#### An introduction to (Generalized) Linear and Nonlinear Mixed Models

#### Dealing with Hierarchical, Nested and Temporal Dependencies in Data:

#### I. Introduction to random effects in statistical models

Fixed vs. Random effects

Random intercepts, random slopes and covariance matrix structures

Cross-level interactions

#### II. The multilevel model for change

Repeated measurements as a clustering level within individuals

Growth curve models

Models of piece-wise and nonlinear growth

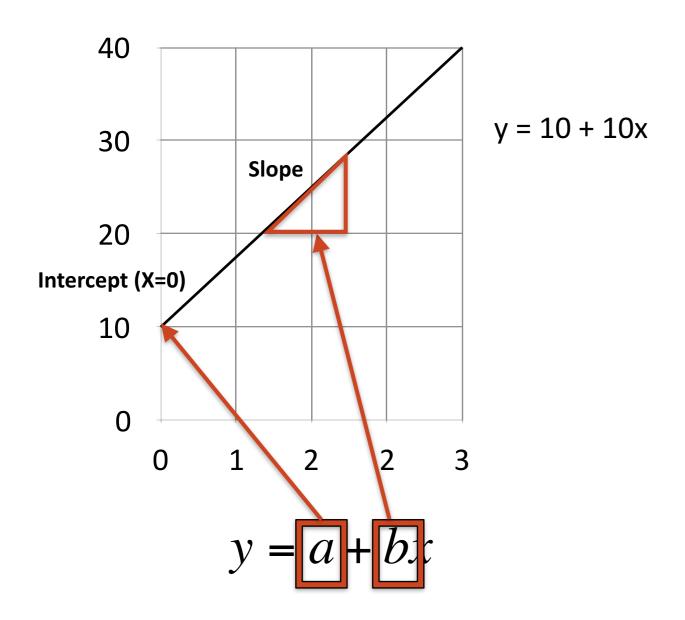
#### III. Advanced models

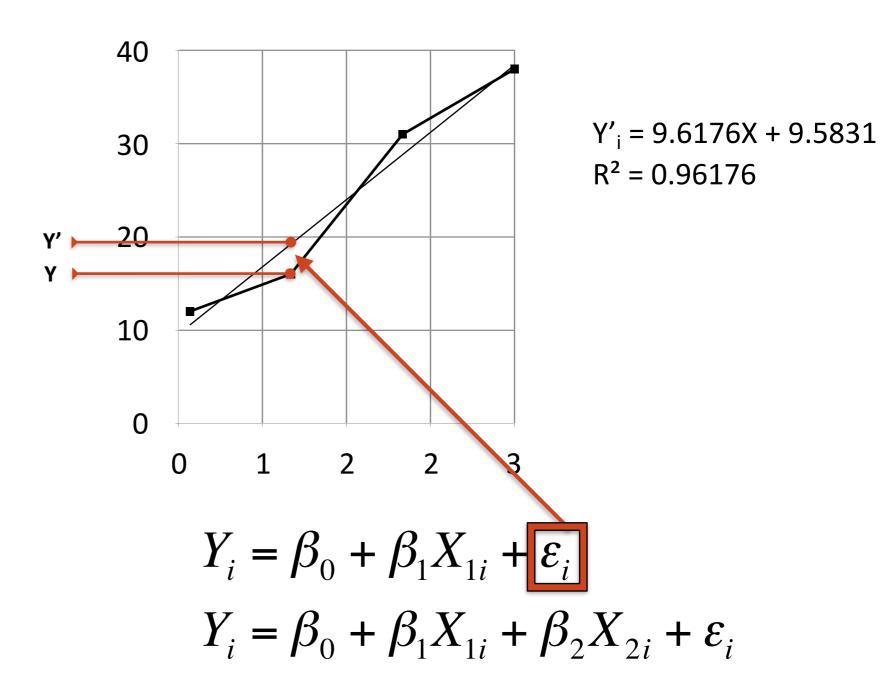
The generalized linear mixed model for binary outcomes and count data

Cross-classified and multiple membership models

Multivariate-multilevel models

#### Statistical models versus equations





#### The line as a model: Multiple linear regression

X<sub>i</sub> as linear additive independent source(s) of variance in Y<sub>i</sub>
 Linear prediction: Y' >> Y<sub>i</sub> from X<sub>i</sub>

- Variance *not* 'explained' by  $X_i$  is captured by an error term  $\mathcal{E}_i$ Residual variance:  $Y' - Y_i$
- Model parameters are estimated using a 'least-squares' method:
   Minimise residual variance: smallest squared differences between observed and predicted scores.
- Remember Ordinary Least Squares (OLS) assumptions?

#### **OLS** regression: Assumptions for being BLUE

Hypothesised model: 
$$Y_i = \sum_{i=1}^k \beta_j X_{ij} + \varepsilon_i$$
 for  $i = 1,...,n$ 

: unobserved, non-random parameter  $\beta_i$ 

: observed, (non-) random variable

: observed, random variable

 $\boldsymbol{\varepsilon}_i$ : unobserved random statistical error  $(E(Y_i) - Y_i)$ , or: 'true' - observation, population - sample)

#### Weak set of assumptions (Gauss-Markov Theorem):

1. 
$$E(Y_i) = \sum_{i=1}^k \beta_j X_{ij}$$
  $E(\varepsilon_i) = 0$  for  $i = 1,...,n$ 

**Strong assumptions (Gauss-Markov Theorem + Central Limit Theorem):** 

 $1 + 2 + 3 + Y_i$  is normally distributed

The model represents a TRUE linear relationship so the expected value of the statistical errors is 0.

2. 
$$V(\varepsilon_i) = \sigma^2 < \infty$$
 for  $i = 1,...,n$ 

Assumption  $1 + 2 = homoskedastic error variance (also for variance of <math>Y_i$ )

3. No dependencies (correlations) among errors (also no correlations between  $Y_i$  and  $\varepsilon_i$ )

# **Assumptions for simple OLS regression**

Single level model														
	1	2	3	4	5	6	7	8	9	10	11	12	13	
1	1	0	0	0	0	0	0	0	0	0	0	0	0	
2	0	1	0	0	0	0	0	0	0	0	0	0	0	
3	0	0	1	0	0	0	0	0	0	0	0	0	0	
4	0	0	0	1	0	0	0	0	0	0	0	0	0	
5	0	0	0	0	1	0	0	0	0	0	0	0	0	
6	0	0	0	0	0	1	0	0	0	0	0	0	0	
7	0	0	0	0	0	0	1	0	0	0	0	0	0	
8	0	0	0	0	0	0	0	1	0	0	0	0	0	
9	0	0	0	0	0	0	0	0	1	0	0	0	0	
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#### **Correlation matrix V**



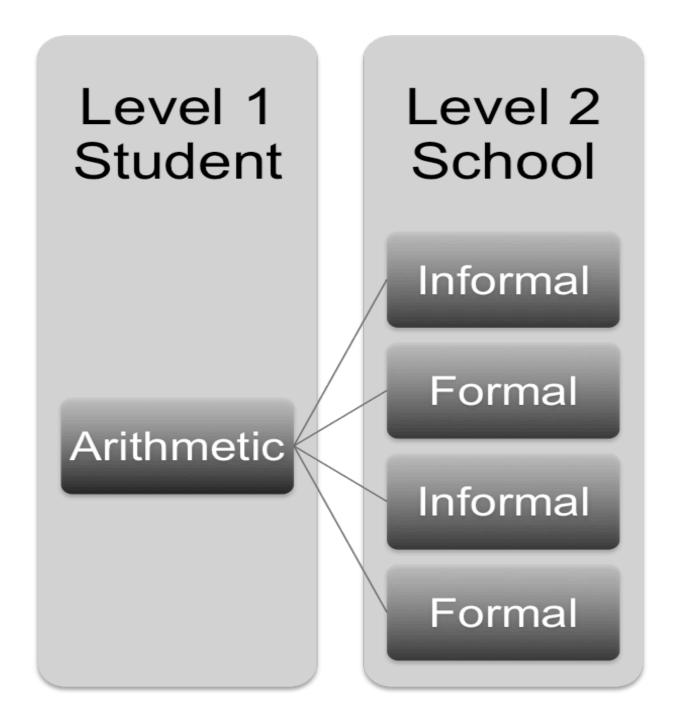
#### **Example: Sources of variability**

- Consider the following dataset:
  - 20 students from 4 different schools
  - X = arithmetic test at the beginning of the school year
  - Y = arithmetic test at the end of the school year
  - 2 schools have a formal teaching style
  - 2 schools have an informal teaching style

- Suppose we want to predict Y from X, what sources of variance are there?
- Students Arithmetic scores
   Schools Teaching style



