd(Intercept) 1.77 0.45 1.03 2.80 1360 1.00
##
## ~text (Number of levels: 6)
## Estimate Est.Error 1-95% CI u-95% CI Eff.Sample Rhat
## sd(Intercept) 0.49 0.42 0.04 1.42 1087 1.00
##
## Population-Level Effects:
## Estimate Est.Error 1-95% CI u-95% CI Eff.Sample
## Intercept[1] 0.87 0.47 0.07 1.90 1692
## Intercept[2] 2.31 0.60 1.33 3.62 1469
## Intercept[3] 3.61 0.80 2.30 5.37 1496
## Intercept[4] 4.30 0.96 2.70 6.42 1578
## mean_boredom -0.28 0.38 -1.12 0.38 2566
## sd_boredom -0.04 0.32 -0.66 0.62 3397
## mean_boredom:sd_boredom 0.42 0.27 -0.04 1.00 2599
## Rhat
## Intercept[1] 1.00
## Intercept[2] 1.01
## Intercept[3] 1.01
## Intercept[4] 1.00
## mean_boredom 1.00
## sd_boredom 1.00
## mean_boredom:sd_boredom 1.00
##
## Samples were drawn using sampling(NUTS). For each parameter, Eff.Sample
## is a crude measure of effective sample size, and Rhat is the potential
## scale reduction factor on split chains (at convergence, Rhat = 1).
```

```
# plots
plot(marginal_effects(m1_bmeanxsd,"mean_boredom:sd_boredom", categorical = F), points = T, po
int_args = c(alpha = 0.8))
```



```
## plot densities and CIs of interaction-effect
m1_bmeanxsd_ggs <- ggs(m1_bmeanxsd) # transforms the brms output into a longformat tibble (used to make different types of plots)
ggplot(filter(m1_bmeanxsd_ggs, Parameter == "b_mean_boredom:sd_boredom", Iteration>1000), aes(x=value)) +
  geom_density(fill = "orange", alpha = .5) + geom_vline(xintercept = 0, col="red", size=1) +
  scale_x_continuous(name="Value", limits=c(-1, 2)) +
  labs(title="Posterior density of interaction-effect") +
  geom_vline(xintercept = summary(m1_bmeanxsd)$fixed[7,3:4], col="blue", linetype=2) # 95% CrI
```

Posterior density of interaction-effect



