

10-Factor Experimental Design on Simulated Data

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12/4/2021

1. Introduction

For this project, a secret model was created with ten independent variables and one dependent variable. The dependent variable is determined by an unknown function: $Y = f(a, b, \dots, j) + \epsilon$. The value of each dependent variable, x_i , satisfies $-1 \leq x_i \leq 1$. To discern information about the model, we are tasked to create an experiment design with a number of run that each specify a particular setting of each dependent variable. Each run is then assigned a cost. A run where every variable is set to ± 1 costs 1 point, a run where each variable $x_i \in \{0, \pm 1\}$ costs 3 points, etc. If additional designs need to be run, points would be deducted. Finally, points would be awarded to correctly identifying the presence (or lack thereof) of each independent variable within the hidden model. This scenario calls for a **factor screening** experiment. Since the main factor settings will be ± 1 , our starting point is a **two level factorial** design. With ten possible variables, a full 2^{10} design would require 1028 runs. This would cost over 1000 points and cost almost more than the total amount available on offer. Reducing this necessitates a **two-level fractional factorial** design, reducing the number of runs.

While a two-level fractional factorial (I will refer to it as **2FrF** going forward) design may allow us to detect linear or approximately-linear effects of factors and their interaction terms, with simply this design we are vulnerable to omitting significant factors that offer second-order (or even) effects. For example, if the unknown function $f(a, b, \dots, j)$ contains a factor x^2 or $\cos(x)$, there exists a possibility that it will remain undetected in a 2FrF design. To this end, it would be useful to add **center runs** and **axial runs**. The combination of the two suggests a **central composite design (CCD)**. Ideally, we would create a **small composite design (SCD)** for 10 factors, however I was only able to find a 9-factor SCD. Instead, I worked with a **hybrid design** by adding center and axis runs to a 2FrF design. Because of the points penalty for variable values that are not (-1), 0, or 1, these axial runs will be **face-centered**. This is a very important detail to note. The consequence of face-centered axial runs is that inaccuracy is introduced when modeling the coefficients corresponding to the a factor effect. However, this is a worthwhile trade off. Our priority is to *detect* the factor effects. Note also, however, that certain methodology (such as linear modeling) will be less accurate, create lower R^2 values, and have lower p-values requiring extra caution.

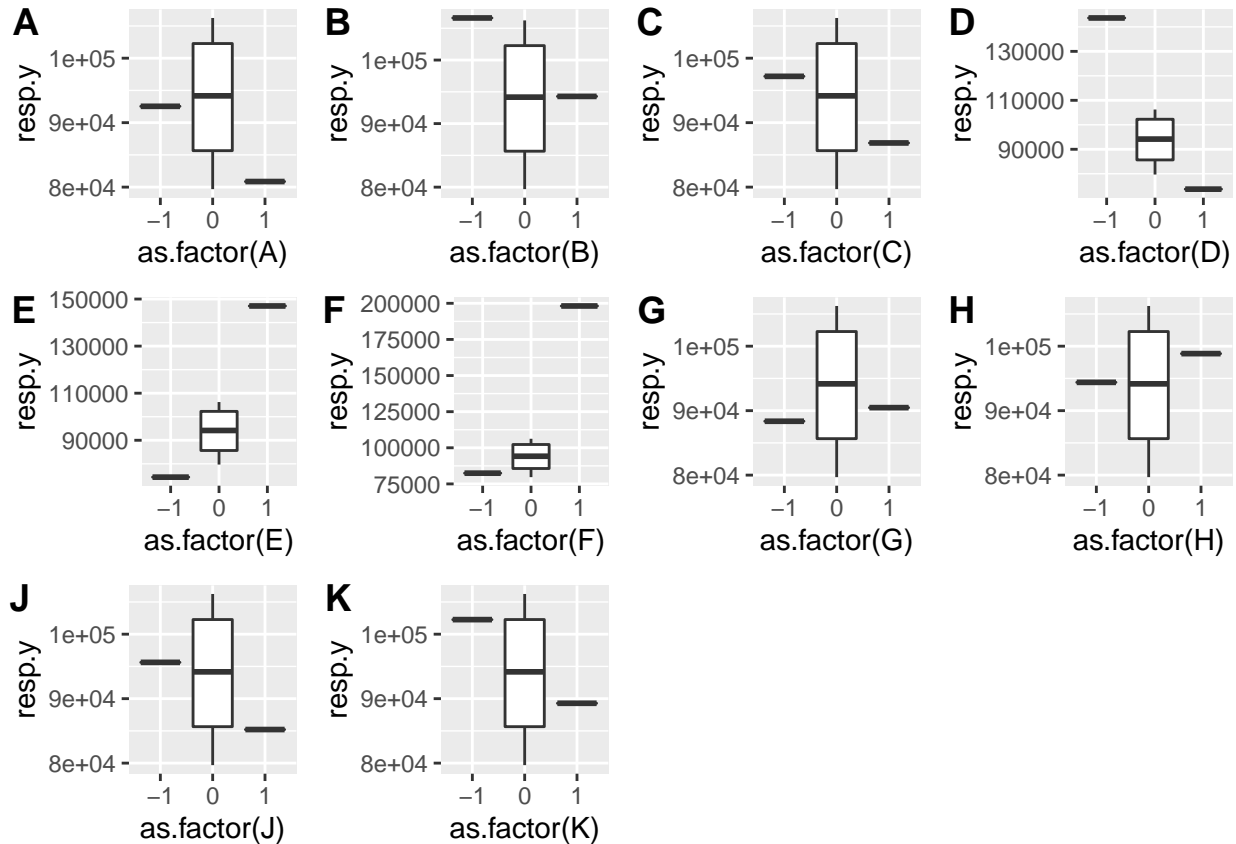
NOTE: FOR THE REMAINDER OF THIS PROJECT, I WILL REFER TO THE INDEPENDENT VARIABLES AS $A, B, C, D, E, F, G, H, J, K$, OMITTING I AS WAS USED IN THE ASSIGNMENT DOCUMENTS AS I IS USUALLY RESERVED FOR THE INTERCEPT EFFECT.

2. Methods

The design I decided upon was a hybrid design. A 2_{V}^{10-3} design with the axial runs of a face-centered central composite design. This design was produced in R with the help of the **FrF2** and **DoE** packages. Analysis performed on them was aided by the **RcmdrPlugin.DoE** package. The design's generators are: $H = ABCDE, J = ABCFG, K = ABCDF$. The creation procedure is detailed in **Appendix 6.1**. The full design documentation (including matrix and attributes) is available in **Appendix 6.2**.

3. Analysis

The central and axial runs are incredibly valuable. In an axial run, all variables except one are set to 0. This means that all main and interaction effects are wiped with the exception of the main effect of the non-zero variable. My first step was to compare the effects of each individual variable on the function output compared with a baseline provided by the center runs. While there is a significant limitation to this (only one run exists for each variable at -1 and 1), if a difference is significant enough, it will mean the variable should be concluded. An additional benefit is that these axial runs bypass the problem of aliasing. However, because they are face-centered, we are likely to get inaccurate estimates for linear parameters.



A visual analysis of the means suggests that *D*, *E*, and *F* have significant main effects in the model. There is a possibility of outliers because of the limited sample size, but all of the effects are quite extreme. There is also the possibility that their effects are non-linear. *F* especially has a possibly exponential effect considering the asymmetry in effect on the mean from between its two settings. Below are the p-values for the F-tests on each variable's fitted model.

```
fstats <- summary(lm(respy ~ D, data = design_centers))$fstatistic
1 - pf(fstats[1], fstats[2], fstats[3])

##          value
## 0.05591169

summary(lm(log(respy) ~ D, data = design_centers, na.action = na.omit))

##
## Call:
## lm.default(formula = log(respy) ~ D, data = design_centers,
##            na.action = na.omit)
```

```
##
## Residuals:
##      Min       1Q   Median       3Q      Max
## -0.26318 -0.09899 -0.02553  0.05884  0.71727
##
## Coefficients:
##              Estimate Std. Error t value Pr(>|t|)
## (Intercept) 11.47955    0.03806 301.594  <2e-16 ***
## D           -0.33353    0.13985  -2.385   0.025  *
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
##
## Residual standard error: 0.1978 on 25 degrees of freedom
## Multiple R-squared:  0.1853, Adjusted R-squared:  0.1528
## F-statistic: 5.688 on 1 and 25 DF,  p-value: 0.02498
fstats <- summary(lm(resp.y ~ E, data = design_centers))$fstatistic
1 - pf(fstats[1], fstats[2], fstats[3])

##      value
## 0.04609538
summary(lm(log(resp.y) ~ E, data = design_centers, na.action = na.omit))

##
## Call:
## lm.default(formula = log(resp.y) ~ E, data = design_centers,
##      na.action = na.omit)
##
## Residuals:
##      Min       1Q   Median       3Q      Max
## -0.27157 -0.09899 -0.02553  0.06192  0.71727
##
## Coefficients:
##              Estimate Std. Error t value Pr(>|t|)
## (Intercept) 11.47955    0.03786 303.178  <2e-16 ***
## E           0.34108    0.13912   2.452   0.0215  *
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
##
## Residual standard error: 0.1967 on 25 degrees of freedom
## Multiple R-squared:  0.1938, Adjusted R-squared:  0.1616
## F-statistic: 6.011 on 1 and 25 DF,  p-value: 0.02155
fstats <- summary(lm(resp.y ~ F, data = design_centers))$fstatistic
1 - pf(fstats[1], fstats[2], fstats[3])

##      value
## 0.000624253
summary(lm(log(resp.y) ~ F, data = design_centers, na.action = na.omit))

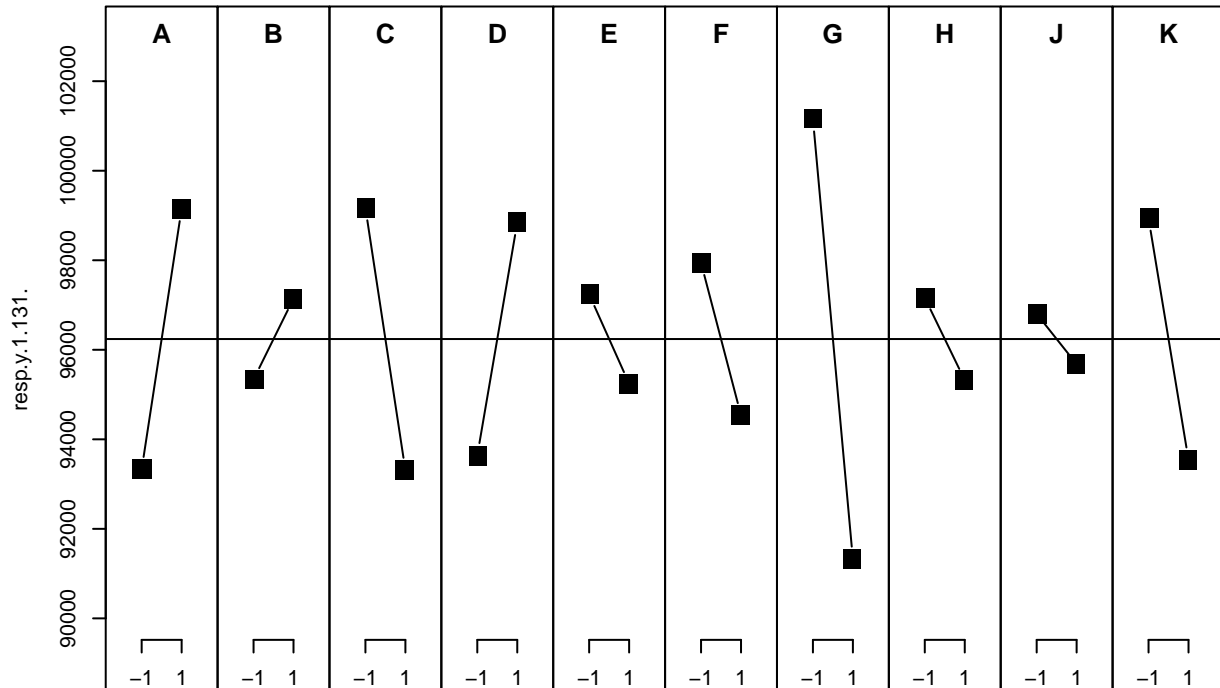
##
## Call:
## lm.default(formula = log(resp.y) ~ F, data = design_centers,
##      na.action = na.omit)
##
```

```
## Residuals:
##      Min       1Q   Median       3Q      Max
## -0.27157 -0.09899 -0.02553  0.06192  0.41898
##
## Coefficients:
##              Estimate Std. Error t value Pr(>|t|)
## (Intercept) 11.47955    0.03476 330.212 < 2e-16 ***
## F           0.43855    0.12773   3.433  0.00209 **
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
##
## Residual standard error: 0.1806 on 25 degrees of freedom
## Multiple R-squared:  0.3204, Adjusted R-squared:  0.2932
## F-statistic: 11.79 on 1 and 25 DF,  p-value: 0.002086
```

To test for potential transformations, a logarithmic transformation was applied to the responses and another model was fitted. This pushed D and E in a direction to suggests a likelier presence of their main effects, F got pushed closer to the rejection region. All three are likely significant and present within the model. If their effects are transformed, they may not show up when analyzing the full experimental results. However, because this is a factor screening experiment, it is safer to include an effect that can then be further trialed than to remove one that to reject one too early.

Next we analyze the experimental results as a whole. A good first step is to look at a main effects plot, identifying potential effects. From the below plot, we can see that the effects are $G, C, A, K, D, F, E, H, B, J$ (roughly in order from most to least significance).

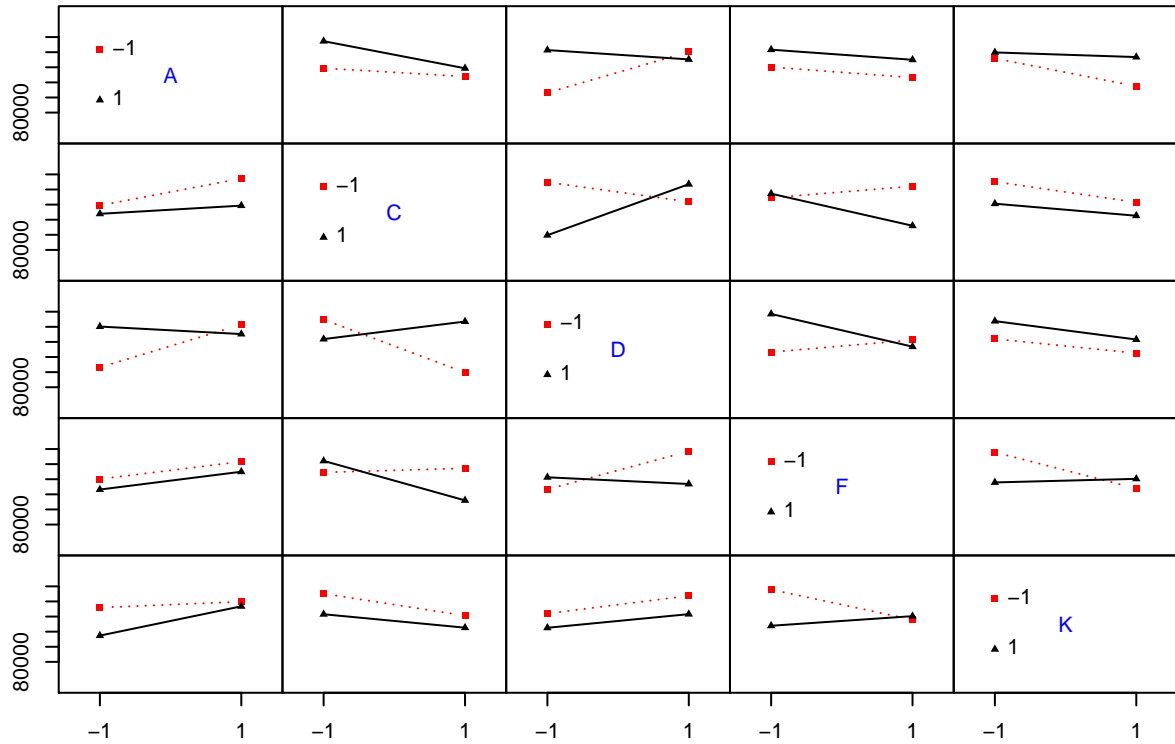
Main effects plot for resp.y.1.131.