#### **Example: Sources of variability**



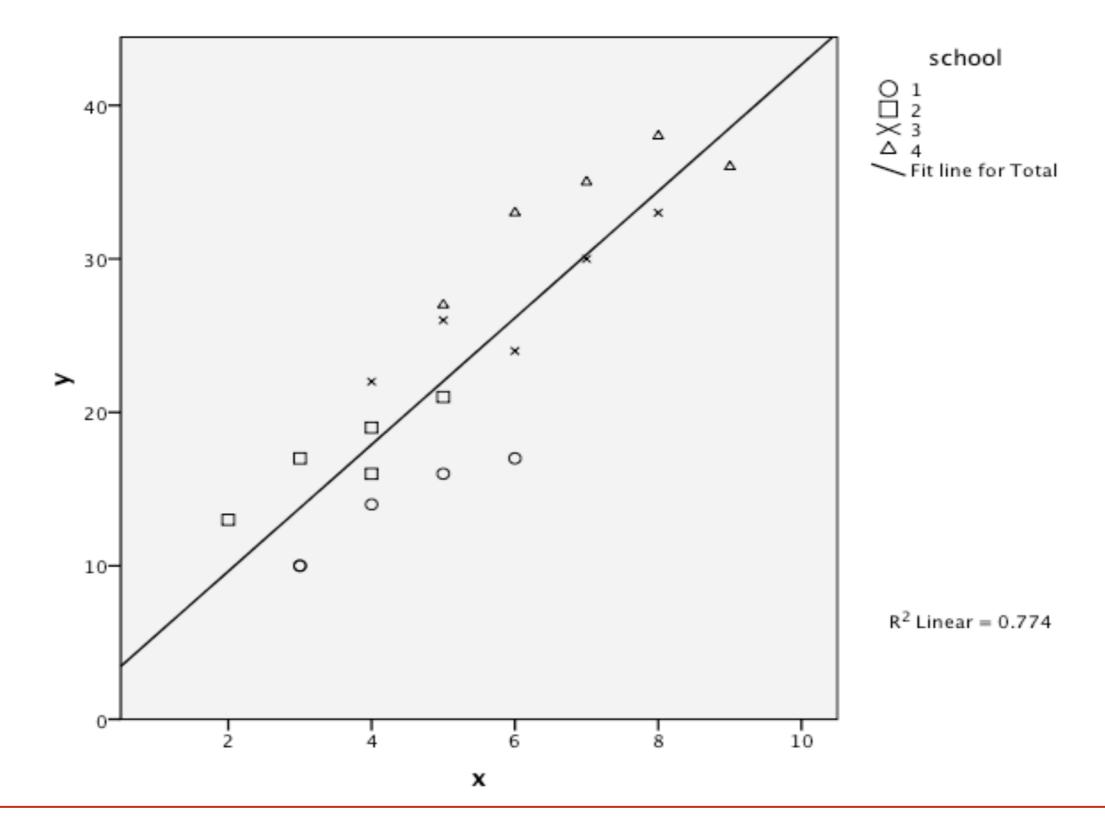
_		-	
	10	-	
_			•

					Formal Teaching
	student	school	x	У	Style
1	1	1	3	10	0
2	2	1	3	10	0
3	3	1	4	14	0
4	4	1	5	16	0
5	5	1	6	17	0
6	6	2	2	13	1
7	6	2	3	17	1
8	8	2	4	16	1
9	9	2	4	19	1
10	10	2	5	21	1
11	11	3	4	22	0
12	12	3	5	26	0
13	13	3	6	24	0
14	14	3	7	30	0
15	15	3	8	33	0
16	16	4	5	27	1
17	17	4	6	33	1
18	18	4	7	35	1
19	19	4	8	38	1
20	20	4	9	36	1

These data clearly vary on multiple levels

 To understand how multilevel analysis works we'll start by taking a classical multiple regression approach to analyse these data

Start with the entire group, pretend there are no levels



Regression analysis at the student level, ignoring school:

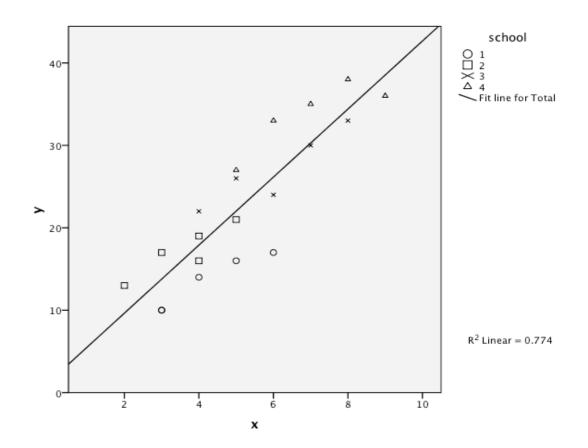
$$Y_i = \beta_0 + \beta_1 X_i + \varepsilon_i$$

#### Coefficients<sup>a</sup>

	Unstandardized Coefficients		Standardized Coefficients			
Model		В	Std. Error	Beta	t	Sig.
1	(Constant)	1.389	2.902		.479	.638
	x	4.127	.525	.880	7.855	.000

a. Dependent Variable: y

$$Y'_i = 1.389 + 4.127X$$



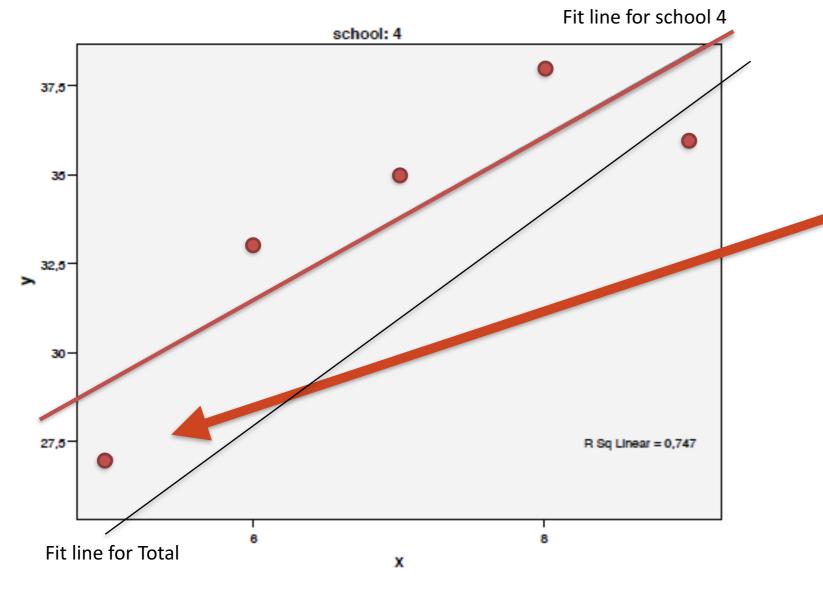
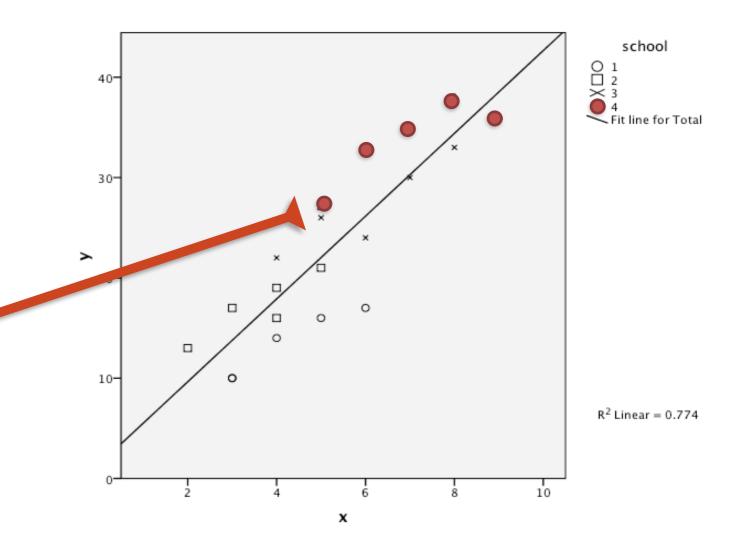


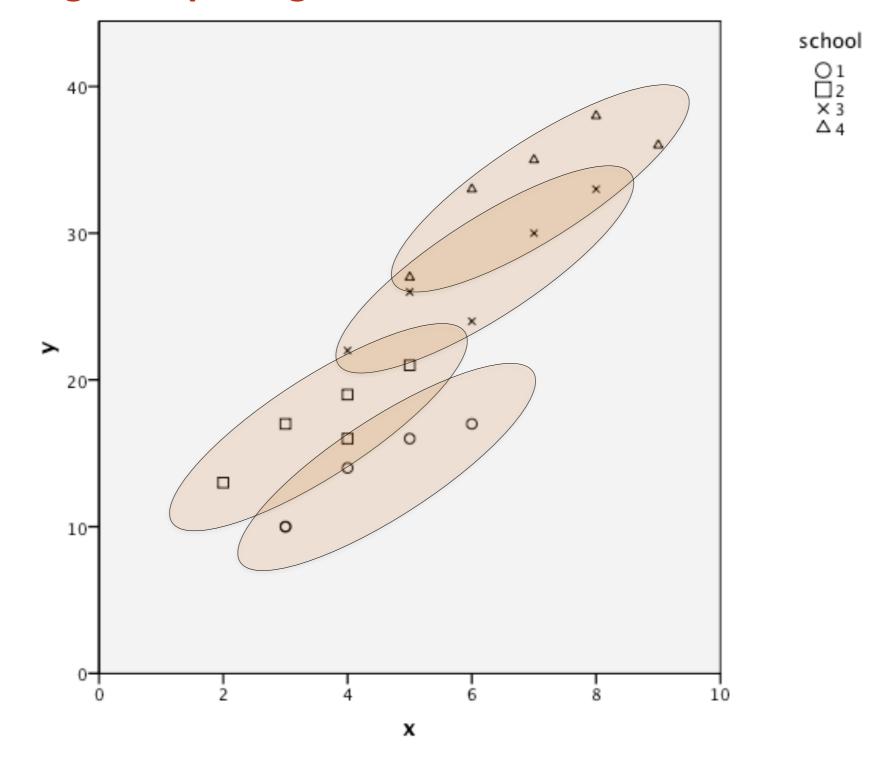
Figure 2: Variation at the student level within a particular school.

$$Y_{i4} = \beta_{04} + \beta_{14}X_{i4} + \varepsilon_{i4}$$



- Different intercept for school 4
- Different slope for school 4?
- Different residuals for pupils relative best fit for school 4





Regression analysis including a possibly different intercept for each school, but a common slope:

$$Y_{ij} = \beta_{0j} + \beta_1 X_{ij} + \epsilon_{ij}$$

#### Model Summaryb

Model	R	R Square	Adjusted R Square	Std. Error of the Estimate
1	.987 <sup>a</sup>	.973	.966	1.647

a. Predictors: (Constant), school3, x, school1, school2

#### Coefficients<sup>a</sup>

			lardized icients	Standardized Coefficients		
Model		В	Std. Error	Beta	t	Sig.
1	(Constant)	16.738	2.167		7.725	.000
	x	2.438	.291	.520	8.373	.000
	schooll	13.575	1.323	674	-10.265	.000
	school2	-8.313	1.437	413	-5.786	.000
	school3	-4.363	1.081	217	-4.034	.001

a. Dependent Variable: y

#### Variables in dataset:

Shool1 = 1, rest 0

Shool2 = 1, rest 0

Shool3 = 1, rest 0

Consequence:

All relative to school 4

5 parameters to be estimated



b. Dependent Variable: y

The analysis includes three dummy variables to represent the four schools; school 4 used as the reference category.

