



Least squares
identification

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Least squares identification

For white noise

Least square
estimation

Model order
increasing
algorithm

recursive least
square

Problem
discussion

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The difference equation of a SISO system is

$$y_k + a_1 y_{k-1} + \cdots + a_n y_{k-n} = b_0 u_k + \cdots + b_n u_{k-n} + \xi_k$$

at time $k = n + 1, n + 2, \cdots, n + N$, there is

$$y_{n+1} + a_1 y_n + \cdots + a_n y_1 = b_0 u_{n+1} + \cdots + b_n u_1 + \xi_{n+1}$$

$$y_{n+2} + a_1 y_{n+1} + \cdots + a_n y_2 = b_0 u_{n+2} + \cdots + b_n u_2 + \xi_{n+2}$$

...

$$y_{n+N} + a_1 y_{n+N-1} + \cdots + a_n y_N = b_0 u_{n+N} + \cdots + b_n u_N + \xi_{n+N}$$



$$Y = \Phi\theta + \xi$$

$$Y = [y_{n+1} \quad y_{n+2} \quad \cdots \quad y_{n+N}]^T$$

$$\Phi = \begin{bmatrix} -y_n & \cdots & -y_1 & u_{n+1} & \cdots & u_1 \\ -y_{n+1} & \cdots & -y_2 & u_{n+2} & \cdots & u_2 \\ \vdots & & \vdots & \vdots & & \vdots \\ -y_{n+N-1} & \cdots & -y_N & u_{n+N} & \cdots & u_N \end{bmatrix}$$

$$\theta = [a_1 \quad \cdots \quad a_n \quad b_0 \quad \cdots \quad b_n]^T$$

$$\xi = [\xi_{n+1} \quad \xi_{n+2} \quad \cdots \quad \xi_{n+N}]^T$$



Identification criterion: least square sum of residuals.

$$\begin{aligned} J &= \sum_{k=n+1}^{n+N} e^2(k) \\ &= (Y - \Phi \hat{\theta})^T (Y - \Phi \hat{\theta}) \\ \hat{\theta}_{LS} &= \arg \min_{\hat{\theta}} J \end{aligned}$$



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$$\begin{aligned}\frac{\partial J}{\partial \hat{\theta}_k} &= \frac{\partial \sum_i (Y_i - \sum_m \Phi_{i,m} \hat{\theta}_m)^2}{\partial \hat{\theta}_k} \\&= 2 \sum_i (Y_i - \sum_m \Phi_{i,m} \hat{\theta}_m) \frac{\partial (Y_i - \sum_m \Phi_{i,m} \hat{\theta}_m)}{\partial \hat{\theta}_k} \\&= 2 \sum_i (Y_i - \sum_m \Phi_{i,m} \hat{\theta}_m) \frac{\partial (-\sum_m \Phi_{i,m} \hat{\theta}_m)}{\partial \hat{\theta}_k} \\&= -2 \sum_i (Y_i - \sum_m \Phi_{i,m} \hat{\theta}_m) \Phi_{i,k} \\ \frac{\partial J}{\partial \hat{\theta}} &= (-2(Y - \Phi \hat{\theta})^T \Phi)^T \\&= -2\Phi^T (Y - \Phi \hat{\theta})\end{aligned}$$

Basic least square method: solution



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$$-2\Phi^T(Y - \Phi\hat{\theta}_{LS}) = 0$$

$$\Phi^T Y - \Phi^T \Phi \hat{\theta}_{LS} = 0$$

$$\Phi^T Y = \Phi^T \Phi \hat{\theta}_{LS}$$

$$\hat{\theta}_{LS} = (\Phi^T \Phi)^{-1} \Phi^T Y$$

Basic least square method: two order derivative



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$$\begin{aligned}\frac{\partial^2 J}{\partial \hat{\theta}^2} &= \frac{\partial(-2\Phi^T(Y - \Phi\hat{\theta}))}{\partial \hat{\theta}} \\ \frac{\partial \frac{\partial J}{\partial \hat{\theta}}}{\partial \hat{\theta}_s} &= \frac{\partial(-2 \sum_i (Y_i - \sum_m \Phi_{i,m} \hat{\theta}_m) \Phi_{i,k})}{\partial \hat{\theta}_s} \\ &= 2 \sum_i \frac{\partial \sum_m \Phi_{i,m} \hat{\theta}_m}{\partial \hat{\theta}_s} \Phi_{i,k} \\ &= 2 \sum_i \Phi_{i,s} \Phi_{i,k} \\ \frac{\partial^2 J}{\partial \hat{\theta}^2} &= 2\Phi^T \Phi\end{aligned}$$

