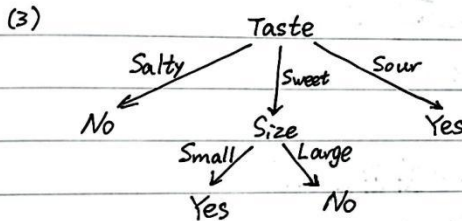


## Homework 2

2. (1)  $H(t) = -\sum_c p(c|t) \log_2 p(c|t) = -\left(\frac{1}{2} \log_2 \frac{1}{2} + \frac{1}{2} \log_2 \frac{1}{2}\right) = 1$

(2)  $\text{Info Gain}_{\text{split}} = H(t) - \sum_{k=1}^3 \frac{n_k}{n} H(k)$   
 $= 1 - \left(\frac{3}{10} \times (-1 \cdot \log_2 1) + \frac{4}{10} \times \left(-\frac{1}{2} \log_2 \frac{1}{2} - \frac{1}{2} \log_2 \frac{1}{2}\right) + \frac{3}{10} \times (-1 \cdot \log_2 1)\right)$   
 $= 1 - \frac{2}{5}$   
 $= \frac{3}{5}$



3. (1)  $L(\theta) = \prod_{i=1}^n \frac{1}{\sqrt{2\pi}\sigma} e^{-\frac{(x_i - \mu)^2}{2\sigma^2}} = \frac{1}{\sqrt{2\pi}\sigma^n} e^{-\frac{\sum_{i=1}^n (x_i - \mu)^2}{2\sigma^2}}$

$\ln L(\theta) = -\frac{n}{2} \ln 2\pi - \frac{n}{2} \ln \sigma^2 - \frac{1}{2\sigma^2} \sum_{i=1}^n (x_i - \mu)^2$

We need  $\ln L(\theta)$  to differentiate to  $\mu$  and  $\sigma^2$  and equal to 0.

$\frac{\partial \ln L(\theta)}{\partial \mu} = \frac{1}{\sigma^2} \sum_{i=1}^n (x_i - \mu) = 0 \Rightarrow \sum_{i=1}^n x_i - n\mu = 0 \Rightarrow \hat{\mu} = \bar{x}$

$\frac{\partial \ln L(\theta)}{\partial \sigma^2} = -\frac{n}{2\sigma^2} + \frac{1}{2\sigma^4} \sum_{i=1}^n (x_i - \mu)^2 = 0 \Rightarrow \hat{\sigma}^2 = \frac{1}{n} \sum_{i=1}^n (x_i - \mu)^2 = \frac{1}{n} \sum_{i=1}^n (x_i - \bar{x})^2$

(2)  $E(\hat{\mu}) = E\left(\frac{1}{n} \sum_{i=1}^n x_i\right) = \frac{1}{n} \sum_{i=1}^n E(x_i) = \frac{1}{n} \cdot n \cdot \mu = \mu$

$\hat{\sigma}^2 = \frac{1}{n} \sum_{i=1}^n (x_i - \bar{x})^2 = \frac{1}{n} \sum_{i=1}^n \left(x_i - \frac{1}{n} \sum_{j=1}^n x_j\right)^2 = \frac{1}{n} \sum_{i=1}^n x_i^2 - \frac{1}{n^2} \sum_{i=1}^n \sum_{j=1}^n x_i x_j$

Suppose  $\mu - x_i = \Delta_i$

$\hat{\sigma}^2 = \frac{1}{n} \sum_{i=1}^n (\mu - \Delta_i)^2 - \frac{1}{n^2} \sum_{i=1}^n \sum_{j=1}^n (\mu - \Delta_i)(\mu - \Delta_j)$   
 $= \mu^2 - \frac{2\mu}{n} \sum_{i=1}^n \Delta_i + \frac{1}{n} \sum_{i=1}^n \Delta_i^2 - \mu^2 + \frac{\mu}{n^2} \sum_{i=1}^n \sum_{j=1}^n (\Delta_i + \Delta_j) - \frac{1}{n^2} \sum_{i=1}^n \sum_{j=1}^n \Delta_i \Delta_j$

$E(\Delta_i) = 0, E(\Delta_i^2) = \sigma^2, E(\Delta_i \Delta_j) = 0$

$E(\hat{\sigma}^2) = -\frac{2\mu}{n} \cdot 0 + \frac{1}{n} \sum_{i=1}^n \sigma^2 + \frac{\mu}{n^2} \cdot 0 + \frac{1}{n^2} \sum_{i=1}^n \sigma^2 = \frac{n-1}{n} \sigma^2$

Hence,  $E\left(\frac{n}{n-1} \hat{\sigma}^2\right) = \sigma^2$ , the conclusion holds.