The regression equations for separate schools are:

School 4: 
$$Y' = 16.738 + 2.438X$$

School 1: 
$$Y' = (16.378 - 13.575) + 2.438X = 3.163 + 2.438X$$

School 2: 
$$Y' = (16.378 - 8.313) + 2.438X = 8.425 + 2.438X$$

School 3: 
$$Y' = (16.378 - 4.363) + 2.438X = 12.375 + 2.438X$$
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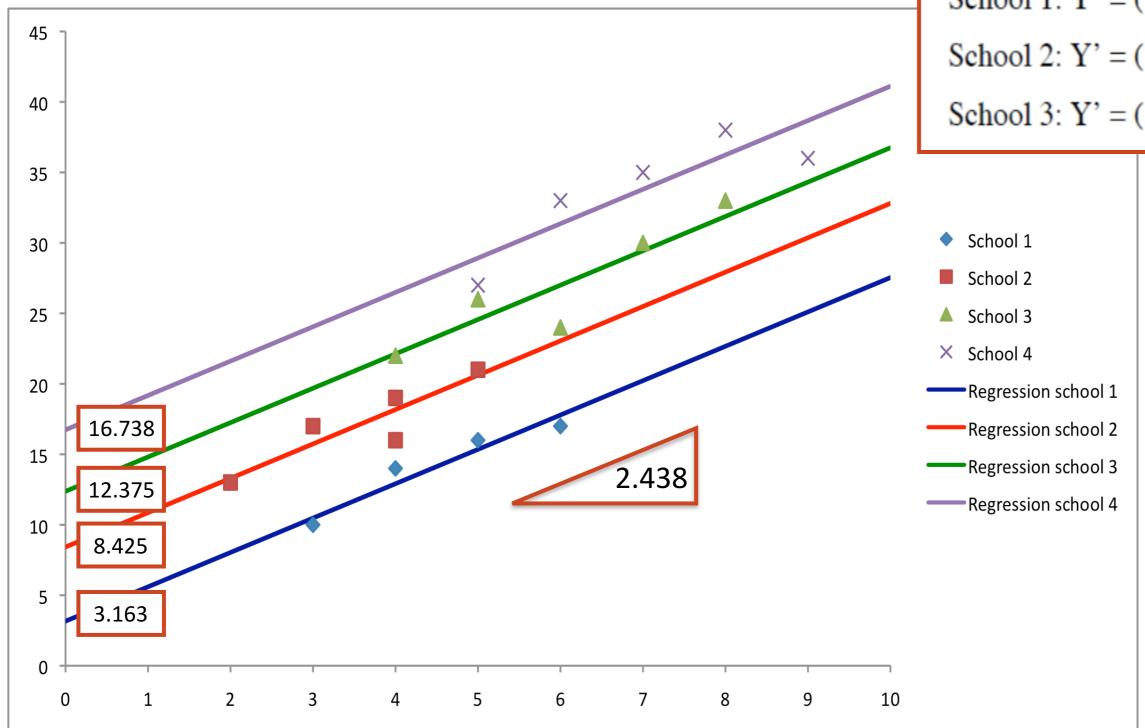
			lardized icients	Standardized Coefficients		
Model		В	Std. Error	Beta	t	Sig.
1	(Constant)	16.738	2.167		7.725	.000
	X	2.438	.291	.520	8.373	.000
	school1	-13.575	1.323	674	-10.265	.000
	school2	-8.313	1.437	413	-5.786	.000
	school3	-4.363	1.081	217	-4.034	.001

a. Dependent Variable: y

#### Note:

- 1. The slope of X is now 2.438, but when ignoring school the slope is 4.127.
- 2. Rsquare increased from .774 to .973. School differences do matter.

#### **Variations between schools (intercept)**



School 4: 
$$Y' = 16.738 + 2.438X$$

School 1: 
$$Y' = (16.378 - 13.575) + 2.438X = 3.163 + 2.438X$$

School 2: 
$$Y' = (16.378 - 8.313) + 2.438X = 8.425 + 2.438X$$

School 3: 
$$Y' = (16.378 - 4.363) + 2.438X = 12.375 + 2.438X$$
.

#### 5 parameters:

- 4 intercepts
- 1 slope



#### Schools also vary in teaching style

- Disregard previously found school differences.
- Same procedure:
  - 2 levels, so 1 dummy variable: Formal style = 1, rest 0.
  - Results relative to Informal style

#### Coefficients<sup>a</sup>

		Unstandardized Coefficients		Standardized Coefficients		
Model		В	Std. Error	Beta	t	Sig.
1	(Constant)	20.606	1.198		17.208	.000
	Formal style	4.488	1.695	.257	2.648	.017
	x centered	4.062	.456	.866	8.917	.000

a. Dependent Variable: y

Schools with informal style: Y' = 20.606 + 4.062X

Schools with formal style: Y' = (20.606 + 4.488) + 4.062X.

#### 3 parameters to be estimated

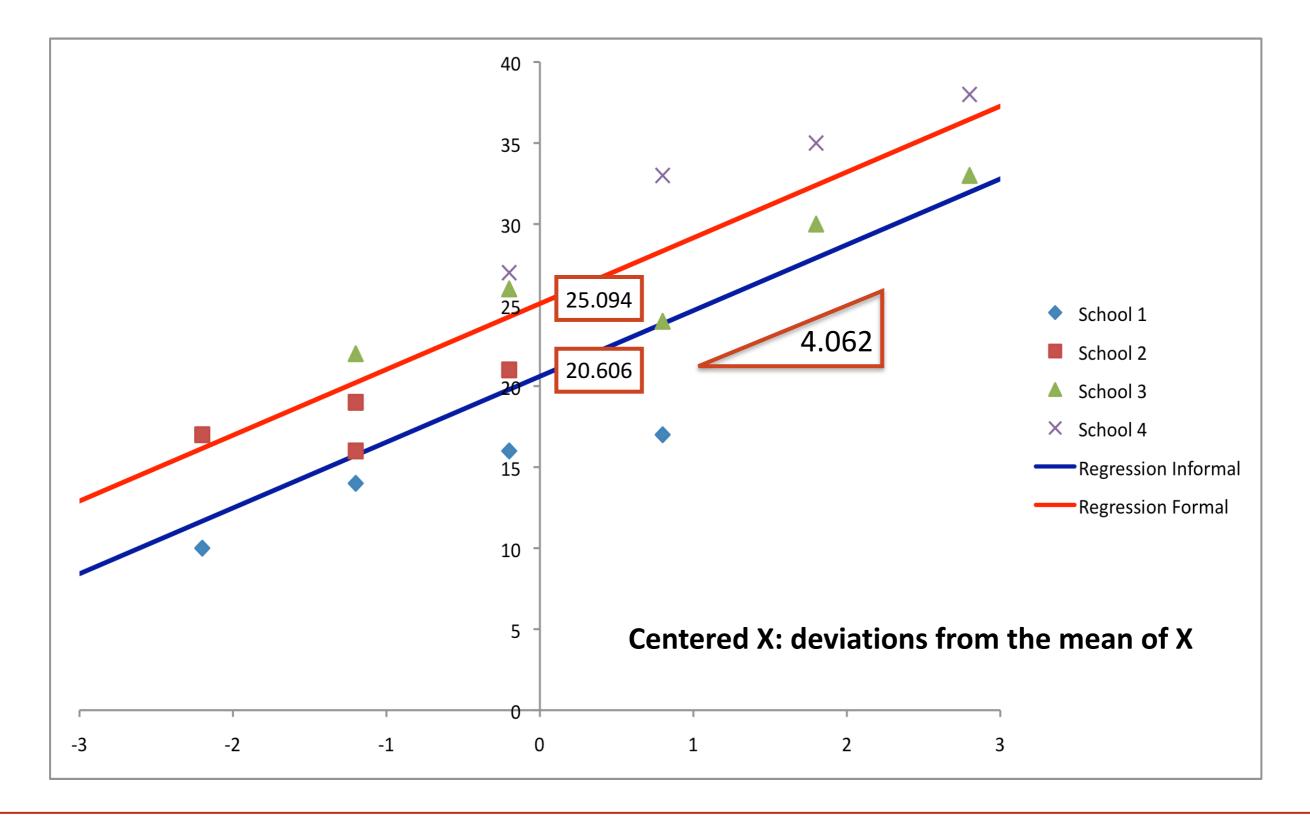
$$R^{2}_{general} = .77$$

$$R^{2}_{schools} = .93$$

$$R^2_{\text{style}} = .84$$



#### Schools also vary in teaching style



#### Teaching style and school differences in 1 model

- Add two dummy variables:
  - D1 = difference between informal schools 1 (-1) and 3 (1), rest 0
  - D2 = difference between formal schools 2 (-1) and 4 (1), rest 0

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	_	Unstandardized Coefficients		Standardized Coefficients		
Model		В	Std. Error	Beta	t	Sig.
1	(Constant)	20.444	.522		39.198	.000
	x centered	2.438	.291	.520	8.373	.000
	Formal style	4.813	.739	.276	6.515	.000
	d1	4.606	.583	.373	7.902	.000
	<b>d</b> 2	4.156	.718	.337	5.786	.000

a. Dependent Variable: y

#### 5 parameters

straightforward

(rearrange equations)

$$R^2_{general} = .77$$
 $R^2_{schools} = .93$ 
 $R^2_{style} = .84$ 
 $R^2_{both} = .97$ 
Interpretation is not

#### Teaching style and school differences: Fixed vs. Random effects

Compare the multiple regression results with the results of a multilevel analysis (fixed parameter estimates only):

	b	SE(b)	t	p
Intercept	20.45	3.02	8.57	.000
X	2.53	0.28	9.11	.000
Formal style	4.79	4.27	1.12	.378

The parameter estimates don't differ very much; the difference is mainly in the standard errors.

Both analyses take school differences into account. The multiple regression with dummy variables to represent schools treats 'school' as a fixed effect; the multilevel analysis treats 'school' as a random effect.

		Coefficient					
	Unstandardized Coefficients						
	В	Std. Error					
(Constant)	20.444	.522					
x centered	2.438	.291					
Formal style	4.813	.739					
<b>d</b> 1	4.606	.583					
<b>d</b> 2	4.156	.718					

# Random intercepts, random slopes and covariance matrix structures

Multilevel model as a regression model

Random intercepts

Random Slopes

Intercept-Slope covariance

# Why multilevel models?

- Take clustered structure of data into account... what does that mean?
- Assumption of independent observations is not satisfied... what does that mean?
- SE are underestimated... DANGER: infer a relationship exists when none is present... (check SE formula)
- In standard regression: only individual level random variability
- Solution: Treat variability at higher levels as random variability
- <u>Assumption</u>: Level 2 variable is a random selection of groups/clusters out of a population who are normallly distributed around the population mean.