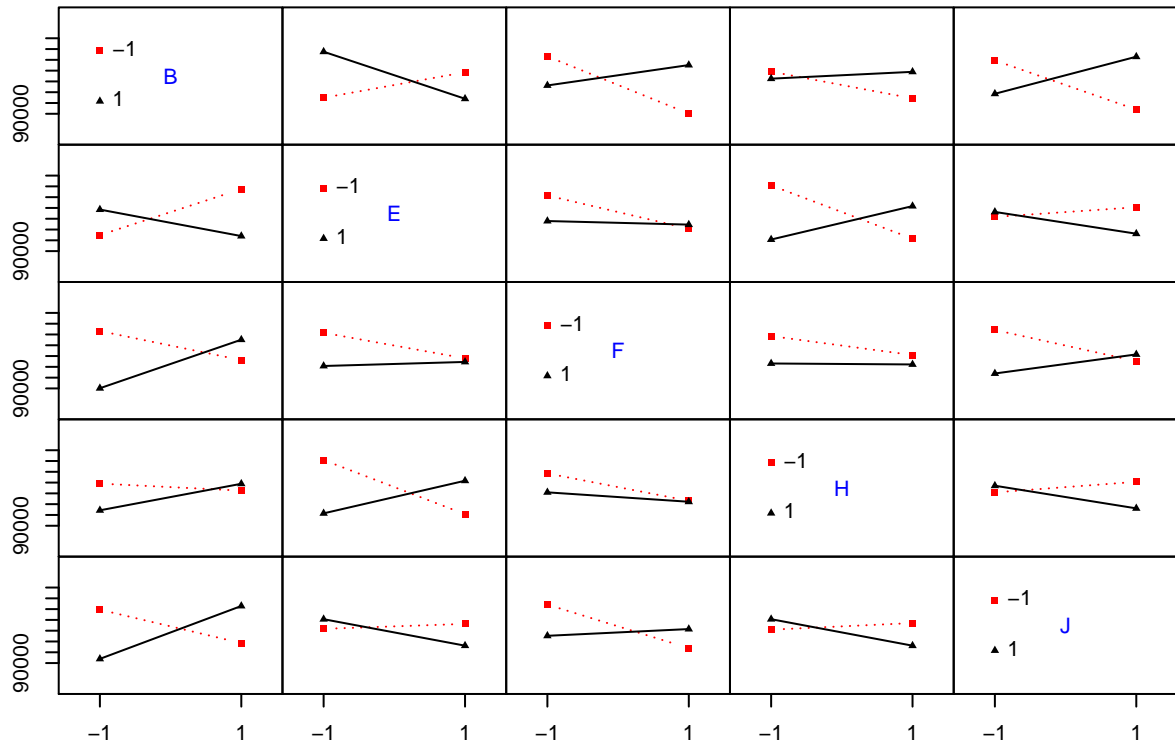
There is already enough evidence to keep D , E , and F within the model, we should further investigate the remaining variables. The least significant variables in the main effects plot will only exist in the model if

they have strong interactions with other variables. Since our design is resolution five, the lowest interaction term that can alias with a single interaction is a four-factor interaction. By scarcity of effects, we can conclude that if any variable has a weak main effect and weak two-factor interactions, it is unlikely to be present within the model.

Interaction plot matrix for resp.y.1.131.



Interaction plot matrix for resp.y.1.131.



```
LinearModel.1 <- lm(resp.y ~ (A + B + C + D + E + F + G + H + J + K)^2 +
  I(A^2) + I(B^2) + I(C^2) + I(D^2) + I(E^2) + I(F^2) + I(G^2) + I(H^2) +
  I(J^2) + I(K^2), data=design)
summary(aov(LinearModel.1))
```

	Df	Sum Sq	Mean Sq	F value	Pr(>F)
## A	1	1.004e+09	1.004e+09	0.923	0.3392
## B	1	8.224e+07	8.224e+07	0.076	0.7840
## C	1	1.144e+09	1.144e+09	1.052	0.3078
## D	1	5.333e+08	5.333e+08	0.490	0.4856
## E	1	2.324e+07	2.324e+07	0.021	0.8841
## F	1	8.085e+07	8.085e+07	0.074	0.7857
## G	1	3.030e+09	3.030e+09	2.787	0.0986
## H	1	9.894e+07	9.894e+07	0.091	0.7636
## J	1	5.226e+07	5.226e+07	0.048	0.8270
## K	1	9.895e+08	9.895e+08	0.910	0.3427
## I(A^2)	1	3.513e+08	3.513e+08	0.323	0.5712
## I(B^2)	1	9.793e+07	9.793e+07	0.090	0.7648
## I(C^2)	1	1.531e+07	1.531e+07	0.014	0.9058
## I(D^2)	1	3.128e+08	3.128e+08	0.288	0.5931
## I(E^2)	1	2.674e+08	2.674e+08	0.246	0.6212
## I(F^2)	1	2.672e+09	2.672e+09	2.457	0.1206
## I(G^2)	1	5.239e+08	5.239e+08	0.482	0.4894
## I(H^2)	1	1.066e+08	1.066e+08	0.098	0.7550
## I(J^2)	1	3.025e+08	3.025e+08	0.278	0.5992
## I(K^2)	1	7.038e+07	7.038e+07	0.065	0.7998

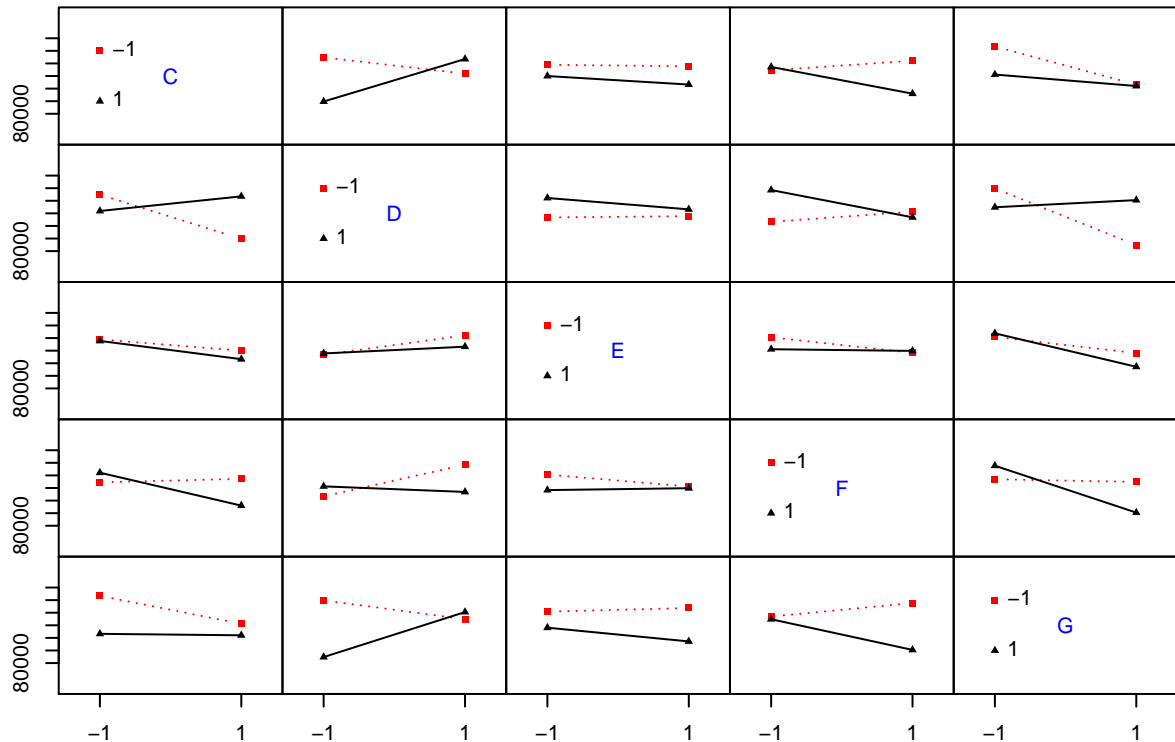
```

## A:B      1 2.002e+07 2.002e+07 0.018 0.8924
## A:C      1 3.105e+08 3.105e+08 0.285 0.5945
## A:D      1 2.206e+09 2.206e+09 2.028 0.1579
## A:E      1 2.152e+08 2.152e+08 0.198 0.6575
## A:F      1 1.446e+05 1.446e+05 0.000 0.9908
## A:G      1 5.848e+05 5.848e+05 0.001 0.9816
## A:H      1 8.883e+07 8.883e+07 0.082 0.7757
## A:J      1 1.365e+09 1.365e+09 1.256 0.2655
## A:K      1 4.779e+08 4.779e+08 0.439 0.5091
## B:C      1 5.547e+06 5.547e+06 0.005 0.9432
## B:D      1 4.934e+07 4.934e+07 0.045 0.8318
## B:E      1 1.446e+09 1.446e+09 1.330 0.2519
## B:F      1 1.655e+09 1.655e+09 1.522 0.2206
## B:G      1 1.837e+09 1.837e+09 1.689 0.1971
## B:H      1 3.096e+08 3.096e+08 0.285 0.5950
## B:J      1 2.045e+09 2.045e+09 1.881 0.1737
## B:K      1 1.620e+08 1.620e+08 0.149 0.7004
## C:D      1 4.361e+09 4.361e+09 4.010 0.0483 *
## C:E      1 6.123e+07 6.123e+07 0.056 0.8130
## C:F      1 1.665e+09 1.665e+09 1.531 0.2192
## C:G      1 8.917e+08 8.917e+08 0.820 0.3676
## C:H      1 5.525e+07 5.525e+07 0.051 0.8222
## C:J      1 1.106e+09 1.106e+09 1.017 0.3159
## C:K      1 6.416e+07 6.416e+07 0.059 0.8086
## D:E      1 2.041e+08 2.041e+08 0.188 0.6659
## D:F      1 1.763e+09 1.763e+09 1.621 0.2063
## D:G      1 5.151e+09 5.151e+09 4.736 0.0322 *
## D:H      1 1.463e+07 1.463e+07 0.013 0.9079
## D:J      1 2.118e+07 2.118e+07 0.019 0.8893
## D:K      1 1.517e+07 1.517e+07 0.014 0.9063
## E:F      1 2.412e+08 2.412e+08 0.222 0.6388
## E:G      1 3.825e+08 3.825e+08 0.352 0.5546
## E:H      1 2.068e+09 2.068e+09 1.902 0.1714
## E:J      1 2.712e+08 2.712e+08 0.249 0.6188
## E:K      1 5.500e+06 5.500e+06 0.005 0.9435
## F:G      1 2.469e+09 2.469e+09 2.270 0.1354
## F:H      1 8.856e+07 8.856e+07 0.081 0.7760
## F:J      1 6.952e+08 6.952e+08 0.639 0.4261
## F:K      1 1.389e+09 1.389e+09 1.277 0.2615
## G:H      1 1.741e+07 1.741e+07 0.016 0.8996
## G:J      1 1.840e+09 1.840e+09 1.692 0.1967
## G:K      1 3.201e+08 3.201e+08 0.294 0.5888
## H:J      1 3.018e+08 3.018e+08 0.278 0.5996
## H:K      1 3.008e+08 3.008e+08 0.277 0.6002
## J:K      1 1.080e+08 1.080e+08 0.099 0.7534
## Residuals 89 9.678e+10 1.087e+09
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1

IAPlot(design_2, abbrev=4, show.alias=FALSE, select=c(3,4,5,6,7))

```

Interaction plot matrix for resp.y.1.131.



Both our preliminary visual inspection and our preliminary linear model, which samples all main effects and two-factor interaction terms (including quadratic effects), suggest that A, B, E, F, H, J, K are unlikely to have two-factor interactions. Since A, B, H, J, K have no significant one- or two-factor effects, it is likely they are not present anywhere in the model. By running a stepwise variable selection, we fall onto the model $y \sim D + G + D:G$.

```
stepwise(LinearModel.1, direction='backward/forward', criterion='BIC', trace = 0)
```

```
##
## Direction: backward/forward
## Criterion: BIC
##
## Call:
## lm.default(formula = resp.y ~ D + G + D:G, data = design)
##
## Coefficients:
## (Intercept)          D          G          D:G
##      96754       2025      -4828       6343
```

Next, let's examine the potential of the C variable. While visually, its main effect appears larger than that of most other variables, it is still statistically insignificant when quantitatively analyzed. However, it has an interaction term with D that is very close to significance. By fitting a linear model to the formula suggested by our stepwise procedure, $y \sim D + G + D:G$, we can add terms to it and conduct an F-test on the nested values to evaluate if specific terms should or shouldn't be included. This procedure, conducted below, suggests that $C : D$ is a worthwhile addition to the model. However, it does not suggest including E and F despite our previous evidence.


```

LinearModel.2 <- lm(resp.y ~ D * G, data=design)
LinearModel.3 <- lm(resp.y ~ D*G + C:D, data=design)
LinearModel.4 <- lm(resp.y ~ D*G + C:D + E + F, data=design)

anova(LinearModel.2, LinearModel.3)

## Analysis of Variance Table
##
## Model 1: resp.y ~ D * G
## Model 2: resp.y ~ D * G + C:D
##   Res.Df      RSS Df Sum of Sq      F Pr(>F)
## 1      151 1.3789e+11
## 2      150 1.3353e+11  1 4.361e+09 4.8989 0.02838 *
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1

anova(LinearModel.3, LinearModel.4)

```

```

## Analysis of Variance Table
##
## Model 1: resp.y ~ D * G + C:D
## Model 2: resp.y ~ D * G + C:D + E + F
##   Res.Df      RSS Df Sum of Sq      F Pr(>F)
## 1      150 1.3353e+11
## 2      148 1.3343e+11  2 104089003 0.0577 0.9439

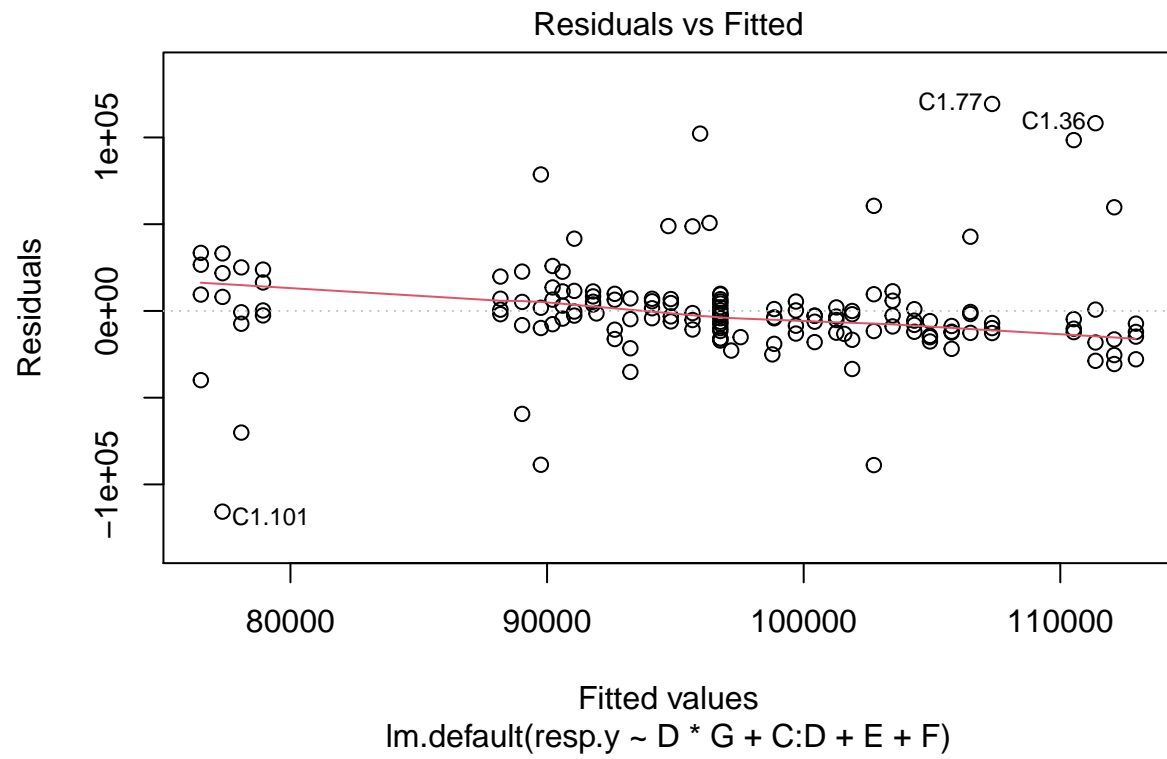
```

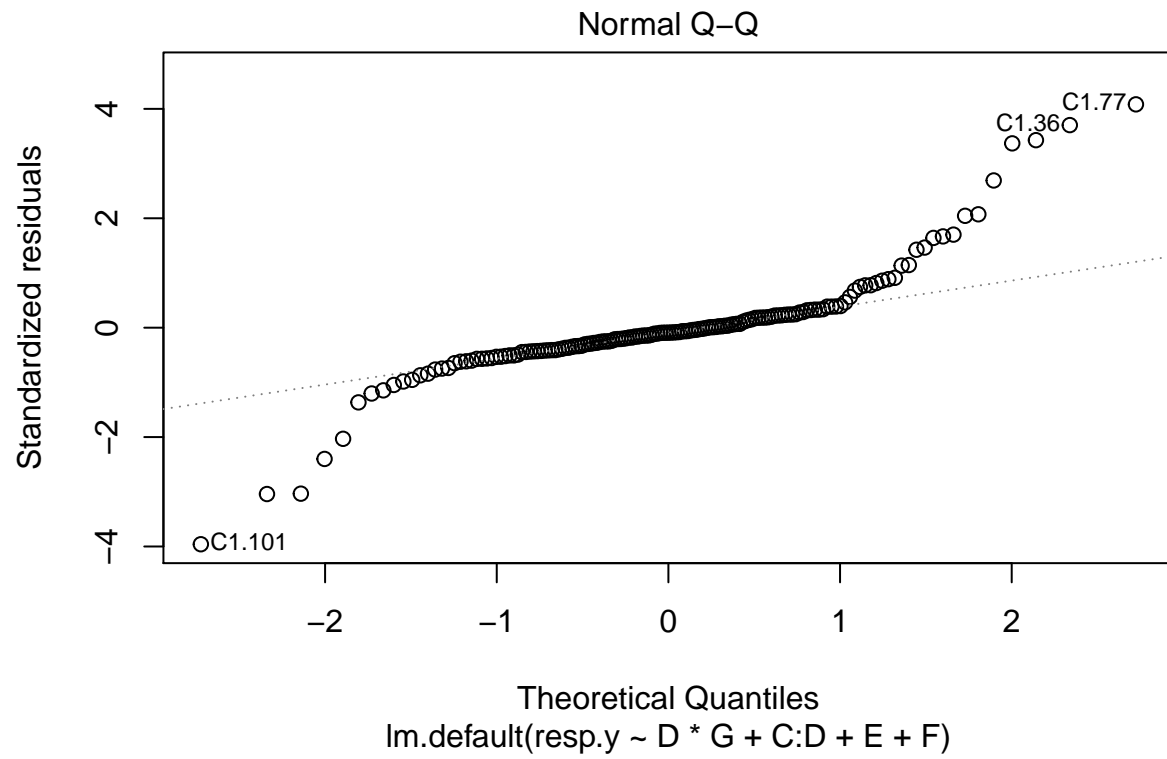
The metrics by which we would normally wish to evaluate our model suggest we are on perilous ground. However I will demonstrate later why chasing after our usual metrics can lead us into unsavory territory, especially given our design's priorities and limitations.

