

Tutorial 1: Multidimensional Random Variables

Ex. 1 $P((X, Y) = (x, y)) = 0$ if $y > x$, $P((X, Y) = (x, y)) = P(W_1 = x, W_2 = y) + P(W_1 = y, W_2 = x) = 2/f$ else. Where W_1 and W_2 are the results of the dices and f the number of faces of the dices (usually $f \geq 4$). Of course X and Y are not independent, $P(X = 6, Y = 1) = 0 \neq P(X = 6)P(Y = 1)$.

This is an example of *order statistic* that is often used and has lots of good properties.

Ex. 2 The density writes $\mathbf{1}_{(x,y) \in D}$, where D is the disk of center 0 and radius 1. To compute the means of x and y we just need to see that the distribution is symmetric w.r.t 0, and thus $\mathbb{E}[X] = \mathbb{E}[y] = 0$.

For the covariance we need to compute:

$$\begin{aligned} \mathbb{E}[(X - \mathbb{E}[X])(Y - \mathbb{E}[Y])] &= \mathbb{E}[XY] = \int xy \mathbf{1}_{(x,y) \in D} dx dy \\ &= \pi \int_{x=-1}^1 \int_{y=-\sqrt{1-x^2}}^{\sqrt{1-x^2}} xy dx dy \\ &= \int_{-1}^1 0 dx = 0. \end{aligned}$$

So the variables are uncorrelated.

Nevertheless they are not independent, as $P((X, Y) \in [0.9, 0.9]^2) = 0 \neq P(X \geq 0.9)P(Y \geq 0.9)$.

Ex. 3 To determine the distribution of $Z = X/Y$ we need to compute $\mathbb{E}[f(Z)]$ for any f . Note that the value when $Y = 0$ is not important as the densities are defined almost everywhere. We use the change of variable $(x, y) \mapsto (x/y, y) = (z, w)$ which is a C^1 diffeomorphism. This transformation has Jacobian $|w|$

$$\begin{aligned} \mathbb{E}[f(Z)] &= \frac{1}{2\pi} \int_{\mathbb{R}^2} f(x/y) \exp(-x^2/2 - y^2/2) dx dy \\ &= \frac{1}{2\pi} \int_{\mathbb{R}^2} f(z) |w| \exp(-x^2/2 - y^2/2) dx dy. \\ &= \frac{1}{2\pi} \int_{\mathbb{R}^2} f(z) |w| \exp(-w^2(z^2 + 1)/2) dz dw. \end{aligned}$$

We recognize that w follows conditionally on z a normal distribution with variance $1/\sqrt{z^2 + 1}$, we just need to add the constants so that it integrates to 1.

$$\begin{aligned}
&= \frac{1}{2\pi} \int_{\mathbb{R}} f(z) \int_{\mathbb{R}} \frac{1}{\sqrt{2\pi(z^2+1)}} |w| \exp(-w^2(z^2+1)/2) dw dz \\
&= \int_{\mathbb{R}} f(z) \frac{1}{2\pi} \frac{1}{\sqrt{2\pi(z^2+1)}} \frac{2\sqrt{2}}{\sqrt{\pi(z^2+1)}} dz \\
&= \int_{\mathbb{R}} f(z) \frac{1}{\pi(z^2+1)} dz.
\end{aligned}$$

Which corresponds to the standard Cauchy distribution.

Ex. 4 Similarly as the previous exercise, we use the change of variable $(x, y) = (r \cos(\theta), r \sin(\theta))$, which has jacobian r .

$$\begin{aligned}
&\int_{\mathbb{R}^2} f(\sqrt{x^2+y^2}) \exp(-(x^2+y^2)/2) 2\pi dx dy \\
&= \int_{\mathbb{R}^+ \times [0, 2\pi]} f(r) \exp(-r^2/2) r (2\pi) dr d\theta \\
&= \int_{\mathbb{R}^+} f(r) r \exp(-r^2/2) dr.
\end{aligned}$$

Which corresponds to the law $\chi(n)$ (that is the square root of the χ^2).

To see if X and R are independent, it suffices to say that knowing R , $X \in [-R, R]$, while marginally X has value on \mathbb{R} . We can compute the conditional distribution of X knowing R using once again the polar coordinates.