# Why multilevel models

• Dependencies in the data:  $\frac{-u_0}{\sigma^2 + \sigma^2}$ 

 Similarity between individuals in the same group: Intra class correlation (ICC) / Proportion of <u>residual</u> variation due to differences between groups: variance partition coefficient (VPC)

Between 0 and 1: 0.3 is large!

#### The multilevel model as a regression model

• The multiple regression model includes only 1 random source of variability. Individual differences expressed as residuals around the regression line:

$$Y_i = \beta_0 + \beta_1 X_i + \beta_2 \text{ Style}_i + \beta_3 d1_i + \beta_4 d2_i + \varepsilon_i$$

In a multilevel model we can include random sources at two or more levels:

$$Y_{ij} = \beta_{0j} + \beta_1 X_{ij} + \beta_2 \, Style_j + \epsilon_{ij}$$
 -> Residuals around regression line within a school

$$eta_{0j} = eta_{00} + \mathbf{u_j}$$
 -> Variation in school intercepts

Or: 
$$Y_{ij} = \beta_{00} + \beta_1 X_{ij} + \beta_2 \text{ Style}_j + (\epsilon_{ij} + u_j)$$



#### In words: variance to be explained

- All multilevel modelling starts by defining an "empty model" in which you define the levels
  you think are present in the data.
- In the case of our school example:

The empty model divides the total variance in arithmetic performance into two components:

Level 2: variance of school means (around the grand mean)

Level 1: variance of individual scores within a school

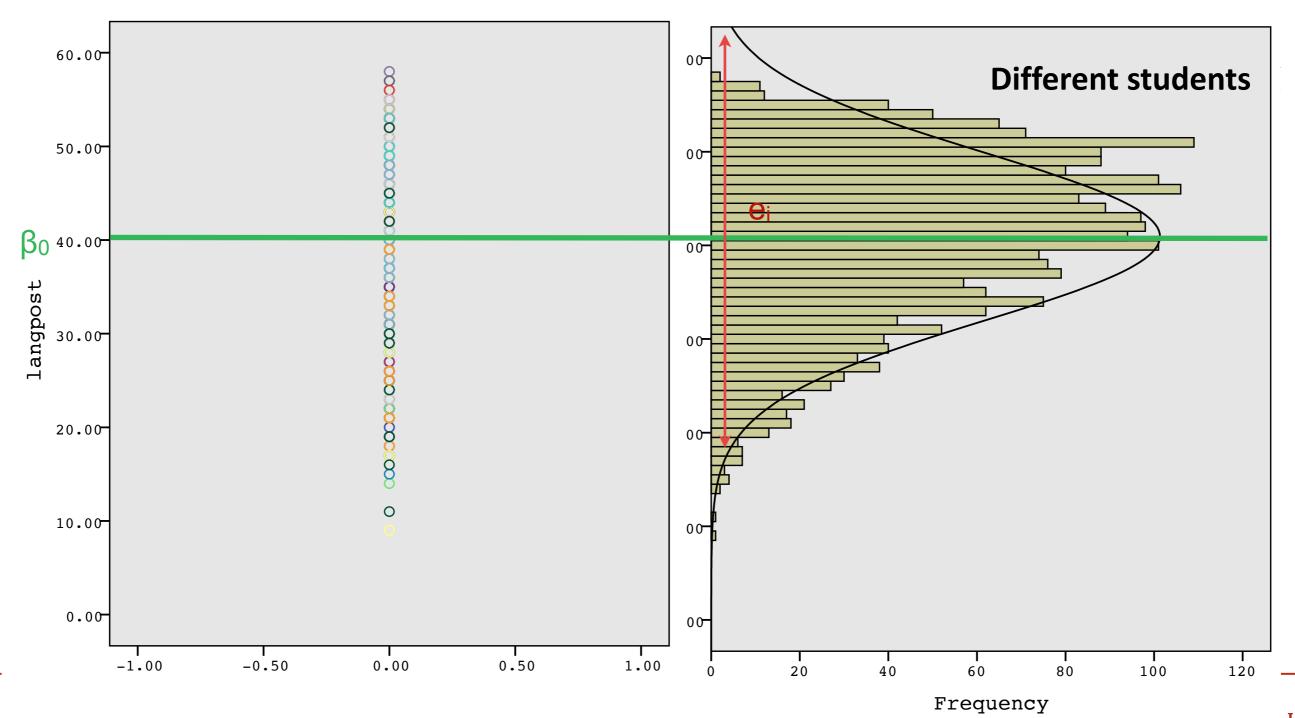
• This is random variance which needs to be explained, by explanatory variables. For instance, the variable teaching style might be able to explain variance between school means (and hence individuals)

Or: 
$$Y_{ij} = \beta_{00} + \beta_1 X_{ij} + \beta_2 \text{ Style}_j + (\epsilon_{ij} + \mathbf{u}_j)$$



# The single level model, no covariate:

$$Y_i = \beta_{0i}X_0 + e_{0i}$$



#### The single level model, no covariate:



Single level model														
	1	2	3	4	5	6	7	8	9	10	11	12	13	
1	1	0	0	0	0	0	0	0	0	0	0	0	0	
2	0	1	0	0	0	0	0	0	0	0	0	0	0	
3	0	0	1	0	0	0	0	0	0	0	0	0	0	
4	0	0	0	1	0	0	0	0	0	0	0	0	0	
5	0	0	0	0	1	0	0	0	0	0	0	0	0	
6	0	0	0	0	0	1	0	0	0	0	0	0	0	
7	0	0	0	0	0	0	1	0	0	0	0	0	0	
8	0	0	0	0	0	0	0	1	0	0	0	0	0	
9	0	0	0	0	0	0	0	0	1	0	0	0	0	
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#### **Correlation matrix V**

### The 'empty' model / Variance components

# **Equations** Note: This is the MLwiN interface

Estimation of one common variance at each level: Distribution of level units around a mean fixed at 0

### Only 3 parameters to be estimated:

 $\beta_0$ : Mean of dependent variable

 $\sigma^2_{u0}$ : Between-group differences

 $\sigma^2_{e0}$ : Within-group differences



#### The 'empty' model / Variance components

L2		1	1	1	1	2	2	3	3	3	3	
	L1	1	2	3	4	1	2	1	2	3	4	
1	1	1	ρ	ρ	ρ	0	0	0	0	0	0	
1	2	ρ	1	ρ	ρ	0	0	0	0	0	0	
1	3	ρ	ρ	1	ρ	0	0	0	0	0	0	
1	4	ρ	ρ	ρ	1	0	0	0	0	0	0	
2	1	0	0	0	0	1	ρ	0	0	0	0	
2	2	0	0	0	0	ρ	1	0	0	0	0	
3	1	0	0	0	0	0	0	1	ρ	ρ	ρ	
3	2	0	0	0	0	0	0	ρ	1	ρ	ρ	
3	3	0	0	0	0	0	0	ρ	ρ	1	ρ	
3	4	0	0	0	0	0	0	ρ	ρ	ρ	1	
	:	:	:	:	:	:	:	:	:	:	:	٠.

Values belonging to the same level (school) can be correlated.

The correlation is the same for each level:  $\rho$ 

Intra Class Correlation (ICC)
Variance Partition Coefficient (VPC)

This is a simple ratio of the variances estimated at each level

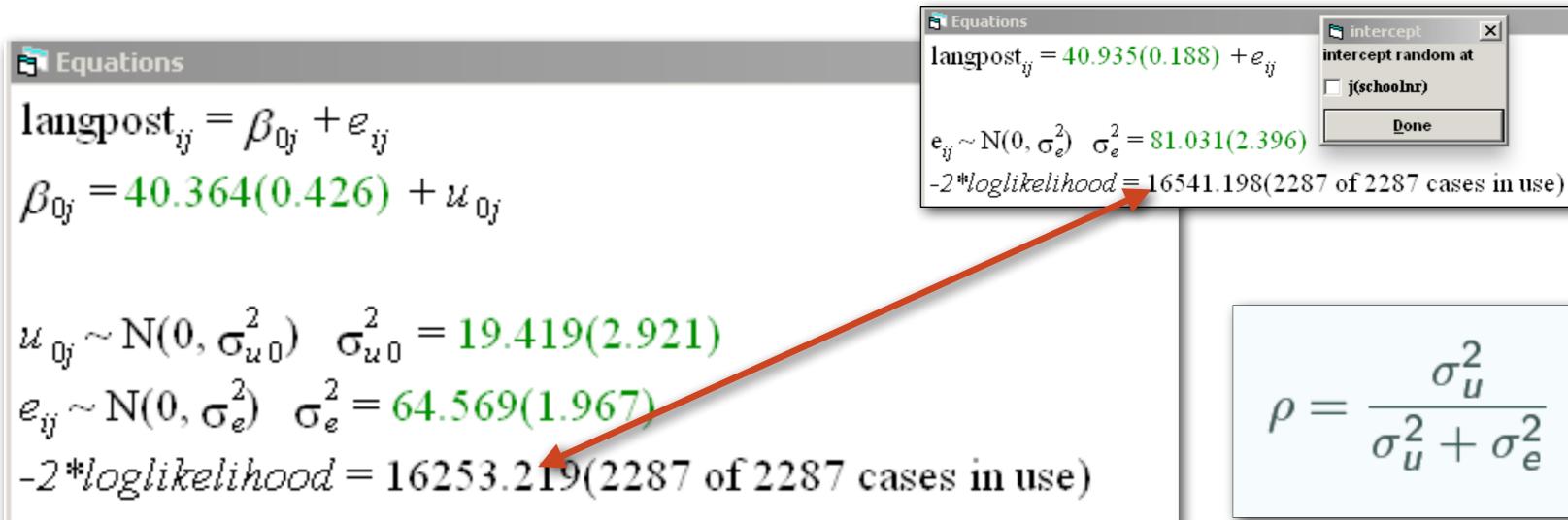
$$\rho = \frac{\sigma_u^2}{\sigma_u^2 + \sigma_e^2}$$

#### **Correlation matrix V**



#### Is a multilevel model necessary?

Variance Components / intercept only model; how much variance exists at the various levels? ICC / VPC. Is a multilevel model necessary?



