

md example

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Problem 3

A simple linear regression model is given by $Y = \beta_0 + \beta_1 X + \epsilon$, where $\epsilon \sim N(0, \sigma^2)$.

- a. Use the method of least squares, as discussed in class to derive the least squares estimators for β_0 and β_1 . The multiple regression model was derived in class.

Solution

Suppose we have $i = 1, \dots, n$ data points where x_1, \dots, x_n is the independent variable and y_1, \dots, y_n is the dependent variable. We want $y_i = \beta_0 + \beta_1 x_i + \epsilon_i$ where $\epsilon_i \sim N(0, \sigma^2)$, but we obtain $\hat{y}_i = \hat{\beta}_0 + \hat{\beta}_1 x_i + \epsilon_i$. We need to try to minimize the sum of the squared residuals defined as follows:

$$L = \sum_{i=1}^n (y_i - \hat{y}_i)^2 \quad (1)$$

We want to minimize this function L with respect to β_0 and β_1 . So, our problem becomes

$$\min_{\beta_0, \beta_1} \sum_{i=1}^n (y_i - \hat{y}_i)^2 \quad (2)$$

which simplifies to

$$\min_{\beta_0, \beta_1} \sum_{i=1}^n (y_i - \hat{\beta}_0 - \hat{\beta}_1 x_i)^2 \quad (3)$$

We get our key equations by taking the partial derivative with respect to each of the parameters we want to minimize and setting them equal to zero:

$$\frac{\partial L}{\partial \hat{\beta}_0} = \sum_{i=1}^n -2(y_i - \hat{\beta}_0 - \hat{\beta}_1 x_i) = 0 \quad (4)$$

$$\frac{\partial L}{\partial \hat{\beta}_1} = \sum_{i=1}^n -2x_i(y_i - \hat{\beta}_0 - \hat{\beta}_1 x_i) = 0 \quad (5)$$

Solving (4), we find that:

$$\sum_{i=1}^n -2(y_i - \hat{\beta}_0 - \hat{\beta}_1 x_i) = 0 \quad (6)$$

$$\sum_{i=1}^n (y_i - \hat{\beta}_0 - \hat{\beta}_1 x_i) = 0 \quad (7)$$

$$\text{Recall that } \sum_{i=1}^n y_i = n\bar{y} \text{ since } \frac{\sum_{i=1}^n y_i}{n} = \bar{y} \quad (8)$$

$$n\bar{y} - n\hat{\beta}_0 - n\hat{\beta}_1 \bar{x} = 0 \quad (9)$$

$$n\hat{\beta}_0 = n\bar{y} - n\hat{\beta}_1 \bar{x} \quad (10)$$

$$\hat{\beta}_0 = \bar{y} - \hat{\beta}_1 \bar{x} \quad (11)$$

Solving (5), we see that:

$$\sum_{i=1}^n -2x_i(y_i - \hat{\beta}_0 - \hat{\beta}_1 x_i) = 0 \quad (12)$$

$$\sum_{i=1}^n x_i(y_i - \hat{\beta}_0 - \hat{\beta}_1 x_i) = 0 \quad (13)$$

$$\sum_{i=1}^n x_i y_i - \sum_{i=1}^n \hat{\beta}_0 x_i - \sum_{i=1}^n \hat{\beta}_1 x_i^2 = 0 \quad (14)$$

$$\sum_{i=1}^n x_i y_i - \sum_{i=1}^n (\bar{y} - \hat{\beta}_1 \bar{x}) x_i - \sum_{i=1}^n \hat{\beta}_1 x_i^2 = 0 \quad (15)$$

$$\sum_{i=1}^n x_i y_i - \bar{y} \sum_{i=1}^n x_i + \hat{\beta}_1 \bar{x} \sum_{i=1}^n x_i - \sum_{i=1}^n \hat{\beta}_1 x_i^2 = 0 \quad (16)$$

$$\sum_{i=1}^n x_i y_i - n\bar{x}\bar{y} + n\hat{\beta}_1 \bar{x}^2 - \hat{\beta}_1 \sum_{i=1}^n x_i^2 = 0 \quad (17)$$

$$\sum_{i=1}^n \hat{\beta}_1 x_i^2 - n\hat{\beta}_1 \bar{x}^2 = \sum_{i=1}^n x_i y_i - n\bar{x}\bar{y} \quad (18)$$

$$\hat{\beta}_1 (\sum_{i=1}^n x_i^2 - n\bar{x}^2) = \sum_{i=1}^n x_i y_i - n\bar{x}\bar{y} \quad (19)$$

$$\hat{\beta}_1 = \frac{\sum_{i=1}^n x_i y_i - n\bar{x}\bar{y}}{\sum_{i=1}^n x_i^2 - n\bar{x}^2} \quad (20)$$

Using the fact that

$$\sum_{i=1}^n (x_i - \bar{x})(y_i - \bar{y}) = \sum_{i=1}^n x_i y_i - n\bar{x}\bar{y}$$

and

$$\sum_{i=1}^n (x_i - \bar{x})^2 = \sum_{i=1}^n x_i^2 - n\bar{x}^2$$

we can rewrite this in the difference of means form:

$$\hat{\beta}_1 = \frac{\sum_{i=1}^n (x_i - \bar{x})(y_i - \bar{y})}{\sum_{i=1}^n (x_i - \bar{x})^2}$$

- b. Now, forcing the intercept coefficient to be zero, re-derive the least squares estimator for the slope. Under what circumstances do you think it would be useful?

Solution

By forcing the intercept to zero, we want to model $y_i = \beta x_i + \epsilon_i$ where $\epsilon_i \sim N(0, \sigma^2)$, but we obtain $\hat{y}_i = \hat{\beta} x_i + \epsilon_i$. This time, we are minimizing the function

$$L = \sum_{i=1}^n (y_i - \hat{\beta} x_i)^2 \quad (21)$$

We want to minimize this function L with respect to β . So, our problem becomes

$$\min_{\hat{\beta}} \sum_{i=1}^n (y_i - \hat{\beta} x_i)^2 \quad (22)$$

Taking the derivative and setting the equation equal to zero, we have our first order condition:

$$\frac{dL}{d\hat{\beta}_0} = \sum_{i=1}^n -2x_i(y_i - \hat{\beta} x_i) = 0 \quad (23)$$

Solving (23), we find

$$\sum_{i=1}^n x_i(y_i - \hat{\beta} x_i) = 0 \quad (24)$$

$$\sum_{i=1}^n x_i y_i - \hat{\beta} \sum_{i=1}^n x_i^2 = 0 \quad (25)$$

$$\hat{\beta} \sum_{i=1}^n x_i^2 = \sum_{i=1}^n x_i y_i \quad (26)$$

$$\hat{\beta} = \frac{\sum_{i=1}^n x_i y_i}{\sum_{i=1}^n x_i^2} \quad (27)$$

$$(28)$$

Most of the time it does not make sense to use a model without an intercept in practice, but it can be helpful if there are physical constraints on the variables. For example, if we want to use the speed a car is traveling to predict the stopping distance of the car, then it might make sense to force the intercept coefficient to zero.