# Lecture 10

# Logistic regression: categorical variables

逻辑回归:分类变量

# Logistic regression

### 逻辑回归实现了从基本(基于最小二乘)一般线性模型到中级/高级一般线性模型的过渡

• Logistic regression makes the transition from the basic (least-square-based) *general linear model* to the intermediate/advanced *generalised linear model* 

### 广义线性模型将线性回归模型扩展到非正态分布变量和非线性关系

• The generalised linear model extends linear regression models to variables not normally distributed, and to non-linear relationships

### 例如,我们可能希望使用回归技术来预测二元响应

- For example, we may want to use regression techniques to predict *binary* responses:
  - we may want to predict the probability that someone is dead or alive, voted or did not vote in the last election etc. as a function of other variables (age, smoking, income etc.)

### 换句话说,我们仍然希望使用回归,二元概率的结果

• In other words, we still want to use a regression:

Probability of binary outcome =  $a + b_1X_1 + b_2X_2...+b_nX_n = a + \Sigma b_iX_i$ 

#### with

a = intercept 截距

b<sub>i</sub> = regression coefficients 回归系数

 $X_i$  = independent variables (continuous or categorical) 自变量(连续变量或分类变量)

# Applications

逻辑回归是一种几乎普遍使用的分类方法,他以概率p预测结果是否发生(二元结果yes/no)

- Logistic regression is a classification method used almost universally
  - it predicts whether an outcome happens or not (binary outcome yes/no) with a probability *p*

### 他经常用于预测二元结果

- It is frequently applied to predict binary outcomes (yes or no)
  - business: costumer choice (purchasing, being late for bills etc.)
  - medicine, pubic health (will develop a condition etc)
  - insurance (risk of event, credit decisions)
  - etc.

### 逻辑回归与神经网络和机器学习密切相关

• Logistic regression is closely linked to neural networks and machine learning

# Odds and log(odds)

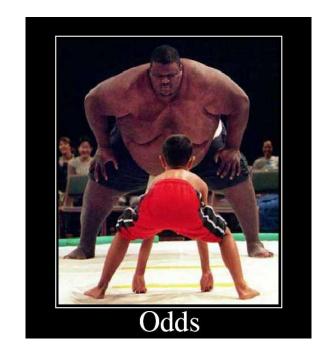
### 要理解逻辑回归,首先我们需要理解概率和概率比

• To understand logistic regressions, first we need to understand *odds* and *odds* ratios

#### 概率odd和事件的概率并不相同

• Important: odds are not the same as the *probability* of the event!

odds of event =  $\frac{probability of event occurring}{probability of event not occurring}$ 



# Odds and log(odds)

• Example: what is the *probability* of your birthday falling on a weekday this year?

- odds of weekday = (5/7) / (2/7) = 5/2 = 2.5 = p/(1-p)
- $\ln(\text{odds of weekday}) = \log(2.5) = 0.91$   $= \log(p/(1-p))$

#### 非事件的概率

• And the probability of the non-event, i.e. weekend day?

```
probability of weekend day = 2/7=0.29
odds of weekend day = 2/5 =0.4
ln(odds of weekend day)= -0.91
=1-p = (1-p)/p = ln((1-p)/p)
```

### Exercises

### Calculate:

- Tossing a fair coin:
  - Probability of heads?
  - Odds of heads?
  - Odds of tails?
  - log(odds of heads)



• Probability of 1? 1/6 1/5 5

log(0.2)

- Odds of 1?
- Odds of *not 1*?
- log(odds of 1)?





### Odds ratio

- Now imagine you must choose between betting on coins (bet on 'heads') or dice (bet on '1'); what are the odds of winning in each?
  - odds of heads = 1/1 = 1
  - odds of a 1 = 1/5 = 0.2
- So it is easier to win a coin toss; but how much easier?
- We can calculate the odds ratio of success in coins vs. dice

• Odds ratio = 
$$\frac{odds \ of \ heads}{odds \ of \ 1} = \frac{1}{0.2} = 5$$

• This means you are 5 times more likely to win by tossing a coin than throwing a die

# Summary

概率p介于0和1之间

• probability p is always between 0 and 1

优势和优势比是从0到无穷大

• odds and odds ratio: 0 to  $+\infty$ 

•  $\log(\text{odds})$  and  $\log(\text{odds ratio})$ :  $-\infty$  to  $+\infty$ 

# Odd and probabilities

• If odds = p/(1-p), then:

