Multivariate Data Analysis

Grégoire Clarté

Winter-Spring 2024



Welcome

- ► Grégoire Clarté [ˈgʁegwaʁ̞ ˈklaʁ̞te]
- ▶ Lecturer in Statistics
- ▶ Office 5608. Office hours friday 2:30-3:00.
- ► Contact: gclarte@ed.ac.uk
- ► GitHub: https://github.com/GClarte/MultivariateDataAnalysis
- ► All the material will be on Learn

Course structure

- ► Lectures: Monday 14:00-16:00;
- ► Workshops: Friday every other week two sessions per day: 10:00:10:50 and 15:10-16:00

Course Structure

- ► Assessment: 85%, written exam 2h;
- ► Coursework: 15% two assignments to hand over on gradescope, the first one (23/2) is theoretical the second (12/4) practical.

Syllabus

Five parts:

- ► Intro to multivariate analysis;
- ▶ Principal component analysis (PCA);
- ▶ Discriminant Analysis (DA)
- Distance and clustering analysis;
- ► Correspondance analysis (CA).

Textbooks and softwares

- ▶ Applied Multivariate Data Analysis, Everitt & Dunn;
- ► Applied Multivariate Statistical Analysis, Johnson & Wichern;
- ► An introduction to Applied Multivariate Analysis with R, Everitt & Hothorn;
- ► Multivariate Statistical Methods, Manly & Navarro Alberto.
- ▶ R (RStudio).

Introduction

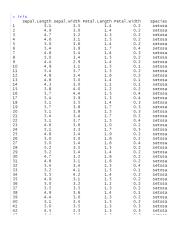
Mathematical context

Observations in \mathbb{R}^d :

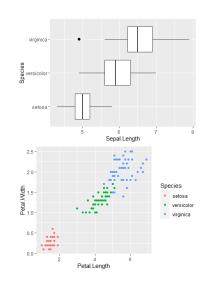
$$X = (X_1, \dots, X_n) = \overbrace{\begin{pmatrix} x_{11} & \cdots & x_{1n} \\ \vdots & & \vdots \\ x_{d1} & \cdots & x_{dn} \end{pmatrix}}^{n \text{ individuals}}$$

Attention: the notation can change in some textbooks

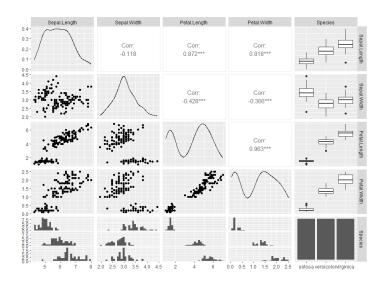
Example in R



Sepal.Length	Sepal.width	Petal.Length	Petal.Width	Species
Min. :4.300	Min. :2.000	Min. :1.000	Min. :0.100	setosa :50
1st Qu.:5.100	1st Qu.:2.800	1st Qu.:1.600	1st Qu.:0.300	versicolor:50
Median :5.800	Median :3.000	Median :4.350	Median :1.300	virginica :50
Mean :5.843	Mean :3.057	Mean :3.758	Mean :1.199	
3rd Qu.:6.400	3rd Qu.:3.300	3rd Qu.:5.100	3rd Qu.:1.800	
Max. :7.900	Max. :4.400	Max. :6.900	Max. :2.500	

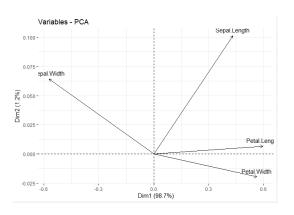


Example in R

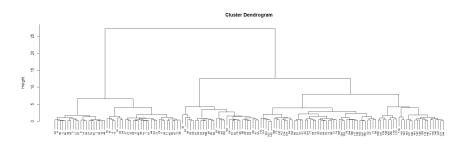


 11

Example in R



Examples in R



Which method to choose?

Depends on the goals:

- ▶ What is the nature of the variables?
- ▶ what are the dependencies?
- ▶ what are the variables of interest?

Several techniques are of general use:

- ▶ Dimension reduction, variable selection (finance, ecology, etc.);
- correlations (i.e. relationships) between variables;
- ► classification/prediction.