```
# 10-fold cross validation: interaction model compared to restricted model
m0b <- add_criterion(m0b, criterion = "kfold", folds = "grouped", group = "participant")
m1_bmeanxsd <- add_criterion(m1_bmeanxsd, criterion = "kfold", folds = "grouped", group = "participant")
print(loo_compare(m0b, m1_bmeanxsd, criterion = "kfold"), simplify = T) ## Estimating out-of
sample predictions (via 10-fold cross validation) of the interaction model, compared to a model
with no predictors yielded better results for the model without the interaction. Accordingly,
we consider this potential interaction effect as irrelevant.
```

##	elpd_diff	se_diff
## m0b	0.0	0.0
## m1_bmeanxsd	-0.8	2.2

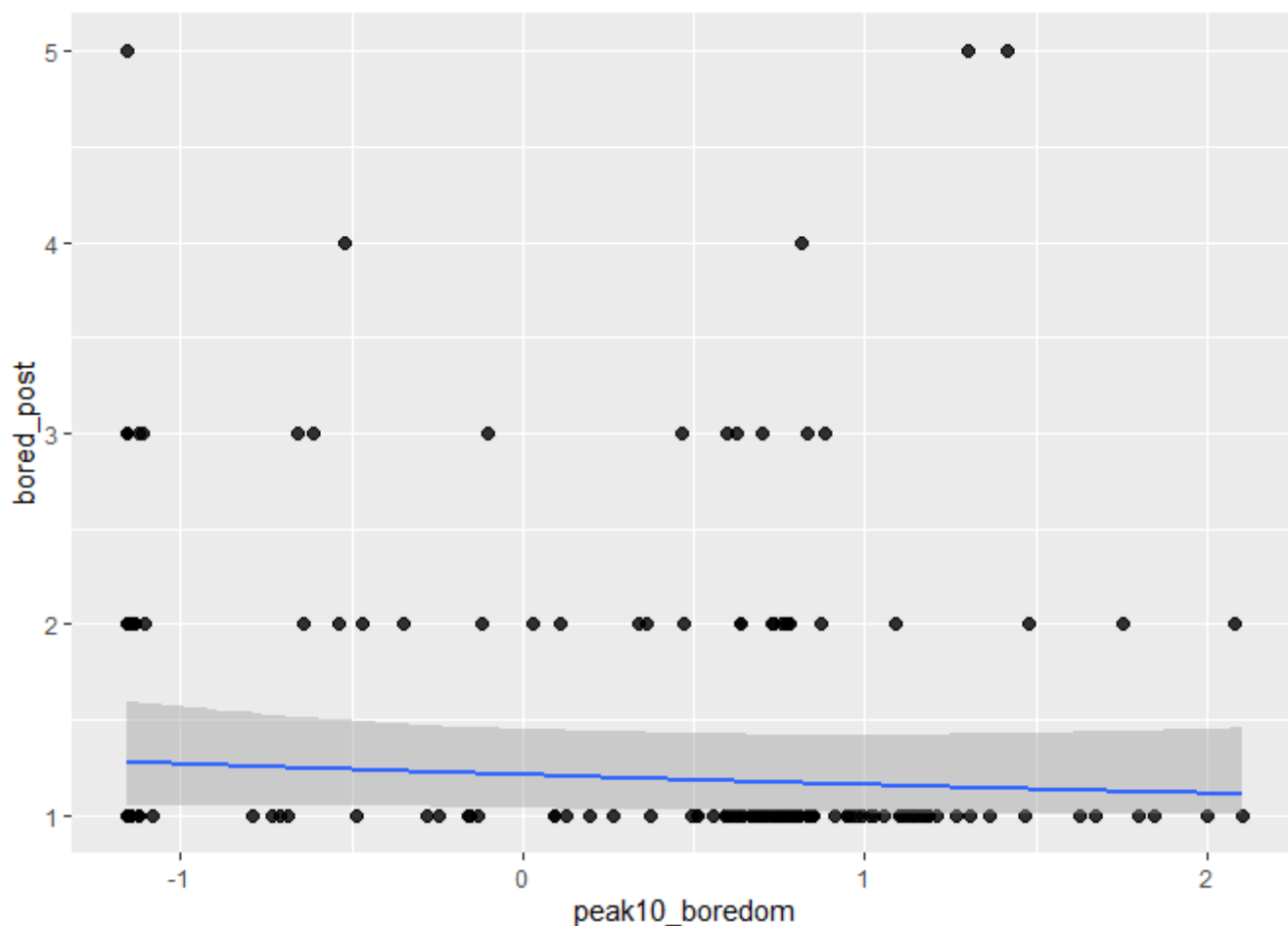
## BOREDOM mean of peaks

```
m1_bpeak <- update(m0b, formula. = ~ . + peak10_boredom,
  prior = c(prior(normal(0, 10), class = b),
    prior(cauchy(0, 10), class = sd)),
  newdata = dfsubb,
  save_all_pars = T)