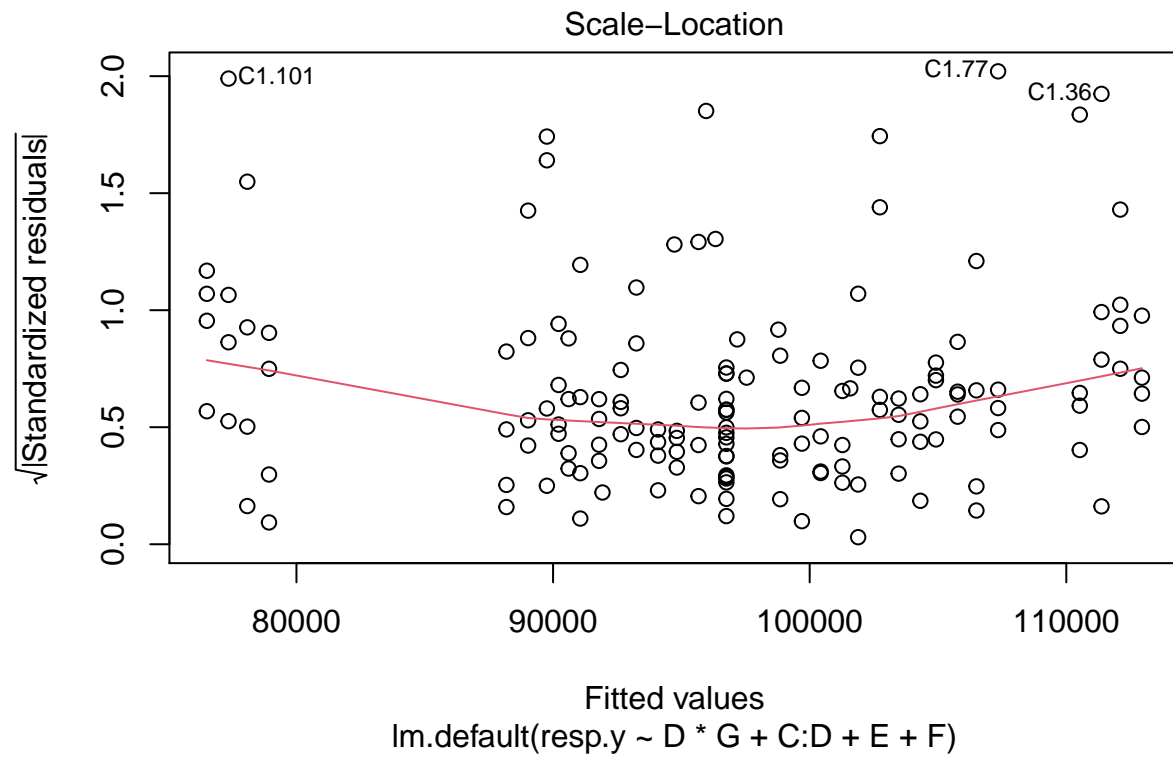
```

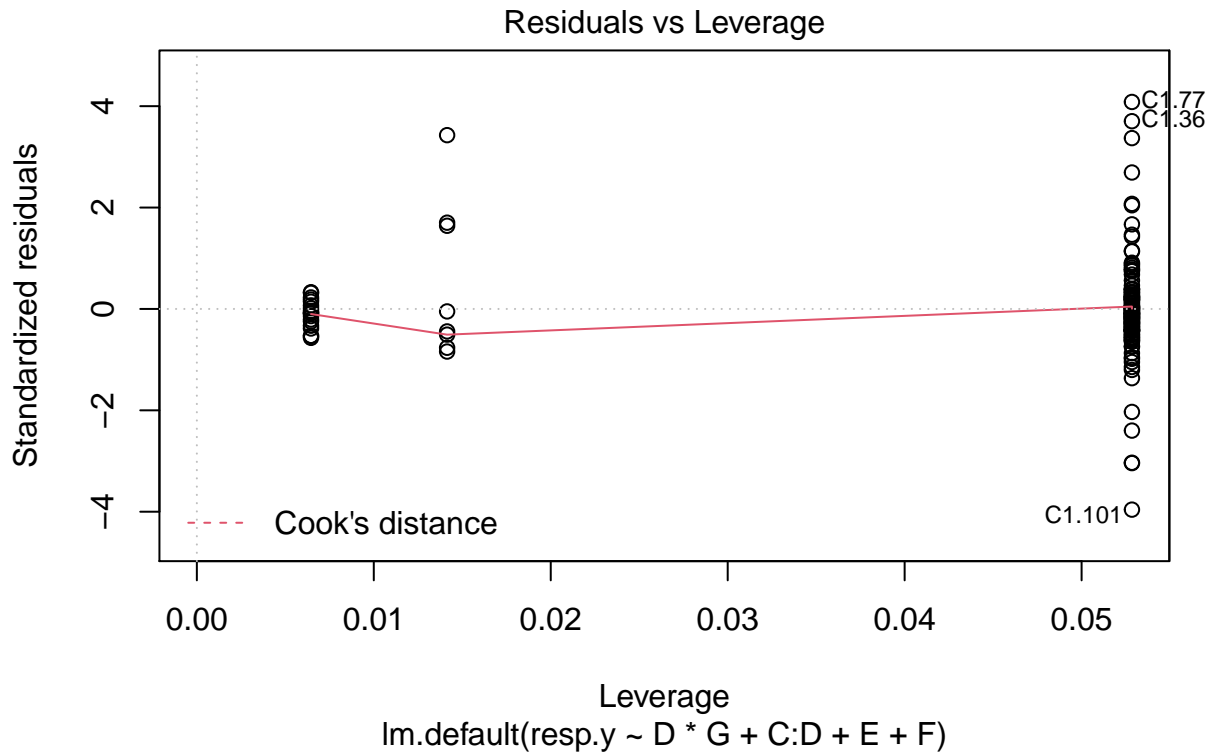
#oldpar <- par(oma=c(0,0,3,0), mfrow=c(2,2))
plot(LinearModel.4)

```







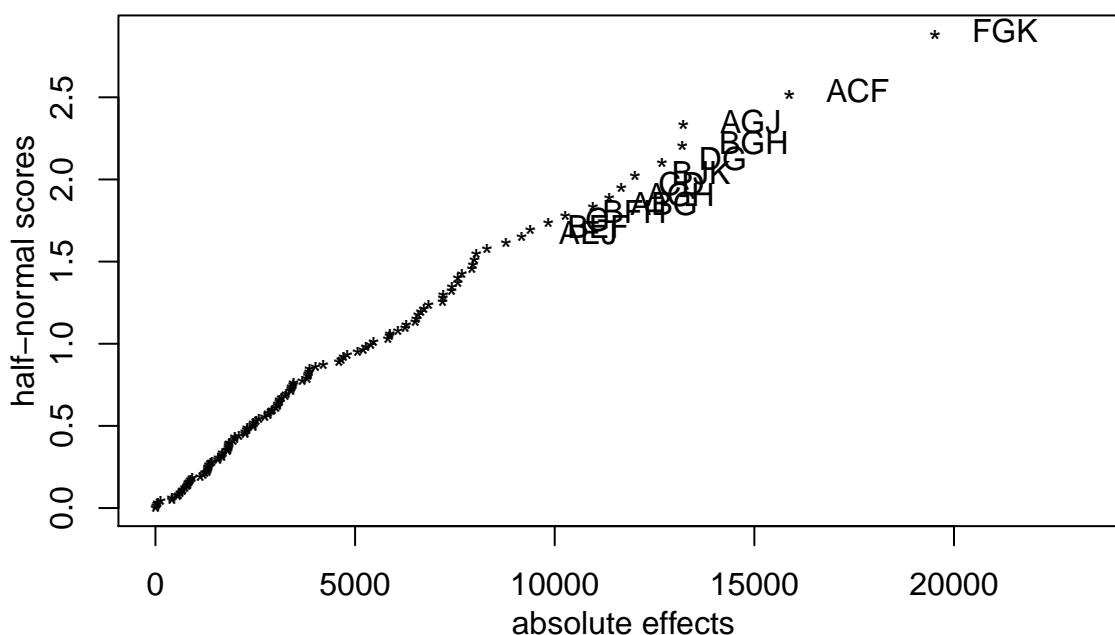


```
#par(oldpar)
```

Next, I will examine the half-normal plot for potential three-factor interactions. By the resolution of my design, some three-factor interactions are aliased with some two-factor interactions. *FGK*, *ACF*, *AGJ*, *BGH*, and *DG* appear as significant outliers. *DG* is already present within our model. *FGK* includes both *F* and *G* which are also already present within our model. *FGK* aliases down to *FG* if *K* isn't present within the model. Similarly logic can be applied to *ACF*. *AGJ* and *BGH* are likely aliases of interactions of terms already included in our model. So, let us consider adding *F : C* and *F : G* to our model.

```
DanielPlot(design_2, code=TRUE, autolab=TRUE, alpha=0.05, half=TRUE,
           response="resp.y.1.131.")
```

Half Normal Plot for resp.y.1.131., alpha=0.05



A = A , B = B , C = C , D = D , E = E , F = F , G = G , H = H , J = J , K = K

Doing so and evaluating an F-test on the nested models shows that, although not definitively statistically significant, we may consider including them. However, the term *CDG* offers no improvement to our model.

```
LinearModel.5 <- lm(resp.y ~ D*G + C:D + E + F:(1 + C + G), data=design)
summary(LinearModel.5)
```

```
##
## Call:
## lm.default(formula = resp.y ~ D * G + C:D + E + F:(1 + C + G),
##           data = design)
##
## Residuals:
##      Min       1Q   Median       3Q      Max
## -108474  -10830   -3136    7025   112094
##
## Coefficients:
##              Estimate Std. Error t value Pr(>|t|)
## (Intercept)   96753.6     2382.9  40.604  <2e-16 ***
## D              2025.5     2601.9   0.778   0.4375
## G             -4828.1     2601.9  -1.856   0.0655 .
## E              -422.8     2601.9  -0.163   0.8711
## D:G            6343.4     2622.2   2.419   0.0168 *
## D:C            5837.0     2622.2   2.226   0.0275 *
## C:F           -3606.6     2622.2  -1.375   0.1711
## G:F           -4391.5     2622.2  -1.675   0.0961 .
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
```

```
##
## Residual standard error: 29670 on 147 degrees of freedom
## Multiple R-squared: 0.1175, Adjusted R-squared: 0.07552
## F-statistic: 2.797 on 7 and 147 DF, p-value: 0.009219
```

```
anova(LinearModel.3, LinearModel.5)
```

```
## Analysis of Variance Table
##
## Model 1: resp.y ~ D * G + C:D
## Model 2: resp.y ~ D * G + C:D + E + F:(1 + C + G)
##   Res.Df      RSS Df Sum of Sq      F Pr(>F)
## 1      150 1.3353e+11
## 2      147 1.2937e+11  3 4156718477 1.5744 0.1981
```

```
#BJK alias CBDG -> CDG (assuming B isn't present)
LinearModel.6 <- lm(resp.y ~ D*G + C:D + E + F:(1 + C + G) + C:D:G,
                    data=design)
anova(LinearModel.5, LinearModel.6)
```

```
## Analysis of Variance Table
##
## Model 1: resp.y ~ D * G + C:D + E + F:(1 + C + G)
## Model 2: resp.y ~ D * G + C:D + E + F:(1 + C + G) + C:D:G
##   Res.Df      RSS Df Sum of Sq      F Pr(>F)
## 1      147 1.2937e+11
## 2      146 1.2927e+11  1 107968722 0.1219 0.7274
```

I would like to demonstrate the significance of the axial runs (despite their extra cost in the model). The below is an ANOVA table on the subset of runs that are exclusive to the two-factor fractional factorial portion of the utilized design. From this ANOVA table, we can find no evidence to suggest keeping *E* and *F* within our model despite significant evidence from the axial runs that they should both be factors within the model.

```
# Importance of axial runs (not center runs)
LinearModel.7 <- lm(resp.y.1.131. ~ (A + B + C + D + E + F + G + H + J +
                                   K)^2, data=design_2)
summary(aov(LinearModel.7))
```

```
##           Df      Sum Sq   Mean Sq F value Pr(>F)
## A           1 1.087e+09 1.087e+09   0.980 0.3253
## B           1 1.047e+08 1.047e+08   0.094 0.7595
## C           1 1.101e+09 1.101e+09   0.993 0.3223
## D           1 8.675e+08 8.675e+08   0.782 0.3792
## E           1 1.274e+08 1.274e+08   0.115 0.7356
## F           1 3.721e+08 3.721e+08   0.336 0.5641
## G           1 3.098e+09 3.098e+09   2.795 0.0988
## H           1 1.085e+08 1.085e+08   0.098 0.7552
## J           1 4.051e+07 4.051e+07   0.037 0.8489
## K           1 9.336e+08 9.336e+08   0.842 0.3618
## A:B         1 2.002e+07 2.002e+07   0.018 0.8935
## A:C         1 3.105e+08 3.105e+08   0.280 0.5983
## A:D         1 2.206e+09 2.206e+09   1.989 0.1625
## A:E         1 2.152e+08 2.152e+08   0.194 0.6608
## A:F         1 1.446e+05 1.446e+05   0.000 0.9909
## A:G         1 5.848e+05 5.848e+05   0.001 0.9817
## A:H         1 8.883e+07 8.883e+07   0.080 0.7779
```

```

## A:J      1 1.365e+09 1.365e+09 1.231 0.2707
## A:K      1 4.779e+08 4.779e+08 0.431 0.5135
## B:C      1 5.547e+06 5.547e+06 0.005 0.9438
## B:D      1 4.934e+07 4.934e+07 0.045 0.8335
## B:E      1 1.446e+09 1.446e+09 1.304 0.2570
## B:F      1 1.655e+09 1.655e+09 1.493 0.2256
## B:G      1 1.837e+09 1.837e+09 1.657 0.2020
## B:H      1 3.096e+08 3.096e+08 0.279 0.5988
## B:J      1 2.045e+09 2.045e+09 1.845 0.1785
## B:K      1 1.620e+08 1.620e+08 0.146 0.7033
## C:D      1 4.361e+09 4.361e+09 3.933 0.0510 .
## C:E      1 6.123e+07 6.123e+07 0.055 0.8149
## C:F      1 1.665e+09 1.665e+09 1.502 0.2243
## C:G      1 8.917e+08 8.917e+08 0.804 0.3727
## C:H      1 5.525e+07 5.525e+07 0.050 0.8240
## C:J      1 1.106e+09 1.106e+09 0.998 0.3211
## C:K      1 6.416e+07 6.416e+07 0.058 0.8106
## D:E      1 2.041e+08 2.041e+08 0.184 0.6692
## D:F      1 1.763e+09 1.763e+09 1.590 0.2113
## D:G      1 5.151e+09 5.151e+09 4.645 0.0343 *
## D:H      1 1.463e+07 1.463e+07 0.013 0.9089
## D:J      1 2.118e+07 2.118e+07 0.019 0.8904
## D:K      1 1.517e+07 1.517e+07 0.014 0.9072
## E:F      1 2.412e+08 2.412e+08 0.218 0.6422
## E:G      1 3.825e+08 3.825e+08 0.345 0.5587
## E:H      1 2.068e+09 2.068e+09 1.865 0.1761
## E:J      1 2.712e+08 2.712e+08 0.245 0.6224
## E:K      1 5.500e+06 5.500e+06 0.005 0.9440
## F:G      1 2.469e+09 2.469e+09 2.226 0.1399
## F:H      1 8.856e+07 8.856e+07 0.080 0.7782
## F:J      1 6.952e+08 6.952e+08 0.627 0.4310
## F:K      1 1.389e+09 1.389e+09 1.253 0.2666
## G:H      1 1.741e+07 1.741e+07 0.016 0.9006
## G:J      1 1.840e+09 1.840e+09 1.659 0.2017
## G:K      1 3.201e+08 3.201e+08 0.289 0.5926
## H:J      1 3.018e+08 3.018e+08 0.272 0.6034
## H:K      1 3.008e+08 3.008e+08 0.271 0.6040
## J:K      1 1.080e+08 1.080e+08 0.097 0.7559
## Residuals 75 8.316e+10 1.109e+09
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1