The requirement of input signal by least square method :
 $[Y_{N \times n} \quad U_{N \times (n+1)}]$



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where:

$$Y_{N \times n} = \begin{bmatrix} -y_n & \cdots & -y_1 \\ -y_{n+1} & \cdots & -y_2 \\ \vdots & & \vdots \\ y_{n+N-1} & \cdots & -y_N \end{bmatrix}$$

$$U_{N \times (n+1)} = \begin{bmatrix} u_{n+1} & \cdots & u_1 \\ u_{n+2} & \cdots & u_2 \\ \vdots & & \vdots \\ u_{n+N} & \cdots & u_N \end{bmatrix}$$

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The requirement of input signal by least square method

$$\begin{bmatrix} Y_{N \times n} & U_{N \times (n+1)} \end{bmatrix}$$



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$$(Y_{N \times n}^T Y_{N \times n})_{i,j} = \sum_{k=1}^{N-1+\min\{i,j\}} y_{n-i+k} y_{n-j+k}$$

$$(Y_{N \times n}^T U_{N \times (n+1)})_{i,j} = - \sum_{k=1}^{N-1+\min\{i,j-1\}} y_{n-i+k} u_{n+1-j+k}$$

$$(U_{N \times (n+1)}^T Y_{N \times n}^T)_{i,j} = - \sum_{k=1}^{N-1+\min\{j,i-1\}} y_{n-j+k} u_{n+1-i+k}$$

$$(U_{N \times (n+1)}^T U_{N \times (n+1)})_{i,j} = \sum_{k=1}^{N-2+\min\{i,j\}} u_{n+1-i+k} u_{n+1-j+k}$$

The requirement of input signal by least square method

$$\begin{bmatrix} R_y & R_{yu} \\ R_{uy} & R_u \end{bmatrix}$$

$$\begin{aligned} \lim_{N \rightarrow \infty} \frac{\Phi^T \Phi}{N} &= \frac{1}{N} \begin{bmatrix} Y_{N \times n}^T Y_{N \times n} & Y_{N \times n}^T U_{N \times (n+1)} \\ U_{N \times (n+1)}^T Y_{N \times n} & U_{N \times (n+1)}^T U_{N \times (n+1)} \end{bmatrix} \\ &= \begin{bmatrix} R_y & R_{yu} \\ R_{uy} & R_u \end{bmatrix} \end{aligned}$$

where:

$$\begin{aligned} R_y &= \begin{bmatrix} R_y(0) & R_y(1) & \cdots & R_y(n-1) \\ R_y(1) & R_y(0) & \cdots & R_y(n-2) \\ \vdots & \vdots & & \vdots \\ R_y(n-1) & R_y(n-2) & \cdots & R_y(0) \end{bmatrix} \\ R_{yu} &= \begin{bmatrix} -R_{yu}(1) & -R_{yu}(0) & \cdots & -R_{yu}(1-n) \\ -R_{yu}(2) & -R_{yu}(1) & \cdots & -R_{yu}(2-n) \\ \vdots & \vdots & & \vdots \\ -R_{yu}(n) & -R_{yu}(n-1) & \cdots & -R_{yu}(0) \end{bmatrix} \end{aligned}$$



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The requirement of input signal by least square method:

$$\begin{bmatrix} R_y & R_{yu} \\ R_{uy} & R_u \end{bmatrix}$$



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$$\begin{aligned} R_{uy} &= R_{yu}^T \\ R_{uu} &= \begin{bmatrix} R_u(0) & R_u(1) & \cdots & R_u(n) \\ R_u(1) & R_u(0) & \cdots & R_u(n-1) \\ \vdots & \vdots & & \vdots \\ R_u(n) & R_u(n-1) & \cdots & R_u(0) \end{bmatrix} \end{aligned}$$

$(n+1)$ order continuous excitation signal



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- definition: $\{u(k)\}$ is called $(n + 1)$ order continuous excitation signal if $(n+1)$ order matrix R_u of series $\{u(k)\}$ is positive definite, ◦
- The requirement of least square method for input signal is: $\{u(k)\}$ is $(n + 1)$ order continuous excitation signal
- R_u is positive definite if R_u is an strongly diagonally dominant matrix. The following signals can satisfy the requirement of positive definite of R_u .
 - White noise sequence ;
 - Pseudo random two bit noise sequence ;
 - Colored noise random signal sequence ◦
- "Pseudo random two bit noise sequence" and "colored noise random signal sequence" are often used as input signals in Engineering ◦



There are four main aspects of the probability property of least squares estimation,

- unbiasedness of the estimation ;
- Consistency of estimates ;
- Validity of estimation ;
- Asymptotic normality of estimators .



