

## 课后题答案 1-5 章（部分）

### 第一章

1.13

$$\begin{cases} U = X + Y & U \in (0, +\infty) \\ V = \frac{X}{X + Y} & V \in (0, 1) \end{cases}$$

可得

$$\begin{aligned} X &= UV \\ Y &= U(1 - V) \end{aligned}$$

计算二维雅可比变换的雅可比因子

$$J = \begin{vmatrix} \frac{\partial x}{\partial u} & \frac{\partial x}{\partial v} \\ \frac{\partial y}{\partial u} & \frac{\partial y}{\partial v} \end{vmatrix} = \begin{vmatrix} v & u \\ 1-v & -u \end{vmatrix} = -uv - u + uv = -u$$

$$\begin{aligned} f(u, v) &= f(x, y) |J| \\ &= \lambda^2 e^{-\lambda(uv + u - uv)} \cdot |-u| \\ &= \lambda^2 u e^{-\lambda u} \end{aligned}$$

$$f(u, v) = \begin{cases} \lambda^2 u e^{-\lambda u} & u > 0, 0 < v < 1 \\ 0 & \text{其他} \end{cases}$$

$$f(u) = \int_0^1 \lambda^2 u e^{-\lambda u} dv = \lambda^2 u e^{-\lambda u}$$

$$f(u) = \begin{cases} \lambda^2 u e^{-\lambda u} & u > 0 \\ 0 & \text{其他} \end{cases}$$

$$f(v) = \int_0^\infty \lambda^2 u e^{-\lambda u} du = 1$$

$$f(v) = \begin{cases} 1 & 0 < v < 1 \\ 0 & \text{其他} \end{cases}$$

1.18

$$f(x, y) = f(x)f(y) = \frac{1}{2\sigma^2} \exp\left\{-\frac{x^2 + y^2}{2\sigma^2}\right\}, -\infty < x < \infty, -\infty < y < \infty$$

(1) 由  $Z = \sqrt{X^2 + Y^2}$ , 可得  $X = \pm\sqrt{Z^2 - Y^2}$

$$\begin{aligned} F_Z(z) &= P\{Z \leq z\} = P\{\sqrt{X^2 + Y^2} \leq z\} \\ &= \int_{-z}^z \int_{-\sqrt{z^2 - y^2}}^{\sqrt{z^2 - y^2}} f(x, y) dx dy \\ &= 4 \int_0^z \int_0^{\sqrt{z^2 - y^2}} \frac{1}{2\pi\sigma^2} \exp\left\{-\frac{x^2 + y^2}{2\sigma^2}\right\} dx dy \\ f(z) &= \frac{dF_Z(z)}{dz} = 4 \int_0^z \frac{1}{2\pi\sigma^2} \exp\left\{-\frac{z^2}{2\sigma^2}\right\} (\sqrt{z^2 - y^2})' dy \\ &= \frac{2z}{\pi\sigma^2} \exp\left\{-\frac{z^2}{2\sigma^2}\right\} \int_0^z \frac{1}{\sqrt{z^2 - y^2}} dy \\ &= \frac{2z}{\pi\sigma^2} \exp\left\{-\frac{z^2}{2\sigma^2}\right\} \left(\arcsin \frac{y}{z}\right) \Big|_0^z \\ &= \frac{z}{\sigma^2} \exp\left\{-\frac{z^2}{2\sigma^2}\right\} \end{aligned}$$

$$\text{所以 } f_Z(z) = \begin{cases} \frac{z}{\sigma^2} \exp\left\{-\frac{z^2}{2\sigma^2}\right\} & z \geq 0 \\ 0 & \text{其他} \end{cases}$$

$$\begin{aligned}
F_W(w) &= P\left\{\frac{X}{Y} \leq w\right\} \\
&= P\{X \leq wY, Y > 0\} + P\{X \geq wY, Y < 0\} \\
&= \int_{y=0}^{y=\infty} \int_{x=-\infty}^{x=wy} f_{XY}(x, y) dx dy + \int_{y=-\infty}^{y=0} \int_{x=wy}^{x=\infty} f_{XY}(x, y) dx dy \\
&= 2 \int_{y=0}^{y=\infty} \int_{x=-\infty}^{x=wy} f_{XY}(x, y) dx dy
\end{aligned}$$

$$\begin{aligned}
f_W(w) &= \frac{dF_W(w)}{dw} = 2 \int_0^\infty y f_{XY}(yw, y) dy \\
&= 2 \int_0^\infty y \frac{1}{2\pi\sigma^2} \exp\left\{-\frac{(w^2+1)y^2}{2\sigma^2}\right\} dy \\
&= \frac{1}{\pi(w^2+1)}
\end{aligned}$$

$$f_W(w) = \frac{1}{\pi(w^2+1)} \quad (-\infty < w < \infty)$$

$$\begin{aligned}
F_{\Theta}(\theta) &= P\{\Theta \leq \theta\} = P\left\{\frac{X}{Y} \leq \tan \theta\right\} \\
&= P\{X \leq Y \tan \theta, Y > 0\} + P\{X \geq Y \tan \theta, Y < 0\} \\
&= 2 \int_0^{\infty} \int_{-\infty}^{y \tan \theta} f_{XY}(x, y) dx dy \\
&= 2 \int_0^{\infty} \int_{-\infty}^{y \tan \theta} \frac{1}{2\pi\sigma^2} \exp\left\{-\frac{x^2 + y^2}{2\sigma^2}\right\} dx dy \\
f_{\Theta}(\theta) &= \frac{dF_{\Theta}}{d\theta} = 2 \int_0^{\infty} \frac{y}{\cos^2 \theta} \exp\left\{-\frac{(y \tan \theta)^2 + y^2}{2\sigma^2}\right\} dy \\
&= \frac{2}{\pi} \int_0^{\infty} \frac{y}{2\sigma^2 \cos^2 \theta} \exp\left\{-\frac{y^2}{2\sigma^2 \cos^2 \theta}\right\} dy \\
&= \frac{1}{\pi} \int_0^{\infty} \exp\left\{-\frac{y^2}{2\sigma^2 \cos^2 \theta}\right\} d\left(\frac{y^2}{2\sigma^2 \cos^2 \theta}\right) \\
&= \frac{1}{\pi} \int_0^{\infty} \exp\{-u\} du \\
&= \frac{1}{\pi} \\
f_{\Theta}(\theta) &= \begin{cases} \frac{1}{\pi} & -\frac{\pi}{2} < \theta < \frac{\pi}{2} \\ 0 & \text{其他} \end{cases}
\end{aligned}$$