```

Lastly, I would like to demonstrate the dangers of chasing after metrics that are commonly used in **model selection** when running an experiment designed primarily for **factor screening**. Consider the below model `LinearModel.8`. It incorporates nine of the ten dependent variables and up to seven-factor interaction terms. However, it achieves a much higher correlation coefficient and an F-test between it and our working model `LinearModel.5` suggests the former to be superior. However, several issues present. Firstly, with such complex interaction terms, there is a significant risk that the true terms are hidden behind aliases. A seven-factor interaction term, for example, has aliases with lower-factor interaction terms on all levels. Ideally, we could work backwards from this model to sieve through the possible aliases and identify the most likely simplest aliases. However, I was unable to find software tools to help me accomplish this. This is one of the largest limitations of this project.

```

# Ludicrous model, chasing R^2
LinearModel.8 <- lm(resp.y ~ G + C:D + D:G + E:H + B:J + G:B:J + C:G:E:H +

```



```

C:D:G:E:H + C:D:G:E:H:B:J, data=design)
summary(LinearModel.8)

```

```

##
## Call:
## lm.default(formula = resp.y ~ G + C:D + D:G + E:H + B:J + G:B:J +
##           C:G:E:H + C:D:G:E:H + C:D:G:E:H:B:J, data = design)
##
## Residuals:
##      Min       1Q   Median       3Q      Max
## -80267 -11932  -1152    8489 101404
##
## Coefficients:
##              Estimate Std. Error t value Pr(>|t|)
## (Intercept)    96754      2112   45.812 < 2e-16 ***
## G             -4828       2306   -2.094 0.038034 *
## C:D            5837       2324    2.512 0.013116 *
## G:D            6343       2324    2.729 0.007130 **
## E:H            4019       2324    1.729 0.085860 .
## B:J            3997       2324    1.720 0.087570 .
## G:B:J          7938       2324    3.416 0.000826 ***
## G:C:E:H        9764       2324    4.201 4.62e-05 ***
## G:C:D:E:H     -5480       2324   -2.358 0.019710 *
## G:C:D:E:H:B:J  6610       2324    2.844 0.005099 **
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
##
## Residual standard error: 26290 on 145 degrees of freedom
## Multiple R-squared:  0.3162, Adjusted R-squared:  0.2738
## F-statistic:  7.45 on 9 and 145 DF,  p-value: 6.594e-09

```

```

anova(LinearModel.5, LinearModel.8)

```

```

## Analysis of Variance Table
##
## Model 1: resp.y ~ D * G + C:D + E + F:(1 + C + G)
## Model 2: resp.y ~ G + C:D + D:G + E:H + B:J + G:B:J + C:G:E:H + C:D:G:E:H +
##           C:D:G:E:H:B:J
##      Res.Df      RSS Df Sum of Sq      F      Pr(>F)
## 1         147 1.2937e+11
## 2         145 1.0025e+11  2 2.9126e+10 21.065 9.309e-09 ***
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1

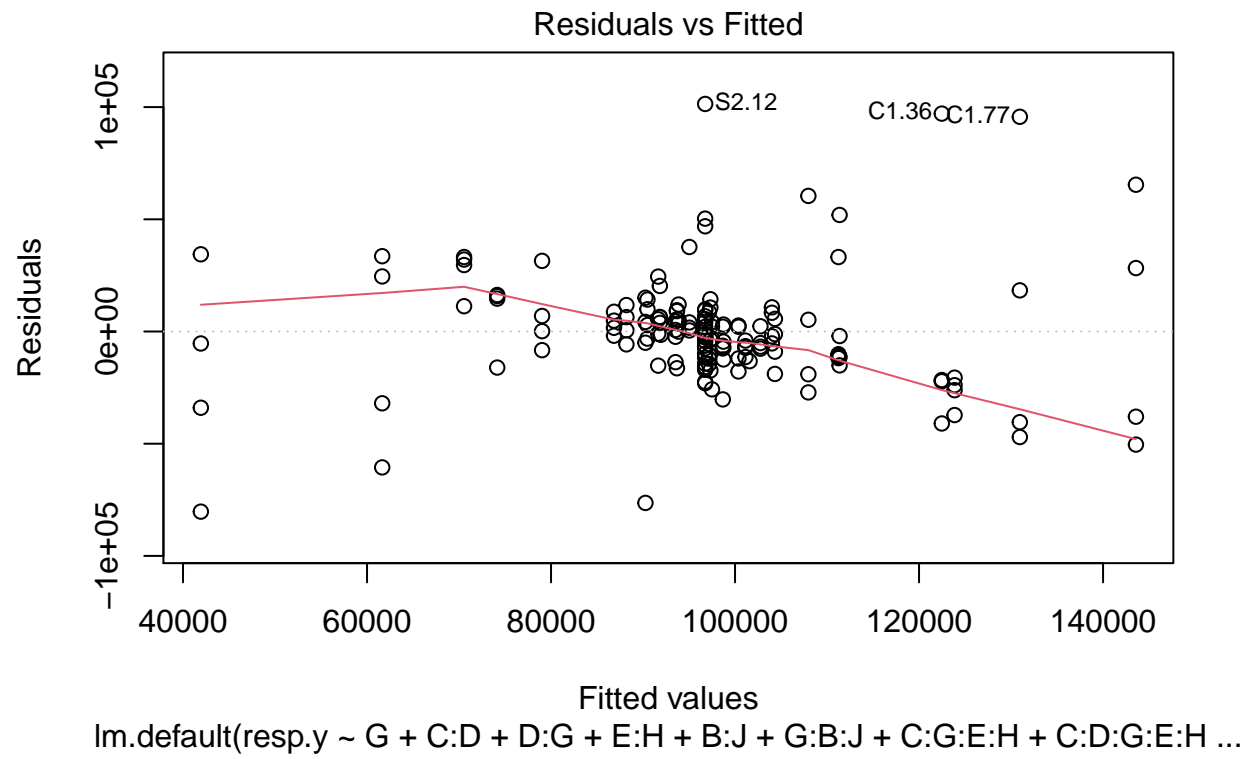
```

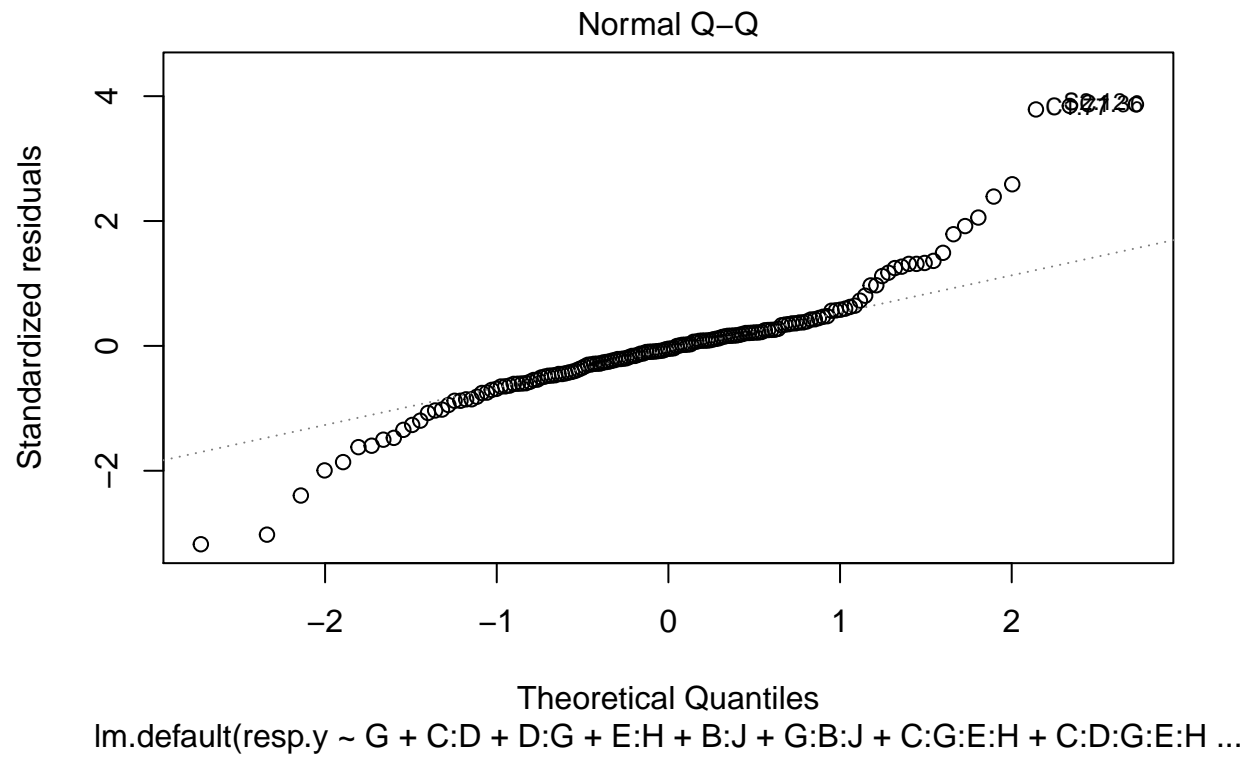
The risks posed by this sort of model selection also presents itself in diagnostic plots. Compare these plots to those of `LinearModel.5`. This model is possibly oversaturated and skewed.

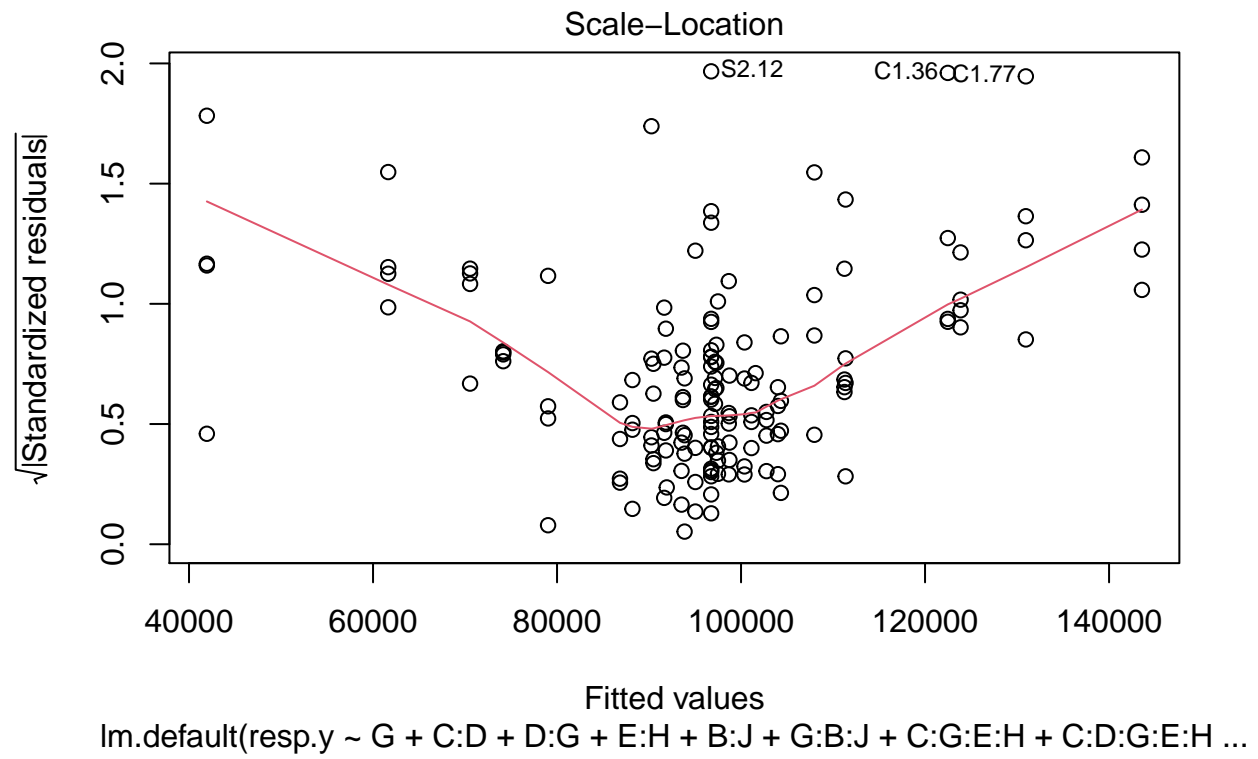
```

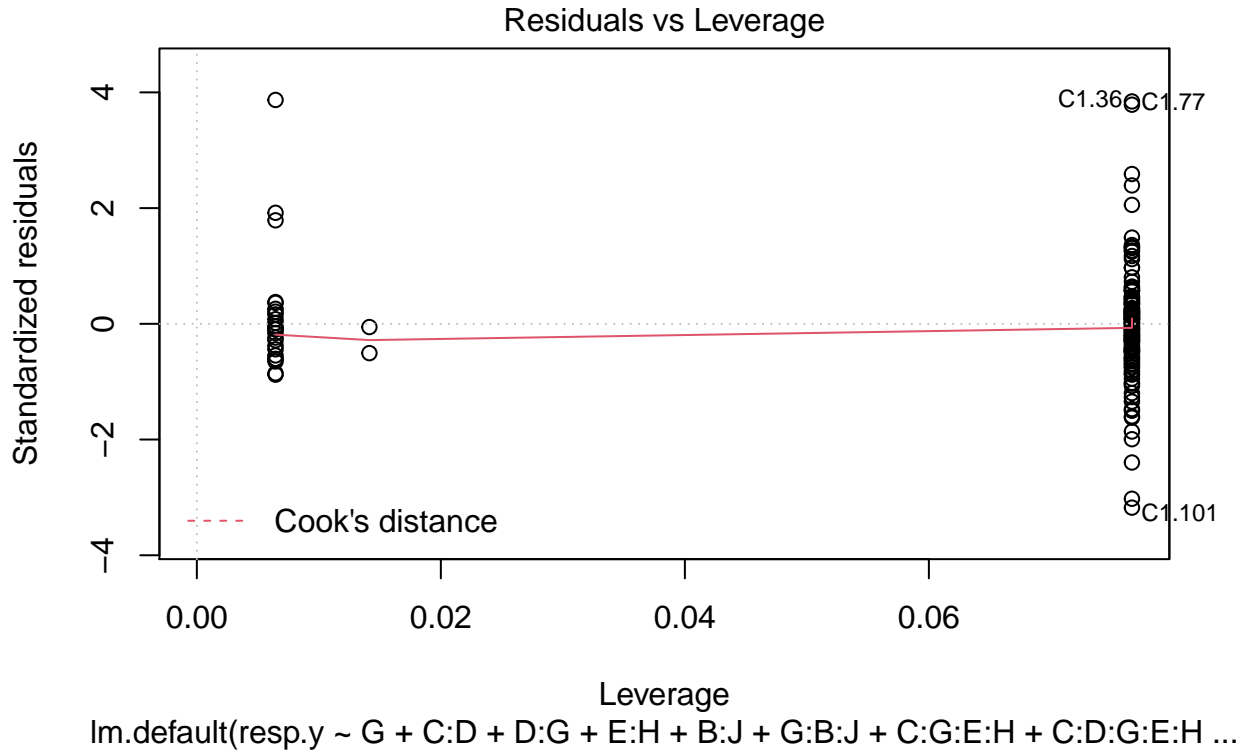
plot(LinearModel.8)

```









4. Results

The results of the above analysis demonstrates that the C, D, E, F, G factors are likely present within the model. The remaining factors most likely aren't. The below table summarizes these results.

Variable	Model
A	Out
B	Out
C	In
D	In
E	In
F	In
G	In
H	Out
J (I)	Out
K (J)	Out

ANOVA tables of other considered models were shown in the **Analysis** section. The final model I propose is $Y \sim D + G + E + F + D:G + C:D + C:F + F:G$.