$\hat{\theta}$ is referred as unbiased estimation of parameter θ if $E\{\hat{\theta}\} = \theta$.

$$\begin{aligned} Y &= \Phi\theta + \xi \\ \hat{\theta} &= (\Phi^T\Phi)^{-1}\Phi^TY \\ E[\hat{\theta}] &= E[(\Phi^T\Phi)^{-1}\Phi^TY] \\ &= E[(\Phi^T\Phi)^{-1}\Phi^T(\Phi\theta + \xi)] \\ &= E[(\Phi^T\Phi)^{-1}\Phi^T\Phi\theta + (\Phi^T\Phi)^{-1}\Phi^T\xi] \\ &= E[\theta + (\Phi^T\Phi)^{-1}\Phi^T\xi] \end{aligned}$$

The necessary and sufficient conditions of least squares estimation for unbiased estimation is:

$$E[(\Phi^T\Phi)^{-1}\Phi^T\xi] = 0$$

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$$\lim_{N \rightarrow \infty} P\{|\hat{\theta} - \theta\} = 1$$

Suppose $\xi\{(k)\}$ is random sequence with zero mean and independent distribution uncorrelated with $\{u(k)\}$:

$$\begin{aligned} E(\hat{\theta} - \theta)^2 &= E[(\Phi^T \Phi)^{-1} \Phi^T \xi \xi^T \Phi (\Phi^T \Phi)^{-1}] \\ &= E\left[\frac{1}{N^2} \left(\frac{1}{N} \Phi^T \Phi\right)^{-1} \Phi^T \xi \xi^T \Phi \left(\frac{1}{N} \Phi^T \Phi\right)^{-1}\right] \\ \lim_{N \rightarrow \infty} E(\hat{\theta} - \theta)^2 &= \frac{1}{N^2} R^{-1} E[\Phi^T \xi \xi^T \Phi] R^{-1} \\ &= \frac{1}{N^2} R^{-1} \sigma^2 E[\Phi^T \Phi] R^{-1} \\ &= \frac{1}{N^2} R^{-1} \sigma^2 N R R^{-1} \\ &= \frac{\sigma^2}{N} R^{-1} \\ &= 0 \end{aligned}$$

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Cramer-Rao inequality:

$$D\hat{\theta} = M^{-1}$$

其中:

$$M = E \left[\left(\frac{\partial \ln p(y|\theta)}{\partial \theta} \right)^T \left(\frac{\partial \ln p(y|\theta)}{\partial \theta} \right) \right]$$

The validity of the estimation



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$$y = \Phi\theta + \xi$$

$$y \sim N(\Phi\theta, \sigma^2 I)$$

$$p(y|\theta) = (2\pi\sigma^2)^{-\frac{N}{2}} \exp \left[-\frac{1}{2\sigma^2} (y - \Phi\theta)^T (y - \Phi\theta) \right]$$

$$\frac{\partial \ln p(y|\theta)}{\partial \theta} = -\frac{1}{\sigma^2} (y - \Phi\theta)^T \Phi$$

$$M = E \left[\frac{1}{\sigma^4} \Phi^T (y - \Phi\theta) (y - \Phi\theta)^T \Phi \right]$$

$$= \frac{1}{\sigma^4} E[\Phi^T \xi \xi^T \Phi]$$

$$\lim_{N \rightarrow \infty} M^{-1} = \sigma^4 (\sigma^2 E[\Phi^T \Phi])^{-1}$$

$$= \frac{\sigma^2}{N} R^{-1}$$



Suppose $\{\xi(k)\}$ is white noise sequence with zero mean and normal distribution. then:

$$y = \Phi\theta + \xi$$

$$y \sim N(\Phi\theta, \sigma^2 I)$$

$$\hat{\theta} = (\Phi^T \Phi)^{-1} \Phi^T y$$



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- recursive algorithm based on model order n ;
- suitable for unknown model order n
- The identification accuracy is the same as that of the basic least square method
- The identification speed is greatly improved than the basic least square method
- It is not necessary to compute the inverse of higher order matrices