## model parameter
summary(m1_bpeak)
```

```
## Family: cumulative
## Links: mu = cloglog; disc = identity
## Formula: bored_post ~ (1 | participant) + (1 | text) + peak10_boredom
## Data: dfsubb (Number of observations: 204)
## Samples: 4 chains, each with iter = 4000; warmup = 2000; thin = 1;
##           total post-warmup samples = 8000
##
## Group-Level Effects:
## ~participant (Number of levels: 102)
##           Estimate Est.Error 1-95% CI u-95% CI Eff.Sample Rhat
## sd(Intercept)      1.56      0.38   0.91   2.44      1494 1.00
##
## ~text (Number of levels: 6)
##           Estimate Est.Error 1-95% CI u-95% CI Eff.Sample Rhat
## sd(Intercept)      0.43      0.32   0.03   1.23      2184 1.00
##
## Population-Level Effects:
##           Estimate Est.Error 1-95% CI u-95% CI Eff.Sample Rhat
## Intercept[1]        0.49      0.33  -0.11   1.22      3405 1.00
## Intercept[2]        1.84      0.44   1.09   2.83      2350 1.00
## Intercept[3]        3.05      0.62   2.00   4.44      2128 1.00
## Intercept[4]        3.66      0.77   2.37   5.37      2257 1.00
## peak10_boredom     -0.16      0.19  -0.56   0.19      4057 1.00
##
## Samples were drawn using sampling(NUTS). For each parameter, Eff.Sample
## is a crude measure of effective sample size, and Rhat is the potential
## scale reduction factor on split chains (at convergence, Rhat = 1).
```

```
# plots
plot(marginal_effects(m1_bpeak, "peak10_boredom", categorical = F), points = T, point_args = c
(alpha = 0.8))
```



## 4. VALENCE

stronger priors for valence, as issues with convergence. `## VALENCE restricted model`

```
## select complete cases of relevant variables
dfsubv <- df %>% select(participant, valence_post, mean_valence, sd_valence, peak10_valence_p
os, peak10_valence_neg, text) %>% drop_na()

# restricted model
## probit-model
m0v_probit <- brm(
  valence_post ~ 1 + (1|participant) + (1|text),
  family = cumulative("probit"),
  prior = c(prior(normal(0, 1), class = Intercept),
            prior(cauchy(0, 1), class = sd)),
  iter = 4000, warmup = 2000, chains = 4, cores = 4,
  control = list(adapt_delta = 0.999, max_treedepth = 15),
  inits = 0,
  data = dfsubv,
  save_all_pars = T)
summary(m0v_probit)
```

```
## Family: cumulative
## Links: mu = probit; disc = identity
## Formula: valence_post ~ 1 + (1 | participant) + (1 | text)
## Data: dfsubv (Number of observations: 193)
## Samples: 4 chains, each with iter = 4000; warmup = 2000; thin = 1;
##           total post-warmup samples = 8000
##
## Group-Level Effects:
## ~participant (Number of levels: 97)
##           Estimate Est.Error 1-95% CI u-95% CI Eff.Sample Rhat
## sd(Intercept)      1.31      0.20   0.94   1.73      1496 1.00
##
## ~text (Number of levels: 6)
##           Estimate Est.Error 1-95% CI u-95% CI Eff.Sample Rhat
## sd(Intercept)      0.55      0.57   0.01   1.98       761 1.00
##
## Population-Level Effects:
##           Estimate Est.Error 1-95% CI u-95% CI Eff.Sample Rhat
## Intercept[1]      -3.12      0.48  -4.07  -2.15      2105 1.00
## Intercept[2]      -2.82      0.43  -3.59  -1.87      1674 1.00
## Intercept[3]      -2.65      0.42  -3.36  -1.72      1437 1.00
## Intercept[4]      -2.35      0.42  -3.02  -1.38      1196 1.00
## Intercept[5]      -1.30      0.43  -1.92  -0.26      1005 1.00
## Intercept[6]      -0.18      0.45  -0.79   0.92       912 1.00
## Intercept[7]       1.33      0.49   0.67   2.51       876 1.00
## Intercept[8]       2.85      0.55   2.01   4.12       921 1.00
##
## Samples were drawn using sampling(NUTS). For each parameter, Eff.Sample
## is a crude measure of effective sample size, and Rhat is the potential
## scale reduction factor on split chains (at convergence, Rhat = 1).
```