LinearModel.5

```
##
## Call:
## lm.default(formula = resp.y ~ D * G + C:D + E + F:(1 + C + G),
```

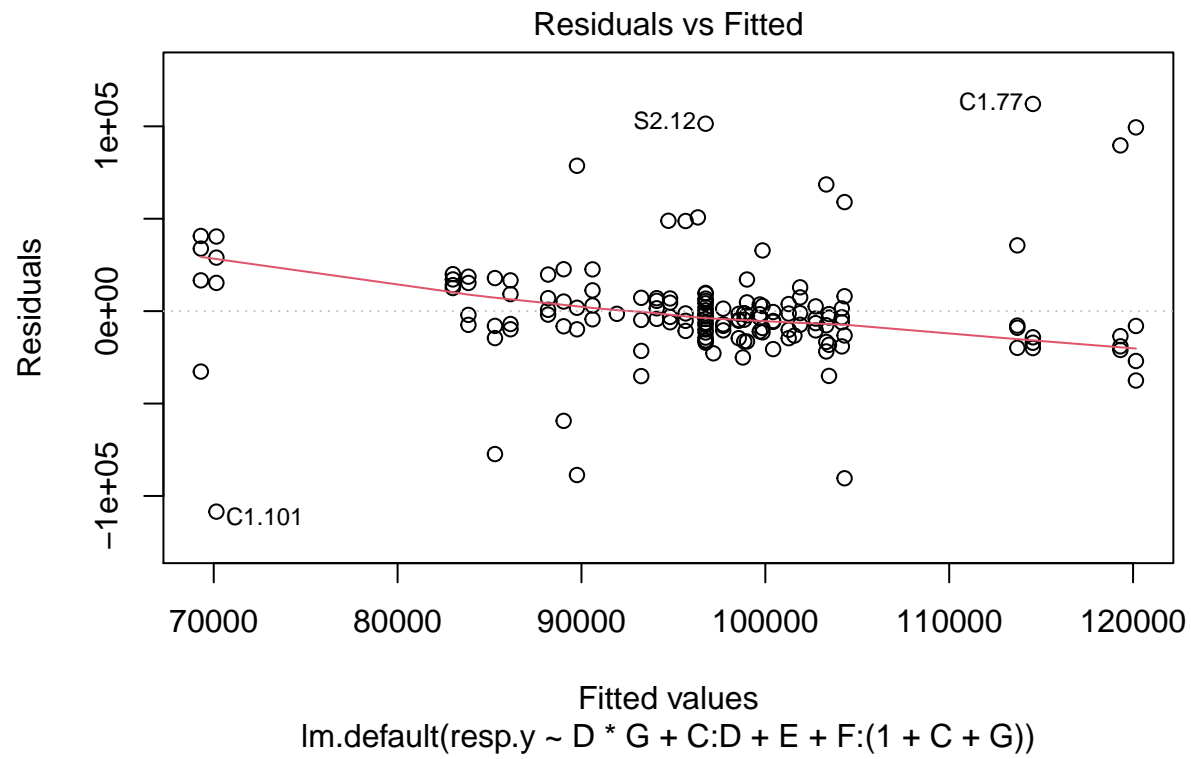
```
##      data = design)
##
## Coefficients:
## (Intercept)          D          G          E          D:G          D:C
##      96753.6      2025.5     -4828.1     -422.8      6343.4      5837.0
##          C:F          G:F
##      -3606.6      -4391.5
```

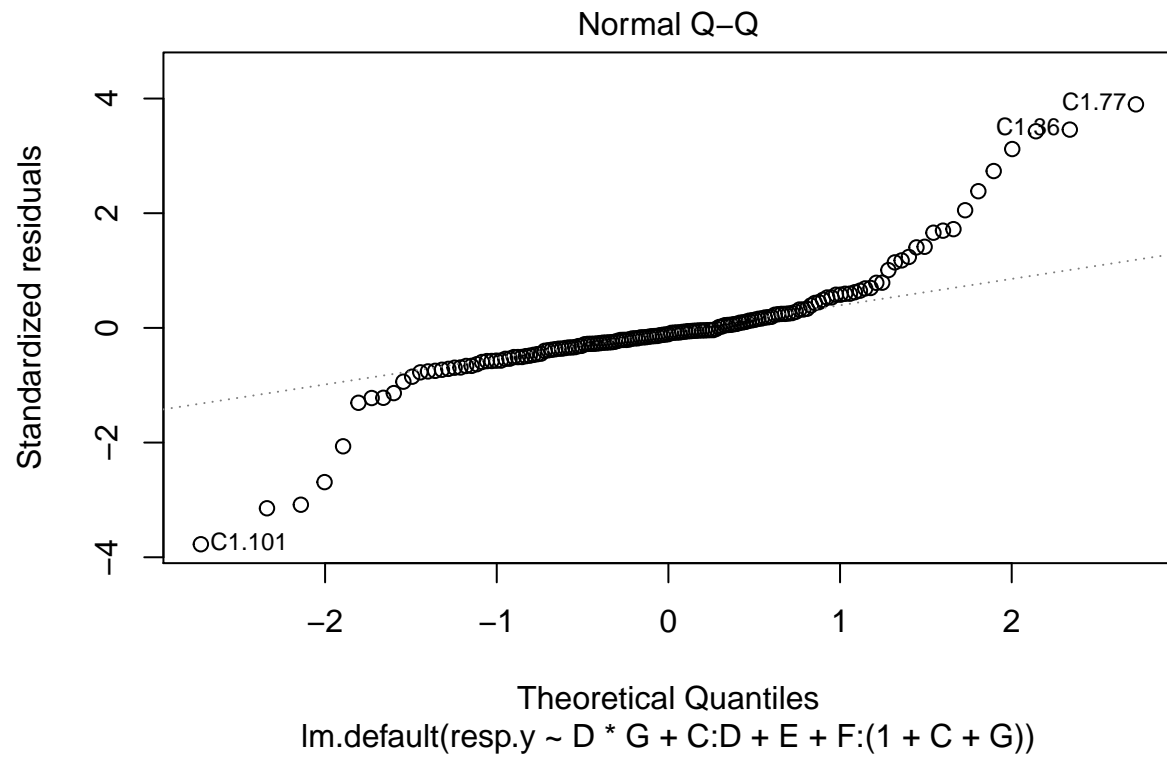
```
summary(anova(LinearModel.5))
```

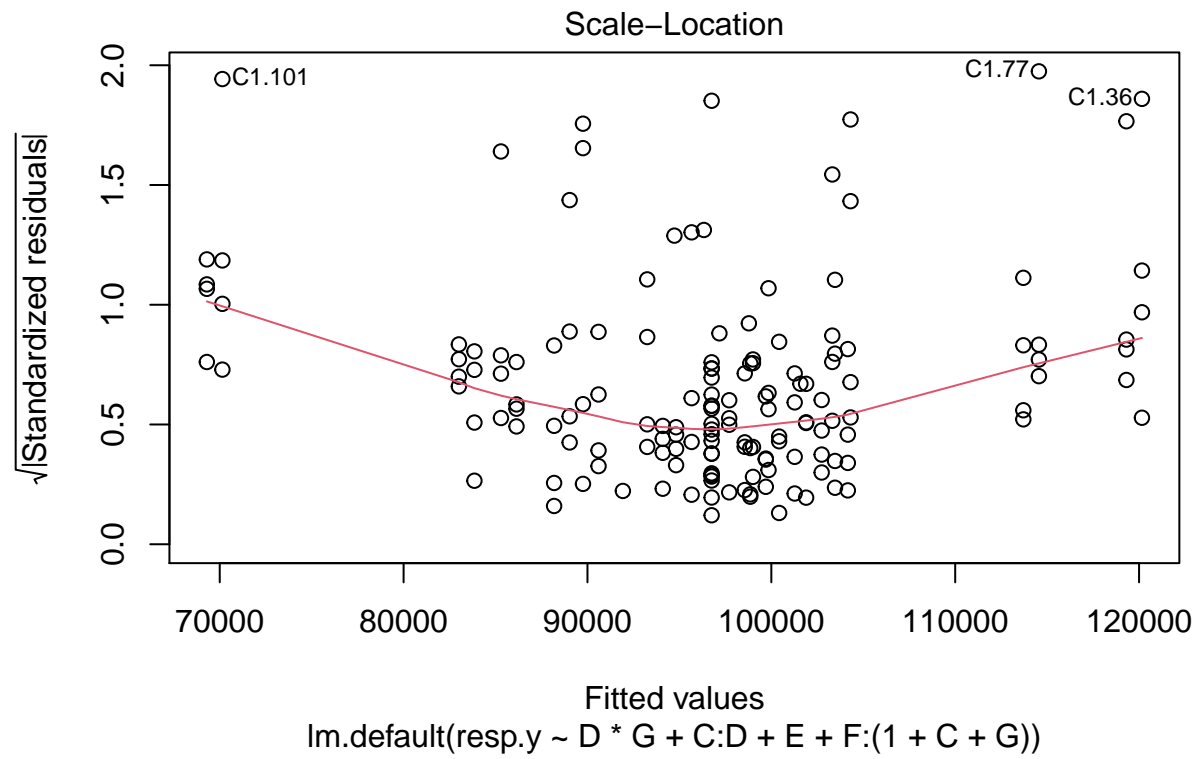
```
##          Df          Sum Sq          Mean Sq          F value
## Min.      : 1.00   Min.      :2.324e+07   Min.      :2.324e+07   Min.      :0.02641
## 1st Qu.: 1.00   1st Qu.:1.382e+09   1st Qu.:7.934e+08   1st Qu.:1.24890
## Median : 1.00   Median :2.749e+09   Median :2.067e+09   Median :2.80486
## Mean    : 19.25   Mean    :1.833e+10   Mean     :2.264e+09   Mean     :2.79713
## 3rd Qu.: 1.00   3rd Qu.:4.558e+09   3rd Qu.:3.363e+09   3rd Qu.:4.19925
## Max.    :147.00   Max.    :1.294e+11   Max.     :5.151e+09   Max.     :5.85234
##
## NA's      :1
```

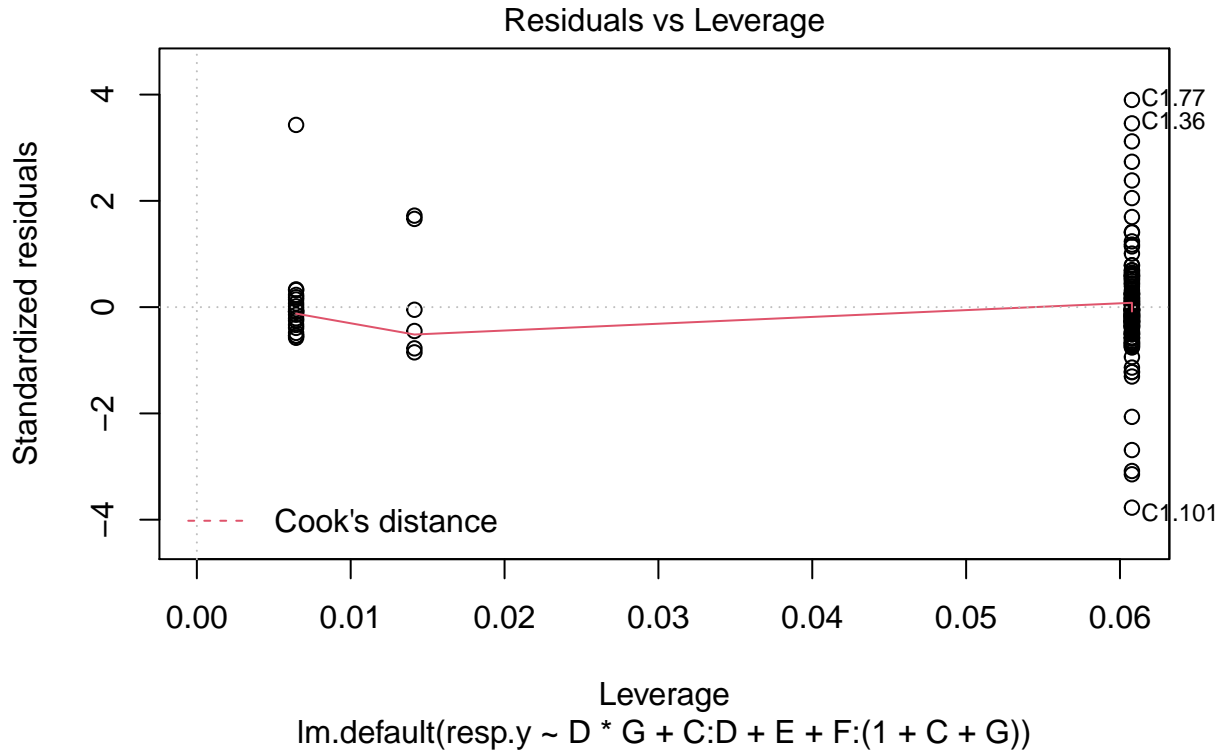
```
##          Pr(>F)
## Min.      :0.01678
## 1st Qu.:0.04652
## Median :0.09611
## Mean     :0.24081
## 3rd Qu.:0.30432
## Max.     :0.87113
## NA's      :1
```

```
plot(LinearModel.5)
```









5. Conclusions and Discussion

This experiment had many limitations. Firstly, full modeling (as opposed to identifying presence) of factor effects is difficult. The design was made to be non-rotatable to minimize the cost of the experiment which reduced modeling accuracy. Identifying factor effect nature (linear, polynomial, exponential, etc.) was another limitation. Ideally, we'd like to identify transformations when developing a statistical model. **Dispersion effects** were also not analyzed. I believe they are outside scope of the problem, but dispersion effects are likely to occur and real life and must also be modeled. Improvements that can be made: **fold-over** design, **sequential experimentation**, and With lower costs a 3^k design may be considered. A **rotatable** central composite design may be useful, as well as finding / creating a SCD for 10 factors.

6. Appendix

6.1 Design Creation

```
design_centers <- design[c("C1.129", "C1.130", "C1.131"), 2:12]
design_centers <- rbind(design_centers, design[132:155, 2:12])

cs <- c("C1.129", "C1.130", "C1.131")

all_0 <- design_centers$A == 0 & design_centers$B == 0 &
  design_centers$C == 0 & design_centers$D == 0 &
  design_centers$E == 0 & design_centers$F == 0 &
  design_centers$G == 0 & design_centers$H == 0 &
  design_centers$J == 0 & design_centers$K == 0

As <- xor(design_centers$A != 0, all_0)
Bs <- xor(design_centers$B != 0, all_0)
Cs <- xor(design_centers$C != 0, all_0)
Ds <- xor(design_centers$D != 0, all_0)
Es <- xor(design_centers$E != 0, all_0)
Fs <- xor(design_centers$F != 0, all_0)
Gs <- xor(design_centers$G != 0, all_0)
Hs <- xor(design_centers$H != 0, all_0)
Js <- xor(design_centers$J != 0, all_0)
Ks <- xor(design_centers$K != 0, all_0)
```

6.2 Design

```
print(design)
```

##	Block.ccd	A	B	C	D	E	F	G	H	J	K	resp.y
##	C1.129	1	0	0	0	0	0	0	0	0	0	94162.434
##	C1.102	1	1	-1	1	-1	-1	1	1	-1	-1	85430.443
##	C1.30	1	1	-1	1	1	1	-1	-1	-1	-1	114789.788
##	C1.82	1	1	-1	-1	-1	1	-1	1	-1	-1	168404.075
##	C1.95	1	-1	1	1	1	1	-1	1	-1	1	93864.418
##	C1.31	1	-1	1	1	1	1	-1	-1	-1	-1	100804.436
##	C1.39	1	-1	1	1	-1	-1	1	-1	-1	1	91181.800
##	C1.16	1	1	1	1	1	-1	-1	-1	-1	-1	105319.684
##	C1.53	1	-1	-1	1	-1	1	1	-1	-1	-1	95109.129
##	C1.18	1	1	-1	-1	-1	1	-1	-1	-1	1	95674.325
##	C1.88	1	1	1	1	-1	1	-1	1	-1	-1	7964.088
##	C1.15	1	-1	1	1	1	-1	-1	-1	1	-1	92309.467
##	C1.75	1	-1	1	-1	1	-1	-1	1	-1	-1	84964.605
##	C1.50	1	1	-1	-1	-1	1	1	-1	-1	-1	105787.870
##	C1.59	1	-1	1	-1	1	1	1	-1	1	-1	96727.560
##	C1.103	1	-1	1	1	-1	-1	1	1	-1	-1	99113.416
##	C1.64	1	1	1	1	1	1	1	-1	1	-1	99986.205
##	C1.60	1	1	1	-1	1	1	1	-1	-1	1	103747.233
##	C1.111	1	-1	1	1	1	-1	1	1	1	-1	93352.191
##	C1.12	1	1	1	-1	1	-1	-1	-1	1	-1	76430.851
##	C1.17	1	-1	-1	-1	-1	1	-1	-1	1	-1	81494.616
##	C1.77	1	-1	-1	1	1	-1	-1	1	-1	-1	226646.146
##	C1.47	1	-1	1	1	1	-1	1	1	1	-1	112367.609