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$$Y = \Phi_n \theta_n + \xi$$

$$\Phi_n = \begin{bmatrix} u_{n+1} & -y_n & u_n & \cdots & -y_1 & u_1 \\ u_{n+2} & -y_{n+1} & u_{n+1} & \cdots & -y_2 & u_2 \\ \vdots & \vdots & \vdots & & \vdots & \vdots \\ u_{n+N} & -y_{n+N-1} & u_{n+N-1} & \cdots & -y_N & u_N \end{bmatrix}$$

$$= \begin{bmatrix} X_1 & \cdots & X_{2n+1} \end{bmatrix}$$

$$\theta_n = \begin{bmatrix} b_0 & a_1 & b_1 & \cdots & a_n & b_n \end{bmatrix}^T$$

$$\xi = \begin{bmatrix} \xi_{n+1} & \cdots & \xi_{n+N} \end{bmatrix}^T$$

$$Y = \begin{bmatrix} y_{n+1} & \cdots & y_{n+N} \end{bmatrix}^T$$



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$$\begin{aligned}\Phi_0 &= X_1 \\ \hat{\theta}_0 &= (\Phi_0^T \Phi_0)^{-1} \Phi_0^T Y \\ &= \frac{\sum_{i=n+1}^{n+N} u_i y_i}{\sum_{i=n+1}^{n+N} u_i^2}\end{aligned}$$

from n to $n + 1$



Identification result of model order $n + 1$ is obtained based on result of model order n . The solution is divided into two steps. First \tilde{P}_n is solved, and then P_{n+1} is solved.

$$\begin{aligned}\Phi_{n+1} &= \begin{bmatrix} \Phi_n & X_{2n+2} & X_{2n+3} \end{bmatrix} \\ &= \begin{bmatrix} \tilde{\Phi}_n & X_{2n+3} \end{bmatrix} \\ \tilde{\Phi}_n &\triangleq \begin{bmatrix} \Phi_n & X_{2n+2} \end{bmatrix} \\ P_n &\triangleq (\Phi_n^T \Phi_n)^{-1} \\ \tilde{P}_n &\triangleq (\tilde{\Phi}_n^T \tilde{\Phi}_n)^{-1} \\ P_{n+1} &= (\Phi_{n+1}^T \Phi_{n+1})^{-1}\end{aligned}$$

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from n to $n + 1$: P_{n+1}



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$$\begin{aligned}P_{n+1} &= \begin{bmatrix} \tilde{\Phi}_n^T \tilde{\Phi}_n & \tilde{\Phi}_n^T X_{2n+3} \\ X_{2n+3}^T \tilde{\Phi}_n & X_{2n+3}^T X_{2n+3} \end{bmatrix}^{-1} \\&= \begin{bmatrix} A_{11} & A_{12} \\ A_{21} & A_{22} \end{bmatrix} \\A_{22} &= (X_{2n+3}^T X_{2n+3} - X_{2n+3}^T \tilde{\Phi}_n \tilde{P}_n \tilde{\Phi}_n^T X_{2n+3})^{-1} \\A_{12} &= A_{21}^T \\&= -\tilde{P}_n \tilde{\Phi}_n^T X_{2n+3} A_{22} \\A_{11} &= \tilde{P}_n - A_{12} X_{2n+3}^T \tilde{\Phi}_n \tilde{P}_n^T\end{aligned}$$

from n to $n + 1$: \tilde{P}_n



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$$\begin{aligned}\tilde{P}_n &= \begin{bmatrix} \Phi_n^T \Phi_n & \Phi_n^T X_{2n+2} \\ X_{2n+2}^T \Phi_n & X_{2n+2}^T X_{2n+2} \end{bmatrix}^{-1} \\ &= \begin{bmatrix} B_{11} & B_{12} \\ B_{21} & B_{22} \end{bmatrix} \\ B_{22} &= (X_{2n+2}^T X_{2n+2} - X_{2n+2}^T \Phi_n P_n \Phi_n^T X_{2n+2})^{-1} \\ B_{12} &= B_{21}^T \\ &= -P_n \Phi_n^T X_{2n+2} B_{22} \\ B_{11} &= P_n - B_{12} X_{2n+2}^T \Phi_n P_n^T\end{aligned}$$



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- Initialize, compute $P_0 = (\Phi_0^T \Phi_0)^{-1}$
- Compute $\hat{\theta}_0 = P_0 \Phi_0^T Y$
- iterate
 - Compute \tilde{P}_n based on P_n
 - Compute P_{n+1} based on \tilde{P}_n
 - Compute $\hat{\theta}_{n+1} = P_{n+1} \Phi_{n+1}^T Y$