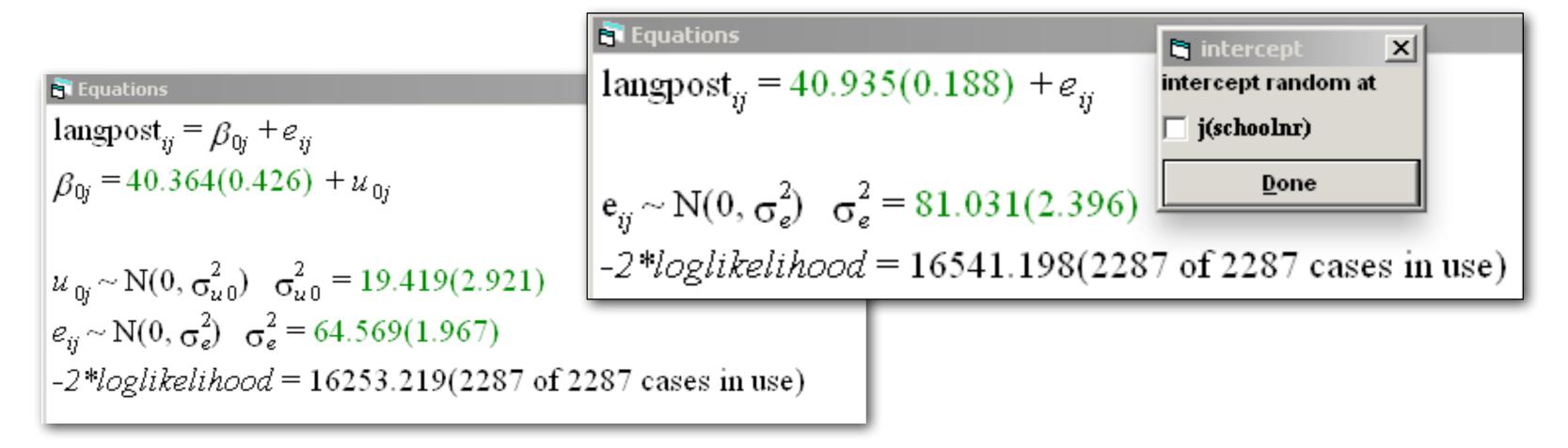
$$\rho = \frac{\sigma_u^2}{\sigma_u^2 + \sigma_e^2}$$

$$ICC = VPC = 19.4 / (64.6 + 19.4) = 0.23$$



#### Is a multilevel model necessary?

Variance Components / intercept only model; how much variance exists at the various levels? ICC / VPC.



**Deviance test for nested models:** 1 level - 2 levels

 $16541.198 - 16253.219 = \chi^{2}(1) = 287.97$  [df = # of extra model parameters]

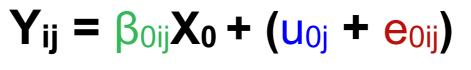


# Variance components model: Estimating random variance "to be explained"

- 'Empty' model: Analyse variance of a dependent variable into two variance components. Level 1: Individuals, Level 2: Groups
- Three parameters:
- Fixed: Grand mean
- Random Level 2: Variance of group means
- Random Level 1: Variance of individuals within groups



# The 'empty' model: 1 mean + 2 random effects



X<sub>0</sub>: A variable containing values of 1 for each case (constant)



$$y_{ij} \sim N(XB, \Omega)$$

$$y_{ij} = \beta_{0ij} x_0$$

$$\beta_{0ij} \neq \beta_0 + u_{0j} + e_{0ij}$$

$$\begin{bmatrix} u_{0j} \end{bmatrix} \sim N(0, \Omega_u) : \Omega_u = \begin{bmatrix} \sigma_{u0}^2 \end{bmatrix}$$

$$\begin{bmatrix} e_{0ij} \end{bmatrix} \sim N(0, \Omega_e) : \Omega_e = \begin{bmatrix} \sigma_{e0}^2 \end{bmatrix}$$

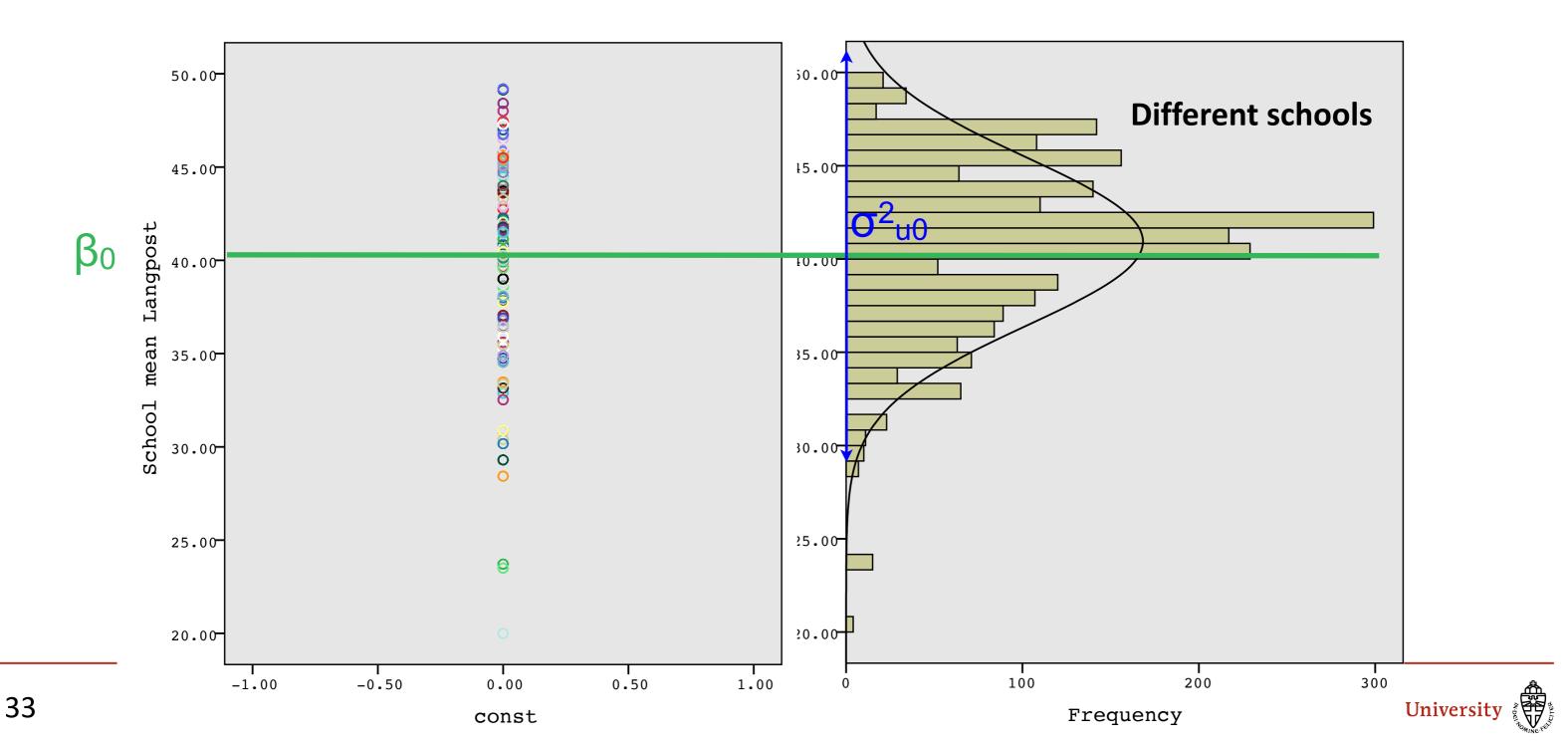
 $\beta_0$ : Mean of dependent variable

 $\sigma^2_{u0}$ : Between-group differences

 $\sigma^2_{e0}$ : Within-group differences

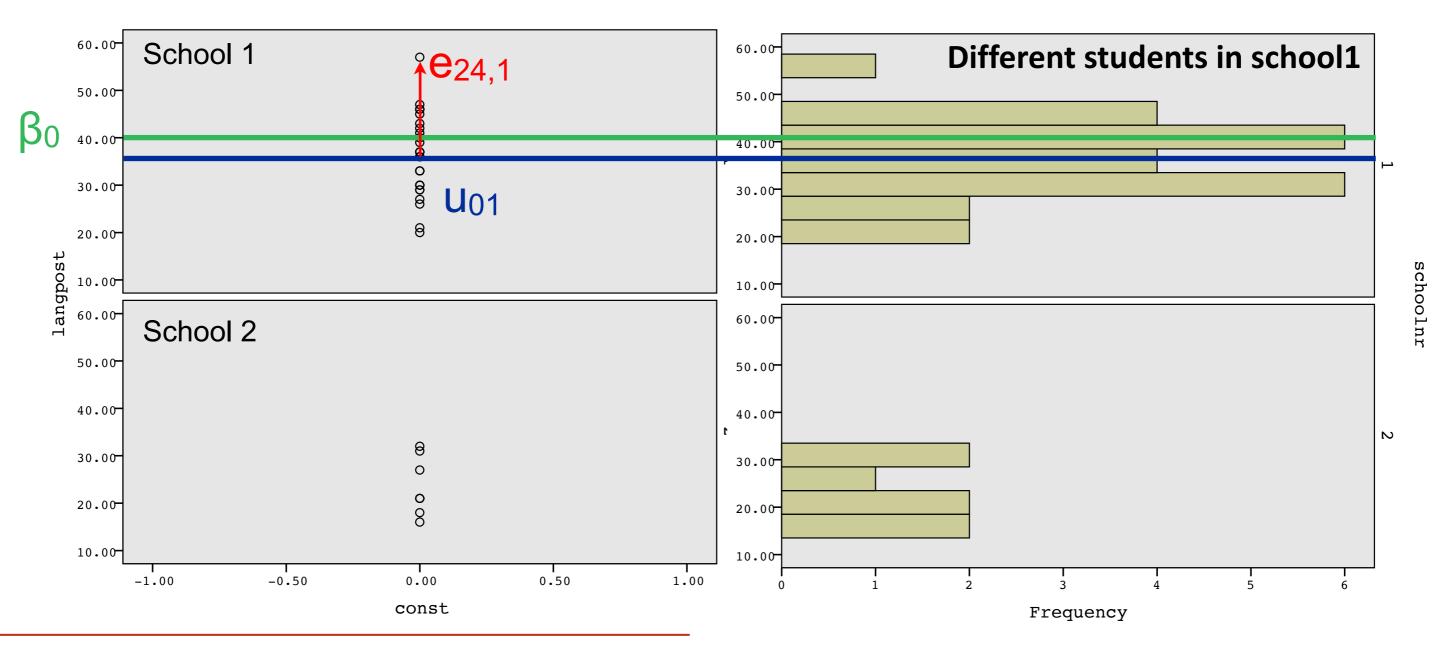
# The Variance components at 2 levels: Between schools