## C1.104	1	1	1	1	-1	-1	1	1	1	1	-1	110516.131
## C1.76	1	1	1	-1	1	-1	-1	1	1	1	-1	144420.676
## C1.113	1	-1	-1	-1	-1	1	1	1	1	-1	-1	86303.755
## C1.79	1	-1	1	1	1	-1	-1	1	1	1	1	94593.098
## C1.14	1	1	-1	1	1	-1	-1	-1	1	-1	1	98709.237
## C1.19	1	-1	1	-1	-1	1	-1	-1	-1	1	-1	86662.415
## C1.128	1	1	1	1	1	1	1	1	1	1	1	90553.555
## C1.8	1	1	1	1	-1	-1	-1	-1	1	1	1	96021.619
## C1.108	1	1	1	-1	1	-1	1	1	1	-1	1	89902.874
## C1.125	1	-1	-1	1	1	1	1	1	1	1	1	99059.651
## C1.57	1	-1	-1	-1	1	1	1	-1	-1	1	1	116120.847
## C1.112	1	1	1	1	1	-1	1	1	-1	1	1	83919.177
## C1.34	1	1	-1	-1	-1	-1	1	-1	1	-1	1	82615.131
## C1.44	1	1	1	-1	1	-1	1	-1	1	1	1	102608.606
## C1.56	1	1	1	1	-1	1	1	-1	-1	-1	-1	99934.242
## C1.100	1	1	1	-1	-1	-1	1	1	-1	-1	-1	29662.972
## C1.43	1	-1	1	-1	1	-1	1	-1	-1	-1	-1	132699.236
## C1.7	1	-1	1	1	-1	-1	-1	-1	-1	-1	-1	88741.567
## C1.27	1	-1	1	-1	1	1	-1	-1	1	1	1	103021.241
## C1.58	1	1	-1	-1	1	1	1	-1	1	-1	-1	82569.428
## C1.55	1	-1	1	1	-1	1	1	-1	1	1	1	79884.555
## C1.40	1	1	1	1	-1	-1	1	-1	1	-1	-1	86638.669
## C1.6	1	1	-1	1	-1	-1	-1	-1	-1	-1	-1	98046.813
## C1.46	1	1	-1	1	1	-1	1	-1	1	1	-1	91143.263
## C1.68	1	1	1	-1	-1	-1	-1	1	-1	1	1	101842.398
## C1.126	1	1	-1	1	1	1	1	1	-1	-1	-1	89751.308
## C1.51	1	-1	1	-1	-1	1	1	-1	-1	-1	1	98313.993
## C1.26	1	1	-1	-1	1	1	-1	-1	1	1	1	95502.066
## C1.116	1	1	1	-1	-1	1	1	1	1	-1	-1	95214.509
## C1.45	1	-1	-1	1	1	-1	1	-1	-1	-1	1	13873.347
## C1.120	1	1	1	1	-1	1	1	1	-1	1	-1	85945.034
## C1.35	1	-1	1	-1	-1	-1	1	-1	1	-1	1	112136.092
## C1.87	1	-1	1	1	-1	1	-1	1	1	1	-1	103181.283
## C1.3	1	-1	1	-1	-1	-1	-1	-1	1	1	-1	85075.300
## C1.54	1	1	-1	1	-1	1	1	-1	1	1	1	94614.184
## C1.127	1	-1	1	1	1	1	1	1	-1	-1	-1	87317.877
## C1.99	1	-1	1	-1	-1	-1	1	1	1	1	1	80831.781
## C1.23	1	-1	1	1	-1	1	-1	-1	1	-1	-1	82462.249
## C1.114	1	1	-1	-1	-1	1	1	1	-1	1	1	107984.822
## C1.110	1	1	-1	1	1	-1	1	1	1	-1	-1	97085.162
## C1.1	1	-1	-1	-1	-1	-1	-1	-1	-1	-1	1	100846.001
## C1.80	1	1	1	1	1	-1	-1	1	-1	-1	-1	97449.680
## C1.130	1	0	0	0	0	0	0	0	0	0	0	103535.347
## C1.124	1	1	1	-1	1	1	1	1	-1	-1	1	58087.345
## C1.109	1	-1	-1	1	1	-1	1	1	-1	1	1	93794.126
## C1.86	1	1	-1	1	-1	1	-1	1	1	1	-1	77307.856
## C1.24	1	1	1	1	-1	1	-1	-1	-1	1	1	97606.967
## C1.11	1	-1	1	-1	1	-1	-1	-1	-1	1	1	102499.323
## C1.107	1	-1	1	-1	1	-1	1	1	-1	1	-1	99658.902
## C1.115	1	-1	1	-1	-1	1	1	1	-1	1	1	88919.933
## C1.33	1	-1	-1	-1	-1	-1	1	-1	-1	1	-1	93185.520
## C1.97	1	-1	-1	-1	-1	-1	1	1	-1	-1	-1	94222.527
## C1.96	1	1	1	1	1	1	-1	1	1	-1	-1	105889.966
## C1.42	1	1	-1	-1	1	-1	1	-1	-1	-1	-1	88367.429

## C1.48	1	1	1	1	1	-1	1	-1	-1	-1	1	163305.110
## C1.13	1	-1	-1	1	1	-1	-1	-1	-1	1	-1	96258.288
## C1.118	1	1	-1	1	-1	1	1	1	1	-1	1	109966.571
## C1.41	1	-1	-1	-1	1	-1	1	-1	1	1	1	90709.220
## C1.78	1	1	-1	1	1	-1	-1	1	1	1	1	100392.014
## C1.9	1	-1	-1	-1	1	-1	-1	-1	1	-1	-1	81829.286
## C1.72	1	1	1	1	-1	-1	-1	1	1	-1	1	76338.298
## C1.70	1	1	-1	1	-1	-1	-1	1	-1	1	-1	102775.417
## C1.106	1	1	-1	-1	1	-1	1	1	-1	1	-1	95640.559
## C1.21	1	-1	-1	1	-1	1	-1	-1	-1	1	1	97722.043
## C1.65	1	-1	-1	-1	-1	-1	-1	1	-1	1	1	86187.382
## C1.94	1	1	-1	1	1	1	-1	1	-1	1	1	104711.034
## C1.69	1	-1	-1	1	-1	-1	-1	1	1	-1	1	79184.857
## C1.61	1	-1	-1	1	1	1	1	-1	1	-1	1	85257.873
## C1.71	1	-1	1	1	-1	-1	-1	1	-1	1	-1	95347.517
## C1.66	1	1	-1	-1	-1	-1	-1	1	1	-1	-1	113205.611
## C1.117	1	-1	-1	1	-1	1	1	1	-1	1	-1	36601.612
## C1.122	1	1	-1	-1	1	1	1	1	1	1	-1	88500.360
## C1.36	1	1	1	-1	-1	-1	1	-1	-1	1	-1	219584.156
## C1.101	1	-1	-1	1	-1	-1	1	1	1	1	-1	-38330.113
## C1.28	1	1	1	-1	1	1	-1	-1	-1	-1	-1	97079.587
## C1.22	1	1	-1	1	-1	1	-1	-1	1	-1	-1	94218.992
## C1.20	1	1	1	-1	-1	1	-1	-1	1	-1	1	171852.401
## C1.90	1	1	-1	-1	1	1	-1	1	1	-1	1	101682.224
## C1.93	1	-1	-1	1	1	1	-1	1	1	-1	-1	149288.073
## C1.81	1	-1	-1	-1	-1	1	-1	1	1	1	1	1099.033
## C1.91	1	-1	1	-1	1	1	-1	1	1	-1	1	91683.539
## C1.29	1	-1	-1	1	1	1	-1	-1	1	1	-1	109344.586
## C1.123	1	-1	1	-1	1	1	1	1	1	1	-1	71726.271
## C1.119	1	-1	1	1	-1	1	1	1	1	-1	1	103145.467
## C1.63	1	-1	1	1	1	1	1	-1	-1	1	-1	101862.642
## C1.121	1	-1	-1	-1	1	1	1	1	-1	-1	1	100456.861
## C1.2	1	1	-1	-1	-1	-1	-1	-1	1	1	-1	98139.545
## C1.25	1	-1	-1	-1	1	1	-1	-1	-1	-1	-1	100165.042
## C1.84	1	1	1	-1	-1	1	-1	1	1	1	1	91582.757
## C1.62	1	1	-1	1	1	1	1	-1	-1	1	-1	68414.445
## C1.10	1	1	-1	-1	1	-1	-1	-1	-1	1	1	99097.288
## C1.37	1	-1	-1	1	-1	-1	1	-1	1	-1	-1	105099.346
## C1.52	1	1	1	-1	-1	1	1	-1	1	1	-1	100279.993
## C1.89	1	-1	-1	-1	1	1	-1	1	-1	1	-1	88796.367
## C1.98	1	1	-1	-1	-1	-1	1	1	1	1	1	111703.672
## C1.74	1	1	-1	-1	1	-1	-1	1	-1	-1	1	94434.997
## C1.85	1	-1	-1	1	-1	1	-1	1	-1	-1	1	70707.961
## C1.92	1	1	1	-1	1	1	-1	1	-1	1	-1	99382.176
## C1.105	1	-1	-1	-1	1	-1	1	1	1	-1	1	101115.650
## C1.67	1	-1	1	-1	-1	-1	-1	1	1	-1	-1	93660.576
## C1.32	1	1	1	1	1	1	-1	-1	1	1	-1	94525.042
## C1.38	1	1	-1	1	-1	-1	1	-1	-1	1	1	99980.520
## C1.83	1	-1	1	-1	-1	1	-1	1	-1	-1	-1	79924.220
## C1.5	1	-1	-1	1	-1	-1	-1	-1	1	1	1	103294.958
## C1.49	1	-1	-1	-1	-1	1	1	-1	1	1	-1	208988.648
## C1.4	1	1	1	-1	-1	-1	-1	-1	-1	-1	1	105614.053
## C1.73	1	-1	-1	-1	1	-1	-1	1	1	1	-1	90420.589
## C1.131	1	0	0	0	0	0	0	0	0	0	0	80801.145

```
## S2.14      2  0  0  0  0  0  0  1  0  0  0  90467.237
## S2.2       2  1  0  0  0  0  0  0  0  0  0  80866.497
## S2.23      2  0  0  0  0  0  0  0  0  0  0 101005.709
## S2.16      2  0  0  0  0  0  0  0  1  0  0  98846.582
## S2.12      2  0  0  0  0  0  1  0  0  0  0 198157.287
## S2.10      2  0  0  0  0  1  0  0  0  0  0 147050.845
## S2.7       2  0  0  0 -1  0  0  0  0  0  0 143636.681
## S2.19      2  0  0  0  0  0  0  0  0  0 -1 102259.860
## S2.24      2  0  0  0  0  0  0  0  0  0  0  79687.058
## S2.17      2  0  0  0  0  0  0  0  0 -1  0  95628.330
## S2.21      2  0  0  0  0  0  0  0  0  0  0 106239.812
## S2.3       2  0 -1  0  0  0  0  0  0  0  0 106621.398
## S2.4       2  0  1  0  0  0  0  0  0  0  0  94279.241
## S2.13      2  0  0  0  0  0  0 -1  0  0  0  88357.682
## S2.9       2  0  0  0  0 -1  0  0  0  0  0  74337.010
## S2.22      2  0  0  0  0  0  0  0  0  0  0  90513.420
## S2.20      2  0  0  0  0  0  0  0  0  0  1  89284.997
## S2.18      2  0  0  0  0  0  0  0  0  1  0  85209.962
## S2.15      2  0  0  0  0  0  0  0 -1  0  0  94391.556
## S2.8       2  0  0  0  1  0  0  0  0  0  0  73716.290
## S2.1       2 -1  0  0  0  0  0  0  0  0  0  92533.426
## S2.5       2  0  0 -1  0  0  0  0  0  0  0  97186.295
## S2.6       2  0  0  1  0  0  0  0  0  0  0  86852.827
## S2.11      2  0  0  0  0  0 -1  0  0  0  0  82431.671
```

```
## class=design, type= ccd
```

```
attributes(design)
```

```
## $row.names
```

```
## [1] "C1.129" "C1.102" "C1.30" "C1.82" "C1.95" "C1.31" "C1.39" "C1.16"
## [9] "C1.53" "C1.18" "C1.88" "C1.15" "C1.75" "C1.50" "C1.59" "C1.103"
## [17] "C1.64" "C1.60" "C1.111" "C1.12" "C1.17" "C1.77" "C1.47" "C1.104"
## [25] "C1.76" "C1.113" "C1.79" "C1.14" "C1.19" "C1.128" "C1.8" "C1.108"
## [33] "C1.125" "C1.57" "C1.112" "C1.34" "C1.44" "C1.56" "C1.100" "C1.43"
## [41] "C1.7" "C1.27" "C1.58" "C1.55" "C1.40" "C1.6" "C1.46" "C1.68"
## [49] "C1.126" "C1.51" "C1.26" "C1.116" "C1.45" "C1.120" "C1.35" "C1.87"
## [57] "C1.3" "C1.54" "C1.127" "C1.99" "C1.23" "C1.114" "C1.110" "C1.1"
## [65] "C1.80" "C1.130" "C1.124" "C1.109" "C1.86" "C1.24" "C1.11" "C1.107"
## [73] "C1.115" "C1.33" "C1.97" "C1.96" "C1.42" "C1.48" "C1.13" "C1.118"
## [81] "C1.41" "C1.78" "C1.9" "C1.72" "C1.70" "C1.106" "C1.21" "C1.65"
## [89] "C1.94" "C1.69" "C1.61" "C1.71" "C1.66" "C1.117" "C1.122" "C1.36"
## [97] "C1.101" "C1.28" "C1.22" "C1.20" "C1.90" "C1.93" "C1.81" "C1.91"
## [105] "C1.29" "C1.123" "C1.119" "C1.63" "C1.121" "C1.2" "C1.25" "C1.84"
## [113] "C1.62" "C1.10" "C1.37" "C1.52" "C1.89" "C1.98" "C1.74" "C1.85"
## [121] "C1.92" "C1.105" "C1.67" "C1.32" "C1.38" "C1.83" "C1.5" "C1.49"
## [129] "C1.4" "C1.73" "C1.131" "S2.14" "S2.2" "S2.23" "S2.16" "S2.12"
## [137] "S2.10" "S2.7" "S2.19" "S2.24" "S2.17" "S2.21" "S2.3" "S2.4"
## [145] "S2.13" "S2.9" "S2.22" "S2.20" "S2.18" "S2.15" "S2.8" "S2.1"
## [153] "S2.5" "S2.6" "S2.11"
```

```
##
```

```
## $desnum
```

```
##      Block.ccd      A      B      C      D      E      F
## C1.129      0  0.000000  0.000000  0.000000  0.000000  0.000000  0.000000
## C1.102      0  1.000000 -1.000000  1.000000 -1.000000 -1.000000  1.000000
## C1.30       0  1.000000 -1.000000  1.000000  1.000000  1.000000 -1.000000
```

## C1.82	0	1.000000	-1.000000	-1.000000	-1.000000	1.000000	-1.000000
## C1.95	0	-1.000000	1.000000	1.000000	1.000000	1.000000	-1.000000
## C1.31	0	-1.000000	1.000000	1.000000	1.000000	1.000000	-1.000000
## C1.39	0	-1.000000	1.000000	1.000000	-1.000000	-1.000000	1.000000
## C1.16	0	1.000000	1.000000	1.000000	1.000000	-1.000000	-1.000000
## C1.53	0	-1.000000	-1.000000	1.000000	-1.000000	1.000000	1.000000
## C1.18	0	1.000000	-1.000000	-1.000000	-1.000000	1.000000	-1.000000
## C1.88	0	1.000000	1.000000	1.000000	-1.000000	1.000000	-1.000000
## C1.15	0	-1.000000	1.000000	1.000000	1.000000	-1.000000	-1.000000
## C1.75	0	-1.000000	1.000000	-1.000000	1.000000	-1.000000	-1.000000
## C1.50	0	1.000000	-1.000000	-1.000000	-1.000000	1.000000	1.000000
## C1.59	0	-1.000000	1.000000	-1.000000	1.000000	1.000000	1.000000
## C1.103	0	-1.000000	1.000000	1.000000	-1.000000	-1.000000	1.000000
## C1.64	0	1.000000	1.000000	1.000000	1.000000	1.000000	1.000000
## C1.60	0	1.000000	1.000000	-1.000000	1.000000	1.000000	1.000000
## C1.111	0	-1.000000	1.000000	1.000000	1.000000	-1.000000	1.000000
## C1.12	0	1.000000	1.000000	-1.000000	1.000000	-1.000000	-1.000000
## C1.17	0	-1.000000	-1.000000	-1.000000	-1.000000	1.000000	-1.000000
## C1.77	0	-1.000000	-1.000000	1.000000	1.000000	-1.000000	-1.000000
## C1.47	0	-1.000000	1.000000	1.000000	1.000000	-1.000000	1.000000
## C1.104	0	1.000000	1.000000	1.000000	-1.000000	-1.000000	1.000000
## C1.76	0	1.000000	1.000000	-1.000000	1.000000	-1.000000	-1.000000
## C1.113	0	-1.000000	-1.000000	-1.000000	-1.000000	1.000000	1.000000
## C1.79	0	-1.000000	1.000000	1.000000	1.000000	-1.000000	-1.000000
## C1.14	0	1.000000	-1.000000	1.000000	1.000000	-1.000000	-1.000000
## C1.19	0	-1.000000	1.000000	-1.000000	-1.000000	1.000000	-1.000000
## C1.128	0	1.000000	1.000000	1.000000	1.000000	1.000000	1.000000
## C1.8	0	1.000000	1.000000	1.000000	-1.000000	-1.000000	-1.000000
## C1.108	0	1.000000	1.000000	-1.000000	1.000000	-1.000000	1.000000
## C1.125	0	-1.000000	-1.000000	1.000000	1.000000	1.000000	1.000000
## C1.57	0	-1.000000	-1.000000	-1.000000	1.000000	1.000000	1.000000
## C1.112	0	1.000000	1.000000	1.000000	1.000000	-1.000000	1.000000
## C1.34	0	1.000000	-1.000000	-1.000000	-1.000000	-1.000000	1.000000
## C1.44	0	1.000000	1.000000	-1.000000	1.000000	-1.000000	1.000000
## C1.56	0	1.000000	1.000000	1.000000	-1.000000	1.000000	1.000000
## C1.100	0	1.000000	1.000000	-1.000000	-1.000000	-1.000000	1.000000
## C1.43	0	-1.000000	1.000000	-1.000000	1.000000	-1.000000	1.000000
## C1.7	0	-1.000000	1.000000	1.000000	-1.000000	-1.000000	-1.000000
## C1.27	0	-1.000000	1.000000	-1.000000	1.000000	1.000000	-1.000000
## C1.58	0	1.000000	-1.000000	-1.000000	1.000000	1.000000	1.000000
## C1.55	0	-1.000000	1.000000	1.000000	-1.000000	1.000000	1.000000
## C1.40	0	1.000000	1.000000	1.000000	-1.000000	-1.000000	1.000000
## C1.6	0	1.000000	-1.000000	1.000000	-1.000000	-1.000000	-1.000000
## C1.46	0	1.000000	-1.000000	1.000000	1.000000	-1.000000	1.000000
## C1.68	0	1.000000	1.000000	-1.000000	-1.000000	-1.000000	-1.000000
## C1.126	0	1.000000	-1.000000	1.000000	1.000000	1.000000	1.000000
## C1.51	0	-1.000000	1.000000	-1.000000	-1.000000	1.000000	1.000000
## C1.26	0	1.000000	-1.000000	-1.000000	1.000000	1.000000	-1.000000
## C1.116	0	1.000000	1.000000	-1.000000	-1.000000	1.000000	1.000000
## C1.45	0	-1.000000	-1.000000	1.000000	1.000000	-1.000000	1.000000
## C1.120	0	1.000000	1.000000	1.000000	-1.000000	1.000000	1.000000
## C1.35	0	-1.000000	1.000000	-1.000000	-1.000000	-1.000000	1.000000
## C1.87	0	-1.000000	1.000000	1.000000	-1.000000	1.000000	-1.000000
## C1.3	0	-1.000000	1.000000	-1.000000	-1.000000	-1.000000	-1.000000