```
# chosen response distribution (link function)
m0v <- m0v_probit
```

## VALENCE mean

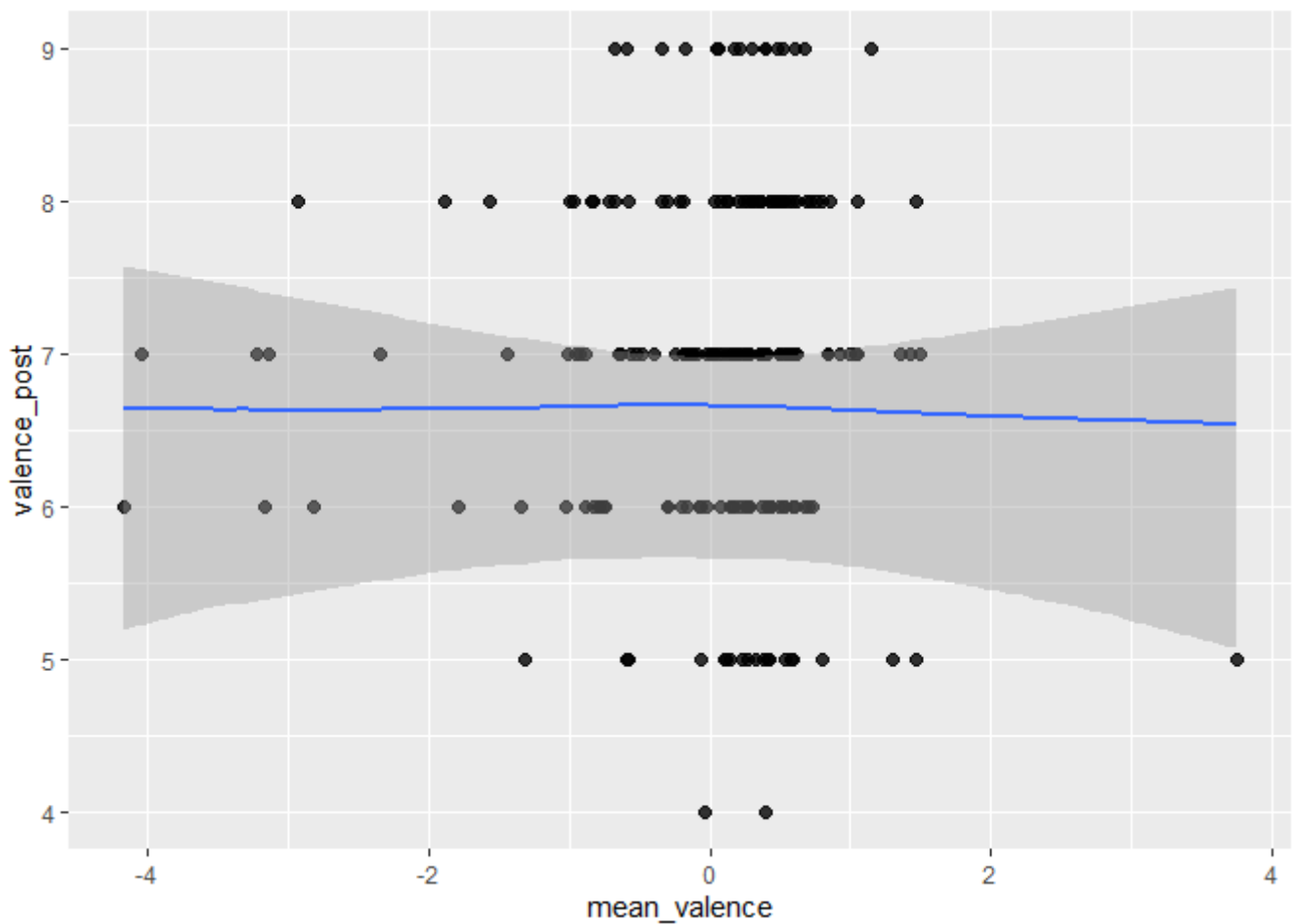
```
# full model including FaceReader's estimate as predictor
m1_vmean <- update(m0v, formula. = ~ . + mean_valence,
                  prior = c(prior(normal(0, 1), class = Intercept),
                           prior(normal(0, 1), class = b),
                           prior(cauchy(0, 1), class = sd)),
                  newdata = dfsubv,
                  save_all_pars = T,
                  seed = 19) # for reproducibility (on the same machine) - to avoid "Stan model x does not contain samples." which sometimes occurred