$$Y_{ij} = \beta_{0ij}X_0 + (u_{0j} + e_{0ij})$$

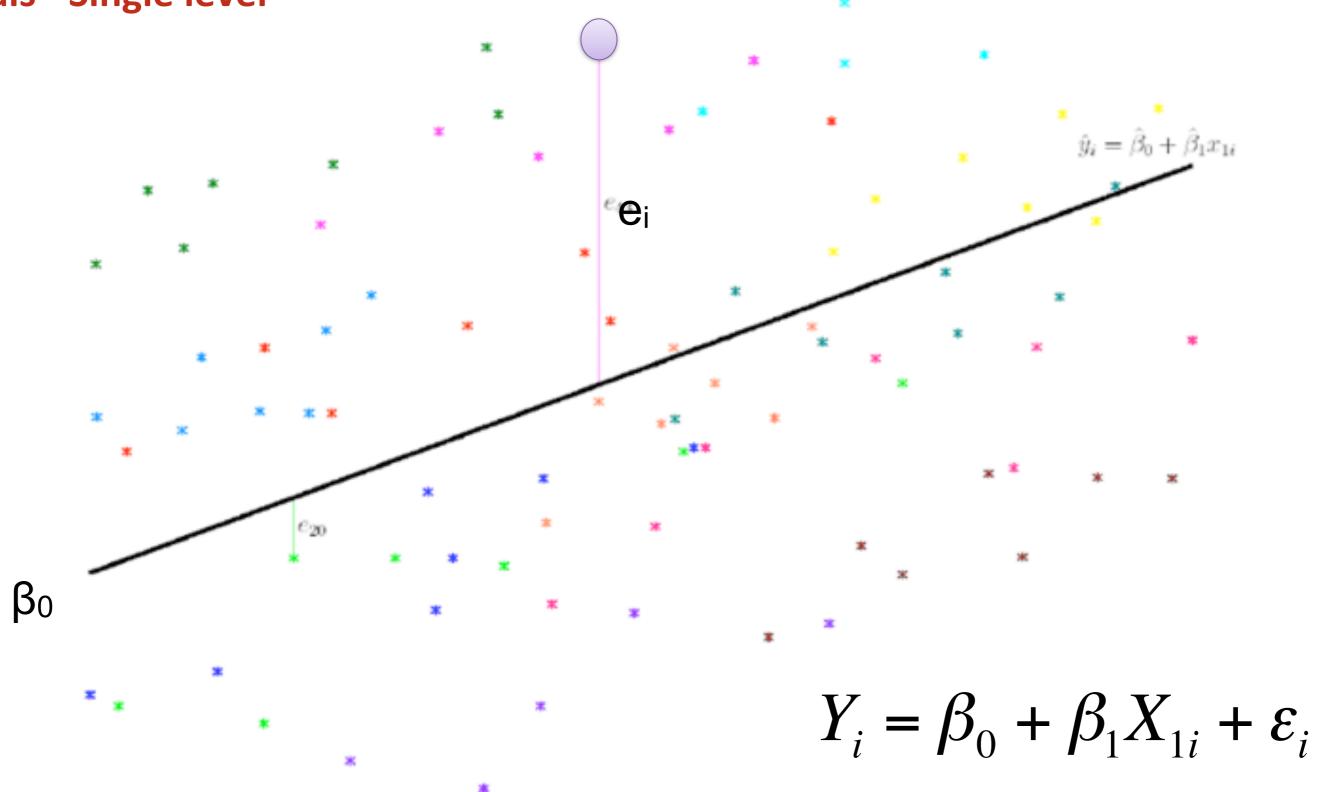


# Variance components at 2 levels: Within schools

$$Y_{ij} = \beta_{0ij}X_0 + (u_{0j} + e_{0ij})$$



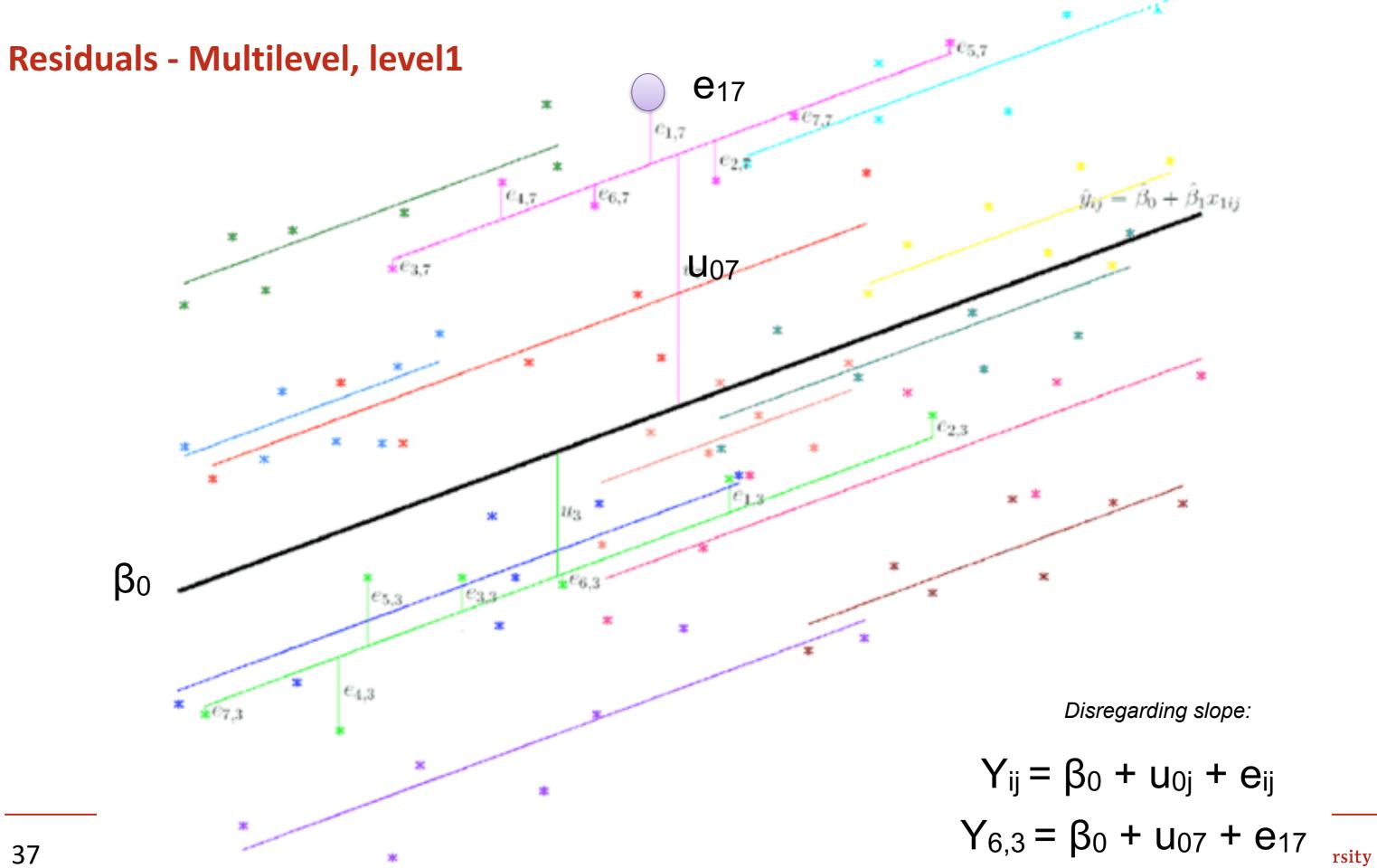
# **Residuals - Single level**





# Residuals - Multilevel, level2 **U**07 $\beta_0 + u_{07}$ $\beta_0$ Disregarding slope: $y_{ij} = \beta_{0j} + e_{ij}$ $\beta_{0j} = \beta_0 + u_{0j}$ 36

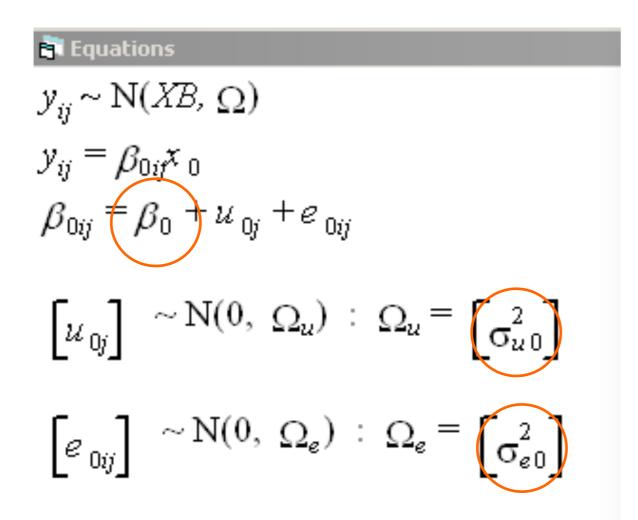






## The 'empty' model: 1 mean + 2 random effects

$$Y_{ij} = \beta_{0ij}X_0 + (u_{0j} + e_{0ij})$$



#### OK, NOW WHAT?

This is "variance to be explained"

Add predictors/covariates: "fixed" effects

See if this reduces variance at different levels Test if it provides a better model-fit