- p = odds(1-p)
- p = odds odds\*p
- p + odds\*p = odds
- p(1 + odds) = odds
- p = odds/(1 + odds)

### Therefore

• p = 
$$\frac{1}{1 + \frac{1}{odds}}$$

(这一页大概就是推导了概率p的计算方法

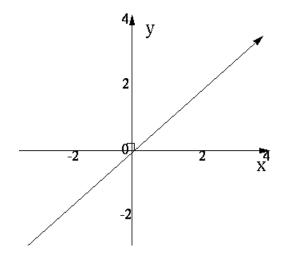
# Break

### 逻辑函数

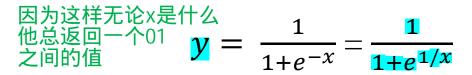
# Logistic function 我们希望使用回归模型从一组预测中计算 二元事件的概率

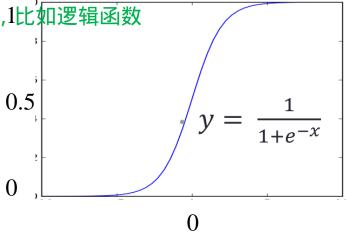
• Back to logistic regression: we want to use a regression model to calculate the probability of binary events (dead/alive, head/tail etc.) from a set of predictors:

$$y = a + b_1X_1 + b_2X_2...+ b_nX_n = a + \sum b_iX_i$$



- Problem: 问题:线性回归预测y在正负无穷之间,但概率p在01之间
  - linear regression predicts y between  $-\infty$  and  $+\infty$
  - but probability is always between 0 and 1!
- Solution: 解决方案:我们信息网概率是通过一个模型来估计的,1比如逻辑函数
  - we want our probabilities to be estimated by a model such as the logistic function
  - why? Because whatever x, it will always return a value between 0 and 1





# Link function: Logit

• We need a link function f to be the x in the  $\frac{\text{\text{\text{\text{\text{\text{\text{\text{\text{to}}}}}}}{\text{\text{wap}}}}{\text{\text{map}}}$  logistic function  $y = \frac{1}{1 + e^{-x}}$  and calculate y as

probability p:

$$p = \frac{1}{1 + e^{-f}} = \frac{1}{1 + \frac{1}{e^{f}}}$$

- But  $p = \frac{1}{1 + \frac{1}{odds}}$
- Therefore  $e^f = odds$ ; or f = log(odds)
- The link function f is called **logit p**:  $f = logit p = log(odds) = log(\frac{p}{1-p})$

### another derivation:

- If we want  $p = \frac{1}{1+e^{-f}}$ , then:
- $p = \frac{e^f}{e^f + 1}$
- $p(e^f+1) = e^f$
- $pe^f + p = e^f$
- $p = e^f pe^f$
- $p = e^{f}(1 p)$
- $e^f = \frac{p}{1-p}$
- $\log(e^f) = \log(\frac{p}{1-p})$
- $f = \log(\frac{p}{1-p})$

note: logit is always
natural log (i.e. log on base
e=2.71)

# f = logit = log(odds of event)

### 范围正负无穷

- f = logit or log(odds) range from  $-\infty$  to  $+\infty$
- therefore we can predict logits with a linear regression on our  $X_1$ ,  $X_2$  etc. variables 因此我们可以通过X1X2等变量的线性回归来预测I ogi ts

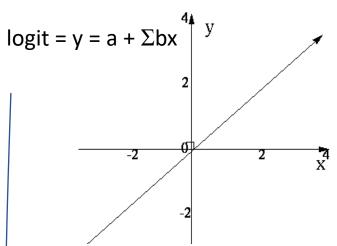
• The *logistic regression model* is thus

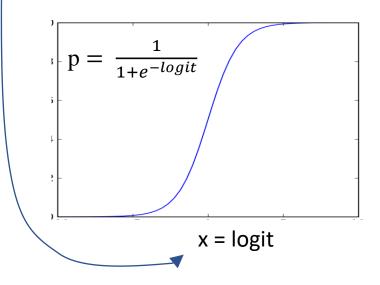
$$f = logit = log(\frac{p}{1-p}) = log(odds) = a + \sum b_i X_i$$

然后用逻辑回归中的logit来获得0和1之间的概率

• Then we use the logits in the logistic regression to obtain probabilities between 0 and 1!

