Statistics 305: Introduction to Biostatistical Methods for Health Sciences

Chapter 14: Inference for Proportions

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Inference for Proportions (Chapter 14)

- ▶ Instead of quantitative measurements, we classify each sampled individual into one of two categories.
 - success, failure (canonical)
 - breast cancer, no breast cancer
 - Etc.
- Mant to make inference about the proportion p of successes in a population, or about the difference between the proportions p_1 and p_2 of successes in two populations.

Example: Women's Health Initiative (WHI)

- A randomized controlled trial, called the Women's Health Initiative, randomized 16,608 post-menopausal women aged 50-79 years to receive either hormone replacement therapy in the form of estrogen plus progestin (EP; $n_1 = 8506$), or a placebo ($n_2 = 8102$).
- ▶ After five years, 166 of those in the EP group had developed invasive breast cancer, compared to 122 in the placebo group.
- ► The populations to compare are postmenopausal women aged 50-79 years at the start of the trial, who are taking EP (population 1) or placebo (population 2).
- ► The estimated populaton proportions are $\hat{p}_1 = 166/8506 = 0.0195$ and $\hat{p}_2 = 122/8102 = 0.0151$
- ▶ It looks like the EP group has a higher risk of breast cancer, but could this difference be due to chance?

Outline of Approach

- Similar approach to inference as inference of population means (quantitative data), with some minor differences.
- ▶ Inference is based on the sampling distribution of $\hat{p}_1 \hat{p}_2$
- ► Two-stage reasoning:
 - ▶ 1. Transform $\hat{p}_1 \hat{p}_2$ into an initial pivotal quantity, Z_1 , whose denominator depends on the unknown p_i 's
 - Get a final pivotal quantity, Z, by replacing the unknown p_i's in the denominator of Z₁ with estimates.
- Confidence intervals and hypothesis tests follow from the approximate sampling distribution of the final pivotal quantity, Z.
- ▶ Note: We will not cover *Inference for a single proportion* in the text (Sections 14.2 14.5).

Sampling Distribution of $\hat{p}_1 - \hat{p}_2$

- Assume that we have independent simple random samples (SRSs) of size n_1 and n_2 , from the two parent populations.
- ▶ Then the distribution of $\hat{p}_1 \hat{p}_2$ has
 - ▶ mean $p_1 p_2$ and
 - ► SD

$$\sqrt{p_1(1-p_1)/n_1+p_2(1-p_2)/n_2}$$
.

- ▶ If the sample sizes are "large", the CLT tells us that the shape of this distribution is approximately normal.
 - ► (The CLT applies because it turns out that a proportion is an average . . . of 0's and 1's.)

Initial Transformation

- For large samples, $\hat{p}_1 \hat{p}_2$ is **approximately** normally distributed with
 - ▶ mean $p_1 p_2$, and
 - ► SD

$$\sqrt{p_1(1-p_1)/n_1+p_2(1-p_2)/n_2}$$

So the distribution of

$$Z_1 = rac{(\hat{p}_1 - \hat{p}_2) - (p_1 - p_2)}{\sqrt{p_1(1 - p_1)/n_1 + p_2(1 - p_2)/n_2}}$$

is approximately N(0,1).

Z_1 with Estimated SDs

▶ Inserting the estimates \hat{p}_1 and \hat{p}_2 for the unknown parameters p_1 and p_2 into Z_1 above gives

$$Z = \frac{(\hat{p}_1 - \hat{p}_2) - (p_1 - p_2)}{\sqrt{\hat{p}_1(1 - \hat{p}_1)/n_1 + \hat{p}_2(1 - \hat{p}_2)/n_2}}$$

- ▶ What is the distribution of *Z*?
 - ▶ Under certain conditions (see below) it is reasonable to assume that *Z* is approximately normal.

Rule of Thumb for Normal Approximation

- ► The normal approximation to the distribution of Z is considered reliable when the sample sizes n₁ and n₂ are "large"
- ▶ The definition of large depends on the underlying p_j 's. The text (page 324) suggests:
 - ▶ $n_1p_1 \ge 5$ and $n_1(1-p_1) \ge 5$ and
 - ▶ $n_2p_2 \ge 5$ and $n_2(1-p_2) \ge 5$
- ▶ The parameters p_1 and p_2 are not known so we insert the estimates $\hat{p}_1 = (\text{number of successes in sample } 1)/n_1$ and $\hat{p}_2 = (\text{number of successes in sample } 2)/n_2$.
- ► After inserting estimates, one can simplify the requirements to the following rule:
 - ► The normal approximation is reliable when there are at least 5 successes and 5 failures in both sample 1 and sample 2.

Checking rule of thumb for WHI Data

- At least 5 successes and failures in both samples.
 - ► True: 166 cancer, 8340 cancer-free in the EP group; 122 cancer, 7980 cancer-free in the placebo group.