Recursive algorithm derivation: Model



Assuming that the input and output data with length of N have been obtained, the least squares estimation is

$$\begin{aligned}Y_N &= \Phi_N \theta + \xi_N \\ \hat{\theta}_N &= (\Phi_N^T \Phi_N)^{-1} \Phi_N^T Y_N \\ \tilde{\theta}_N &= \theta - \hat{\theta}_N \\ &= -(\Phi_N^T \Phi_N)^{-1} \Phi_N^T \xi_N\end{aligned}$$

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After obtaining new data u_{n+N+1}, y_{n+N+1} ,

$$\begin{aligned}y_{(n+N+1)} &= \Psi^T \theta + \xi_{(n+N+1)} \\ y_{N+1} &= \Psi^T \theta + \xi_{N+1} \\ \Psi_i &= [-y_{(n+i-1)} \quad \cdots \quad -y_{(i)} \quad u_{(n+i)} \quad \cdots \quad u_{(i)}]^T \\ \begin{bmatrix} Y_N \\ y_{N+1} \end{bmatrix} &= \begin{bmatrix} \Phi_N \\ \Psi_{N+1}^T \end{bmatrix} \theta + \begin{bmatrix} \xi_N \\ \xi_{N+1} \end{bmatrix}\end{aligned}$$

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Recursive algorithm derivation : P_{N+1}

$$\begin{aligned}\hat{\theta}_{N+1} &= \left(\begin{bmatrix} \Phi_N \\ \Psi_{N+1}^T \end{bmatrix}^T \begin{bmatrix} \Phi_N \\ \Psi_{N+1}^T \end{bmatrix} \right)^{-1} \begin{bmatrix} \Phi_N \\ \Psi_{N+1}^T \end{bmatrix}^T \begin{bmatrix} Y_N \\ y_{N+1} \end{bmatrix} \\ &= (\Phi_N^T \underbrace{\Phi_N}_{N,2n+1} + \underbrace{\Psi_{N+1} \Psi_{N+1}^T}_{2n+1,1})^{-1} (\Phi_N^T \underbrace{Y_N}_{N,1} + \Psi_{N+1} \underbrace{y_{N+1}}_{1,1}) \\ \hat{\theta}_{N+1} &= P_{N+1} (\Phi_N^T Y_N + \Psi_{N+1} y_{N+1})\end{aligned}$$

其中：

$$\begin{aligned}P_{N+1} &= (P_N^{-1} + \Psi_{N+1} \Psi_{N+1}^T)^{-1} \\ P_N &= (\Phi_N^T \Phi_N)^{-1}\end{aligned}$$



If the inverse of the corresponding matrix exists , then:

$$(A + BC^T)^{-1} = A^{-1} - A^{-1}B(I + C^T A^{-1}B)^{-1}C^T A^{-1}$$

therefore:

$$\begin{aligned}P_{N+1} &= (P_N^{-1} + \Psi_{N+1}\Psi_{N+1}^T)^{-1} \\&= P_N - P_N\Psi_{N+1}(1 + \Psi_{N+1}^T P_N \Psi_{N+1})^{-1}\Psi_{N+1}^T P_N \\ \hat{\theta}_{N+1} &= A + B \\ A &= P_{N+1}\Phi_N^T Y_N \\ B &= P_{N+1}\Psi_{N+1}Y_{N+1} \\ i &= 1 + \Psi_{N+1}^T P_N \Psi_{N+1}\end{aligned}$$

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Recursive algorithm derivation: Simplification

$$\begin{aligned} A &= (P_N - P_N \Psi_{N+1} i^{-1} \Psi_{N+1}^T P_N) \Phi_N^T Y_N \\ &= P_N \Phi_N^T Y_N - P_N \Psi_{N+1} i^{-1} \Psi_{N+1}^T P_N \Phi_N^T Y_N \\ &= \hat{\theta}_N - P_N \Psi_{N+1} i^{-1} \Psi_{N+1}^T \hat{\theta}_N \end{aligned}$$