#### Random intercept model, 1 covariate, fixed slope

Random intercepts (for each school)
 plus a covariate with a fixed slope

Compare to ANCOVA: Groups still a random factor

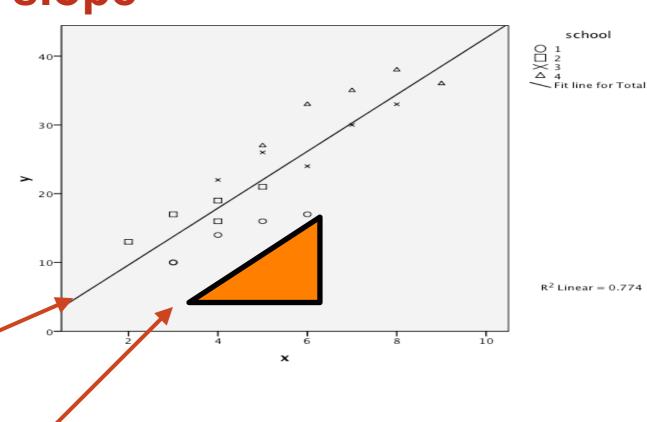
Four parameters:

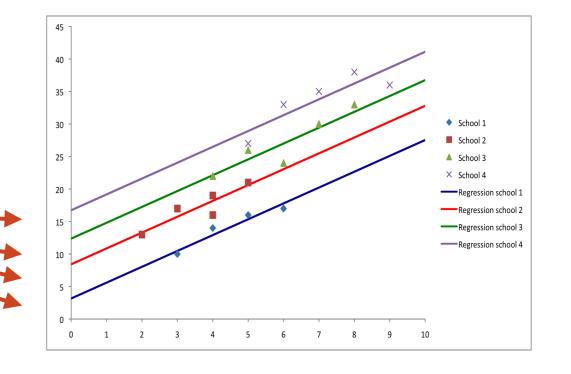
- Fixed: Average intercept

- Fixed: Pooled within-group slope of covariate

- Random Level 2: Variance of intercepts

- Random Level 1: Residual variance within groups







#### Random intercept model, 1 covariate, fixed slope

$$Y_{ij} = \beta_{0ij}X_0 + \beta_1X_{1ij} + (u_{0j} + e_{0ij})$$

# 🚉 Equations

$$y_{ij} \sim N(XB, \Omega)$$

$$y_{ij} = \beta_{0ij}x_0 + \beta_1x_{ij}$$

$$\beta_{0ij} = \beta_0 + u_{0j} + e_{0ij}$$

$$\beta_{0ij} = \beta_0 + u_{0j} + e_{0ij}$$

$$\begin{bmatrix} u_{0j} \end{bmatrix} \sim N(0, \Omega_u) : \Omega_u = \begin{bmatrix} \sigma_{u0}^2 \end{bmatrix}$$

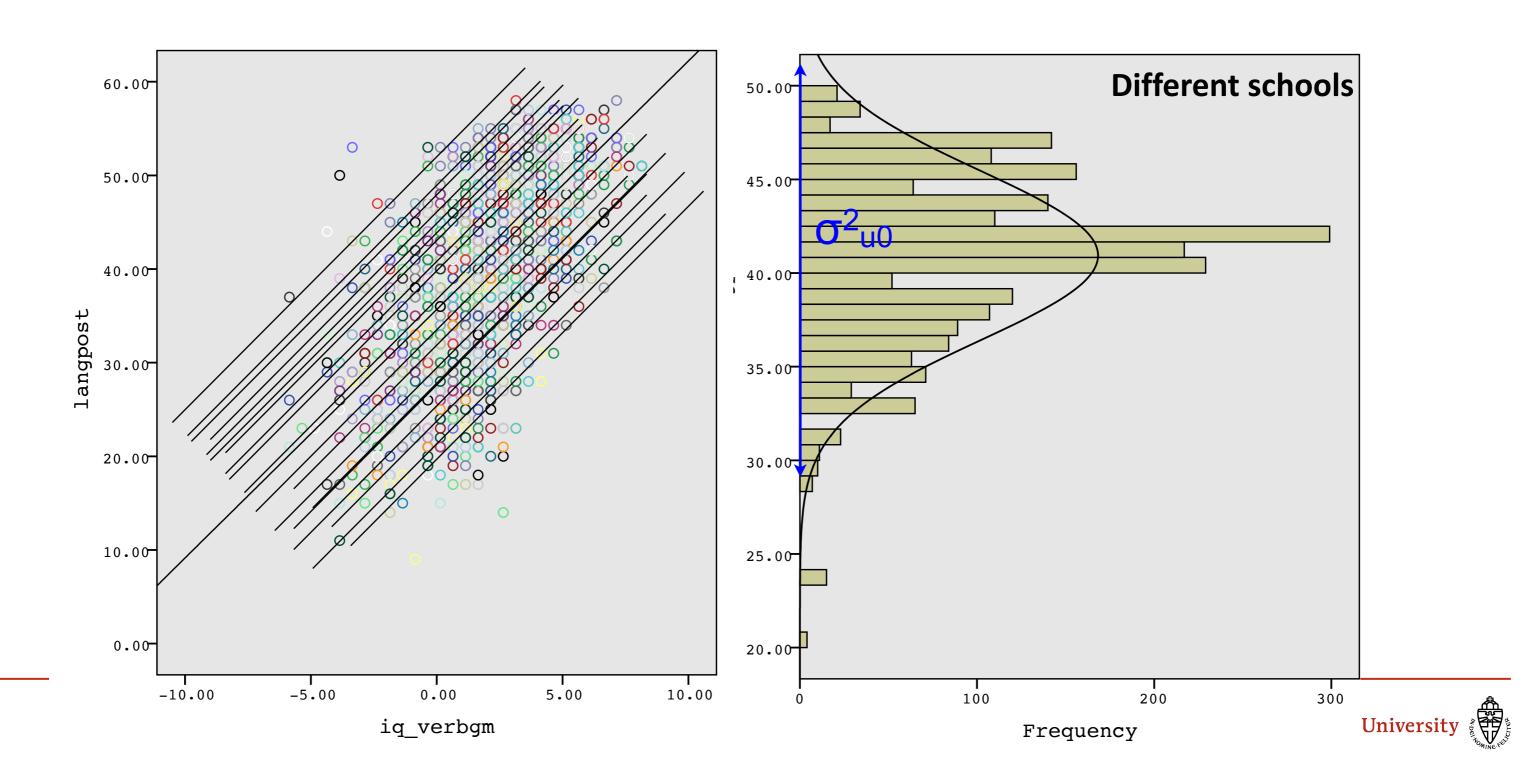
$$\begin{bmatrix} e_{0ij} \end{bmatrix} \sim N(0, \Omega_e) : \Omega_e = \begin{bmatrix} \sigma_{e0}^2 \end{bmatrix}$$

β<sub>1</sub>: Fixed slope

# Random intercept model, 1 covariate, fixed slope

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$$Y_{ij} = \beta_{0ij}X_0 + \beta_1X_{1ij} + (u_{0j} + e_{0ij})$$



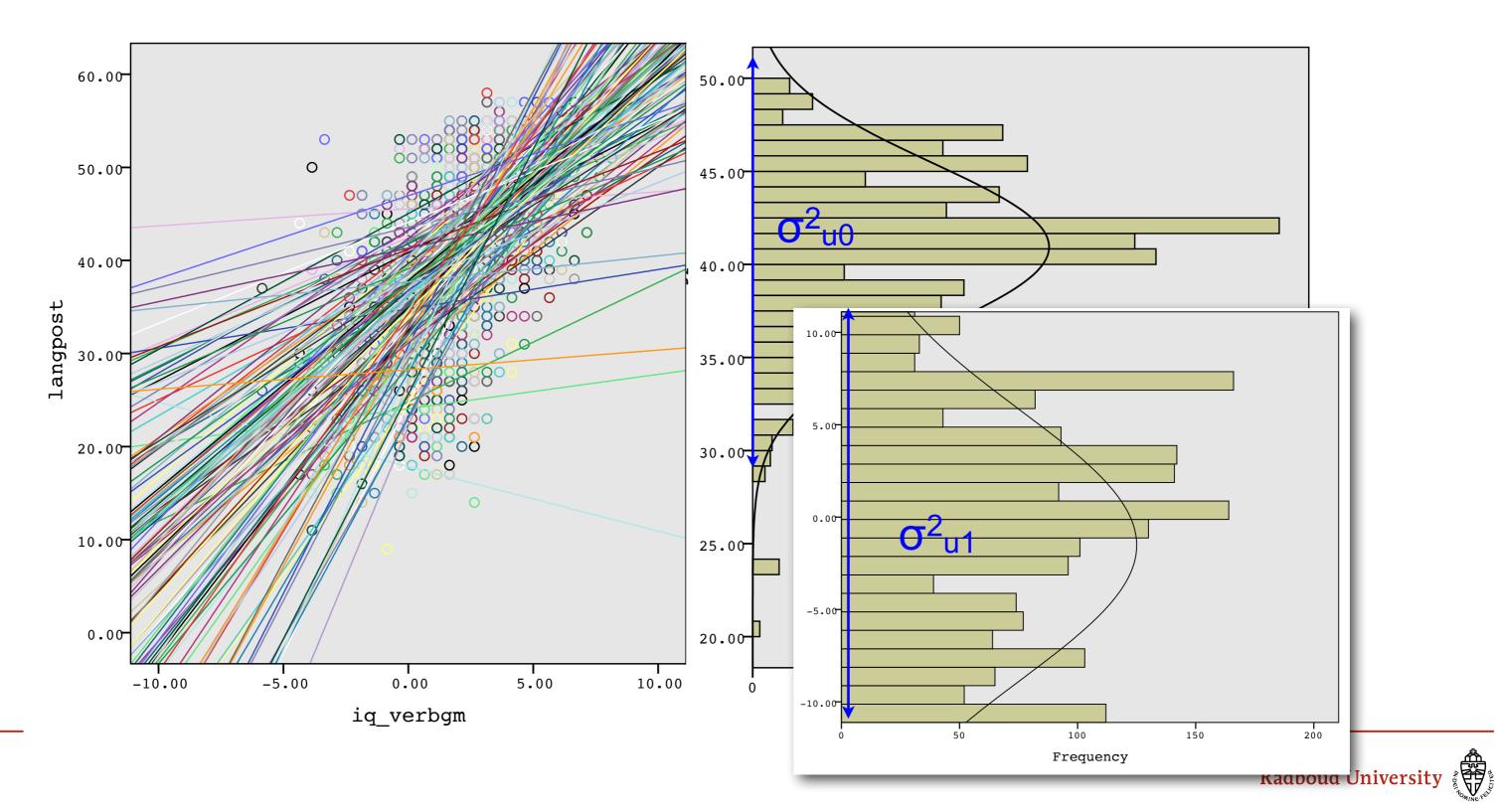
#### Random intercept, 1 covariate, random slope