$$p = \frac{1}{1+e^{-f}} = \frac{1}{1+e^{-\log(odds)}}$$





# Fitting logistic regression

参数a和b市通过MML方法估计的,不是通过最小二乘法

- The parameters a and  $b_i$  are estimated by MML (method of maximum likelihood), not by least squares
  - (we can't expand on MML in this course)

因此,统计显著性和拟合度不是基于方差,而是基于观测值和预测值之间"偏差"的度量 =个体案例预测正误的对比

- For this reason, statistical significance and goodness of fit are based not on variance, but on measures of 'deviance' between observed and predicted values
  - = comparison between right and wrong predictions of individual cases

但在线性回归中,估计的参数(系数、截距)具有决定其显著性的p值 基于与t和正态分布相关的z分布的显著性检验,解释为t-test或者F-test

- But as in linear regression, estimated parameters (coefficients, intercept) have a *P*-value that determines their significance
  - significance test based on a *z*-distribution related to *t* and normal distributions
  - interpreted like t-tests or F-tests. i.e. parameter is significant if P<0.05; 95% confidence intervals are provided etc. 举例:如果p<0.05则参数显著,提供95%置信区间等

### 逻辑回归:分类变量

# Logistic regression: categorical variable

Example: let's say we want to test the effect of smoking (X, yes or no) on hypertension (Y, also yes or no)

- Y=0: no hypertension; Y=1: hypertension Y=0无高血压, Y=1高血压 X=0非吸烟者(基线组)
- X=0: non-smoker (baseline group); X=1: smoker (exposure group) X=1吸烟者(暴露组)

### 逻辑回归的模型是

• Logistic regression model is then:

f = logit p = log(odds of hypertension) = a + b\*X

In baseline group, X=0; Therefore

• log(odds of having hypertension when X=0) = a + b\*0 = a

=the intercept = baseline = reference level (that is, the level of hypertension for nonsmokers X = 0截距=基线=参考水平

### 对a取幂,得到基线的概率

If we exponentiate a, we obtain odds at baseline

- e<sup>a</sup> = odds of hypertension for non-smokers
- $p = \frac{1}{1+e^{-a}} = \frac{odds \ of \ non-smokers}{1+odds \ of \ non-smokers} =$ probability of hypertension for non-smokers



### 这些是基线值

 Those are the baseline values, i.e. the odds and probabilities for groups without exposure (when all



 $X_i=0$ , i.e. even if nobody smoked)

# a + b = log (odds in the exposure group)

• Now the odds for smokers:

• 
$$f = logit = a + bX = a + b.1 = a + b$$



### a + b = log(odds of hypertension for smokers)

 $e^{a+b} = e^a e^b =$ the odds of hypertension for smokers

$$p = \frac{1}{1 + e^{-(a+b)}} = \frac{odds \ of \ smokers}{1 + odds \ of \ smokers} = probability \ of \ hypertension \ for \ smokers$$

Those are the results for the *exposure group* (smokers)

# Important: b=log(odds ratio)

So what is b then?

How likely is hypertension if you are a smoker compared a non-smoker?

• answer: it is the odds ratio (of hypertension in smokers to non-smokers)!

So: b=log(概率比)

odds ratio = odds(hypert. in smokers)/odds(hypert. in non-smokers)=  $e^a e^{b/}e^a = e^b$ 

And:

 $log(odds\ ratio) = log(e^b) = b$ 

逻辑回归中的系数b是log(暴露组中的高血压相对于基线)

- The coefficient b in the logistic regression is the log(odds of hypertension in exposure group *relative to baseline*)
  - in logistic regression, we test for significance of coefficient b
    - same as in linear regression! 逻辑回归中检验b的显著性和线性回归相同,对于变量的显著性,需要b不同于0,p<0.05
    - for a significant effect of variable, we need b different from 0 (i.e. P value < 0.05)
  - if b=0 (non-significant) 如果b=0,则不显著,暴露/基线=1,对于暴露和基线概率是相同的
     odds ratio for exposure vs. baseline = e<sup>b</sup> = e<sup>0</sup> = 1 =变量x对事件y的概率是没有影响的
    - odds ratio for exposure vs. baseline =  $e^b = e^0 = 1$
    - = the odds are the same for exposure and baseline (1 to 1),
    - = the variable X has no effect on probability of event Y

### Odds ratio b

• Let's add some *hypothetical* numbers to the example:

- odds of hypertension for smokers ( $=e^{a+b}$ ) = 0.3
- odds of hypertension for non-smokers ( $=e^a$ ) = 0.1
- The odds of hypertension in smokers would be three times higher in smokers
  - *odds ratio* = odds smokers/odds non smokers = 0.3/0.1 = 3