$$\begin{aligned} B &= (P_N - P_N \Psi_{N+1} i^{-1} \Psi_{N+1}^T P_N) \Psi_{N+1} y_{N+1} \\ &= i^{-1} (P_N (1 + \Psi_{N+1}^T P_N \Psi_{N+1}) - P_N \Psi_{N+1} \Psi_{N+1}^T P_N) \Psi_{N+1} y_{N+1} \\ &= i^{-1} (P_N + P_N \Psi_{N+1}^T P_N \Psi_{N+1} - P_N \Psi_{N+1} \Psi_{N+1}^T P_N) \Psi_{N+1} y_{N+1} \\ &= i^{-1} (P_N \Psi_{N+1} + P_N \Psi_{N+1}^T P_N \Psi_{N+1} \Psi_{N+1} \\ &\quad - P_N \Psi_{N+1} \Psi_{N+1}^T P_N \Psi_{N+1}) y_{N+1} \\ &= i^{-1} (P_N \Psi_{N+1} + P_N \Psi_{N+1} \Psi_{N+1}^T P_N \Psi_{N+1} \\ &\quad - P_N \Psi_{N+1} \Psi_{N+1}^T P_N \Psi_{N+1}) y_{N+1} \\ &= i^{-1} P_N \Psi_{N+1} y_{N+1} \end{aligned}$$

Recursive algorithm derivation: result



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$$\begin{aligned}\hat{\theta}_{N+1} &= \hat{\theta}_N - P_N \Psi_{N+1} i^{-1} \Psi_{N+1}^T \hat{\theta}_N + i^{-1} P_N \Psi_{N+1} y_{N+1} \\ &= \hat{\theta}_N + i^{-1} P_N \Psi_{N+1} (-\Psi_{N+1}^T \hat{\theta}_N + y_{N+1}) \\ &= \hat{\theta}_N + K_{N+1} (y_{N+1} - \Psi_{N+1}^T \hat{\theta}_N) \\ K_{N+1} &= P_N \Psi_{N+1} (1 + \Psi_{N+1}^T P_N \Psi_{N+1})^{-1} \\ P_{N+1} &= P_N - K_{N+1} \Psi_{N+1}^T P_N\end{aligned}$$

Obtain initial value:

- Basic least squares estimation
- $\hat{\theta}_0 = 0, P_0 = c^2 I$, where c is a sufficient large constant.



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$$\begin{aligned}P_N &= (P_{N-1}^{-1} + \Psi_N \Psi_N^T)^{-1} \\P_N^{-1} &= P_{N-1}^{-1} + \Psi_N \Psi_N^T \\P_{N-1}^{-1} &= P_{N-2}^{-1} + \Psi_{N-1} \Psi_{N-1}^T \\P_{N-2}^{-1} &= P_{N-3}^{-1} + \Psi_{N-2} \Psi_{N-2}^T \\P_{N-3}^{-1} &= P_{N-4}^{-1} + \Psi_{N-3} \Psi_{N-3}^T \\&\vdots \\P_1^{-1} &= P_0^{-1} + \Psi_1 \Psi_1^T \\P_N^{-1} &= P_0^{-1} + \sum_{i=1}^N \Psi_i \Psi_i^T\end{aligned}$$



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$$\Phi_N = \begin{bmatrix} \Psi_1^T \\ \Psi_2^T \\ \vdots \\ \Psi_N^T \end{bmatrix}$$

$$P_N^{-1} = \frac{1}{c^2} I + \begin{bmatrix} \Psi_1 & \Psi_2 & \cdots & \Psi_N \end{bmatrix} \begin{bmatrix} \Psi_1^T \\ \Psi_2^T \\ \vdots \\ \Psi_N^T \end{bmatrix}$$

$$= \frac{1}{c^2} I + \Phi^T \Phi$$

$$\lim_{c \rightarrow \infty} P_N^{-1} = \Phi_N^T \Phi_N$$

$$\hat{\theta}_N = P_N \Phi_N^T Y_N$$

$$= (\Phi_N^T \Phi_N)^{-1} \Phi_N^T Y_N$$

The relationship between residual and innovation



Innovation $\tilde{y}_i = y_i - \Psi_i^T \hat{\theta}_{i-1}$ is used to describe prediction error at time i. residual $\varepsilon_i = y_i - \Psi_i^T \hat{\theta}_i$ is used to describe the output bias at time i.