- Random intercepts plus a covariate with a random slope
- Compare to heterogeneous (non-parallel) regression (or factor covariate interaction)
- Six parameters:
- Fixed: Average intercept
- Fixed: Average pooled within-group slope of covariate
- Random Level 2: Variance of intercepts
- Random Level 2: Variance of slopes
- Random Level 2: Intercept-slope covariance
- Random Level 1: Residual variance within groups



# Random intercept, 1 covariate, random slope

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#### 👸 Equations

$$y_{ij} \sim N(XB, \Omega)$$

$$y_{ij} = \beta_{0ij}x_0 + \beta_{1j}x_{1j}$$

$$\beta_{0ij} = \beta_0 + u_{0j} + e_{0ij}$$

$$\beta_{1j} = \beta_1 + u_{1j}$$

 $\beta_{1j}$ : Mean pooled within-group slope

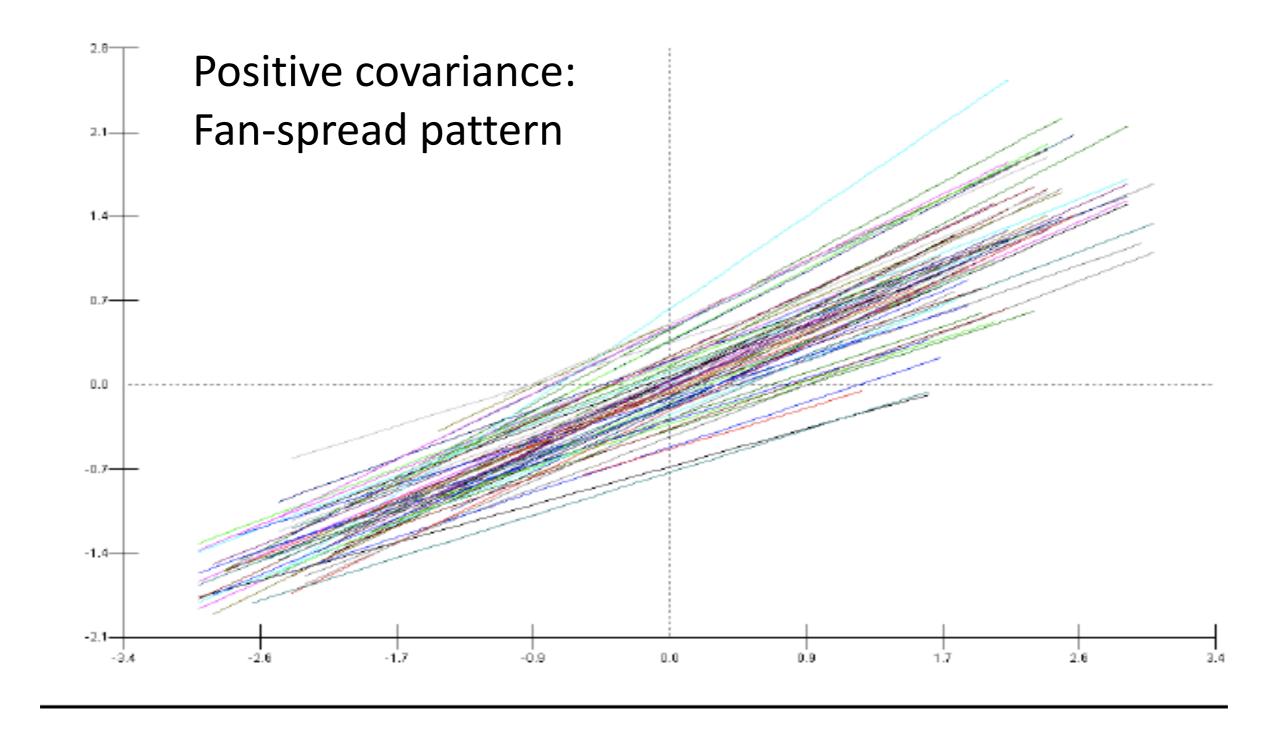
$$\begin{bmatrix} u_{0j} \\ u_{1j} \end{bmatrix} \sim \mathbf{N}(0, \Omega_u) : \Omega_u = \begin{bmatrix} \sigma_{u0}^2 \\ \sigma_{u01} \end{bmatrix}$$

$$\begin{bmatrix} e_{0ij} \end{bmatrix} \sim N(0, \Omega_e) : \Omega_e = \begin{bmatrix} \sigma_{e0}^2 \end{bmatrix}$$

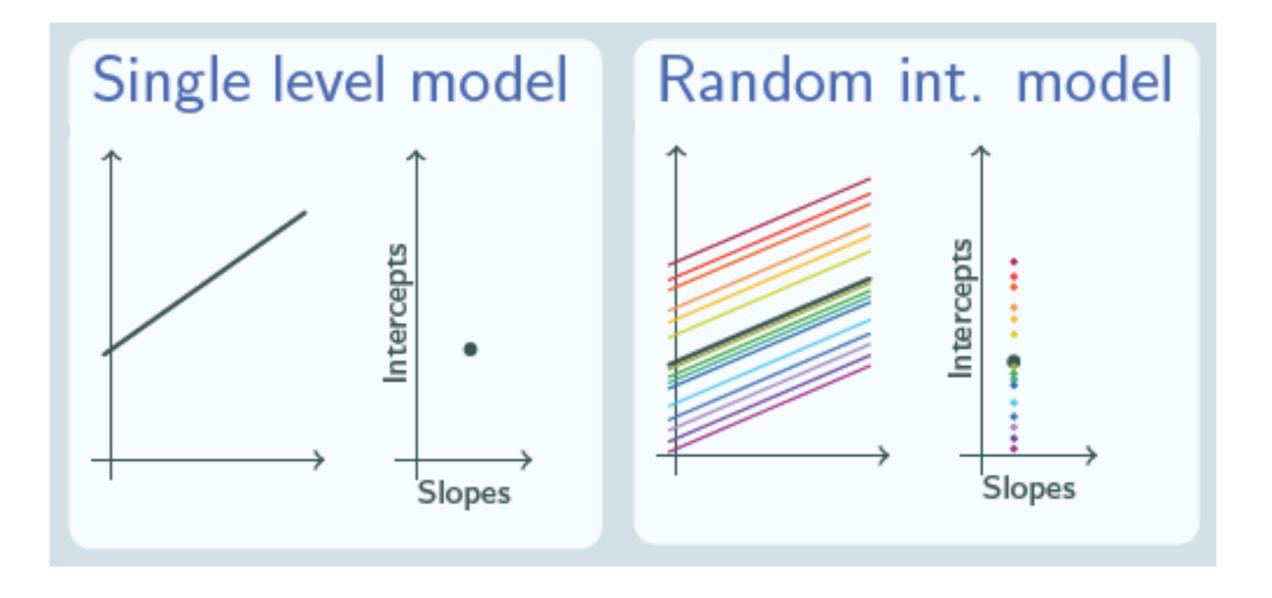
 $\sigma^2_{u1}$ : Variance of slopes

 $\sigma^2_{u01}$ : Intercept-slope covariance

#### **Intercept-Slope covariance**

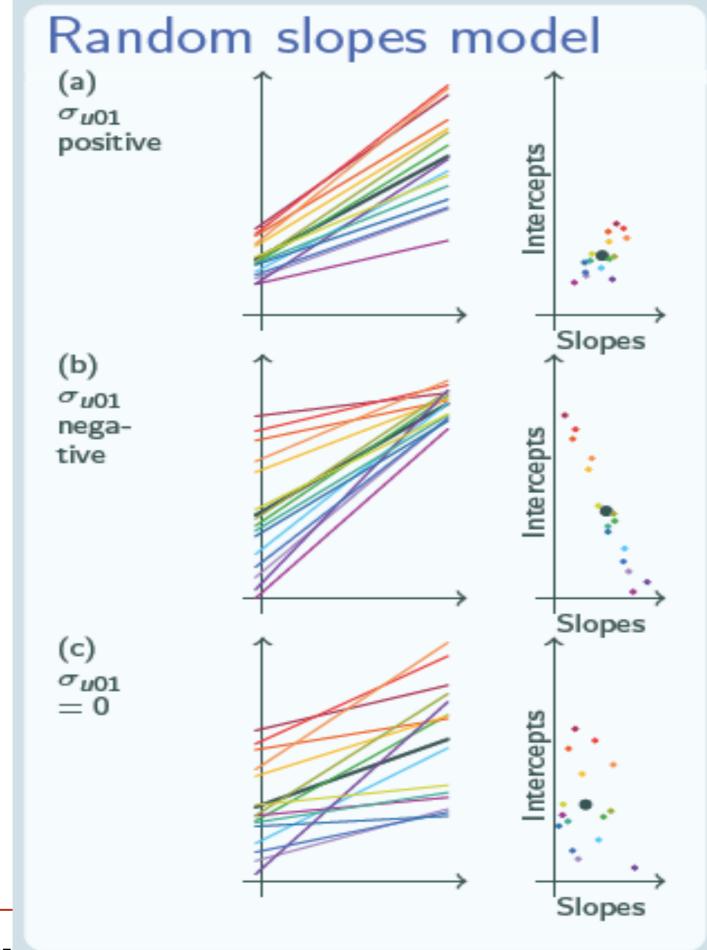


#### Interpret $\sigma_{u0}^2$ , $\sigma_{u1}^2$ and $\sigma_{u01}$ together



For single level or random intercept models,  $\sigma_{u01}$  is not defined (there is no variation in slopes)





For random slope models,

- $\sigma_{u01}$  positive means a pattern of fanning out
- $\sigma_{u01}$  negative means a pattern of fanning in
- $\sigma_{u01} = 0$  means no pattern

#### **Residuals**

In multilevel models, residuals exist at every level

random-intercept model:

$$\begin{split} Y_{ij} &= \beta_{0j} + \beta_1 X_{ij} + \epsilon_{ij} \\ &= (\beta_0 + u_{0j}) + \beta_1 X_{ij} + \epsilon_{ij} \\ &= (\beta_0 + \beta_1 X_{ij}) + u_{0j} + \epsilon_{ij} \end{split}$$

= predicted value + level-2 residual + level-1 residual

#### Reasiduals: Caterpillar plot