## model parameter
summary(m1_vmean)
```



```
## Family: cumulative
## Links: mu = probit; disc = identity
## Formula: valence_post ~ (1 | participant) + (1 | text) + mean_valence
## Data: dfsubv (Number of observations: 193)
## Samples: 4 chains, each with iter = 4000; warmup = 2000; thin = 1;
##           total post-warmup samples = 8000
##
## Group-Level Effects:
## ~participant (Number of levels: 97)
##           Estimate Est.Error l-95% CI u-95% CI Eff.Sample Rhat
## sd(Intercept)      1.34      0.20    0.96    1.75      1812 1.00
##
## ~text (Number of levels: 6)
##           Estimate Est.Error l-95% CI u-95% CI Eff.Sample Rhat
## sd(Intercept)      0.55      0.58    0.01    2.07       686 1.00
##
## Population-Level Effects:
##           Estimate Est.Error l-95% CI u-95% CI Eff.Sample Rhat
## Intercept[1]      -3.13      0.48   -4.08   -2.14      2212 1.00
## Intercept[2]      -2.83      0.43   -3.61   -1.89      1597 1.00
## Intercept[3]      -2.67      0.42   -3.40   -1.75      1400 1.00
## Intercept[4]      -2.36      0.42   -3.04   -1.42      1152 1.00
## Intercept[5]      -1.31      0.43   -1.92   -0.27       886 1.00
## Intercept[6]      -0.19      0.45   -0.78    0.92       797 1.00
## Intercept[7]       1.34      0.49    0.67    2.53       797 1.00
## Intercept[8]       2.86      0.56    2.01    4.16       918 1.00
## mean_valence      -0.02      0.15   -0.31    0.27      3917 1.00
##
## Samples were drawn using sampling(NUTS). For each parameter, Eff.Sample
## is a crude measure of effective sample size, and Rhat is the potential
## scale reduction factor on split chains (at convergence, Rhat = 1).
```

```
# plots
plot(marginal_effects(m1_vmean, "mean_valence", categorical = F), points = T, point_args = c
(alpha = 0.8))
```



## VALENCE mean\*SD

```
# full model including FaceReader's estimates as predictor
m1_vmeanxsd <- update(m0v, formula. = ~ . + mean_valence*sd_valence,
  prior = c(prior(normal(0, 1), class = Intercept),
    prior(normal(0, 1), class = b),
    prior(cauchy(0, 1), class = sd)),
  control = list(adapt_delta = 0.999, max_treedepth = 15),
  inits = 0,
  newdata = dfsbv,
  save_all_pars = T,
  seed = 21) # for reproducibility (on the same machine) - to avoid diverge
nt transitions which sometimes occurred