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$$\begin{aligned}\varepsilon &= y_i - \Psi_i^T \hat{\theta}_i \\&= y_i - \Psi_i^T (\hat{\theta}_{i-1} + K_i \tilde{y}_i) \\&= \tilde{y}_i - \Psi_i^T K_i \tilde{y}_i \\&= (1 - \Psi_i^T K_i) \tilde{y}_i \\&= (1 - \Psi_i^T P_{i-1} \Psi_i (\Psi_i^T P_{i-1} \Psi_i + 1)^{-1}) \tilde{y}_i \\&= \frac{\Psi_i^T P_{i-1} \Psi_i + 1 - \Psi_i^T P_{i-1} \Psi_i}{\Psi_i^T P_{i-1} \Psi_i + 1} \tilde{y}_i \\&= \frac{\tilde{y}_i}{\Psi_i^T P_{i-1} \Psi_i + 1}\end{aligned}$$

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Model order
increasing
algorithm

recursive least
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$$\begin{aligned} J_i &= (Y_i - \Phi_i \theta_i)^T (Y_i - \Phi_i \theta_i) \\ J_{i-1} &= (Y_{i-1} - \Phi_{i-1} \theta_{i-1})^T (Y_{i-1} - \Phi_{i-1} \theta_{i-1}) \\ Y_i - \Phi_i \theta_i &= Y_i - \Phi_i (\hat{\theta}_{i-1} + K_i \tilde{y}_i) \\ &= \begin{bmatrix} Y_{i-1} \\ y_i \end{bmatrix} - \begin{bmatrix} \Phi_{i-1} \\ \Psi_i^T \end{bmatrix} (\hat{\theta}_{i-1} + K_i \tilde{y}_i) \\ &= \begin{bmatrix} Y_{i-1} - \Phi_{i-1} \hat{\theta}_{i-1} \\ \tilde{y}_i \end{bmatrix} - \begin{bmatrix} \Phi_{i-1} \\ \Psi_i^T \end{bmatrix} K_i \tilde{y}_i \end{aligned}$$



$$\begin{aligned} J_i &= J_{i-1} - 2K_i^T \Phi_{i-1}^T (Y_{i-1} - \Phi_{i-1} \hat{\theta}_{i-1}) \tilde{y}_i + K_i^T \Phi_{i-1}^T \Phi_{i-1} K_i \tilde{y}_i^2 \\ &\quad + (1 - 2K_i^T \Psi_i + K_i^T \Psi_i \Psi_i^T K_i) \tilde{y}_i^2 \end{aligned}$$

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$$\begin{aligned} &= J_{i-1} - 2K_i^T (\Phi_{i-1}^T Y_{i-1} - \Phi_{i-1}^T \Phi_{i-1} \hat{\theta}_{i-1}) \tilde{y}_i \\ &\quad + (1 - 2K_i^T \Psi_i + K_i^T \Phi_i \Phi_i^T K_i) \tilde{y}_i^2 \\ &= J_{i-1} + (1 - 2K_i^T \Psi_i + K_i^T \Phi_i \Phi_i^T K_i) \tilde{y}_i^2 \\ &= J_{i-1} + (1 - 2K_i^T \Psi_i + K_i^T P_i^{-1} K_i) \tilde{y}_i^2 \\ &= J_{i-1} + (1 - 2K_i^T \Psi_i + K_i^T \Psi_i) \tilde{y}_i^2 \\ &= J_{i-1} + (1 - K_i^T \Psi_i) \tilde{y}_i^2 \\ &= J_{i-1} + (1 - \Psi_i^T P_{i-1} \Psi_i (\Psi_i^T P_{i-1} \Psi_i + 1)^{-1}) \tilde{y}_i^2 \\ &= J_{i-1} + \frac{\Psi_i^T P_{i-1} \Psi_i + 1 - \Psi_i^T P_{i-1} \Psi_i}{\Psi_i^T P_{i-1} \Psi_i + 1} \tilde{y}_i^2 \\ &= J_{i-1} + \frac{\tilde{y}_i^2}{\Psi_i^T P_{i-1} \Psi_i + 1} \end{aligned}$$



When there is error δK_i in K_i :

$$\delta P_i = \delta K_i \Psi_i^T P_{i-1}$$

Compute new form of P_i :

$$\begin{aligned} P_i &= (I - K_i \Psi_i^T) P_{i-1} \\ &= (I - K_i \Psi_i^T) P_{i-1} - P_{i-1} \Psi_i K_i^T + P_{i-1} \Psi_i K_i^T \\ &= (I - K_i \Psi_i^T) P_{i-1} - P_{i-1} \Psi_i K_i^T + K_i (\Psi_i^T P_{i-1} \Psi_i + 1) K_i^T \\ &= (I - K_i \Psi_i^T) P_{i-1} - (I - K_i \Psi_i^T) P_{i-1} \Psi_i K_i^T + K_i K_i^T \\ &= (I - K_i \Psi_i^T) (P_{i-1} - P_{i-1} \Psi_i K_i^T) + K_i K_i^T \\ &= (I - K_i \Psi_i^T) P_{i-1} (I - \Psi_i K_i^T) + K_i K_i^T \end{aligned}$$

