

## SECTION 6.3: METHOD OF TRANSFORMATIONS

*Exercise (6.1).* Let  $Y$  be a random variable with probability density function given by

$$f(y) = \begin{cases} 2(1-y), & 0 \leq y \leq 1, \\ 0, & \text{elsewhere.} \end{cases}$$

- (a) Find the density function of  $U_1 = 2Y - 1$ .

*Solution:*

$$U_1 = 2Y - 1$$

$$\frac{dy}{du_1} = \frac{1}{2}$$

$$2Y = U_1 + 1$$

$$f_{u_1}(u_1) = 2\left(1 - \frac{u_1 + 1}{2}\right)\frac{1}{2}$$

$$Y = \frac{U_1 + 1}{2}$$

$$= \frac{1 - u_1}{2}$$

$$f_{U_1} = \begin{cases} \frac{1+u_1}{2}, & -1 \leq u_1 \leq 1 \\ 0, & \text{elsewhere} \end{cases}$$

□

- (b) Find the density function of  $U_2 = 1 - 2Y$ .

*Solution:*

$$U_2 = 1 - 2Y$$

$$\frac{dy}{du_2} = -\frac{1}{2}$$

$$2Y = 1 - U_2$$

$$f_{U_2} = 2\left(1 - \frac{1 - u_2}{2}\right) * \frac{1}{2}$$

$$Y = \frac{1 - U_2}{2}$$

$$= \frac{1 + u_2}{2}$$

$$f_{U_2}(u_2) = \begin{cases} \frac{1+u_2}{2}, & -1 \leq u_2 \leq 1 \\ 0, & \text{elsewhere} \end{cases}$$

□

- (c) Find the density function of  $U_3 = Y^2$ .

*Solution:*

$$U_3 = Y^2$$

$$Y = \sqrt{U_3}$$

$$\frac{dy}{du_3} = \frac{1}{2\sqrt{u_3}}$$

$$f_{U_3}(u_3) = 2(1 - \sqrt{u_3}) \cdot \frac{1}{2\sqrt{u_3}} \\ = \frac{1 - \sqrt{u_3}}{\sqrt{u_3}}$$

$$f_{U_3}(u_3) = \begin{cases} \frac{1-\sqrt{u_3}}{\sqrt{u_3}}, & 0 \leq u_3 \leq 1 \\ 0, & \text{elsewhere} \end{cases}$$

□

- (d) Find  $E(U_1)$ ,  $E(U_2)$ , and  $E(U_3)$  by using the derived density functions for these random variables.

*Solution:*

$$E(U_1) = \int_{-1}^1 u_1 \cdot \frac{1-u_1}{2} du_1 = \frac{1}{2} \left[ \frac{u_1^2}{2} - \frac{u_1^3}{3} \right]_{-1}^1 = \frac{1}{2} \left[ -\frac{2}{3} \right] = -\frac{1}{3}$$

$$E(U_2) = \int_{-1}^1 u_2 \cdot \frac{1+u_2}{2} du_2 = \frac{1}{2} \left[ \frac{u_2^2}{2} + \frac{u_2^3}{3} \right]_{-1}^1 = \frac{1}{2} \left[ \frac{2}{3} \right] = \frac{1}{3}$$

$$E(U_3) = \int_0^1 u_3 \cdot \frac{1-\sqrt{u_3}}{\sqrt{u_3}} du_3 = \int_0^1 (\sqrt{u_3} - u_3) du_3 = \left[ \frac{2u_3^{3/2}}{3} - \frac{u_3^2}{2} \right]_0^1 = \frac{2}{3} - \frac{1}{2} = \frac{1}{6}$$

□

- (e) Find  $E(U_1)$ ,  $E(U_2)$ , and  $E(U_3)$  by the methods of Chapter 4.

$$\text{Solution: First, } E(Y) = \int_0^1 y \cdot 2(1-y) dy = 2 \left[ \frac{y^2}{2} - \frac{y^3}{3} \right]_0^1 = \frac{1}{3}$$

$$E(U_1) = E(2Y - 1) = 2E(Y) - 1 = 2 \cdot \frac{1}{3} - 1 = -\frac{1}{3}$$

$$E(U_2) = E(1 - 2Y) = 1 - 2E(Y) = 1 - \frac{2}{3} = \frac{1}{3}$$

$$E(U_3) = E(Y^2) = \int_0^1 y^2 \cdot 2(1-y) dy = 2 \left[ \frac{y^3}{3} - \frac{y^4}{4} \right]_0^1 = \frac{2}{3} - \frac{1}{2} = \frac{1}{6}$$