## model parameter
summary(m1_vmeanxsd)
```

```

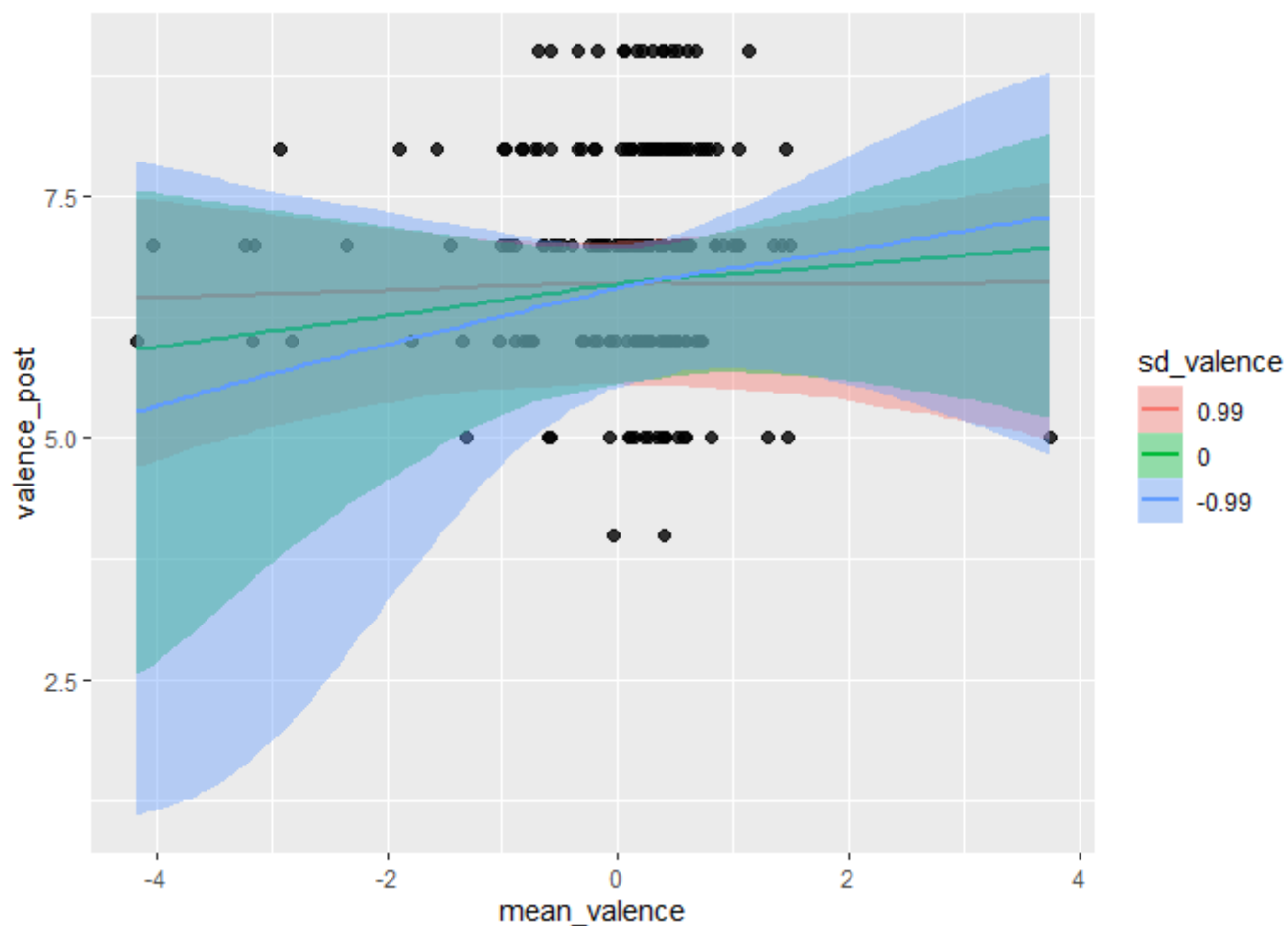
## Family: cumulative
## Links: mu = probit; disc = identity
## Formula: valence_post ~ (1 | participant) + (1 | text) + mean_valence + sd_valence + mean_valence:sd_valence
## Data: dfsubv (Number of observations: 193)
## Samples: 4 chains, each with iter = 4000; warmup = 2000; thin = 1;
##           total post-warmup samples = 8000
##
## Group-Level Effects:
## ~participant (Number of levels: 97)
##           Estimate Est.Error 1-95% CI u-95% CI Eff.Sample Rhat
## sd(Intercept)      1.36      0.21   0.98   1.78      1594 1.00
##
## ~text (Number of levels: 6)
##           Estimate Est.Error 1-95% CI u-95% CI Eff.Sample Rhat
## sd(Intercept)      0.60      0.62   0.01   2.15       704 1.01
##
## Population-Level Effects:
##           Estimate Est.Error 1-95% CI u-95% CI Eff.Sample
## Intercept[1]      -3.06      0.49  -4.05  -2.08      2020
## Intercept[2]      -2.77      0.44  -3.55  -1.83      1473
## Intercept[3]      -2.60      0.43  -3.34  -1.65      1309
## Intercept[4]      -2.29      0.44  -3.01  -1.30      1119
## Intercept[5]      -1.23      0.46  -1.91  -0.18       866
## Intercept[6]      -0.10      0.48  -0.76   1.03       790
## Intercept[7]       1.44      0.53   0.71   2.66       760
## Intercept[8]       2.99      0.59   2.07   4.32       827
## mean_valence       0.16      0.25  -0.32   0.65      2869
## sd_valence         0.03      0.14  -0.25   0.31      4093
## mean_valence:sd_valence -0.14      0.15  -0.44   0.15      3355
##
##           Rhat
## Intercept[1]      1.00
## Intercept[2]      1.00
## Intercept[3]      1.00
## Intercept[4]      1.00
## Intercept[5]      1.01
## Intercept[6]      1.01
## Intercept[7]      1.01
## Intercept[8]      1.01
## mean_valence      1.00
## sd_valence         1.00
## mean_valence:sd_valence 1.00
##
## Samples were drawn using sampling(NUTS). For each parameter, Eff.Sample
## is a crude measure of effective sample size, and Rhat is the potential
## scale reduction factor on split chains (at convergence, Rhat = 1).

```