## C1.54	0	1.000000	-1.000000	1.000000	-1.000000	1.000000	1.000000
## C1.127	0	-1.000000	1.000000	1.000000	1.000000	1.000000	1.000000
## C1.99	0	-1.000000	1.000000	-1.000000	-1.000000	-1.000000	1.000000
## C1.23	0	-1.000000	1.000000	1.000000	-1.000000	1.000000	-1.000000
## C1.114	0	1.000000	-1.000000	-1.000000	-1.000000	1.000000	1.000000
## C1.110	0	1.000000	-1.000000	1.000000	1.000000	-1.000000	1.000000
## C1.1	0	-1.000000	-1.000000	-1.000000	-1.000000	-1.000000	-1.000000
## C1.80	0	1.000000	1.000000	1.000000	1.000000	-1.000000	-1.000000
## C1.130	0	0.000000	0.000000	0.000000	0.000000	0.000000	0.000000
## C1.124	0	1.000000	1.000000	-1.000000	1.000000	1.000000	1.000000
## C1.109	0	-1.000000	-1.000000	1.000000	1.000000	-1.000000	1.000000
## C1.86	0	1.000000	-1.000000	1.000000	-1.000000	1.000000	-1.000000
## C1.24	0	1.000000	1.000000	1.000000	-1.000000	1.000000	-1.000000
## C1.11	0	-1.000000	1.000000	-1.000000	1.000000	-1.000000	-1.000000
## C1.107	0	-1.000000	1.000000	-1.000000	1.000000	-1.000000	1.000000
## C1.115	0	-1.000000	1.000000	-1.000000	-1.000000	1.000000	1.000000
## C1.33	0	-1.000000	-1.000000	-1.000000	-1.000000	-1.000000	1.000000
## C1.97	0	-1.000000	-1.000000	-1.000000	-1.000000	-1.000000	1.000000
## C1.96	0	1.000000	1.000000	1.000000	1.000000	1.000000	-1.000000
## C1.42	0	1.000000	-1.000000	-1.000000	1.000000	-1.000000	1.000000
## C1.48	0	1.000000	1.000000	1.000000	1.000000	-1.000000	1.000000
## C1.13	0	-1.000000	-1.000000	1.000000	1.000000	-1.000000	-1.000000
## C1.118	0	1.000000	-1.000000	1.000000	-1.000000	1.000000	1.000000
## C1.41	0	-1.000000	-1.000000	-1.000000	1.000000	-1.000000	1.000000
## C1.78	0	1.000000	-1.000000	1.000000	1.000000	-1.000000	-1.000000
## C1.9	0	-1.000000	-1.000000	-1.000000	1.000000	-1.000000	-1.000000
## C1.72	0	1.000000	1.000000	1.000000	-1.000000	-1.000000	-1.000000
## C1.70	0	1.000000	-1.000000	1.000000	-1.000000	-1.000000	-1.000000
## C1.106	0	1.000000	-1.000000	-1.000000	1.000000	-1.000000	1.000000
## C1.21	0	-1.000000	-1.000000	1.000000	-1.000000	1.000000	-1.000000
## C1.65	0	-1.000000	-1.000000	-1.000000	-1.000000	-1.000000	-1.000000
## C1.94	0	1.000000	-1.000000	1.000000	1.000000	1.000000	-1.000000
## C1.69	0	-1.000000	-1.000000	1.000000	-1.000000	-1.000000	-1.000000
## C1.61	0	-1.000000	-1.000000	1.000000	1.000000	1.000000	1.000000
## C1.71	0	-1.000000	1.000000	1.000000	-1.000000	-1.000000	-1.000000
## C1.66	0	1.000000	-1.000000	-1.000000	-1.000000	-1.000000	-1.000000
## C1.117	0	-1.000000	-1.000000	1.000000	-1.000000	1.000000	1.000000
## C1.122	0	1.000000	-1.000000	-1.000000	1.000000	1.000000	1.000000
## C1.36	0	1.000000	1.000000	-1.000000	-1.000000	-1.000000	1.000000
## C1.101	0	-1.000000	-1.000000	1.000000	-1.000000	-1.000000	1.000000
## C1.28	0	1.000000	1.000000	-1.000000	1.000000	1.000000	-1.000000
## C1.22	0	1.000000	-1.000000	1.000000	-1.000000	1.000000	-1.000000
## C1.20	0	1.000000	1.000000	-1.000000	-1.000000	1.000000	-1.000000
## C1.90	0	1.000000	-1.000000	-1.000000	1.000000	1.000000	-1.000000
## C1.93	0	-1.000000	-1.000000	1.000000	1.000000	1.000000	-1.000000
## C1.81	0	-1.000000	-1.000000	-1.000000	-1.000000	1.000000	-1.000000
## C1.91	0	-1.000000	1.000000	-1.000000	1.000000	1.000000	-1.000000
## C1.29	0	-1.000000	-1.000000	1.000000	1.000000	1.000000	-1.000000
## C1.123	0	-1.000000	1.000000	-1.000000	1.000000	1.000000	1.000000
## C1.119	0	-1.000000	1.000000	1.000000	-1.000000	1.000000	1.000000
## C1.63	0	-1.000000	1.000000	1.000000	1.000000	1.000000	1.000000
## C1.121	0	-1.000000	-1.000000	-1.000000	1.000000	1.000000	1.000000
## C1.2	0	1.000000	-1.000000	-1.000000	-1.000000	-1.000000	-1.000000
## C1.25	0	-1.000000	-1.000000	-1.000000	1.000000	1.000000	-1.000000

## C1.84	0	1.000000	1.000000	-1.000000	-1.000000	1.000000	-1.000000
## C1.62	0	1.000000	-1.000000	1.000000	1.000000	1.000000	1.000000
## C1.10	0	1.000000	-1.000000	-1.000000	1.000000	-1.000000	-1.000000
## C1.37	0	-1.000000	-1.000000	1.000000	-1.000000	-1.000000	1.000000
## C1.52	0	1.000000	1.000000	-1.000000	-1.000000	1.000000	1.000000
## C1.89	0	-1.000000	-1.000000	-1.000000	1.000000	1.000000	-1.000000
## C1.98	0	1.000000	-1.000000	-1.000000	-1.000000	-1.000000	1.000000
## C1.74	0	1.000000	-1.000000	-1.000000	1.000000	-1.000000	-1.000000
## C1.85	0	-1.000000	-1.000000	1.000000	-1.000000	1.000000	-1.000000
## C1.92	0	1.000000	1.000000	-1.000000	1.000000	1.000000	-1.000000
## C1.105	0	-1.000000	-1.000000	-1.000000	1.000000	-1.000000	1.000000
## C1.67	0	-1.000000	1.000000	-1.000000	-1.000000	-1.000000	-1.000000
## C1.32	0	1.000000	1.000000	1.000000	1.000000	1.000000	-1.000000
## C1.38	0	1.000000	-1.000000	1.000000	-1.000000	-1.000000	1.000000
## C1.83	0	-1.000000	1.000000	-1.000000	-1.000000	1.000000	-1.000000
## C1.5	0	-1.000000	-1.000000	1.000000	-1.000000	-1.000000	-1.000000
## C1.49	0	-1.000000	-1.000000	-1.000000	-1.000000	1.000000	1.000000
## C1.4	0	1.000000	1.000000	-1.000000	-1.000000	-1.000000	-1.000000
## C1.73	0	-1.000000	-1.000000	-1.000000	1.000000	-1.000000	-1.000000
## C1.131	0	0.000000	0.000000	0.000000	0.000000	0.000000	0.000000
## S2.14	1	0.000000	0.000000	0.000000	0.000000	0.000000	0.000000
## S2.2	1	3.424207	0.000000	0.000000	0.000000	0.000000	0.000000
## S2.23	1	0.000000	0.000000	0.000000	0.000000	0.000000	0.000000
## S2.16	1	0.000000	0.000000	0.000000	0.000000	0.000000	0.000000
## S2.12	1	0.000000	0.000000	0.000000	0.000000	0.000000	3.424207
## S2.10	1	0.000000	0.000000	0.000000	0.000000	3.424207	0.000000
## S2.7	1	0.000000	0.000000	0.000000	-3.424207	0.000000	0.000000
## S2.19	1	0.000000	0.000000	0.000000	0.000000	0.000000	0.000000
## S2.24	1	0.000000	0.000000	0.000000	0.000000	0.000000	0.000000
## S2.17	1	0.000000	0.000000	0.000000	0.000000	0.000000	0.000000
## S2.21	1	0.000000	0.000000	0.000000	0.000000	0.000000	0.000000
## S2.3	1	0.000000	-3.424207	0.000000	0.000000	0.000000	0.000000
## S2.4	1	0.000000	3.424207	0.000000	0.000000	0.000000	0.000000
## S2.13	1	0.000000	0.000000	0.000000	0.000000	0.000000	0.000000
## S2.9	1	0.000000	0.000000	0.000000	0.000000	-3.424207	0.000000
## S2.22	1	0.000000	0.000000	0.000000	0.000000	0.000000	0.000000
## S2.20	1	0.000000	0.000000	0.000000	0.000000	0.000000	0.000000
## S2.18	1	0.000000	0.000000	0.000000	0.000000	0.000000	0.000000
## S2.15	1	0.000000	0.000000	0.000000	0.000000	0.000000	0.000000
## S2.8	1	0.000000	0.000000	0.000000	3.424207	0.000000	0.000000
## S2.1	1	-3.424207	0.000000	0.000000	0.000000	0.000000	0.000000
## S2.5	1	0.000000	0.000000	-3.424207	0.000000	0.000000	0.000000
## S2.6	1	0.000000	0.000000	3.424207	0.000000	0.000000	0.000000
## S2.11	1	0.000000	0.000000	0.000000	0.000000	0.000000	-3.424207
##		G	H	J	K	resp.y	
## C1.129	0.000000	0.000000	0.000000	0.000000	94162.434		
## C1.102	1.000000	-1.000000	-1.000000	1.000000	85430.443		
## C1.30	-1.000000	-1.000000	-1.000000	1.000000	114789.788		
## C1.82	1.000000	-1.000000	-1.000000	-1.000000	168404.075		
## C1.95	1.000000	-1.000000	1.000000	1.000000	93864.418		
## C1.31	-1.000000	-1.000000	-1.000000	1.000000	100804.436		
## C1.39	-1.000000	-1.000000	1.000000	1.000000	91181.800		
## C1.16	-1.000000	-1.000000	1.000000	-1.000000	105319.684		
## C1.53	-1.000000	-1.000000	-1.000000	-1.000000	95109.129		

##	C1.18	-1.000000	-1.000000	1.000000	-1.000000	95674.325
##	C1.88	1.000000	-1.000000	-1.000000	1.000000	7964.088
##	C1.15	-1.000000	1.000000	-1.000000	1.000000	92309.467
##	C1.75	1.000000	-1.000000	-1.000000	1.000000	84964.605
##	C1.50	-1.000000	-1.000000	-1.000000	1.000000	105787.870
##	C1.59	-1.000000	1.000000	-1.000000	-1.000000	96727.560
##	C1.103	1.000000	-1.000000	-1.000000	1.000000	99113.416
##	C1.64	-1.000000	1.000000	-1.000000	1.000000	99986.205
##	C1.60	-1.000000	-1.000000	1.000000	1.000000	103747.233
##	C1.111	1.000000	1.000000	-1.000000	-1.000000	93352.191
##	C1.12	-1.000000	1.000000	-1.000000	-1.000000	76430.851
##	C1.17	-1.000000	1.000000	-1.000000	1.000000	81494.616
##	C1.77	1.000000	-1.000000	-1.000000	-1.000000	226646.146
##	C1.47	-1.000000	1.000000	1.000000	-1.000000	112367.609
##	C1.104	1.000000	1.000000	1.000000	-1.000000	110516.131
##	C1.76	1.000000	1.000000	1.000000	-1.000000	144420.676
##	C1.113	1.000000	1.000000	-1.000000	-1.000000	86303.755
##	C1.79	1.000000	1.000000	1.000000	1.000000	94593.098
##	C1.14	-1.000000	1.000000	-1.000000	1.000000	98709.237
##	C1.19	-1.000000	-1.000000	1.000000	-1.000000	86662.415
##	C1.128	1.000000	1.000000	1.000000	1.000000	90553.555
##	C1.8	-1.000000	1.000000	1.000000	1.000000	96021.619
##	C1.108	1.000000	1.000000	-1.000000	1.000000	89902.874
##	C1.125	1.000000	1.000000	1.000000	1.000000	99059.651
##	C1.57	-1.000000	-1.000000	1.000000	1.000000	116120.847
##	C1.112	1.000000	-1.000000	1.000000	1.000000	83919.177
##	C1.34	-1.000000	1.000000	-1.000000	1.000000	82615.131
##	C1.44	-1.000000	1.000000	1.000000	1.000000	102608.606
##	C1.56	-1.000000	-1.000000	-1.000000	-1.000000	99934.242
##	C1.100	1.000000	-1.000000	-1.000000	-1.000000	29662.972
##	C1.43	-1.000000	-1.000000	-1.000000	-1.000000	132699.236
##	C1.7	-1.000000	-1.000000	-1.000000	-1.000000	88741.567
##	C1.27	-1.000000	1.000000	1.000000	1.000000	103021.241
##	C1.58	-1.000000	1.000000	-1.000000	-1.000000	82569.428
##	C1.55	-1.000000	1.000000	1.000000	1.000000	79884.555
##	C1.40	-1.000000	1.000000	-1.000000	-1.000000	86638.669
##	C1.6	-1.000000	-1.000000	-1.000000	-1.000000	98046.813
##	C1.46	-1.000000	1.000000	1.000000	-1.000000	91143.263
##	C1.68	1.000000	-1.000000	1.000000	1.000000	101842.398
##	C1.126	1.000000	-1.000000	-1.000000	-1.000000	89751.308
##	C1.51	-1.000000	-1.000000	-1.000000	1.000000	98313.993
##	C1.26	-1.000000	1.000000	1.000000	1.000000	95502.066
##	C1.116	1.000000	1.000000	-1.000000	-1.000000	95214.509
##	C1.45	-1.000000	-1.000000	-1.000000	1.000000	13873.347
##	C1.120	1.000000	-1.000000	1.000000	-1.000000	85945.034
##	C1.35	-1.000000	1.000000	-1.000000	1.000000	112136.092
##	C1.87	1.000000	1.000000	1.000000	-1.000000	103181.283
##	C1.3	-1.000000	1.000000	1.000000	-1.000000	85075.300
##	C1.54	-1.000000	1.000000	1.000000	1.000000	94614.184
##	C1.127	1.000000	-1.000000	-1.000000	-1.000000	87317.877
##	C1.99	1.000000	1.000000	1.000000	1.000000	80831.781
##	C1.23	-1.000000	1.000000	-1.000000	-1.000000	82462.249
##	C1.114	1.000000	-1.000000	1.000000	1.000000	107984.822
##	C1.110	1.000000	1.000000	-1.000000	-1.000000	97085.162