两组的概率比是一个非常有用的指标,代表了一个因素对事件发生的影响

• The *odds ratio of the two groups (exposure/baseline)* is a very useful representation of the effect of a factor on the occurrence of event

逻辑回归中总是给b或log,不是直接给事件的比值

- Logistic regression always reports b or **log of odds**, not odds of event in exposure group relative to baseline
  - more precisely, R reports *log(odds ratio of event in exposure vs. baseline)*
  - so in this example above, R would report b=log(3)=1.098612
    - We have to exponentiate b to obtain odds ratio = 3

R中总是给log值,需要自己取幂才能得到概率比

# Break

### Example in R: hypertension, smoking, obesity

- File *hypertension* presents data on people with or without hypertension as a function of two factors: smoking and obesity
- Cases coded as 'yes' or 'no' no按字母顺序排在第一位,并作为基线读取
  - 'no' comes first alphabetically and is read as baseline
  - alternatively: 'no'=0, 'yes'=1 (don't use 1 or 2!!!) no=0 yes=1
  - In this example, data are presented as a table
    - (we'll see a different way of presenting data with each case as a line)

### >hypertension

	smoking	obesity	total	hyper	nonhyper
1	no	no	247	40	207
2	yes	no	102	15	87
3	no	yes	59	16	43
4	yes	yes	25	8	17

### Example in R: hypertension, smoking, obesity

- When data are presented as table 表中有高血压和非高血压的数量
  - table has number of positives (hypertension, Y = 1) and negatives (no hypertension, Y = 0) 两个预测变量:吸烟和肥胖
  - two predictors or X variables:  $X_1 = \text{smoking}$ ,  $X_2 = \text{obesity}$ 
    - For both, yes = 1, no = 0
  - this has been done already for you (file *hypnonhyp*)
    - i.e. the dependent variable will be the matrix *hypnonhyp*
- Row 1: non-smoker, non-obese
- Row 2: smoker, non-obese
- Row 3: non-smoker, obese
- Row 4: smoker, obese

	hyper	nonhyper
1	40	207
2	15	87
3	16	43
4	8	17

Note: don't worry about the table!

• You will not be asked to create one!

# Running model

```
> model.hyper <- glm(hypnonhyp ~ smoking+obesity, binomial)
> summary(model.hyper)
Call:
glm(formula = hypnonhyp ~ smoking + obesity, family = binomial)
Deviance Residuals:
                        4
0.1593 -0.2520 -0.2653 0.4018
Coefficients:
                     Estimate Std. Error z value Pr(>|z|)
smokingyes
                     -0.01654 0.27617 -0.060 0.95224
obesityyes
                     0.76005
                              0.28270 2.689 0.00718 **
Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
(Dispersion parameter for binomial family taken to be 1)
 Null deviance: 7.15022 on 3 degrees of freedom
Residual deviance: 0.32067 on 1 degrees of freedom
AIC: 23.935
```

Number of Fisher Scoring iterations: 3

- Logistic regression is an example of generalised linear model
  - function *glm*
- Syntax is simple and similar to linear regression

### 逻辑模型写的像有两个预测因子的多元回归

- Logistic model written like a multiple regression with *two* predictors:
  - hypnonhyp ~ smoking+ obesity
  - (ps. interactions later)

### 自变量一定要添加bi nomi al

- Argument *binomial* sets logistic regression
  - Never forget to add:

family = binomial

Otherwise it fits a linear rather than the logistic regression!!!

否则更符合线性而不是逻辑回归

### Residuals

```
> model.hyper <- glm(hypnonhyp ~ smoking+obesity, binomial)
> summary(model.hyper)
Call:
glm(formula = hypnonhyp ~ smoking + obesity, family = binomial)
Deviance Residuals:
0.1593 -0.2520 -0.2653 0.4018
Coefficients:
                    Estimate Std. Error z value Pr(>|z|)
-0.01654 0.27617 -0.060 0.95224
smokingyes
obesityyes
                    Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
(Dispersion parameter for binomial family taken to be 1)
 Null deviance: 7.15022 on 3 degrees of freedom
Residual deviance: 0.32067 on 1 degrees of freedom
AIC: 23.935
Number of Fisher Scoring iterations: 3
```