```

# plots
plot(marginal_effects(m1_vmeanxsd,"mean_valence:sd_valence"), points = T, point_args = c(alpha = 0.8))

```



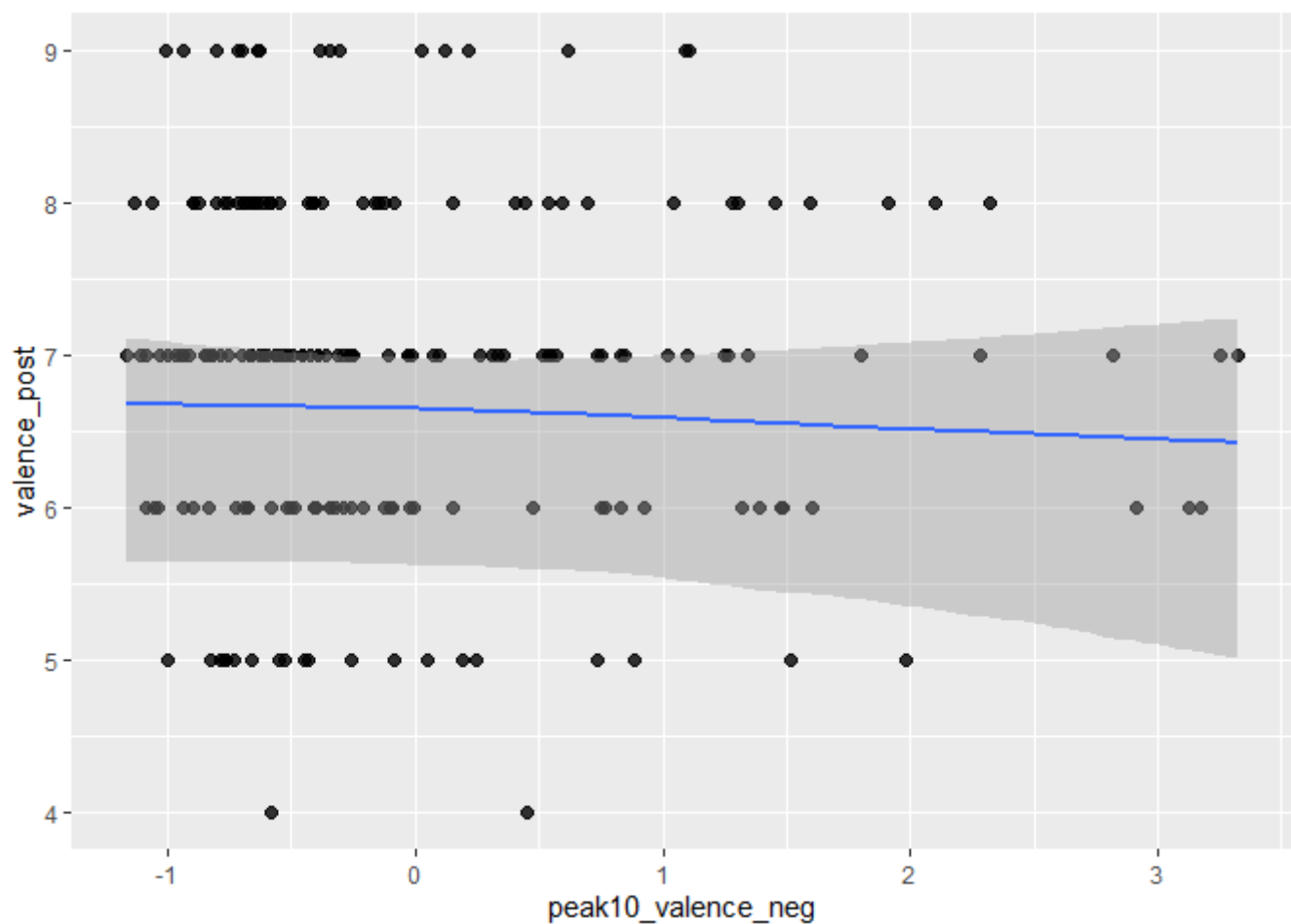
## VALENCE mean of peaks