What is MDA?

Descriptive or inferential?

- ▶ Descriptive: sample proportion, means, standard deviation, correlations, covariances, etc.
- ► Inferential: equivalent of the univariate models but in larger dimension (MANOVA, (generalised) linear regression, etc.)

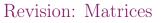
Representation of multivariate data is a challenge.

Main issue: correlation of variables \Rightarrow you (the statistician) have to be aware of that.

MDA is the base of Data Mining.

Goals of MDA

- ➤ Simplify the data: selection of important variables, makes it more readable;
- sorting and grouping: classification of objects;
- studying the dependencies: independent variables, conditionally independent variables, correlated variables, etc.
- ▶ inference/prediction: determine relationships, carry inference on a model, extrapolate from it;
- ▶ hypothesis testing: tests in higher dimension.



Notations

$$d \times 1 \text{ vector } X = \begin{pmatrix} x_1 \\ \vdots \\ x_d \end{pmatrix};$$

- ▶ transposed vector $X' = X^{\top} = (x_1, \dots, x_d);$
- ▶ transposed matrix $A^{\top} = (a_{ji})_{ij}$.

Definitions

- ▶ If p = q the matrix is square;
- ightharpoonup if $\forall i, j, \ a_{ij} = 0, \ A = 0$, the zero (null) matrix;
- ▶ If p = q, and if $a_{ij} = 0, i \neq j$; $a_{ii} = 1, \forall i$ then $A = I_d$ the identity matrix;
- ▶ the set of $p \times q$ matrices with coefficient in \mathbb{R} is written $\mathcal{M}_{p,q}(\mathbb{R})$.

For A a square matrix:

- ightharpoonup A is symmetric iif $A^{\top} = A$;
- ▶ A is antisymmetric (skew-symmetric) iif $A^{\top} = -A$;
- ightharpoonup A is diagonal iif $a_{ij} = 0, \forall i \neq j$.

Definitions

▶ Matrix multiplication: if $A \in \mathcal{M}_{p,q}(\mathbb{R})$, $B \in \mathcal{M}_{q,r}(\mathbb{R})$: $C = AB := (c_{i,j}) \in \mathcal{M}_{p,r}(\mathbb{R})$, where $c_{ij} = \sum_{k=1}^{q} a_{ik} b_{kj}$.

Property: $(AB)^{\top} = B^{\top}A^{\top}, (A^{\top})^{\top} = A.$

For a square matrix A:

- ightharpoonup A is orthogonal iif $AA^{\top} = A^{\top}A = I$;
- ▶ A is idempotent of order ℓ iif $A^{\ell} = \underbrace{A \times \cdots \times A}_{\ell \text{ times}} = A$
- ▶ if B is such that AB = I, then B is the inverse matrix of A, written A^{-1} .

The trace of the matrix A is $\operatorname{Tr}(A) = \sum_{i=1}^{p} a_{ii}$. The trace is linear (i.e. $\operatorname{Tr}(A + \lambda B) = \operatorname{Tr}(A) + \lambda \operatorname{Tr}(B)$); $\operatorname{Tr}(A^{\top}) = \operatorname{Tr}(A)$, and $\operatorname{Tr}(AB) = \operatorname{Tr}(BA)$.

Eigenvalues

Definition 1

 λ is an eigenvalue of A a square matrix if there exists $x \neq 0$ such that:

$$Ax = \lambda x \Leftrightarrow (A - \lambda I)x = 0.$$

We call x the eigenvector, the set of eigenvectors for an eigenvalue is called the eigenspace, and the set of eigenvalues is the spectrum of the matrix.

All the eigenvalues are roots of the characteristic polynomial in λ , $p(\lambda) = \text{Det}(A - \lambda I)$.

 21

Eigenvalues

If we can write:

$$p(\lambda) = \prod_{i=1}^{n} (\lambda - \lambda_i)^{n_i},$$

where n_i is the algebraic multiplicity, N_i the number of unique eigenvalues and $\sum_i n_i = d$.