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When there is error δK_i in K_i :

$$\begin{aligned}
 \delta P_i &= (I - (K_i + \delta K_i)\Psi_i^T)P_{i-1}(I - \Psi_i(K_i + \delta K_i)^T) \\
 &\quad + (K_i + \delta K_i)(K_i + \delta K_i)^T - P_i \\
 &= -\delta K_i\Psi_i^T P_{i-1}(I - \Psi_i K_i^T) + K_i\delta K_i^T \\
 &\quad - (I - K_i\Psi_i^T)P_{i-1}\Psi_i\delta K_i^T + \delta K_i K_i^T \\
 &\quad + \delta K_i\Psi_i^T P_{i-1}\Psi_i\delta K_i^T + \delta K_i\delta K_i^T \\
 &\quad + (I - K_i\Psi_i^T)P_{i-1}(I - \Psi_i K_i^T) + K_i K_i^T - P_i \\
 &= -\delta K_i\Psi_i^T P_{i-1}(I - \Psi_i K_i^T) + K_i\delta K_i^T \\
 &\quad - (I - K_i\Psi_i^T)P_{i-1}\Psi_i\delta K_i^T + \delta K_i K_i^T + O(\delta K_i) \\
 &= -\delta K_i\Psi_i^T P_i^T + \delta K_i K_i^T - P_i\Psi_i\delta K_i^T + K_i\delta K_i^T + O(\delta K_i) \\
 &= -\delta K_i K_i^T + \delta K_i K_i^T - K_i\delta K_i^T + K_i\delta K_i^T + O(\delta K_i) \\
 &= O(\delta K_i)
 \end{aligned}$$

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$$\begin{aligned}y_i &= \Psi_i^T \theta + \xi_i \\ \tilde{\theta}_i &\stackrel{\text{def}}{=} \theta - \hat{\theta}_i \\ &= \theta - [\hat{\theta}_{i-1} + K_i(y_i - \Psi_i^T \hat{\theta}_{i-1})] \\ &= \tilde{\theta}_{i-1} - K_i(y_i - \Psi_i^T \hat{\theta}_{i-1}) \\ &= \tilde{\theta}_{i-1} - K_i(\Psi_i^T \theta + \xi_i - \Psi_i^T \hat{\theta}_{i-1}) \\ &= \tilde{\theta}_{i-1} - K_i(\Psi_i^T \tilde{\theta}_{i-1} + \xi_i) \\ &= (I - K_i \Psi_i^T) \tilde{\theta}_{i-1} - K_i \xi_i \\ &= P_i P_{i-1}^{-1} \tilde{\theta}_{i-1} - K_i \xi_i \\ &= A_i \tilde{\theta}_{i-1} - K_i \xi_i \\ A_i &= P_i P_{i-1}^{-1}\end{aligned}$$



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$$\begin{aligned}A_i x &= \lambda x \\(P_{i-1}^{-1} + \Psi_i \Psi_i^T)^{-1} P_{i-1}^{-1} x &= \lambda x \\P_{i-1}^{-1} x &= [P_{i-1}^{-1} + \Psi_i \Psi_i^T] \lambda x \\(1 - \lambda) P_{i-1}^{-1} x &= \lambda \Psi_i \Psi_i^T x \\(1 - \lambda) x^T P_{i-1}^{-1} x &= \lambda x^T \Psi_i \Psi_i^T x\end{aligned}$$

where: P_{i-1}^{-1} is positive definite and $\Psi_i \Psi_i^T$ is non-negative definite, so $0 < \lambda \leq 1$. that is: $\tilde{\theta}_i \leq \tilde{\theta}_0$.

The relationship between least squares estimation and Kalman filtering



State space model:

$$\begin{aligned}\theta_{i+1} &= \theta_i \\ y_i &= \Psi_i^T \theta_i + \xi_i\end{aligned}$$

Kalman filtering:

$$\begin{aligned}\hat{\theta}_i &= \hat{\theta}_{i-1} + K_i(y_i - \Psi_i^T \hat{\theta}_{i-1}) \\ K_i &= S_i \Psi_i (\Psi_i^T S_i \Psi_i + \sigma^2)^{-1} \\ S_i &= P_{i-1} \\ P_i &= (I - K_i \Psi_i^T) P_{i-1} \\ \hat{\theta}_0 &= 0\end{aligned}$$

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