```
m1_vpeak <- update(m0v, formula. = ~ . + peak10_valence_neg + peak10_valence_pos,
  prior = c(prior(normal(0, 1), class = Intercept),
    prior(normal(0, 1), class = b),
    prior(cauchy(0, 1), class = sd)),
  newdata = dfsbv,
  save_all_pars = T)
```

```
## model parameter
summary(m1_vpeak)
```

```
## Family: cumulative
## Links: mu = probit; disc = identity
## Formula: valence_post ~ (1 | participant) + (1 | text) + peak10_valence_neg + peak10_valence_pos
## Data: dfsubv (Number of observations: 193)
## Samples: 4 chains, each with iter = 4000; warmup = 2000; thin = 1;
## total post-warmup samples = 8000
##
## Group-Level Effects:
## ~participant (Number of levels: 97)
##      Estimate Est.Error 1-95% CI u-95% CI Eff.Sample Rhat
## sd(Intercept)      1.34      0.20    0.97    1.79      1666 1.00
##
## ~text (Number of levels: 6)
##      Estimate Est.Error 1-95% CI u-95% CI Eff.Sample Rhat
## sd(Intercept)      0.56      0.58    0.02    2.05       789 1.00
##
## Population-Level Effects:
##      Estimate Est.Error 1-95% CI u-95% CI Eff.Sample Rhat
## Intercept[1]      -3.13      0.49   -4.09   -2.16      2249 1.00
## Intercept[2]      -2.82      0.43   -3.60   -1.89      1684 1.00
## Intercept[3]      -2.65      0.43   -3.40   -1.71      1541 1.00
## Intercept[4]      -2.35      0.43   -3.05   -1.37      1277 1.00
## Intercept[5]      -1.30      0.45   -1.94   -0.21      1035 1.00
## Intercept[6]      -0.18      0.46   -0.79    0.95       961 1.00
## Intercept[7]       1.36      0.50    0.69    2.56       931 1.00
## Intercept[8]       2.89      0.56    2.04    4.21      1061 1.00
## peak10_valence_neg -0.06      0.14   -0.34    0.23      3692 1.00
## peak10_valence_pos  0.03      0.12   -0.22    0.27      4991 1.00
##
## Samples were drawn using sampling(NUTS). For each parameter, Eff.Sample
## is a crude measure of effective sample size, and Rhat is the potential
## scale reduction factor on split chains (at convergence, Rhat = 1).
```

```
# plots
plot(marginal_effects(m1_vpeak, "peak10_valence_neg", categorical = F), points = T, point_args = c(alpha = 0.8))
```



```
plot(marginal_effects(m1_vpeak,"peak10_valence_pos", categorical = F), points = T, point_args = c(alpha = 0.8))
```