□

*Exercise (6.6).* The joint distribution of amount of pollutant emitted from a smokestack without a cleaning device ( $Y_1$ ) and a similar smokestack with a cleaning device ( $Y_2$ ) was given in Exercise 5.10 to be

$$f(y_1, y_2) = \begin{cases} 1, & 0 \leq y_1 \leq 2, 0 \leq y_2 \leq 1, 2y_2 \leq y_1, \\ 0, & \text{elsewhere.} \end{cases}$$

The reduction in amount of pollutant due to the cleaning device is given by  $U = Y_1 - Y_2$ .

- (a) Find the probability density function for  $U$ .

*Solution:* Let  $U = Y_1 - Y_2$  and  $V = Y_2$ . Then  $Y_1 = U + V$ ,  $Y_2 = V$ , and  $|J| = 1$ .

The region becomes:  $0 \leq v \leq 1$ ,  $2v \leq u + v \leq 2$ , which gives  $v \leq u$  and  $v \leq 2 - u$ .

$$f_U(u) = \int_0^{\min(u, 1, 2-u)} 1 dv$$

$$f_U(u) = \begin{cases} u, & 0 \leq u \leq 1 \\ 2 - u, & 1 < u \leq 2 \\ 0, & \text{elsewhere} \end{cases}$$

□

- (b) Use the answer in part (a) to find  $E(U)$ . Compare your results with those of Exercise 5.78(c).

*Solution:*

$$\begin{aligned} E(U) &= \int_0^1 u \cdot u du + \int_1^2 u(2-u) du \\ &= \left[ \frac{u^3}{3} \right]_0^1 + \left[ u^2 - \frac{u^3}{3} \right]_1^2 \\ &= \frac{1}{3} + \left[ \left( 4 - \frac{8}{3} \right) - \left( 1 - \frac{1}{3} \right) \right] \\ &= \frac{1}{3} + \frac{2}{3} = 1 \end{aligned}$$

□

*Exercise (6.7).* Suppose that  $Z$  has a standard normal distribution.

- (a) Find the density function of  $U = Z^2$ .

*Solution:* For  $Z \sim N(0, 1)$ ,  $f_Z(z) = \frac{1}{\sqrt{2\pi}} e^{-z^2/2}$ .

Since  $u = z^2$  gives  $z = \pm\sqrt{u}$  and  $\left|\frac{dz}{du}\right| = \frac{1}{2\sqrt{u}}$ :

$$\begin{aligned} f_U(u) &= f_Z(\sqrt{u}) \cdot \frac{1}{2\sqrt{u}} + f_Z(-\sqrt{u}) \cdot \frac{1}{2\sqrt{u}} \\ &= 2 \cdot \frac{1}{\sqrt{2\pi}} e^{-u/2} \cdot \frac{1}{2\sqrt{u}} = \frac{1}{\sqrt{2\pi u}} e^{-u/2}, \quad u > 0 \end{aligned}$$

□

(b) Does  $U$  have a gamma distribution? What are the values of  $\alpha$  and  $\beta$ ?

*Solution:* Yes. Rewriting  $f_U(u) = \frac{1}{\sqrt{2\pi}} u^{-1/2} e^{-u/2}$  and using  $\Gamma(1/2) = \sqrt{\pi}$ :

$$f_U(u) = \frac{1}{\Gamma(1/2) \cdot 2^{1/2}} u^{1/2-1} e^{-u/2}$$

Thus  $\alpha = 1/2$  and  $\beta = 2$ .

□

(c) What is another name for the distribution of  $U$ ?

*Solution:* Chi-square distribution with 1 degree of freedom,  $\chi^2(1)$ .

□

*Exercise (6.8).* Assume that  $Y$  has a beta distribution with parameters  $\alpha$  and  $\beta$ .

(a) Find the density function of  $U = 1 - Y$ .

*Solution:* For  $Y \sim \text{Beta}(\alpha, \beta)$ ,  $f_Y(y) = \frac{1}{B(\alpha, \beta)} y^{\alpha-1} (1-y)^{\beta-1}$ .

