

FYS-STK3-4155 Sept 1

statistics, reminder

$$D = \{ (x_0, y_0), (x_1, y_1) \dots (x_m, y_m) \}$$

probability distribution

$$p(x)$$

$$0 \leq p(x) \leq 1$$

$$x \in [a, b]$$

$$\int_a^b p(x) dx = 1 \quad \text{continuous case}$$

$$\sum_{x_i \in X} p(x_i) = 1 \quad \text{discrete case}$$

Expectation values

$$E[x^n] = \int_a^b p(x) x^n dx$$

$$E[x] = \mu = \langle x \rangle = \int_a^b p(x) x dx$$

Discrete:

$$\mu = \sum_{i=0}^{n-1} x_i p(x_i)$$

sample mean

$$\bar{\mu} = \frac{1}{n} \sum_i x_i$$

$\bar{\mu} \neq \mu$ in general

variance

$$\text{var}[x] = \sigma^2$$

$$= \int (x - \mu)^2 p(x) dx$$

$$\left(\text{Discrete} = \sum_i (x_i - \mu)^2 p(x_i) \right)$$

sample variance

$$\bar{\sigma}^2 = \frac{1}{n} \sum_i (x_i - \bar{\mu})^2$$
$$\neq \sigma^2$$

iid = independent and

identically distributed

$$P(x_1, x_2) = p(x_1) p(x_2)$$

$$\text{cov}[x_1, x_2] = \int dx_1 dx_2 (x_1 - \mu_1)(x_2 - \mu_2) p(x_1, x_2)$$

$$\text{iid: cov}[x_1, x_2] =$$

$$\int dx_1 dx_2 (x_1 - \mu)(x_2 - \mu) p(x_1) p(x_2)$$

$$= \int dx_1 (x_1 - \mu) p(x_1) \int dx_2 (x_2 - \mu) p(x_2)$$

$$\left[\int dx p(x) x = \mu \right]$$

$$\int dx p(x) = 1$$

$$= 0.$$

if not

$$\text{cov}[x_1, x_2] = E[x_1 x_2]$$

$$- \mu_1 \mu_2$$

Discrete case

$$\text{cov}[x_1, x_2] = \sum_{i,j} (x_i - \mu_1)(x_j - \mu_2) \times p(x_i, x_j)$$

Joint probability (product rule)

$$p(A, B) = p(A \cap B)$$

$$= p(A|B) p(B)$$

$$= p(B|A) p(A)$$

Marginal probability

$$p(A) = \sum_b p(A|B=b) p(B=b)$$

Conditional prob ($p(B) > 0$)

$$p(A|B) = \frac{p(A, B)}{p(B)}$$

$$= \frac{P(B|A) P(A)}{\sum_a P(B|A=a) P(A=a)}$$

Baye's theorem.

Example

A = output from
a Mammogram
(0,1)

B = have cancer
(0,1)

sensitivity

$$P(A=1|B=1) = 0.8$$

$$P(B|A)$$

prior $P(B) = 0.004$

$$P(A=1 | B=0) = 0.1$$

Bayes's theorem

$$P(B=1 | A=1) =$$

$$\frac{P(A=1 | B=1) P(B=1)}{P(A=1 | B=1) P(B=1) + P(A=1 | B=0) P(B=0)}$$

$$= \frac{0.8 \times 0.004}{0.8 \times 0.004 + 0.1 \times 0.996}$$

$$= 0.03 \Rightarrow$$

3%

$$\mathcal{D} = \{ (x_0, y_0), \dots, (x_{n-1}, y_{n-1}) \}$$

$$P(y | x, \beta) = ?$$

y are iid. \nwarrow not stochastic
 $\sim x, \beta \nwarrow$

$$y = f(x) + \varepsilon \quad \setminus ?$$

$$\varepsilon \sim N(0, 1) \quad \left. \varepsilon \right\} \text{var}_\varepsilon = \sigma_\varepsilon^2$$

$$P(y | x \beta) = \prod_{i=0}^{n-1} P_i$$

$$P_i = P(y_i | x_i \beta)$$

$$x_i \beta = \sum_{j=0}^{p-1} x_{ij} \beta_j$$

Maximum likelihood estimator (MLE)

$$\hat{\beta} = \underset{\beta \in \mathbb{R}^p}{\text{arg max}} \ P(y | x \beta) \quad \uparrow \quad P(\beta)$$

\Rightarrow maximization of

$$\frac{\partial}{\partial \beta} \prod_i P_i(y_i | x_i \beta)$$

$$E[y_i] = x_i \beta$$

$$\text{var}[y_i] = \sigma_e^2$$

$$y_i \sim N(x_i^T \beta, \sigma_e^2)$$

$$C(\beta) = P(y|x\beta)$$

$$C(\beta) = \prod_{i=1}^n P_i$$

$$\ell(\beta) = \sum_{i=1}^n \log P(y_i | x_i^T \beta)$$

$$C(\beta) = -\ell(\beta)$$

$$= - \sum_{i=1}^n \log P(y_i | x_i^T \beta)$$

$$\hat{\beta} = \arg \min_{\beta \in \mathbb{R}^p} C(\beta)$$

$$y_i \sim N(x_i^T \beta, \sigma_e^2)$$

$$= \frac{1}{\sqrt{2\pi\sigma_e^2}} \exp\left(-\frac{(y_i - x_i^T \beta)^2}{2\sigma_e^2}\right)$$

$$= P_i$$

$$\log P_i = -\frac{1}{2} \log(2\pi\sigma_\epsilon^2) - \frac{(y_i - x_i^T \beta)^2}{2\sigma_\epsilon^2}$$

$$C(\beta) = \frac{n}{2} \log(2\pi\sigma_\epsilon^2) + \sum_{i=0}^{n-1} \frac{(y_i - x_i^T \beta)^2}{2\sigma_\epsilon^2}$$

$$\frac{\partial C}{\partial \beta} = 0 = X^T (y - X\beta)$$