When solving $(A - \lambda_i)e_i = 0$, there will be $m_i \leq n_i$ linearly independent solutions, m_i is the *geometric* multiplicity. In general algebraic multiplicity is larger than geometric multiplicity. In general not all matrices have eigenvalues (c.f. rotation matrices).

Example

 $A = \begin{pmatrix} -1 & 2 \\ -3 & 4 \end{pmatrix}$, what are the eigen(value/vector/space)? Characteristic polynomial:

$$p_A(\lambda) = \lambda^2 - 3\lambda + 2 = (\lambda - 1)(\lambda - 2).$$

After solving $\begin{pmatrix} -2 & 2 \\ -3 & 3 \end{pmatrix} x = 0$, and $\begin{pmatrix} -3 & 2 \\ -3 & 2 \end{pmatrix} x = 0$, we get eigenvectors: $e_1 = (1,1)$ and $e_2 = (2,3)$.

$$A = \begin{pmatrix} 1 & -1 \\ 2 & 2 \end{pmatrix}$$

Which has characteristic polynomial: $p_A(\lambda) = \lambda^2 - 3\lambda + 4$. This polynomial is irreducible (in \mathbb{R}).

Properties

- \blacktriangleright A and A^{\top} have the same eigenvalues;
- ▶ A is singular (i.e. non-invertible) iif $0 \in Sp(A)$;
- $\blacktriangleright \operatorname{Tr}(A) = \sum_{\lambda \in Sp(A)} \lambda;$
- $\blacktriangleright \ \mathrm{Det}(A) = \prod_{\lambda \in Sp(A)} \lambda;$
- ▶ if A is invertible, with eigenvalues λ_i , the eigenvalues of A^{-1} are λ_i^{-1} with the same eigenvectors and multiplicities.



Multidimensional RV

Notations:

► Random Vector:
$$X = \begin{pmatrix} X_1 \\ \vdots \\ X_p \end{pmatrix} = (X_1, \dots, X_p)^\top.$$

 $ightharpoonup x \in \mathbb{R}^p$ is a realisation of X.

Reminder:

A random vector is a function from (Ω, \mathcal{F}, P) a probability space to \mathbb{R}^p .

Revisions

Definition

A probability mass function (probability density function) with respect to the Lebesgue measure in \mathbb{R}^d is a function:

$$f: \begin{cases} \mathbb{R}^p \to \mathbb{R} \\ x \mapsto f(x) \end{cases}$$

such that, for any $O \subset \mathbb{R}^p$ measurable subset (e.g. Open),

$$\int_{O} f(x) \mathrm{d}x = P(X \in O).$$

In the discrete case (i.e. with respect to the counting measure on \mathbb{R}^p), f(x) = P(X = x).

Note: Proper definition of these objects require measure theory, the use of these objects do not require measure theory in the simple cases.

Properties

Let f be a pdf.

- $\blacktriangleright \ \forall x \in \mathbb{R}^p, \ f(x) \le 0;$
- ▶ $\int_{\mathbb{R}^p} f(x) dx = 1$, and in the discrete case $\sum_x f(x) = 1$.

If we write $X=(Y,Z), Y\in\mathbb{R}^s, Z\in\mathbb{R}^t, t+s=p$, the marginal pdf of Y and Z are $f_Y(y)=\int_{\mathbb{R}^t}f(y,z)\mathrm{d}z$, and resp. for f_Y . Once again, note the discrete case: $f_Y(y)=\sum_z f(y,z)$.

The cumulative density function (cdf) is defined for a real random variable X of density f by:

$$F_X(t) = \int_{-\infty}^t f(x) dx.$$

Conditional distributions

We write $X = (Y, Z), Y \in \mathbb{R}^s, Z \in \mathbb{R}^t, t + s = p$. The conditional density is *defined* as:

$$f(y \mid z) = \frac{f(y, z)}{f_Z(z)}.$$

In the discrete case:

$$P(Y = y \mid Z = z) = \frac{P(Y = y, Z = z)}{P(Z = z)}$$

Theorem 2

$$f(y,z) = f_Y(y)f(z \mid y).$$

This is the total probabilities theorem.