### 残差是作为偏差(而不是方差)给出的

- Residuals are given
   as deviance (not
   variance) 每组中观察到的和预测的Logit值
  - difference between 是上面表格的yesno)
    observed and predicted
    logit values in each
    group (no/no, no/yes,
    yes/no, yes/yes)
  - residuals in logit scale (neither probability nor cell count)

# Intercept

```
> model.hyper <- glm(hypnonhyp ~ smoking+obesity, binomial)
                                                                 • Intercept a = -1.67
> summary(model.hyper)
Call:
                                                         a=log(基线组,高血压概率)

    a=log(odds of hypertension,

glm(formula = hypnonhyp ~ smoking + obesity, family = binomial)
                                                                   baseline group)
Deviance Residuals:
                       4
                                                                     • =non-smokers (X1=0), non-obese
                                                                      (X2 = 0)
0.1593 -0.2520 -0.2653 0.4018
Coefficients:
                                                                     • e<sup>a</sup> =the odds of hypertension if you're
                    Estimate Std. Error z value Pr(>|z|)
                                                                      non-smoker, non-obese
                    -1.67143  0.16731  -9.990  < 2e-16 ***
(Intercept)
                                                                    • a=0.188
smokingyes
                    -0.01654 0.27617 -0.060 0.95224
                                                        a取幂等于不吸烟不肥胖,得高血压的概率
obesityyes
                     0.76005
                             0.28270 2.689 0.00718 **
                                                  z检验: 截距明显 z-test: intercept is significantly
                                                                   different from 0
                                                   不同干0
Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
                                                  基线中的高血压概率 odds of hypertension in baseline (e<sup>a</sup>)= 样本中基线时高血压not 1
(Dispersion parameter for binomial family taken to be 1)
 Null deviance: 7.15022 on 3 degrees of freedom
                                                   概率不是0.5

    probability of hypertension in baseline

Residual deviance: 0.32067 on 1 degrees of freedom
```

截距a=-1.67

different from 0.5 in the sample

Number of Fisher Scoring iterations: 3

AIC: 23.935

# Effect of smoking

> model.hyper <- glm(hypnonhyp ~ smoking+obesity, binomial)

> summary(model.hyper)

Call:

glm(formula = hypnonhyp ~ smoking + obesity, family = binomial)

**Deviance Residuals:** 

4

0.1593 -0.2520 -0.2653 0.4018

Coefficients:

Estimate Std. Error z value Pr(>|z|)

-1.67143 0.16731 -9.990 < 2e-16 \*\*\* (Intercept)

smokingyes -0.01654 0.27617 -0.060 0.95224

obesityyes 0.76005 0.28270 2.689 0.00718 \*\*

Signif. codes: 0 '\*\*\*' 0.001 '\*\*' 0.01 '\*' 0.05 '.' 0.1 ' ' 1

(Dispersion parameter for binomial family taken to be 1)

Null deviance: 7.15022 on 3 degrees of freedom

Residual deviance: 0.32067 on 1 degrees of freedom

AIC: 23.935

Number of Fisher Scoring iterations: 3

#### 吸烟回归系数b

### Regression coefficient b for smoking:

- smokers (X=1) are shown as smokingyes,
  - variable name plus group ('yes')
- b=log(odds ratio)=-0.0165
- =log odds of hypertension for smokers relative to non-smokers

b=log(吸烟得高血压概率/不吸烟得高血压概率)

- b与0无显著差异 • But P(z) = 0.95!
  - b is not significantly different from 0
  - odds ratio not different from  $e^0=1$

概率比与1无差异

得高血压概率

所以在这个实验中So smokers are not more likely to 吸烟者和不吸烟者have hypertension than nonsmokers in this hypothetical sample

• (don't start smoking because of me!)

# Effect of obesity

0.76005 0.28270 2.689 0.00718 \*\*

---

Signif. codes: 0 '\*\*\*' 0.001 '\*\*' 0.01 '\*' 0.05 '.' 0.1 ' ' 1 (Dispersion parameter for binomial family taken to be 1)

Null deviance: 7.15022 on 3 degrees of freedom Residual deviance: 0.32067 on 1 degrees of freedom

AIC: 23.935

obesityyes

Number of Fisher Scoring iterations: 3

### 肥胖的回归系数b=0.76

- Regression coefficient b for obesity: b=0.76
  - =log odds of hypertension for obese relative to non-obese
- P(z) = 0.00718 b明显不同于0
  - b is significantly different from 0
  - b = ln(odds of hypertension in obese relative to baseline) > 0
  - odds ratio= e<sup>0.76</sup> = 2.14 概率比2.14
    - odds ratio >1; obese at higher risk!