## 第二章

### 2.1

(1) 令  $t_i = 0$ , 则  $X(t_i = 0) = A$ , 有

$$f_X(x; 0) = f_A(x) \left| \frac{dA}{dX} \right| = \begin{cases} 1 & 0 \leq x \leq 1 \\ 0 & \text{其他} \end{cases}$$

令  $t_i = \frac{\pi}{4\omega}$ , 则  $X(t_i = \frac{\pi}{4\omega}) = \frac{\sqrt{2}}{2}A$ , 有

$$f_X(x; \frac{\pi}{4\omega}) = f_A(\sqrt{2}x) \left| \frac{dA}{dX} \right| = \begin{cases} \sqrt{2} & 0 \leq x \leq \frac{\sqrt{2}}{2} \\ 0 & \text{其他} \end{cases}$$

令  $t_i = \frac{3\pi}{4\omega}$ , 则  $X(t_i = \frac{3\pi}{4\omega}) = -\frac{\sqrt{2}}{2}A$ , 有

$$f_X(x; \frac{3\pi}{4\omega}) = f_A(-\sqrt{2}x) \left| \frac{dA}{dX} \right| = \begin{cases} \sqrt{2} & -\frac{\sqrt{2}}{2} \leq x \leq 0 \\ 0 & \text{其他} \end{cases}$$

令  $t_i = \frac{\pi}{\omega}$ , 则  $X(t_i = \frac{\pi}{\omega}) = -A$ , 有

$$f_X(x; \frac{\pi}{\omega}) = f_A(-x) \left| \frac{dA}{dX} \right| = \begin{cases} 1 & -1 \leq x \leq 0 \\ 0 & \text{其他} \end{cases}$$

(2) 令  $t_i' = \frac{\pi}{2\omega}$ , 则  $X(t_i') = 0$ , 有

$$f_X(x) = \delta(x)$$

一维随机变量的函数的分布函数, 公式:  $f_X(x) = f_A(a) \left| \frac{dA}{dX} \right|$

$X(t_i) = A \cos \omega t_i$ , 假设  $0 < t_i < \frac{\pi}{2\omega}$

$$f(x_i) = \begin{cases} \frac{1}{\cos \omega t_i} & 0 \leq x_i \leq \cos \omega t_i \\ 0 & \text{其他} \end{cases}$$

$$\int_0^{\cos \omega t_i} f(x_i) dx_i = 1$$

当  $\cos \omega t_i \rightarrow 0^+$  时,  $f(x_i) \rightarrow \infty$

### 2.13

$$f_{A\Theta}(a, \theta) = f_A(a)f_{\Theta}(\theta) = \begin{cases} \frac{a}{2\pi\sigma^2} \exp(-\frac{a^2}{2\sigma^2}) & a > 0, 0 \leq \theta \leq 2\pi \\ 0 & \text{其他} \end{cases}$$

$$m_X(t) = E[X(t)]$$

$$\begin{aligned} &= \int_{-\infty}^{\infty} [\int_{-\infty}^{\infty} f_A(a) f_{\Theta}(\theta) a \cos(\omega t + \theta) d\theta] da \\ &= \int_0^{\infty} \frac{a^2}{\sigma^2} \exp(-\frac{a^2}{\sigma^2}) da \int_0^{2\pi} \frac{1}{2\pi} \cos(\omega t + \theta) d\theta \\ &= 0 \end{aligned}$$

$m_X(t)$  与时间无关

$$R_X(t_1, t_2) = E[X(t_1)X(t_2)]$$

$$\begin{aligned} &= \int_{-\infty}^{\infty} [\int_{-\infty}^{\infty} a^2 \cos(\omega t_1 + \theta) \cos(\omega t_2 + \theta) f_A(a) f_{\Theta}(\theta) d\theta] da \\ &= \int_0^{\infty} \frac{a^2}{2\sigma^2} \exp(-\frac{a^2}{2\sigma^2}) da^2 \int_0^{2\pi} \frac{\cos[\omega(t_1 + t_2) + 2\theta] + \cos[\omega(t_1 - t_2)]}{2} \cdot \frac{1}{2\pi} d\theta \\ &= \sigma^2 \cos[\omega(t_1 - t_2)] \end{aligned}$$

令  $\tau = t_1 - t_2$ , 则  $R_X(\tau) = \sigma^2 \cos \omega \tau$

又因为  $E[X^2(t)] = R_X(0) = \sigma^2 < \infty$

故  $X(t)$  是平稳随机过程

(1)

$$\begin{aligned}R_z(\tau) &= E[Z(t+\tau)Z(t)] \\&= E[X(t+\tau)Y(t+\tau)X(t)Y(t)] \\&= E[X(t+\tau)X(t)]E[Y(t+\tau)Y(t)] \\&= R_X(\tau) \cdot R_Y(\tau)\end{aligned}$$

$$S_z(\omega) = \frac{1}{2\pi} S_X(\omega) * S_Y(\omega)$$

(2)

$$S_X(\omega) = \frac{\sin^2(\omega/2)}{(\omega/2)^2}$$

$$R_Y(\tau) = \cos \omega_0 \tau, \quad S_Y(\omega) = \frac{\pi}{2} [\delta(\omega - \omega_0) + \delta(\omega + \omega_0)]$$

$$S_z(\omega) = \frac{1}{4} \left\{ \frac{\sin^2[(\omega - \omega_0)/2]}{[(\omega - \omega_0)/2]^2} + \frac{\sin^2[(\omega + \omega_0)/2]}{[(\omega + \omega_0)/2]^2} \right\}$$