(In)dependencies

Two r.v. Y and Z are independent iif their density writes:

$$f(y,z) = f_Y(y)f_Z(z).$$

If Y and Z are independent, then:

$$f(y \mid z) = f_Y(y).$$

Expectation

Let g be a measurable function, we define for f the density of X:

$$\mathbb{E}[g(X)] = \int_{\mathbb{R}^p} g(x)f(x)dx$$

in the continuous (Lebesgue) case, and

$$\mathbb{E}[g(X)] = \sum_{x} g(x)P(X = x)$$

in the discrete (counting) case.

Properties:

- $ightharpoonup \mathbb{E}[k] = k, \ k \in \mathbb{R}^p,$
- ► $E[\alpha g(X) + \beta h(X)] = \alpha \mathbb{E}[g(X)] + \beta \mathbb{E}[h(X)]$, for g and h measurable.

The mean is an element of the space of destination of g. In particular, $E[X] \in \mathbb{R}^p$ is called the *mean vector*.

(Co)variance

If Y is a real random variable, its variance is defined as:

$$\sigma_Y^2 = \mathbb{E}[(Y - \mathbb{E}[Y])^2].$$

Let X = (Y, Z), with $Y, Z \in \mathbb{R}$ (i.e. bidimensional rv).

$$\sigma_{Y,Z} = Cov(Y,Z) = \mathbb{E}[(Y - \mathbb{E}[Z])(Z - \mathbb{E}[Z])].$$

Note that $\sigma_{Y,Z} = \sigma_{Z,Y}$, and that $\sigma_{Y,Z} = \mathbb{E}[YZ] - \mathbb{E}[Y]\mathbb{E}[Z]$. The covariance matrix of X is defined as:

$$\Sigma_X = \begin{pmatrix} \sigma_Y^2 & \sigma_{Y,Z} \\ \sigma_{Z,Y} & \sigma_Z^2 \end{pmatrix}.$$

- \triangleright Σ_X is a square, symmetric, positive semidefinite matrix.
- ► Y and Z are independent $\Rightarrow \sigma_{Y,Z} = 0$.

Properties of Covariance

- $ightharpoonup Cov(Y, \alpha) = 0;$
- ▶ the covariance is bilinear:

$$Cov(\alpha Y + \beta Z, W) = \alpha Cov(Y, W) + \beta Cov(Z, W);$$

Var(Y+Z) = Cov(Y+Z,Y+Z) = V(Z) + V(Y) + 2Cov(Y,Z).

Correlation coefficient

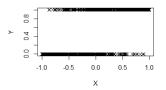
The correlation coefficient is defined as:

$$\rho_{Y,Z} = \frac{\sigma_{Y,Z}}{\sigma_Y \sigma_Z}.$$

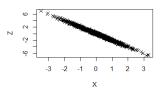
- ► It accounts for the degree of linear relationship between the two r.v.
- ▶ $-1 \le \rho_{Y,Z} \le 1$;
- $\rho = 0$ no linear relation, no correlation
- ightharpoonup
 ho = 1 exact direct linear relationship
- $\rho = -1$ exact inverse linear relationship

Example

▶ If $X_i \sim \mathcal{U}(-1, 1)$, and $Y_i \mid X_i \sim \mathcal{B}\left(\frac{arctan(X_i) - arctan(1)}{arctan(1) - arctan(-1)}\right)$, we have $\rho_{X,Y} \simeq 0.65$.

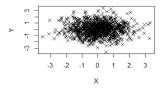


▶ if $X_i \sim \mathcal{N}(0,1)$, and $Z_i \mid X_i \sim \mathcal{N}(-2X_i, 0.2)$, $\rho_{X,Z} \simeq -0.99$.

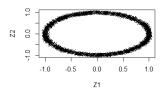


Example

▶ if $X_i, Y_i \sim \mathcal{N}(0, 1)$ independent: $\rho \simeq 0$



▶ $X_i \sim \mathcal{N}(1, 0.02)$, $Y_i \sim \mathcal{U}(0, 2\pi)$ independent of X_i . We define $Z_i^1 = X_i \cos(Y_i)$ and $Z_i^2 = X_i \sin(Y_i)$.



With a correlation coefficient: $\rho \simeq -0.01$

Higher dimension

All these properties and definition can be extended to any dimension problem.