Confidence Intervals

▶ The level-*C* CI for $p_1 - p_2$ is of the form:

estimate \pm margin of error

- ▶ The estimate is $\hat{p}_1 \hat{p}_2$
- ▶ The margin of error is $z^* \times SE$ where
 - z^* is the upper (1-C)/2 critical value of the standard normal distribution.
 - ▶ SE is the estimated SD of $\hat{p}_1 \hat{p}_2$ in the denominator of Z; namely, $SE = \sqrt{\hat{p}_1(1-\hat{p}_1)/n_1 + \hat{p}_2(1-\hat{p}_2)/n_2}$

Example (WHI)

- ▶ Recall: 16,608 women aged 50-79 years randomized to receive either estrogen plus progestin (EP; $n_1 = 8506$), or a placebo ($n_2 = 8102$). After five years, 166 in the EP group developed invasive breast cancer, compared to 122 in placebo group.
- ► For EP, $\hat{p}_1 = 166/8506$ and, for placebo, $\hat{p}_2 = 122/8102$.
- ▶ 95% CI is estimate \pm margin of error, where
 - estimate of $p_1 p_2$ is $\hat{p}_1 \hat{p}_2 = 0.0044$
 - margin of error is a critical value times standard error of difference.
- ▶ The critical value is 1.96 (see R demo).
- ► The standard error is $\sqrt{\hat{p}_1(1-\hat{p}_1)/n_1 + \hat{p}_2(1-\hat{p}_2)/n_2} = 0.002$
- ▶ The margin of error is therefore 1.96 * .002 = 0.00392.
- ▶ Putting it all together, the CI is 0.0044 ± 0.00392 or approximately (0.0005, 0.008).

Test Statistic

- ▶ The null hypothesis is $H_0: p_1 p_2 = 0$.
- Numerator of the test statistic is therefore the estimated difference $(\hat{p}_1 \hat{p}_2)$ minus 0.
- Denominator of the test statistic is

$$SE = \sqrt{\hat{p}_1(1-\hat{p}_1)/n_1 + \hat{p}_2(1-\hat{p}_2)/n_2}$$

- ▶ BUT, under H_0 , we have $p_1 = p_2$. Call this common value p.
- Assuming a common proportion p in the two populations, we pool the 2 samples to obtain an estimate \hat{p} ; i.e.,

$$\hat{p} = (\text{number of cancers in both samples})/(n_1 + n_2).$$

- ► The formula for the SE of $\hat{p}_1 \hat{p}_2$ simplifies to $\sqrt{\hat{p}(1-\hat{p}) \times (1/n_1 + 1/n_2)}$
- ▶ So the statistic for testing $H_0: p_1 p_2 = 0$ is

$$Z = \frac{(\hat{p}_1 - \hat{p}_2)}{\sqrt{\hat{p}(1-\hat{p})(1/n_1 + 1/n_2)}}$$

Example (WHI, continued)

- \vdash $H_0: p_1-p_2=0.$
- ► Test statistic numerator is $\hat{p}_1 \hat{p}_2 = 0.0044$.
- ► Test statistic denominator is the SE based on the pooled estimate of $p = p_1 = p_2$.
 - Pooled estimate of the common population proportion is $\hat{p} = (166 + 122)/(8506 + 8102) = 0.0173$.
 - ▶ So the SE is

$$\sqrt{\hat{p}(1-\hat{p})(1/n_1+1/n_2)}
= \sqrt{0.0173(1-0.0173)(1/8506+1/8102)}
= 0.002.$$

▶ The test statistic value is z = 0.0044/0.002 = 2.2.

p-value

- ▶ The *p*-value is the chance of a value of the test statistic that is as or more extreme than what we did observe in our data, when the null hypothesis is true.
- ► Same logic as we saw earlier for inference of population means (Chapter 11).
- ▶ Let Z be a standard normal random variable and z be the observed value of the test statistic. Then:
 - ▶ For H_a : $p_1 p_2 \neq 0$, p-value = $2P(Z \geq |z|)$.
 - ▶ For $H_a: p_1 p_2 > 0$, p-value = $P(Z \ge z)$.
 - ▶ For H_a : $p_1 p_2 < 0$, p-value = $P(Z \le z)$.

Example (WHI, continued)

- Suppose we wish to test H_0 : $p_1 p_2 = 0$ vs. H_a : $p_1 p_2 \neq 0$ at level $\alpha = 0.05$.
- For an observed value of the test statistic z = 2.2, computer software calculates a pvalue of about 0.03 (see R demo)
- ▶ We therefore reject H_0 at the 5% level: There is statistical evidence that women taking EP have a higher risk of invasive breast cancer than those taking the placebo.
 - Note: We say that the EP group has higher risk than the placebo group because \hat{p}_1 (for EP) is greater than \hat{p}_2 (for placebo), as evidenced by z = 2.2 being greater than zero.

Summary

- ▶ Inference for the difference $p_1 p_2$ between two population proportions is based on a pivotal quantity.
- Confidence intervals are of the form estimate ± margin of error, where
 - estimate is difference between sample means, and
 - margin of error is a critical value times the standard error of the difference in sample proportions
- ▶ To test the null hypothesis $H_0: p_1 p_2 = 0$ against an alternative H_a we calculate a test statistic and the p-value
 - ► The p-value is the chance of seeing a value of the test statistic as or more extreme than the value that was observed in our data, under the null hypothesis.
 - ▶ Compare the p-value to a significance level α to obtain a statistical hypothesis test
- ► The statistical inference is considered reliable when there are at least 5 successes and 5 failures in each sample.