2.29

$$S_X(\omega) = \frac{\omega^2 + 4}{\omega^4 + 10\omega^2 + 9} = \frac{3}{8} \cdot \frac{1}{\omega^2 + 1} + \frac{5}{8} \cdot \frac{1}{\omega^2 + 9}$$

$$R_X(\tau) = \frac{3}{16} e^{-|\tau|} + \frac{5}{48} e^{-3|\tau|}, \quad \psi_X^2 = R_X(0) = \frac{7}{24}$$

$$S_Y(\omega) = \frac{\omega^2}{\omega^4 + 3\omega^2 + 2} = -\frac{1}{\omega^2 + 1} + \frac{2}{\omega^2 + 2}$$

$$R_Y(\tau) = -\frac{1}{2} e^{-|\tau|} + \frac{1}{\sqrt{2}} e^{-\sqrt{2}|\tau|}, \quad \psi_Y^2 = R_Y(0) = \frac{\sqrt{2}-1}{2}$$

2.30

由题图可得  $Y(t) = X(t) + X(t+T)$

$$E[Y(t)Y(t-\tau)] = E\{[X(t) - X(t-T)][X(t-\tau) - X(t-T-\tau)]\}$$

$$= R_X(\tau) - R_X(\tau+T) - R_X(\tau-T) + R_X(\tau)$$

$$= 2R_X(\tau) - R_X(\tau+T) - R_X(\tau-T)$$

$$S_Y(\omega) = \int_{-\infty}^{\infty} [2R_X(\tau) - R_X(\tau+T) - R_X(\tau-T)] e^{-j\omega\tau} d\tau$$

$$= 2S_X(\omega) + \int_{-\infty}^{\infty} R_X(\tau') e^{-j\omega(\tau'-T)} d\tau' + \int_{-\infty}^{\infty} R_X(\tau'') e^{-j\omega(\tau''+T)} d\tau''$$

$$= 2S_X(\omega) + S_X(\omega)(e^{j\omega T} + e^{-j\omega T})$$

$$= 2S_X(\omega)(1 + \cos \omega T)$$



### 第三章

#### 3.3

$$\begin{aligned}
 R_Y(t_1, t_2) &= E[Y(t_1)Y(t_2)] = E\left[a \frac{dX(t_1)}{dt_1} \cdot a \frac{dX(t_2)}{dt_2}\right] \\
 &= a^2 \frac{\partial^2 R_X(t_1, t_2)}{\partial t_1 \partial t_2} \quad (\tau = t_1 - t_2) \\
 &= a^2 \frac{\partial^2 R_X(\tau)}{\partial \tau^2} \cdot \frac{\partial \tau}{\partial t_1} \cdot \frac{\partial \tau}{\partial t_2} \\
 &= -a^2 \frac{\partial^2 R_X(\tau)}{\partial \tau^2} \\
 \frac{\partial R_X(\tau)}{\partial \tau} &= \sigma_X^2 \cdot (-2\alpha^2 \tau) e^{-\alpha^2 \tau^2} \\
 \frac{\partial^2 R_X(\tau)}{\partial \tau^2} &= \sigma_X^2 \cdot (-2\alpha^2 + 4\alpha^4 \tau^2) e^{-\alpha^2 \tau^2} = -2\sigma_X^2 \alpha^2 e^{-\alpha^2 \tau^2} (1 - 2\alpha^2 \tau^2) \\
 R_Y(t_1, t_2) &= 2(1 - 2\alpha^2 \tau^2) a^2 \sigma_X^2 \alpha^2 e^{-\alpha^2 \tau^2}
 \end{aligned}$$

#### 3.4

$$\begin{aligned}
 R_Y(\tau) &= E[Y(t)Y(t-\tau)] \\
 &= E\{[X(t) + \dot{X}(t)][X(t-\tau) + \dot{X}(t-\tau)]\} \\
 &= E\{X(t)X(t-\tau) + X(t)\dot{X}(t-\tau) + \dot{X}(t)X(t-\tau) + \dot{X}(t)\dot{X}(t-\tau)\} \\
 &= R_X(\tau) - R'_X(\tau) + R'_X(\tau) - R''_X(\tau) \\
 &= R_X(\tau) - R''_X(\tau) \\
 &= (3 - 4\tau^2)e^{-\tau^2}
 \end{aligned}$$

#### 3.5

一阶线性微分方程  $\frac{dy}{dx} + p(x)y = Q(x)$  的通解:

$$\begin{aligned} y &= e^{-\int p(x)dx} \left( \int Q(x)e^{\int p(x)dx} dx + C \right) \\ &= Ce^{-\int p(x)dx} + e^{-\int p(x)dx} \int Q(x)e^{\int p(x)dx} dx \end{aligned}$$

求  $E[Y(t)]$ :

对微分方程两边取期望

$$\begin{cases} m_Y'(t) + 2m_Y(t) = 2 & t > 0 \\ m_Y(0) = 1 \end{cases}$$

$$m_Y(t) = Ce^{-2t} + 2e^{-2t} \cdot \frac{e^{2t}}{2} = Ce^{-2t} + 1, \text{ 解得 } C = 0$$

所以  $m_Y(t) = 1$

求  $R_{XY}(t_1, t_2)$ :

改写微分方程为

$$\begin{cases} Y'(t_2) + 2Y(t_2) = X(t_2) \\ Y(0) = 1 \end{cases}$$

$$\begin{cases} E[X(t_1)Y'(t_2)] + 2E[X(t_1)Y(t_2)] = E[X(t_1)X(t_2)] \\ E[X(t_1)Y(0)] = E[X(t_1)] \end{cases}$$

可得

$$\begin{cases} \frac{\partial R_{XY}(t_1, t_2)}{\partial t_2} + 2R_{XY}(t_1, t_2) = R_X(t_1, t_2) = 4 + 2e^{-|t_1 - t_2|} \\ R_{XY}(t_1, 0) = 2 \end{cases}$$