##	C1.1	-1.000000	-1.000000	-1.000000	1.000000	100846.001
##	C1.80	1.000000	-1.000000	-1.000000	-1.000000	97449.680
##	C1.130	0.000000	0.000000	0.000000	0.000000	103535.347
##	C1.124	1.000000	-1.000000	-1.000000	1.000000	58087.345
##	C1.109	1.000000	-1.000000	1.000000	1.000000	93794.126
##	C1.86	1.000000	1.000000	1.000000	-1.000000	77307.856
##	C1.24	-1.000000	-1.000000	1.000000	1.000000	97606.967
##	C1.11	-1.000000	-1.000000	1.000000	1.000000	102499.323
##	C1.107	1.000000	-1.000000	1.000000	-1.000000	99658.902
##	C1.115	1.000000	-1.000000	1.000000	1.000000	88919.933
##	C1.33	-1.000000	-1.000000	1.000000	-1.000000	93185.520
##	C1.97	1.000000	-1.000000	-1.000000	-1.000000	94222.527
##	C1.96	1.000000	1.000000	-1.000000	-1.000000	105889.966
##	C1.42	-1.000000	-1.000000	-1.000000	-1.000000	88367.429
##	C1.48	-1.000000	-1.000000	-1.000000	1.000000	163305.110
##	C1.13	-1.000000	-1.000000	1.000000	-1.000000	96258.288
##	C1.118	1.000000	1.000000	-1.000000	1.000000	109966.571
##	C1.41	-1.000000	1.000000	1.000000	1.000000	90709.220
##	C1.78	1.000000	1.000000	1.000000	1.000000	100392.014
##	C1.9	-1.000000	1.000000	-1.000000	-1.000000	81829.286
##	C1.72	1.000000	1.000000	-1.000000	1.000000	76338.298
##	C1.70	1.000000	-1.000000	1.000000	-1.000000	102775.417
##	C1.106	1.000000	-1.000000	1.000000	-1.000000	95640.559
##	C1.21	-1.000000	-1.000000	1.000000	1.000000	97722.043
##	C1.65	1.000000	-1.000000	1.000000	1.000000	86187.382
##	C1.94	1.000000	-1.000000	1.000000	1.000000	104711.034
##	C1.69	1.000000	1.000000	-1.000000	1.000000	79184.857
##	C1.61	-1.000000	1.000000	-1.000000	1.000000	85257.873
##	C1.71	1.000000	-1.000000	1.000000	-1.000000	95347.517
##	C1.66	1.000000	1.000000	-1.000000	-1.000000	113205.611
##	C1.117	1.000000	-1.000000	1.000000	-1.000000	36601.612
##	C1.122	1.000000	1.000000	1.000000	-1.000000	88500.360
##	C1.36	-1.000000	-1.000000	1.000000	-1.000000	219584.156
##	C1.101	1.000000	1.000000	1.000000	-1.000000	-38330.113
##	C1.28	-1.000000	-1.000000	-1.000000	-1.000000	97079.587
##	C1.22	-1.000000	1.000000	-1.000000	-1.000000	94218.992
##	C1.20	-1.000000	1.000000	-1.000000	1.000000	171852.401
##	C1.90	1.000000	1.000000	-1.000000	1.000000	101682.224
##	C1.93	1.000000	1.000000	-1.000000	-1.000000	149288.073
##	C1.81	1.000000	1.000000	1.000000	1.000000	1099.033
##	C1.91	1.000000	1.000000	-1.000000	1.000000	91683.539
##	C1.29	-1.000000	1.000000	1.000000	-1.000000	109344.586
##	C1.123	1.000000	1.000000	1.000000	-1.000000	71726.271
##	C1.119	1.000000	1.000000	-1.000000	1.000000	103145.467
##	C1.63	-1.000000	-1.000000	1.000000	-1.000000	101862.642
##	C1.121	1.000000	-1.000000	-1.000000	1.000000	100456.861
##	C1.2	-1.000000	1.000000	1.000000	-1.000000	98139.545
##	C1.25	-1.000000	-1.000000	-1.000000	-1.000000	100165.042
##	C1.84	1.000000	1.000000	1.000000	1.000000	91582.757
##	C1.62	-1.000000	-1.000000	1.000000	-1.000000	68414.445
##	C1.10	-1.000000	-1.000000	1.000000	1.000000	99097.288
##	C1.37	-1.000000	1.000000	-1.000000	-1.000000	105099.346
##	C1.52	-1.000000	1.000000	1.000000	-1.000000	100279.993
##	C1.89	1.000000	-1.000000	1.000000	-1.000000	88796.367

```

## C1.98 1.000000 1.000000 1.000000 1.000000 111703.672
## C1.74 1.000000 -1.000000 -1.000000 1.000000 94434.997
## C1.85 1.000000 -1.000000 -1.000000 1.000000 70707.961
## C1.92 1.000000 -1.000000 1.000000 -1.000000 99382.176
## C1.105 1.000000 1.000000 -1.000000 1.000000 101115.650
## C1.67 1.000000 1.000000 -1.000000 -1.000000 93660.576
## C1.32 -1.000000 1.000000 1.000000 -1.000000 94525.042
## C1.38 -1.000000 -1.000000 1.000000 1.000000 99980.520
## C1.83 1.000000 -1.000000 -1.000000 -1.000000 79924.220
## C1.5 -1.000000 1.000000 1.000000 1.000000 103294.958
## C1.49 -1.000000 1.000000 1.000000 -1.000000 208988.648
## C1.4 -1.000000 -1.000000 -1.000000 1.000000 105614.053
## C1.73 1.000000 1.000000 1.000000 -1.000000 90420.589
## C1.131 0.000000 0.000000 0.000000 0.000000 80801.145
## S2.14 3.424207 0.000000 0.000000 0.000000 90467.237
## S2.2 0.000000 0.000000 0.000000 0.000000 80866.497
## S2.23 0.000000 0.000000 0.000000 0.000000 101005.709
## S2.16 0.000000 3.424207 0.000000 0.000000 98846.582
## S2.12 0.000000 0.000000 0.000000 0.000000 198157.287
## S2.10 0.000000 0.000000 0.000000 0.000000 147050.845
## S2.7 0.000000 0.000000 0.000000 0.000000 143636.681
## S2.19 0.000000 0.000000 0.000000 -3.424207 102259.860
## S2.24 0.000000 0.000000 0.000000 0.000000 79687.058
## S2.17 0.000000 0.000000 -3.424207 0.000000 95628.330
## S2.21 0.000000 0.000000 0.000000 0.000000 106239.812
## S2.3 0.000000 0.000000 0.000000 0.000000 106621.398
## S2.4 0.000000 0.000000 0.000000 0.000000 94279.241
## S2.13 -3.424207 0.000000 0.000000 0.000000 88357.682
## S2.9 0.000000 0.000000 0.000000 0.000000 74337.010
## S2.22 0.000000 0.000000 0.000000 0.000000 90513.420
## S2.20 0.000000 0.000000 0.000000 3.424207 89284.997
## S2.18 0.000000 0.000000 3.424207 0.000000 85209.962
## S2.15 0.000000 -3.424207 0.000000 0.000000 94391.556
## S2.8 0.000000 0.000000 0.000000 0.000000 73716.290
## S2.1 0.000000 0.000000 0.000000 0.000000 92533.426
## S2.5 0.000000 0.000000 0.000000 0.000000 97186.295
## S2.6 0.000000 0.000000 0.000000 0.000000 86852.827
## S2.11 0.000000 0.000000 0.000000 0.000000 82431.671
##
## $run.order
## run.no.in.std.order run.no run.no.std.rp
## 1 C1.129 1 C1.129
## 2 C1.102 2 C1.102
## 3 C1.30 3 C1.30
## 4 C1.82 4 C1.82
## 5 C1.95 5 C1.95
## 6 C1.31 6 C1.31
## 7 C1.39 7 C1.39
## 8 C1.16 8 C1.16
## 9 C1.53 9 C1.53
## 10 C1.18 10 C1.18
## 11 C1.88 11 C1.88
## 12 C1.15 12 C1.15
## 13 C1.75 13 C1.75

```

## 14	C1.50	14	C1.50
## 15	C1.59	15	C1.59
## 16	C1.103	16	C1.103
## 17	C1.64	17	C1.64
## 18	C1.60	18	C1.60
## 19	C1.111	19	C1.111
## 20	C1.12	20	C1.12
## 21	C1.17	21	C1.17
## 22	C1.77	22	C1.77
## 23	C1.47	23	C1.47
## 24	C1.104	24	C1.104
## 25	C1.76	25	C1.76
## 26	C1.113	26	C1.113
## 27	C1.79	27	C1.79
## 28	C1.14	28	C1.14
## 29	C1.19	29	C1.19
## 30	C1.128	30	C1.128
## 31	C1.8	31	C1.8
## 32	C1.108	32	C1.108
## 33	C1.125	33	C1.125
## 34	C1.57	34	C1.57
## 35	C1.112	35	C1.112
## 36	C1.34	36	C1.34
## 37	C1.44	37	C1.44
## 38	C1.56	38	C1.56
## 39	C1.100	39	C1.100
## 40	C1.43	40	C1.43
## 41	C1.7	41	C1.7
## 42	C1.27	42	C1.27
## 43	C1.58	43	C1.58
## 44	C1.55	44	C1.55
## 45	C1.40	45	C1.40
## 46	C1.6	46	C1.6
## 47	C1.46	47	C1.46
## 48	C1.68	48	C1.68
## 49	C1.126	49	C1.126
## 50	C1.51	50	C1.51
## 51	C1.26	51	C1.26
## 52	C1.116	52	C1.116
## 53	C1.45	53	C1.45
## 54	C1.120	54	C1.120
## 55	C1.35	55	C1.35
## 56	C1.87	56	C1.87
## 57	C1.3	57	C1.3
## 58	C1.54	58	C1.54
## 59	C1.127	59	C1.127
## 60	C1.99	60	C1.99
## 61	C1.23	61	C1.23
## 62	C1.114	62	C1.114
## 63	C1.110	63	C1.110
## 64	C1.1	64	C1.1
## 65	C1.80	65	C1.80
## 66	C1.130	66	C1.130
## 67	C1.124	67	C1.124

## 68	C1.109	68	C1.109
## 69	C1.86	69	C1.86
## 70	C1.24	70	C1.24
## 71	C1.11	71	C1.11
## 72	C1.107	72	C1.107
## 73	C1.115	73	C1.115
## 74	C1.33	74	C1.33
## 75	C1.97	75	C1.97
## 76	C1.96	76	C1.96
## 77	C1.42	77	C1.42
## 78	C1.48	78	C1.48
## 79	C1.13	79	C1.13
## 80	C1.118	80	C1.118
## 81	C1.41	81	C1.41
## 82	C1.78	82	C1.78
## 83	C1.9	83	C1.9
## 84	C1.72	84	C1.72
## 85	C1.70	85	C1.70
## 86	C1.106	86	C1.106
## 87	C1.21	87	C1.21
## 88	C1.65	88	C1.65
## 89	C1.94	89	C1.94
## 90	C1.69	90	C1.69
## 91	C1.61	91	C1.61
## 92	C1.71	92	C1.71
## 93	C1.66	93	C1.66
## 94	C1.117	94	C1.117
## 95	C1.122	95	C1.122
## 96	C1.36	96	C1.36
## 97	C1.101	97	C1.101
## 98	C1.28	98	C1.28
## 99	C1.22	99	C1.22
## 100	C1.20	100	C1.20
## 101	C1.90	101	C1.90
## 102	C1.93	102	C1.93
## 103	C1.81	103	C1.81
## 104	C1.91	104	C1.91
## 105	C1.29	105	C1.29
## 106	C1.123	106	C1.123
## 107	C1.119	107	C1.119
## 108	C1.63	108	C1.63
## 109	C1.121	109	C1.121
## 110	C1.2	110	C1.2
## 111	C1.25	111	C1.25
## 112	C1.84	112	C1.84
## 113	C1.62	113	C1.62
## 114	C1.10	114	C1.10
## 115	C1.37	115	C1.37
## 116	C1.52	116	C1.52
## 117	C1.89	117	C1.89
## 118	C1.98	118	C1.98
## 119	C1.74	119	C1.74
## 120	C1.85	120	C1.85
## 121	C1.92	121	C1.92

## 122	C1.105	122	C1.105
## 123	C1.67	123	C1.67
## 124	C1.32	124	C1.32
## 125	C1.38	125	C1.38
## 126	C1.83	126	C1.83
## 127	C1.5	127	C1.5
## 128	C1.49	128	C1.49
## 129	C1.4	129	C1.4
## 130	C1.73	130	C1.73
## 131	C1.131	131	C1.131
## 132	S2.14	132	S2.14
## 133	S2.2	133	S2.2
## 134	S2.23	134	S2.23
## 135	S2.16	135	S2.16
## 136	S2.12	136	S2.12
## 137	S2.10	137	S2.10
## 138	S2.7	138	S2.7
## 139	S2.19	139	S2.19
## 140	S2.24	140	S2.24
## 141	S2.17	141	S2.17
## 142	S2.21	142	S2.21
## 143	S2.3	143	S2.3
## 144	S2.4	144	S2.4
## 145	S2.13	145	S2.13
## 146	S2.9	146	S2.9
## 147	S2.22	147	S2.22
## 148	S2.20	148	S2.20
## 149	S2.18	149	S2.18
## 150	S2.15	150	S2.15
## 151	S2.8	151	S2.8
## 152	S2.1	152	S2.1
## 153	S2.5	153	S2.5
## 154	S2.6	154	S2.6
## 155	S2.11	155	S2.11
##			
##	\$design.info		
##	\$design.info\$type		
##	[1] "ccd"		
##			
##	\$design.info\$nruns		
##	[1] 155		
##			
##	\$design.info\$nfactors		
##	[1] 10		
##			
##	\$design.info\$factor.names		
##	\$design.info\$factor.names\$A		
##	[1] -1 1		
##			
##	\$design.info\$factor.names\$B		
##	[1] -1 1		
##			
##	\$design.info\$factor.names\$C		
##	[1] -1 1		

```

##
## $design.info$factor.names$D
## [1] -1  1
##
## $design.info$factor.names$E
## [1] -1  1
##
## $design.info$factor.names$F
## [1] -1  1
##
## $design.info$factor.names$G
## [1] -1  1
##
## $design.info$factor.names$H
## [1] -1  1
##
## $design.info$factor.names$J
## [1] -1  1
##
## $design.info$factor.names$K
## [1] -1  1
##
##
## $design.info$catlg.name
## [1] "catlg"
##
## $design.info$catlg.entry
## Design: 10-3.1
##      128 runs, 10 factors,
##      Resolution V
##      Generating columns: 31 103 43
##      WLP (3plus): 0 0 3 3 , 45 clear 2fis
##
## $design.info$aliased
## $design.info$aliased$legend
## [1] "A=A" "B=B" "C=C" "D=D" "E=E" "F=F" "G=G" "H=H" "J=J" "K=K"
##
## $design.info$aliased[[2]]
## [1] "no aliasing among main effects and 2fis"
##
##
## $design.info$FrF2.version
## [1] "2.2-2"
##
## $design.info$replications
## [1] 1
##
## $design.info$repeat.only
## [1] FALSE
##
## $design.info$randomize
## [1] TRUE
##
## $design.info$seed

```



```

## NULL
##
## $design.info$creator
## $design.info$creator[[1]]
## FrF2(nfactors = 10, resolution = 5, ncenter = 3)
##
## $design.info$creator[[2]]
## ccd.augment(design)
##
##
## $design.info$quantitative
##   A   B   C   D   E   F   G   H   J   K
## TRUE TRUE TRUE TRUE TRUE TRUE TRUE TRUE TRUE TRUE
##
## $design.info$ncube
## [1] 128
##
## $design.info$ncenter
## [1] 3 4
##
## $design.info$coding
## $design.info$coding$x1
## x1 ~ (A - 0)/1
##
## $design.info$coding$x2
## x2 ~ (B - 0)/1
##
## $design.info$coding$x3
## x3 ~ (C - 0)/1
##
## $design.info$coding$x4
## x4 ~ (D - 0)/1
##
## $design.info$coding$x5
## x5 ~ (E - 0)/1
##
## $design.info$coding$x6
## x6 ~ (F - 0)/1
##
## $design.info$coding$x7
## x7 ~ (G - 0)/1
##
## $design.info$coding$x8
## x8 ~ (H - 0)/1
##
## $design.info$coding$x9
## x9 ~ (J - 0)/1
##
## $design.info$coding$x10
## x10 ~ (K - 0)/1
##
##
## $design.info$block.name
## [1] "Block.ccd"

```

```

##
## $design.info$cube.gen
## $design.info$cube.gen[[1]]
## x8 ~ x1 * x2 * x3 * x4 * x5
## <environment: 0x0000000026d771d0>
##
## $design.info$cube.gen[[2]]
## x9 ~ x1 * x2 * x3 * x6 * x7
## <environment: 0x0000000026d83550>
##
## $design.info$cube.gen[[3]]
## x10 ~ x1 * x2 * x4 * x6
## <environment: 0x0000000026d8baa0>
##
##
## $design.info$nstar
## [1] 20
##
## $design.info$response.names
## [1] "resp.y"
##
##
## $names
## [1] "Block.ccd" "A" "B" "C" "D" "E"
## [7] "F" "G" "H" "J" "K" "resp.y"
##
## $class
## [1] "design" "data.frame"
summary(design)

## Multi-step-call:
## [[1]]
## FrF2(nfactors = 10, resolution = 5, ncenter = 3)
##
## [[2]]
## ccd.augment(design)
##
##
## Experimental design of type ccd
## 155 runs
##
## Factor settings (cube):
## A B C D E F G H J K
## 1 -1 -1 -1 -1 -1 -1 -1 -1 -1
## 2 1 1 1 1 1 1 1 1 1
##
## Numbers of cube and star points:
## Cube Star
## 128 20
## Numbers of center points:
## Cube Star
## 3 4
##
## Responses:

```

```
## [1] resp.y
```