Since  $u = 1 - y$ , we have  $y = 1 - u$  and  $\left|\frac{dy}{du}\right| = 1$ :

$$f_U(u) = \frac{1}{B(\alpha, \beta)} (1-u)^{\alpha-1} u^{\beta-1}, \quad 0 \leq u \leq 1$$

□

(b) Identify the density of  $U$  as one of the types we studied in Chapter 4. Be sure to identify any parameter values.

*Solution:*  $U$  has a beta distribution with parameters  $\beta$  and  $\alpha$ :  $U \sim \text{Beta}(\beta, \alpha)$ .

□

(c) How is  $E(U)$  related to  $E(Y)$ ?

*Solution:*  $E(U) = E(1 - Y) = 1 - E(Y)$ , or equivalently  $E(U) = \frac{\beta}{\alpha+\beta}$  and  $E(Y) =$

$$\frac{\alpha}{\alpha+\beta}.$$

□

(d) How is  $V(U)$  related to  $V(Y)$ ?

*Solution:*  $V(U) = V(1 - Y) = V(Y)$ , since variance is unaffected by sign changes.  $\square$

*Exercise (6.9).* Suppose that a unit of mineral ore contains a proportion  $Y_1$  of metal A and a proportion  $Y_2$  of metal B. Experience has shown that the joint probability density function of  $Y_1$  and  $Y_2$  is uniform over the region  $0 \leq y_1 \leq 1$ ,  $0 \leq y_2 \leq 1$ ,  $0 \leq y_1 + y_2 \leq 1$ . Let  $U = Y_1 + Y_2$ , the proportion of either metal A or B per unit. Find

- (a) the probability density function for  $U$ .

*Solution:* The region has area  $1/2$ , so  $f(y_1, y_2) = 2$  on the region. Let  $U = Y_1 + Y_2$

$$\text{and } V = Y_2.$$

Then  $Y_1 = U - V$ ,  $Y_2 = V$ , and  $|J| = 1$ . For  $0 \leq u \leq 1$ ,  $v$  ranges from 0 to  $u$ :

$$f_U(u) = \int_0^u 2 \, dv = 2u, \quad 0 \leq u \leq 1$$

$\square$

- (b)  $E(U)$  by using the answer to part (a).

*Solution:*

$$E(U) = \int_0^1 u \cdot 2u \, du = 2 \left[ \frac{u^3}{3} \right]_0^1 = \frac{2}{3}$$

$\square$

- (c)  $E(U)$  by using only the marginal densities of  $Y_1$  and  $Y_2$ .

*Solution:*

$$f_1(y_1) = \int_0^{1-y_1} 2 \, dy_2 = 2(1 - y_1), \quad 0 \leq y_1 \leq 1$$

Similarly,  $f_2(y_2) = 2(1 - y_2)$ . Then:

$$E(Y_1) = \int_0^1 y_1 \cdot 2(1 - y_1) \, dy_1 = 2 \left[ \frac{y_1^2}{2} - \frac{y_1^3}{3} \right]_0^1 = \frac{1}{3}$$

By symmetry,  $E(Y_2) = \frac{1}{3}$ . Thus  $E(U) = E(Y_1) + E(Y_2) = \frac{2}{3}$ .  $\square$

*Exercise (6.10).* The total time from arrival to completion of service at a fast-food outlet,  $Y_1$ , and the time spent waiting in line before arriving at the service window,  $Y_2$ , were given in Exercise 5.15 with joint density function

$$f(y_1, y_2) = \begin{cases} e^{-y_1}, & 0 \leq y_2 \leq y_1 < \infty, \\ 0, & \text{elsewhere.} \end{cases}$$

Another random variable of interest is  $U = Y_1 - Y_2$ , the time spent at the service window. Find

- (a) the probability density function for  $U$ .

*Solution:* Let  $U = Y_1 - Y_2$  and  $V = Y_2$ . Then  $Y_1 = U + V$ ,  $Y_2 = V$ , and  $|J| = 1$ .

The region  $0 \leq y_2 \leq y_1$  becomes  $0 \leq v$  and  $v \leq u + v$ , so  $u \geq 0$ :

$$f_U(u) = \int_0^\infty e^{-(u+v)} dv = e^{-u} \int_0^\infty e^{-v} dv = e^{-u}, \quad u \geq 0$$

□

- (b)  $E(U)$  and  $V(U)$ . Compare your answers with the results of Exercise 5.108.