分类讨论

$$\text{当 } t_1 > t_2 \text{ 时, 解得 } C = -\frac{2}{3}e^{-t_1}, \quad R_{XY}(t_1, t_2) = 2 + \frac{2}{3}e^{-t_1+t_2} - \frac{2}{3}e^{-t_1-2t_2}$$

$$\text{当 } t_1 < t_2 \text{ 时, 解得 } C = -2e^{t_1}, \quad R_{XY}(t_1, t_2) = 2 + 2(e^{t_1-t_2} + e^{t_1-2t_2})$$

求  $R_Y(t_1, t_2)$  :

将原方程改写为

$$\begin{cases} \frac{\partial R_Y(t_1, t_2)}{\partial t_1} + 2R_Y(t_1, t_2) = R_{XX}(t_1, t_2) \\ R_Y(0, t_2) = 1 \end{cases}$$

$$\text{当 } t_1 > t_2 \text{ 时, 解得 } R_Y(t_1, t_2) = 1 + \frac{2}{3}(e^{-t_1+t_2} - e^{-t_1-2t_2} + e^{-2t_1-2t_2} - e^{-2t_1+t_2})$$

$$\text{当 } t_1 < t_2 \text{ 时, 解得 } R_Y(t_1, t_2) = 1 + \frac{2}{3}(e^{t_1-t_2} - e^{t_1-2t_2} + e^{-2t_1-2t_2} - e^{-2t_1-t_2})$$

$$RC \frac{dy(t)}{dt} + y(t) = x(t)$$

对上式两边同时取傅里叶变换，利用微分性质可得

$$RC \cdot j\omega Y(j\omega) + Y(j\omega) = X(j\omega)$$

则传递函数为

$$H(j\omega) = \frac{Y(j\omega)}{X(j\omega)} = \frac{1}{1 + j\omega RC} = \frac{a}{a + j\omega} \quad (a = \frac{1}{RC})$$

输出过程的物理功率谱密度为

$$F_Y(\omega) = |H(j\omega)|^2 F_X(\omega) = \frac{a^2 N_0}{a^2 + \omega^2} \quad (0 < \omega < \infty)$$

输出过程的自相关函数为

$$R_Y(\tau) = \mathcal{F}^{-1}[\frac{F_Y(\omega)}{2}] = \frac{N_0}{4} a e^{-a|\tau|} \quad (a = \frac{1}{RC})$$

由于  $t_3 > t_2 > t_1$ ，则有

$$\begin{aligned} \frac{R_Y(t_3 - t_2) R_Y(t_2 - t_1)}{R_Y(0)} &= \frac{\frac{N_0 a}{4} e^{-(t_3 - t_2)} \cdot \frac{N_0 a}{4} e^{-(t_2 - t_1)}}{\frac{N_0 a}{4}} \\ &= \frac{N_0 a}{4} e^{-(t_3 - t_1)} = R_Y(t_3 - t_1) \end{aligned}$$

故

$$R_Y(t_3 - t_1) = \frac{R_Y(t_3 - t_2) R_Y(t_2 - t_1)}{R_Y(0)}$$

输出过程的自相关函数为

$$\begin{aligned} R_Y(t_1, t_2) &= E\left[\int_0^{t_1} X(u) du \int_0^{t_2} X(v) dv\right] \\ &= \int_0^{t_1} \int_0^{t_2} E[X(u)X(v)] du dv \end{aligned}$$

当  $t_1 < t_2$  时, 令  $\tau = u - v, d\tau = -dv$ , 则有

$$\begin{aligned} R_Y(t_1, t_2) &= \int_0^{t_1} \left[ \int_u^{t_2} \sigma^2 \delta(\tau) d(-\tau) \right] du \\ &= \int_0^{t_1} \left[ \int_{u-t_2}^u \sigma^2 \delta(\tau) d(\tau) \right] du \\ &= \int_0^{t_1} \sigma^2 du \\ &= \sigma^2 t_1 \end{aligned}$$

当  $t_1 > t_2$  时, 令  $\tau = u - v, d\tau = du$ , 则有

$$\begin{aligned} R_Y(t_1, t_2) &= \int_0^{t_1} \left[ \int_{-\tau}^{t_1-\tau} \sigma^2 \delta(\tau) d\tau \right] dv \\ &= \int_0^{t_1} \left[ \int_{-\tau}^{t_1-\tau} \sigma^2 \delta(\tau) d(\tau) \right] dv \\ &= \int_0^{t_2} \sigma^2 dv \\ &= \sigma^2 t_2 \end{aligned}$$

综上可得

$$R_Y(t_1, t_2) = \sigma^2 \min(t_1, t_2)$$

$$\psi^2 = E[Y^2(t)] = R_Y(t, t) = \sigma^2 t$$

$$R_W(t_1, t_2) = \delta(t_1 - t_2)$$

$$R_X(t_1, t_2) = \delta(t_1 - t_2) \cdot [U(t_1) - U(t_1 - T)] \cdot [U(t_2) - U(t_2 - T)]$$

$$R_X(t_1, t_2) = \begin{cases} \delta(t_1 - t_2) & 0 \leq t_1 \leq T, 0 \leq t_2 \leq T \\ 0 & \text{其他} \end{cases}$$

$$h(t) = e^{-\alpha t} U(t), \alpha > 0$$

输出  $Y(t)$  自相关函数为

$$\begin{aligned} R_Y(t_1, t_2) &= E[Y(t_1)Y(t_2)] = E\left[\int_{-\infty}^{\infty} X(u)h(t_1 - u)du \int_{-\infty}^{\infty} X(v)h(t_2 - v)dv\right] \\ &= \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} E[X(u)X(v)]h(t_1 - u)h(t_2 - v)dudv \\ &= e^{-\alpha(t_1 + t_2)} \int_0^{t_2} \int_0^{t_1} \delta(u - v)e^{\alpha(u+v)} dudv \end{aligned}$$