No.  
Date

4. ① Suppose  $P(y_i) = \prod_{k=1}^c p_k^{I(y_i=k)}$

Likelihood function  $L(p_1, p_2, \dots, p_c) = \prod_{i=1}^n P(y_i) = \prod_{i=1}^n \prod_{k=1}^c p_k^{I(y_i=k)}$

$\ln L = \sum_{i=1}^n \sum_{k=1}^c I(y_i=k) \ln(p_k)$

By Lagrangian multiplier method,

$\ell(p_1, \dots, p_c, \lambda) = \sum_{i=1}^n \sum_{k=1}^c I(y_i=k) \ln p_k - \lambda \left( \sum_{k=1}^c p_k - 1 \right)$

To get MLE estimator of  $p_k$ ,  $\frac{\partial \ell}{\partial p_k} = 0$ .

$\frac{\partial \ell}{\partial p_k} = \sum_{i=1}^n \frac{I(y_i=k)}{p_k} - \lambda = 0 \Rightarrow \hat{p}_k = \frac{\sum_{i=1}^n I(y_i=k)}{n}$

Set  $\lambda = n$ , then  $\hat{p}_k = \frac{1}{n} \sum_{i=1}^n I(y_i=k)$

② Suppose  $P(x_i | y_i) = \prod_{s=1}^S \prod_{k=1}^c p_{sk}^{I(x_i=s, y_i=k)}$

Then  $L(p_{11}, \dots, p_{tc}) = \prod_{i=1}^n P(x_i | y_i) = \prod_{i=1}^n \prod_{s=1}^S \prod_{k=1}^c p_{sk}^{I(x_i=s, y_i=k)}$

$\ell(p_{11}, \dots, p_{tc}) = \ln L = \sum_{i=1}^n \sum_{s=1}^S \sum_{k=1}^c I(x_i=s, y_i=k) \ln p_{sk}$

By Lagrangian multiplier method,

$\ell(p_{11}, \dots, p_{tc}, \lambda) = \sum_{i=1}^n \sum_{s=1}^S \sum_{k=1}^c I(x_i=s, y_i=k) \ln p_{sk} - \lambda \left( \sum_{s=1}^S p_{sk} - 1 \right)$

$\frac{\partial \ell}{\partial p_{sk}} = \sum_{i=1}^n \frac{I(x_i=s, y_i=k)}{p_{sk}} - \lambda = 0 \Rightarrow \hat{p}_{sk} = \frac{1}{\lambda} \sum_{i=1}^n I(x_i=s, y_i=k)$

According to  $\sum_{s=1}^S p_{sk} = 1$ , set  $\lambda = \sum_{i=1}^n I(y_i=k)$

then  $\hat{p}_{sk} = \frac{\sum_{i=1}^n I(x_i=s, y_i=k)}{\sum_{i=1}^n I(y_i=k)}$

5. We can first sample  $n$  unlabeled examples,  $S_x = (x_1, \dots, x_n)$ , according to  $P_x$ , and an additional unlabeled example,  $x \sim P_x$ , then find  $\pi_s(x)$  to be the nearest neighbour of  $x$  in  $S_x$ , and finally sample  $y \sim \eta(x)$  and  $y_{\pi_s(x)} \sim \eta(\pi_s(x))$ . Then

$E(E(f^{INN})) = \sum_{x \sim P_x, x \in P} E_{y \sim \eta(x), y' \sim \eta(\pi_s(x))} (1(y \neq y'))$

$= \sum_{x \sim P_x, x \in P} P_{y \sim \eta(x), y' \sim \eta(\pi_s(x))} (y \neq y')$

$P_{y \sim \eta(x), y' \sim \eta(\pi_s(x))} (y \neq y') = \eta(x') (1 - \eta(x)) + (1 - \eta(x')) \eta(x)$

Suppose is  $x'$   $= 2\eta(x) (1 - \eta(x)) + (\eta(x) - \eta(x')) (2\eta(x) - 1)$

$\eta$  is  $c$ -Lipschitz,  $|\eta(x) - \eta(x')| \leq c \|x - x'\|$ , and  $|2\eta(x) - 1| \leq 1$

$$\mathbb{P}_{y \sim \eta(x), y' \sim \eta(x')} (y \neq y') \leq 2\eta(x)(1 - \eta(x)) + c \|x - x'\|$$

$$\text{Hence, } \mathbb{E}_S [\mathcal{E}(f''')] \leq \mathbb{E}_x [2\eta(x)(1 - \eta(x))] + c \mathbb{E}_{S, x} [\|x - x'\|]$$

$$\mathbb{E}_x [\eta(x)(1 - \eta(x))] \leq \mathbb{E}_x [\min\{\eta(x), 1 - \eta(x)\}] = \mathcal{E}(f^*)$$

$$\text{Then, } \mathbb{E}_S [\mathcal{E}(f''')] \leq \mathcal{E}(f^*) + c \mathbb{E}_{S \sim \mathcal{P}^n} \mathbb{E}_{x \sim \mathcal{P}_n} [\|x - x_{S(x)}\|]$$