In \mathbb{R}^p the covariance matrix Σ is $p \times p$, the mean is a vector in \mathbb{R}^p , etc.

Let A be a matrix of size $m \times p$, and X a random vector in \mathbb{R}^p , linearity writes:

 $\mathbb{E}[AX] = A\mathbb{E}[X]$, and $\Sigma_{AX} = A\Sigma_X A^{\top}$.

Gaussian RV

$$p = 1$$
 $f(x) = \frac{1}{\sigma\sqrt{2\pi}} \exp\left(-\frac{1}{2}\left(\frac{x-m}{\sigma}\right)^2\right)$

$$p > 1$$
 $f(x) = \frac{1}{(2\pi)^{p/2}|\Sigma|^{1/2}} \exp\left(-\frac{1}{2}(x-m)^{\top}\Sigma^{-1}(x-m)\right).$

Note: $(x-m)^{\top} \Sigma^{-1} (x-m)$ is called Mahalanobis distance between x and m associated with Σ .



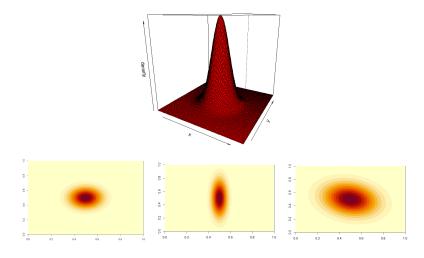
Figure: Prasanta Chandra Mahalanobis (1893-1972)

Gaussian RV

If $X \sim \mathcal{N}(m, \Sigma)$

- ▶ $\mathbb{E}[X] = m$, $Cov(X) = \Sigma$ (thus, Σ is symmetric, definite positive).
- ▶ If the variables are highly correlated: the eigenvalues of Σ are close to 0, $|\Sigma|$ is small.
- Any linear combination of X is Gaussian: $AX + \beta$ is also normal. Compute its mean and covariance (exercice).
- ▶ Special cases if A is $1 \times p$: projection of the vector X, $AX \in \mathbb{R}$.
- ▶ We can reduce any Gaussian RV to a **centered reduced** gaussian distribution (i.e. with mean 0 and with variance I_p) (**exercice**).

Some plots



Related distribution: Chi-square

▶ If
$$X \sim \mathcal{N}(0, I_p)$$
, then $||X||^2 = \sum_{i=1}^p X_i^2 = X^\top X \sim \chi_p^2$.

Conditional distributions

All marginal and conditional distributions are also **normal**.

- ▶ If we take A = (1, 0, ..., 0), AX corresponds to the 1st marginal.
- same for any subset of dimensions $A = (0, \dots, 1, \dots, 1, \dots, 0)$.

Consequence, if X and Y are Gaussian, we can write their joint distribution:

$$\mathbb{E}\left[\begin{pmatrix} X \\ Y \end{pmatrix}\right] = \begin{pmatrix} \mu_X \\ \mu_Y \end{pmatrix}, \ Cov(X,Y) = \begin{pmatrix} \Sigma_{XX} & \Sigma_{XY} \\ \Sigma_{YX} & \Sigma_{YY} \end{pmatrix}.$$

Conditional distributions

Note: For Gaussian distributions, Independence \Leftrightarrow Uncorrelation. That is, if $\Sigma_{XY}=0$ then X and Y are independent.

How to compute the conditional distribution of X and Y?

- we have the joint distribution and the marginal that are all Gaussian.
- ▶ After computation it is multivariate normal:

$$\mathbb{E}[X \mid Y] = \mu_X + \Sigma_{XY} \Sigma_{YY}^{-1} (y - \mu_Y)$$
$$Cov(X, Y) = \Sigma_{XX} - \Sigma_{XY} \Sigma_{YY}^{-1} \Sigma_{YX}$$

Geometric interpretation of the Covariance matrix

- \triangleright Σ is symmetrical definite positive.
- Level lines of the Gaussian density are of the form $x^{\top} \Sigma x = \alpha$.
- ▶ What does $\{x \in \mathbb{R}^p \mid x^{\top} \Sigma x = 1\}$ look like?
- ightharpoonup is symmetric definite positive...