令  $t = \min(t_1, t_2)$ ，则有

(I) 当  $t < 0$  时，  $R_Y(t_1, t_2) = 0$

(II) 当  $0 \leq t \leq T$  时，

$$\text{假设 } t_1 > t_2, \quad \int_0^{t_2} \int_0^{t_1} \delta(u - v)e^{\alpha(u+v)} dudv = \int_0^{t_2} e^{2\alpha v} dv = \frac{e^{2\alpha t_2} - 1}{2\alpha}$$

$$\text{假设 } t_1 < t_2, \quad \int_0^{t_2} \int_0^{t_1} \delta(u - v)e^{\alpha(u+v)} dudv = \int_0^{t_1} e^{2\alpha u} du = \frac{e^{2\alpha t_1} - 1}{2\alpha}$$

$$\text{整理可得 } R_Y(t_1, t_2) = \frac{1}{2\alpha} (e^{2\alpha t} - 1) e^{-\alpha(t_1 + t_2)}$$

$$(III) \text{ 当 } t \geq T \text{ 时, } R_Y(t_1, t_2) = \frac{1}{2\alpha} (e^{2\alpha T} - 1) e^{-\alpha(t_1+t_2)}$$

故

$$R_Y(t_1, t_2) = \begin{cases} 0 & t \leq 0 \\ \frac{1}{2\alpha} (e^{2\alpha t} - 1) e^{-\alpha(t_1+t_2)} & 0 < t \leq T \\ \frac{1}{2\alpha} (e^{2\alpha T} - 1) e^{-\alpha(t_1+t_2)} & t > T \end{cases}$$

式中  $t = \min(t_1, t_2)$

(1) 由系统框图可得

$$Y(t) = X(t) - X(t-T) = X(t) * [\delta(t) - \delta(t-T)]$$

$$Z(t) = \int_{-\infty}^{\infty} Y(\lambda) u(t-\lambda) d\lambda = Y(t) * u(t)$$

故有

$$Z(t) = X(t) * [\delta(t) - \delta(t-T)] * u(t) = X(t) * [u(t) - u(t-T)]$$

$$h(t) = u(t) - u(t-T)$$

$$H(j\omega) = T \text{Sa}\left(\frac{\omega T}{2}\right) e^{-j\frac{\omega T}{2}}$$

$$(2) \quad S_Z(\omega) = |H(j\omega)|^2 S_X(\omega) = S_0 T^2 \text{Sa}^2\left(\frac{\omega T}{2}\right)$$

$$\psi_Z^2 = R_Z(0) = \frac{1}{2\pi} \int_{-\infty}^{\infty} S_0 T^2 \text{Sa}^2\left(\frac{\omega T}{2}\right) d\omega$$

$$= \frac{2}{\pi} S_0 T \int_0^{\infty} \frac{\sin^2 x}{x^2} dx \quad (x = \frac{\omega T}{2})$$

$$= S_0 T$$

3.28



(1) 求  $m_Y(t)$  :

$$m_X(t) = E[X(t)] = E[W(t) * h_1(t)] = E[W(t)] * h_1(t) = 0$$

$$m_Y(t) = E[Y(t)] = E[X(t) - X(t-T)] = 0$$

(2) 求  $\sigma_Y^2$

对于系统 1, 由  $h_1(t) = e^{-\alpha t} U(t), \alpha > 0$  可得  $H_1(j\omega) = \frac{1}{\alpha + j\omega}$

对于系统 2, 由  $Y(t) = X(t) - X(t-T) = X(t) * [\delta(t) - \delta(t-T)]$

$$h_2(t) = \delta(t) - \delta(t-T) \text{ 可得 } H_2(j\omega) = 1 - e^{-j\omega T}$$

故

$$\begin{aligned} S_Y(\omega) &= |H_1(j\omega)|^2 |H_2(j\omega)|^2 S_X(\omega) \\ &= \frac{N_0}{2} \cdot \frac{2 - 2\cos \omega T}{\alpha^2 + \omega^2} \end{aligned}$$

$$\sigma_Y^2 = R_Y(0) = \frac{1}{2\pi} \int_{-\infty}^{\infty} S_Y(\omega) d\omega = \frac{N_0}{2\alpha} (1 - e^{-\alpha T})$$

故

$$P_Y(y) = \frac{1}{\sigma_Y \sqrt{2\pi}} \exp\left\{-\frac{y^2}{2\sigma_Y^2}\right\}$$

式中  $\sigma_Y^2 = \frac{N_0}{2\alpha} (1 - e^{-\alpha T}), \sigma_Y > 0$

## 第四章

### 4.13

$$\begin{aligned}
 R_{xc}(\tau) &= E \left\{ \left[ \hat{X}(t) \cos \omega_0 t - X(t) \sin \omega_0 t \right] \times \left[ X(t-\tau) \cos \omega_0(t-\tau) + \hat{X}(t-\tau) \sin \omega_0(t-\tau) \right] \right\} \\
 &= E \left[ \hat{X}(t) X(t-\tau) \right] \cos \omega_0 t \cos \omega_0(t-\tau) \\
 &\quad + E \left[ \hat{X}(t) \hat{X}(t-\tau) \right] \cos \omega_0 t \sin \omega_0(t-\tau) \\
 &\quad - E \left[ X(t) X(t-\tau) \right] \sin \omega_0 t \cos \omega_0(t-\tau) \\
 &\quad - E \left[ X(t) \hat{X}(t-\tau) \right] \sin \omega_0 t \sin \omega_0(t-\tau) \\
 &= R_{\hat{X}X}(\tau) [\cos \omega_0 t \cos \omega_0(t-\tau) - \sin \omega_0 t \sin \omega_0(t-\tau)] - R_X(\tau) [\sin \omega_0 t \cos \omega_0(t-\tau) - \cos \omega_0 t \sin \omega_0(t-\tau)] \\
 &= R_X(\tau) \cos \omega_0 \tau - R_X(\tau) \sin \omega_0 \tau
 \end{aligned}$$

### 4.19

(1)