For that, we know that knowing R the point is uniformly distributed on the circle (see previous integral). We can proceed similarly, using the change of variable $\theta = \arccos(x/r)$ with jacobian $1/\sqrt{r^2-x^2}$.

$$\begin{aligned}
\mathbb{E}[f(X) \mid R] &= \int_{[0, 2\pi]} f(r \cos(\theta)) d\theta \\
&= \int_{[0, \pi]} f(r \cos(\theta)) d\theta + \int_{[-\pi, 0]} f(r \cos(\theta)) d\theta \\
&= \int_{[-R, R]} f(x) \frac{1}{\sqrt{r^2-x^2}} dx
\end{aligned}$$

Ex. 5 Let's start by computing the marginals:

$$f_X(x) = \int y = x^2 \sqrt{x} 3dy = 3(\sqrt{x}-x^2). \quad (1)$$

Note that this is perfectly symmetric for Y (D is symmetric w.r.t. $x \leftrightarrow y$) and thus that:

$$f_Y(y) = \sqrt{y} - y^2.$$

Similarly, the conditionals will be symmetric:

$$f_{X|Y} = \mathbf{1}_D(x, y) \frac{1}{\sqrt{y} - y^2}.$$

And we can compute the mean:

$$\mathbb{E}[X | Y = y] = \int_R \mathbf{1}_D(x, y) \frac{x}{\sqrt{y} - y^2} = \frac{y - y^4}{\sqrt{y} - y^2}.$$

Ex. 5 This events rewrite $p = P(x \in [0, 1] \cap [y - 1/4, y + 1/4])$. To compute this probability we need to differentiate the case $y \geq 3/4$ and $y \leq 3/4$. In the end we have:

$$\begin{aligned} p &= \int_{y=1/4}^{3/4} 1/2 dy + 2 \int_{y=0}^{1/4} 1/4 + y dy \\ &= 1/4 + 2(1/16 + 1/32) = 1/4 + 3/16 = 7/16. \end{aligned}$$

Ex. 6 We start by writing the joint distribution of (y_1, \dots, y_n, z) where $z = \sum y_i$:

$$P(Y_1 = y_1, \dots, Y_n = y_n, Z = z) = \begin{cases} 0 & \text{if } z \neq \sum y_i \\ P(Y_1 = y_1, \dots, Y_n = y_n) & \text{otherwise} \end{cases}$$

The marginal density of Z is Poisson (famous property) with parameter $\lambda_1 + \dots + \lambda_n$. Then we can simply write the conditonal density:

$$P(Y_1 = y_1, \dots, Y_n = y_n | \sum Y_i = n) = \frac{\lambda_1^{y_1} \dots \lambda_n^{y_n} \exp(-\sum \lambda_i)}{y_1! \dots y_n!} \frac{n!}{(\sum \lambda_i)^n \exp(-\sum \lambda_i)}.$$

Which is exactly what we were looking for.

Ex. 7 The sum of two normally distributed r.v. is normally distributed, which leads to: $x + y \sim \mathcal{N}(1, 2)$, and then $z \sim \mathcal{N}(7/3, 2/9)$.

Ex. 8

1. We need to have $\sum_{y \in \{1, 2, 3\}, x \in \{1, 2, 3, 4\}} c(x + y) = 1$. The sum equals $54c$ and thus $c = 1/54$.
2. The marginals are $P(x) = \sum_y c(x + y) = c(3x + 6)$ and $P(y) = c(10 + 4y)$.
3. The conditional is then $P(x | y) = \frac{x+y}{10+4y}$, and similarly for the other conditional.
4. These variables are clearly not independent, the conditional densities show it.

Ex. 9

1. T is the minimum of two independent Exponential distributions. We have then: $P(T \geq t) = P(X \geq t)P(Y \geq t) = \exp(-2\lambda t)$, thus $T \sim \mathcal{E}(2\lambda)$.

2. We can prove that the exponential distribution is *memoryless* for that it is sufficient to prove that $P(X \geq \ell + t \mid X \geq \ell) = P(X \geq t)$, for that we write the conditional probability: $P(X \geq \ell + t \mid X \geq \ell) = P(X \geq \ell + t \cap X \geq \ell) / P(X \geq \ell) = P(X \geq \ell + t) / P(X \geq \ell) = \exp(-\lambda(\ell + t)) / \exp(-\lambda\ell)$. Knowing that, we conclude that the time U has same distribution as Y , and that the second person will stay twice as long in mean knowing that they did not leave first.

Ex. 10 It suffices to investigate the distribution of $AD - CB$, which has an absolutely continuous distribution, as A, B, C and D have absolutely continuous distributions and are independent. That is $P(AD - CB = 0) = P(AB = CD) + P(AB = 0, CD = 0) = 0$ because everything is independent and is absolutely continuous

Ex. 11

$$Cov(Y, Z) = Cov(X - \rho Y, Y) = Cov(X, Y) - \rho Cov(Y, Y) = \rho - \rho = 0,$$

where we used the fact that if $Var(X) = Var(Y) = 1$ then $Cov(X, Y) = \rho$.