*Solution:*  $U$  is exponential with parameter 1, so  $E(U) = 1$  and  $V(U) = 1$ . □

*Exercise (6.11).* Suppose that two electronic components in the guidance system for a missile operate independently and that each has a length of life governed by the exponential distribution with mean 1 (with measurements in hundreds of hours). Find the

- (a) probability density function for the average length of life of the two components.

*Solution:* Let  $W = Y_1 + Y_2$ . Since  $Y_1, Y_2$  are independent exponential(1),  $W \sim$

Gamma(2, 1) with  $f_W(w) = we^{-w}$  for  $w \geq 0$ .

For  $U = W/2$ , we have  $w = 2u$  and  $\frac{dw}{du} = 2$ :

$$f_U(u) = f_W(2u) \cdot 2 = (2u)e^{-2u} \cdot 2 = 4ue^{-2u}, \quad u \geq 0$$

□

- (b) mean and variance of this average, using the answer in part (a). Check your answer by computing the mean and variance, using Theorem 5.12.

*Solution:*

$$\begin{aligned} E(U) &= \int_0^\infty u \cdot 4ue^{-2u} du = 4 \cdot \frac{2!}{2^3} = 1 \\ E(U^2) &= \int_0^\infty u^2 \cdot 4ue^{-2u} du = 4 \cdot \frac{3!}{2^4} = \frac{3}{2} \\ V(U) &= E(U^2) - [E(U)]^2 = \frac{3}{2} - 1 = \frac{1}{2} \end{aligned}$$

Using Theorem 5.12:  $E(U) = \frac{E(Y_1) + E(Y_2)}{2} = 1$  and  $V(U) = \frac{V(Y_1) + V(Y_2)}{4} = \frac{1}{2}$ . □

*Exercise (6.13).* If  $Y_1$  and  $Y_2$  are independent exponential random variables, both with mean  $\beta$ , find the density function for their sum. (In Exercise 5.7, we considered two independent exponential random variables, both with mean 1 and determined  $P(Y_1 + Y_2 \leq 3)$ .)

*Solution:* Since  $Y_1, Y_2$  are independent exponential with mean  $\beta$ , their sum  $U = Y_1 + Y_2$  has a gamma distribution with  $\alpha = 2$  and parameter  $\beta$ :

$$f_U(u) = \frac{1}{\Gamma(2)\beta^2} u^{2-1} e^{-u/\beta} = \frac{u}{\beta^2} e^{-u/\beta}, \quad u \geq 0$$

□

*Exercise (6.14).* In a process of sintering (heating) two types of copper powder (see Exercise 5.152), the density function for  $Y_1$ , the volume proportion of solid copper in a sample, was given by

$$f_1(y_1) = \begin{cases} 6y_1(1-y_1), & 0 \leq y_1 \leq 1, \\ 0, & \text{elsewhere.} \end{cases}$$

The density function for  $Y_2$ , the proportion of type A crystals among the solid copper, was given as

$$f_2(y_2) = \begin{cases} 3y_2^2, & 0 \leq y_2 \leq 1, \\ 0, & \text{elsewhere.} \end{cases}$$

The variable  $U = Y_1 Y_2$  gives the proportion of the sample volume due to type A crystals. If  $Y_1$  and  $Y_2$  are independent, find the probability density function for  $U$ .

*Solution:* Let  $U = Y_1 Y_2$  and  $V = Y_2$ . Then  $Y_1 = U/V$ ,  $Y_2 = V$ , and  $|J| = \frac{1}{v}$ .

Since  $Y_1, Y_2$  are independent:

$$\begin{aligned} f_{U,V}(u,v) &= f_1(u/v) \cdot f_2(v) \cdot \frac{1}{v} \\ &= 6(u/v)(1-u/v) \cdot 3v^2 \cdot \frac{1}{v} \\ &= 18u - \frac{18u^2}{v} \end{aligned}$$

For  $0 \leq u \leq 1$ ,  $v$  ranges from  $u$  to 1 (from  $u/v \leq 1$  and  $v \leq 1$ ):

$$\begin{aligned} f_U(u) &= \int_u^1 \left( 18u - \frac{18u^2}{v} \right) dv \\ &= 18u(1-u) - 18u^2 [\ln v]_u^1 \\ &= 18u(1-u) + 18u^2 \ln u, \quad 0 \leq u \leq 1 \end{aligned}$$

□