$n(t)$  和  $\hat{n}(t)$  :

$$n(t) = n_c(t) \cos \omega_0 t - n_s(t) \sin \omega_0 t$$

$$\hat{n}(t) = n_c(t) \sin \omega_0 t + n_s(t) \cos \omega_0 t$$

两正交分量  $n_c(t)$  和  $n_s(t)$  :

$$n_c(t) = n(t) \cos \omega_0 t + \hat{n}(t) \sin \omega_0 t$$

$$n_s(t) = \hat{n}(t) \cos \omega_0 t - n(t) \sin \omega_0 t$$

两正交分量的自相关函数  $R_{n_c}(\tau)$  和  $R_{n_s}(\tau)$  :

$$R_{n_c}(\tau) = R_{n_s}(\tau) = R_n(\tau) \cos \omega_0 \tau + \hat{R}_n(\tau) \sin \omega_0 \tau$$

两正交分量的功率谱密度  $G_{n_c}(\omega)$  和  $G_{n_s}(\omega)$  :

$$G_{n_c}(\omega) = G_{n_z}(\omega) = \begin{cases} G_n(\omega - \omega_0) + G_n(\omega + \omega_0) & |\omega| < \omega_c \\ 0 & \text{其他} \end{cases}$$

(2)

两正交分量的互相关函数  $R_{cs}(\tau)$  和  $R_{sc}(\tau)$  :

$$R_{cs}(\tau) = R_n(\tau) \sin \omega_0 \tau - \hat{R}_n(\tau) \cos \omega_0 \tau$$

$$R_{sc}(\tau) = \hat{R}_n(\tau) \cos \omega_0 \tau - R_n(\tau) \sin \omega_0 \tau$$

$$R_{cs}(\tau) = -R_{sc}(\tau)$$

两正交分量的互功率谱密度  $G_{cs}(\omega)$  和  $G_{sc}(\omega)$  :

$$G_{cs}(\omega) = -G_{sc}(\omega) = \begin{cases} -j[G_n(\omega - \omega_0) - G_n(\omega + \omega_0)] & |\omega| < \omega_c \\ 0 & \text{其他} \end{cases}$$

## 4.22

(1)

$$R_{Y_1 Y_2}(\tau) = R_X(\tau) * h_1(\tau) * h_2(-\tau)$$

$$S_{Y_1 Y_2}(\omega) = S_X(\omega) H_1(j\omega) H_2^*(j\omega) = \frac{N_0}{2} H_1(j\omega) H_2^*(j\omega)$$

$R_{Y_1 Y_2}(\tau)$  为偶函数等价于  $S_{Y_1 Y_2}(\omega)$  为偶函数

所以当  $H_1(j\omega) H_2^*(j\omega)$  为实对称函数时, 互相关函数  $R_{Y_1 Y_2}(\tau)$  为偶函数

(2)

$Y_1(t)$ ,  $Y_2(t)$  统计独立等价于  $Y_1(t)$ ,  $Y_2(t)$  不相关, 因此有  $R_{Y_1 Y_2}(\tau) = 0$

因此  $h_1(t)$  和  $h_2(t)$  应满足  $h_1(t) * h_2(-t) = 0$

在频域里  $H_1(j\omega) H_2^*(j\omega) = 0$

即在频域里要求两个系统的通带不混叠

## 第五章

### 5.5

由于  $U$  和  $V$  是统计独立的高斯随机变量，容易知道  $X(t)$  是高斯随机过程，因此欲求其一维和二维概率密度函数，只需求其前两阶矩即可。

先求一维概率密度函数。因为

$$m_X(t) = E\{X(t)\} = E(U)\cos\omega t + E(V)\sin\omega t = 0$$

和

$$\sigma_X^2(t) = E\{X^2(t)\} = E(U^2)\cos^2\omega t + E(V^2)\sin^2\omega t + 2E(UV)\sin\omega t\cos\omega t = \sigma^2$$

故有

$$p_X(x) = \frac{1}{\sqrt{2\pi}\sigma} \exp\left\{-\frac{x^2}{2\sigma^2}\right\}$$

现在求二维概率密度函数。因为

$$\mathbf{a} = \begin{bmatrix} m_X(t_1) \\ m_X(t_2) \end{bmatrix} = \mathbf{0}$$

$$\mathbf{C} = \begin{bmatrix} E\{X^2(t_1)\} & E\{X(t_1)X(t_2)\} \\ E\{X(t_2)X(t_1)\} & E\{X^2(t_2)\} \end{bmatrix}$$

$$p_X(x_1, x_2; \tau) = \frac{1}{2\pi|\mathbf{C}|^{1/2}} \exp\left\{-\frac{1}{2}[x_1, x_2]\mathbf{C}\begin{bmatrix} x_1 \\ x_2 \end{bmatrix}\right\}$$

### 5.6

对于任意时刻  $t$ ,  $Y(t)$  为一个高斯随机变量。这样,  $Y(t)$  的均值和方差分别为

$$\begin{aligned} m_Y(t) &= E\{Y(t)\} = \frac{1}{\varepsilon} [E\{X(t+\varepsilon)\} - E\{X(t)\}] = 0 \\ \sigma_Y^2(t) &= E\{Y^2(t)\} \\ &= \frac{1}{\varepsilon^2} [E\{X^2(t+\varepsilon)\} + E\{X^2(t)\} - 2E\{X(t+\varepsilon)X(t)\}] \\ &= \frac{2}{\varepsilon^2} [R_X(0) - R_X(\varepsilon)] \end{aligned}$$

故有

$$p_Y(y) = \frac{1}{\sqrt{4\pi[R_X(0) - R_X(\varepsilon)]/\varepsilon^2}} \exp\left\{-\frac{\varepsilon^2 y^2}{4[R_X(0) - R_X(\varepsilon)]}\right\}$$

注意到  $R_X(\tau)$  为偶函数, 故  $R'_X(0) = 0$ 。故有

$$\begin{aligned} \lim_{\varepsilon \rightarrow 0} \frac{2[R_X(0) - R_X(\varepsilon)]}{\varepsilon^2} &= \lim_{\varepsilon \rightarrow 0} \frac{1}{\varepsilon} \cdot \frac{1}{\varepsilon} \{[R_X(0) - R_X(\varepsilon)] + [R_X(0) - R_X(\varepsilon)]\} \\ &= \lim_{\varepsilon \rightarrow 0} \frac{-[R'_X(0) + R'_X(\varepsilon)]}{\varepsilon} \\ &= \lim_{\varepsilon \rightarrow 0} \frac{-R'_X(\varepsilon)}{\varepsilon} \\ &= -R''_X(0) \end{aligned}$$