概率比>1,肥胖有危险

在这个样本中 • Obesity more than doubles odds 肥胖使高血压的几率 of hypertension *in this sample* 

### 拟合度

### Goodness of fit

```
> model.hyper <- glm(hypnonhyp ~ smoking+obesity, binomial)
> summary(model.hyper)
Call:
glm(formula = hypnonhyp ~ smoking + obesity, family = binomial)
Deviance Residuals:
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Coefficients:
                       Estimate Std. Error z value Pr(>|z|)
                       -1.67143  0.16731  -9.990  < 2e-16 ***
(Intercept)
smokingyes
                       -0.01654 0.27617 -0.060 0.95224
                                 0.28270 2.689 0.00718 **
obesityyes
                       0.76005
Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
(Dispersion parameter for binomial family taken to be 1)
  Null deviance: 7.15022 on 3 degrees of freedom
```

### MML不用方差来衡量拟合度

- MML does not use variance to measure goodness of fit
  - it includes no 'dispersion parameter', which has to be taken as 1

### 在MML中,偏差取代方差

- In MML, deviance replaces variance 零偏差=当模型仅包括截距时的偏差(=
  - null deviance = deviance when model 因子 includes only intercept (=before 吸烟和 predictors *smoking* and *obesity*) 肥胖之前)

### 剩余偏差是预测因素后无法解释的偏差

 residual deviance is unexplained deviance after predictors

### 零偏差和剩余偏差之间的差异是预测因子对模型的贡献

 difference between null and residual is the contribution of predictors to model

Number of Fisher Scoring iterations: 3

AIC: 23.935

Residual deviance: 0.32067 on 1 degrees of freedom

### Goodness of fit

> model.hyper <- glm(hypnonhyp ~ smoking+obesity, binomial)

> summary(model.hyper)

#### Call:

glm(formula = hypnonhyp ~ smoking + obesity, family = binomial) **Deviance Residuals:** 

0.1593 -0.2520 -0.2653 0.4018

#### Coefficients:

Estimate Std. Error z value Pr(>|z|)

-1.67143 0.16731 -9.990 < 2e-16 \*\*\* (Intercept)

smokingyes -0.01654 0.27617 -0.060 0.95224

obesityyes 

Signif. codes: 0 '\*\*\*' 0.001 '\*\*' 0.01 '\*' 0.05 '.' 0.1 ' ' 1

(Dispersion parameter for binomial family taken to be 1)

Null deviance: 7.15022 on 3 degrees of freedom

Residual deviance: 0.32067 on 1 degrees of freedom

AIC: 23.935

Number of Fisher Scoring iterations: 3

## 因为没有方差,所以拟合度不是用R方来衡量的 我们用AIC(akai ke信息准则)来替代

- Because there is no variance, goodness of fit is not measured by R<sup>2</sup>
  - we use AIC (Akaike Information Criterion) instead

在回归中添加额外的预测因子可能会增加拟合优度,即使

• Remember: adding additional 類測因子 predictors to regression may increase goodness of fit even when predictor is not significant AIC测量拟合优度,同时惩罚使用额外预测因子的模型

- AIC measures goodness of fit while punishing models for use of additional predictors
  - the better and more parsimonious the model, the lower the AIC

### 模型越好、越节约,AIC就会越低

 Models with lowest AIC are selected

选择AIC最低的模型

# Guide to interpretation and calculations:

- a = log(odds of event in baseline group)
- exp(a) = baseline odds of event
- Probability p of event in baseline: baseline odds/(1 + baseline odds)

基线事件概率:基线比值/(1+基线比值)

### Then

### 如果b显著则不同于0

- b = log(odds ratio); if b is significant (different from 0):
- exp(b) = odds ratio

#### 暴露组的比值

- exp(a+b) = exp(a)\*exp(b) = odds(baseline)\*odds ratio = odds in exposure group
- Probability p in exposure group = exposure odds/(1 + exposure odds)

### Exercises

- Since *smoking* is not significant, you must optimise the model by excluding *smoking*, and run model only with variable *obesity* (manually, or with *step* function)
- 1. Is a significant? What does that mean?
- 2. Is b significant? What does that mean?
- Calculate:
- 3. Baseline odds of hypertension
- 4. Odds ratio of hypertension (obese vs. non-obese)
- 5. Odds of hypertension in obese
- 6. Probability of hypertension in non-obese
- 7. Probability of hypertension in obese