于是

$$p_Y(y) \xrightarrow{\varepsilon \rightarrow 0} \frac{1}{\sqrt{2\pi[-R''_X(0)]}} \exp\left\{\frac{y^2}{2R''_X(0)}\right\}$$

问题:

$$\begin{aligned} \lim_{\varepsilon \rightarrow 0} \frac{2[R_X(0) - R_X(\varepsilon)]}{\varepsilon^2} &= \lim_{\varepsilon \rightarrow 0} \frac{1}{\varepsilon} \cdot \frac{1}{\frac{\varepsilon}{2}} \left\{ \left[ R_X(0) - R_X\left(\frac{\varepsilon}{2}\right) \right] + \left[ R_X\left(\frac{\varepsilon}{2}\right) - R_X(\varepsilon) \right] \right\} \\ &= \lim_{\varepsilon \rightarrow 0} \frac{-\left[ R'_X(0) + R'_X\left(\frac{\varepsilon}{2}\right) \right]}{\varepsilon} \\ &= \lim_{\varepsilon \rightarrow 0} \frac{-R'_X\left(\frac{\varepsilon}{2}\right)}{\varepsilon} \\ &= -\frac{1}{2} R''_X(0) \end{aligned}$$

$$E(Y) = E\left\{\int_0^1 X(t)dt\right\} = \int_0^1 E\{X(t)\}dt = 0$$

$$\begin{aligned} E(Y^2) &= E\left\{\int_0^1 X(t)dt \int_0^1 X(\tau)d\tau\right\} \\ &= \int_0^1 \int_0^1 E\{X(t)X(\tau)\}dtd\tau \\ &= \int_0^1 \int_0^1 e^{-|t-\tau|}dtd\tau \\ &= \int_0^1 \left[\int_0^t e^{-(t-\tau)}d\tau + \int_t^1 e^{t-\tau}d\tau\right]dt \\ &= 2e^{-1} \end{aligned}$$

于是可得

$$p_Y(y) = \frac{1}{\sqrt{4\pi e^{-1}}} \exp\left\{-\frac{y^2}{4e^{-1}}\right\}$$

## 5.8

(1)

$Z$  是高斯随机变量, 可得  $E(Z) = 0$ ,  $E(Z^2) = \sigma^2$

$$p_Z(z) = \frac{1}{\sqrt{2\pi}\sigma} \exp\left\{-\frac{z^2}{2\sigma^2}\right\}$$

(2)

先计算随机变量  $R$  的分布函数

$$\begin{aligned} F_R(r) &= P(R \leq r) = P\left(\sqrt{X^2 + Y^2} \leq r\right) \\ &= \iint_{\sqrt{x^2 + y^2} \leq r} p(x, y)dx dy \\ &= \int_0^{2\pi} \int_0^r \frac{1}{2\pi\sigma^2} \exp\left\{-\frac{r^2}{2\sigma^2}\right\} dr d\theta \\ &= \int_0^r \frac{r}{\sigma^2} \exp\left\{-\frac{r^2}{2\sigma^2}\right\} dr \end{aligned}$$

相应的概率密度函数为

$$p_R(r) = F'_R(r) = \begin{cases} \frac{r}{\sigma^2} \exp\left\{-\frac{r^2}{2\sigma^2}\right\} & r \geq 0 \\ 0 & \text{其他} \end{cases}$$

(3)

根据题意, 该样本是

$$Z(0) = X, \quad Z\left(\frac{1}{4}\right) = Y, \quad Z\left(\frac{1}{2}\right) = -X$$

故  $Z(0)$  和  $Z\left(\frac{1}{4}\right)$  统计独立, 而  $Z\left(\frac{1}{2}\right)$  是  $Z(0)$  的线性函数, 由  $p_z(z) = \frac{1}{\sqrt{2\pi}\sigma} \exp\left\{-\frac{z^2}{2\sigma^2}\right\}$

可得

$$p_{z_1 z_2}(z_1, z_2) = \frac{1}{2\pi\sigma^2} \exp\left\{-\frac{z_1^2 + z_2^2}{2\sigma^2}\right\}$$

$$p(z_3 | z_1 z_2) = p(z_3 | z_1) = \delta(z_3 + z_1)$$

故

$$p(z_1, z_2, z_3) = \frac{1}{2\pi\sigma^2} \exp\left\{-\frac{z_1^2 + z_2^2}{2\sigma^2}\right\} \delta(z_1 + z_3)$$

## 5.13

设线性变换为

$$\begin{aligned} Y_1 &= L_{11}X_1 + L_{12}X_2 + \cdots + L_{1n}X_n \\ Y_2 &= L_{21}X_1 + L_{22}X_2 + \cdots + L_{2n}X_n \\ &\vdots \\ Y_n &= L_{n1}X_1 + L_{n2}X_2 + \cdots + L_{nn}X_n \end{aligned}$$

线性变换矩阵形式为

$$\mathbf{Y} = \mathbf{LX}$$

式中

$$\mathbf{L} = \begin{bmatrix} L_{11} & L_{12} & \cdots & L_{1n} \\ L_{21} & L_{22} & \cdots & L_{2n} \\ \vdots & \vdots & \ddots & \vdots \\ L_{n1} & L_{n2} & \cdots & L_{nn} \end{bmatrix}$$

$$\mathbf{Y} = [Y_1, Y_2, \cdots, Y_n]^T, \quad \mathbf{X} = [X_1, X_2, \cdots, X_n]^T$$

$$f_x(x) = \frac{1}{(2\pi)^{\frac{n}{2}} |\mathbf{C}|^{\frac{1}{2}}} \exp\left\{-\frac{1}{2}(\mathbf{x}-\mathbf{a})^T \mathbf{C}^{-1}(\mathbf{x}-\mathbf{a})\right\}$$

为了方便推导, 假定其均值矢量为零  $\mathbf{a} = \mathbf{0}$

令  $\mathbf{\Gamma} = \mathbf{L}^{-1}$ , 则有

$$\mathbf{X} = \mathbf{\Gamma Y}$$

$$p_{\mathbf{Y}}(\mathbf{y}) = p(\mathbf{I}\mathbf{y})|\mathbf{J}|$$

$$|\mathbf{J}| = \left| \frac{\partial \mathbf{I}\mathbf{Y}}{\partial \mathbf{Y}} \right| = |\mathbf{I}| = 1/|\mathbf{L}|$$

于是有

$$\begin{aligned} f_{\mathbf{Y}}(\mathbf{y}) &= \frac{1}{(2\pi)^{\frac{n}{2}} \left( |\mathbf{L}|^2 |\mathbf{C}| \right)^{\frac{1}{2}}} \exp \left\{ -\frac{1}{2} (\mathbf{I}\mathbf{y})^{\top} \mathbf{C}^{-1} (\mathbf{I}\mathbf{y}) \right\} \\ &= \frac{1}{(2\pi)^{\frac{n}{2}} \left( |\mathbf{L}|^2 |\mathbf{C}| \right)^{\frac{1}{2}}} \exp \left\{ -\frac{1}{2} \mathbf{y}^{\top} (\mathbf{I}^{\top} \mathbf{C}^{-1} \mathbf{I}) \mathbf{y} \right\} \end{aligned}$$

$$\text{令 } \mathbf{F}^{-1} = \mathbf{I}^{\top} \mathbf{C}^{-1} \mathbf{I} = (\mathbf{L}^{\top})^{-1} \mathbf{C}^{-1} \mathbf{L}^{-1} = (\mathbf{L} \mathbf{C} \mathbf{L}^{\top})^{-1}, \quad |\mathbf{F}| = |\mathbf{L}| |\mathbf{C}| |\mathbf{L}^{\top}| = |\mathbf{L}|^2 |\mathbf{C}|$$

所以有

$$f_{\mathbf{Y}}(\mathbf{y}) = \frac{1}{(2\pi)^{\frac{n}{2}} |\mathbf{F}|^{\frac{1}{2}}} \exp \left\{ -\frac{1}{2} (\mathbf{y} - \mathbf{b})^{\top} \mathbf{F}^{-1} (\mathbf{y} - \mathbf{b}) \right\}$$

5.15

$$\begin{aligned} E[A(t)] &= \int_{-\infty}^{\infty} A p(A) dA = \int_0^{\infty} A \frac{A}{\sigma_x^2} e^{-\frac{A^2}{2\sigma_x^2}} dA = \int_0^{\infty} -A d e^{-\frac{A^2}{2\sigma_x^2}}, (A > 0) \\ &= -A e^{-\frac{A^2}{2\sigma_x^2}} \Big|_0^{\infty} + \int_0^{\infty} e^{-\frac{A^2}{2\sigma_x^2}} dA = 0 + \sqrt{2\pi} \int_0^{\infty} \frac{1}{\sqrt{2\pi}\sigma_x} e^{-\frac{A^2}{2\sigma_x^2}} dA = \sqrt{2\pi}\sigma_x \times \frac{1}{2} = \sqrt{\frac{\pi}{2}}\sigma_x \end{aligned}$$

$$\begin{aligned} E[A^2(t)] &= \int_0^{\infty} A^2 \frac{A}{\sigma_x^2} e^{-\frac{A^2}{2\sigma_x^2}} dA = 2\sigma_x^2 \int_0^{\infty} \frac{A^2}{2\sigma_x^2} e^{-\frac{A^2}{2\sigma_x^2}} d \frac{A^2}{2\sigma_x^2} = 2\sigma_x^2 \int_0^{\infty} x d e^{-x}, (x > 0) \\ &= 2\sigma_x^2 \end{aligned}$$

$$D[A(t)] = E[A^2(t)] - E^2[A(t)] = \left( 2 - \frac{\pi}{2} \right) \sigma_x^2$$

5.16



$$R_T(\tau) = E[Y(t)Y(t-\tau)] = E[X^2(t)X^2(t-\tau)]$$

$$E[X_1^n X_2^k] = (-j)^{n+k} \frac{\partial^{n+k} \phi_X(v_1, v_2)}{\partial v_1^n \partial v_2^k} \bigg|_{v_1=v_2=0}$$

随机变量的各阶矩可以通过对特征函数求导得到

$$E[X^2(t)X^2(t-\tau)] = (-j)^4 \frac{\partial^4 \phi_X(v_1, v_2; t, t-\tau)}{\partial v_1^2 \partial v_2^2} \bigg|_{v_1=v_2=0}$$

已知  $X(t)$  为均值为 0 的高斯平稳随机过程，特征函数为

$$\phi_X(v_1, v_2; \tau) = \exp \left\{ j \mathbf{a}^T \mathbf{v} - \frac{1}{2} \mathbf{v}^T \mathbf{C} \mathbf{v} \right\}$$

式中

$$\mathbf{a} = \begin{bmatrix} a_1 \\ a_2 \end{bmatrix} = \begin{bmatrix} 0 \\ 0 \end{bmatrix}, \quad \mathbf{v} = \begin{bmatrix} v_1 \\ v_2 \end{bmatrix}, \quad \mathbf{C} = \begin{bmatrix} R_X(0) & R_X(\tau) \\ R_X(\tau) & R_X(0) \end{bmatrix}$$

所以

$$\phi_X(v_1, v_2; \tau) = \exp \left\{ -\frac{1}{2} [R_X(0)v_1^2 + 2R_X(\tau)v_1v_2 + R_X(0)v_2^2] \right\}$$

故

$$\begin{aligned} R_T(\tau) &= (-j)^4 \frac{\partial^4 \exp \left\{ -\frac{1}{2} [R_X(0)v_1^2 + 2R_X(\tau)v_1v_2 + R_X(0)v_2^2] \right\}}{\partial v_1^2 \partial v_2^2} \bigg|_{v_1=v_2=0} \\ &= [R_X(0)]^2 + 2[R_X(\tau)]^2 \end{aligned}$$