Statistical learning homework 4

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1 ESL 7.2

Let
$$p = P\{\hat{G}(x_0) = G(x_0) | x_0\}$$
 we have

$$\operatorname{Err}(x_0) = P\{Y \neq \hat{G}(x_0) | X = x_0\}$$

$$= P\{Y \neq G(x_0) | X = x_0\} \cdot p + P\{Y = G(x_0) | X = x_0\} \cdot (1 - p)$$

$$= P\{Y \neq G(x_0) | X = x_0\} - P\{Y \neq G(x_0) | X = x_0\} \cdot (1 - p) + P\{Y = G(x_0) | X = x_0\} \cdot (1 - p)$$

$$= \operatorname{Err}_{B}(x_0) + [P\{Y = G(x_0) | X = x_0\} - P\{Y \neq G(x_0) | X = x_0\}] \cdot (1 - p)$$

$$= \operatorname{Err}_{B}(x_0) + |2f(x_0) - 1|P\{\hat{G}(x_0) \neq G(x_0) | X = x_0\}.$$

Now let $\sigma^2 = \text{Var}[f(x_0)]$ and $\mu = \mathbb{E}[\hat{f}(x_0)]$ we have

$$\begin{split} P\{\hat{G}(x_0) \neq G(x_0), X = x_0\} &= P\{\hat{f}(x_0) < 0.5, G(x_0) = 1\} + P\{\hat{f}(x_0) > 0.5 | G(x_0) = 0\} \\ &\approx \Phi\left(\frac{0.5 - \mu}{\sigma}\right) P\{G(x_0) = 1\} + \left[1 - \Phi\left(\frac{0.5 - \mu}{\sigma}\right)\right] P\{G(x_0) = 0\} \\ &= \Phi\left(\frac{0.5 - \mu}{\sigma}\right) P\{G(x_0) = 1\} + \Phi\left(\frac{\mu - 0.5}{\sigma}\right) P\{G(x_0) = 0\} \\ &= \Phi\left(\frac{\mathrm{sign}(0.5 - f(x_0))(\mu - 0.5)}{\sigma}\right), \end{split}$$

hence we conclude the proof.

2 ESL 7.6

We have

$$\hat{Y}(x) = \frac{1}{k} \sum_{x_i \in N_k(x)} y_i = Sy,$$

here S is a binary matrix with values 1/k or 0 such that the sum of each row sums is 1. Moreover, the diagonal entries are all 1/k, since x_i should be the neighbour of x_i 's, hence

$$df(\mathbf{S}) = tr(\mathbf{S}) = \frac{N}{k}.$$

3 ESL 7.7

Let $x = \operatorname{tr}(S)/N$ we can calculate that

$$GCV = \frac{1}{N} \sum_{i}^{N} \left[\frac{y_i - \hat{y}_i}{1 - \operatorname{tr}(S)/N} \right]^2 \approx \frac{1}{N} \sum_{i}^{N} (y_i - \hat{y}_i)^2 \left[1 + 2 \frac{\operatorname{tr}(S)}{N} \right]$$
$$= \frac{1}{N} \sum_{i}^{N} (y_i - \hat{y}_i)^2 + 2 \frac{d}{N^2} \sum_{i}^{N} (y_i - \hat{y}_i)^2$$
$$\approx \overline{\operatorname{err}} + 2 \frac{d}{N} \sigma_{\epsilon}^2$$
$$= C_p.$$

4 ESL 10.2

The goal is to find

$$f^*(x) = \arg\min_{f(x)} \mathbb{E}_{Y|x} \left[e^{-Yf(x)} \right],$$

taking the derivative of f(x) and setting it to zero we have

$$\frac{\partial}{\partial f} E_{Y|x} \left[e^{-Yf(x)} \right] = E_{Y|x} \left[-Ye^{-Yf(x)} \right] = 0,$$

thus

$$e^{f(x)}P\{Y = -1|x\} - e^{-f(x)}P\{Y = 1|x\} = 0,$$

hence

$$f(x) = \frac{1}{2} \log \left(\frac{P\{Y = 1|x\}}{P\{Y = -1|x\}} \right).$$

5 ESL 10.5

5.1 a

The optimization problem can be written as

$$\min_{f} \quad \mathbb{E}\left[\exp\left(-\frac{1}{K}Y^{T}f\right)\right]$$
 s.t.
$$\mathbf{1}^{T}f = 0,$$

the Lagrange function is

$$\mathcal{L}(f, \nu) = \mathbb{E}\left[\exp\left(-\frac{1}{K}Y^T f\right)\right] + \nu\left(\mathbf{1}^T f\right),$$

then

$$\frac{\partial \mathcal{L}}{\partial f} = -\frac{1}{K} \mathbb{E} \left[Y \exp \left(-\frac{1}{K} Y^T f \right) \right] + \nu \mathbf{1}^T = 0,$$

on the other hand we have

$$\mathbb{E}\left[Y\exp\left(-\frac{1}{K}Y^Tf\right)\right] = \sum_{k=1}^K P\{G = \mathcal{G}_k\}Y_k\exp\left(\frac{1}{K(K-1)}\sum_{i=1}^K f_i - \frac{1}{K-1}f_k\right),$$

hence the entries of the above are all equal, which means

$$P\{G = \mathcal{G}_k\} \exp\left(-\frac{1}{K-1}f_k\right) = P\{G = \mathcal{G}_t\} \exp\left(-\frac{1}{K-1}f_t\right), \quad \forall k, t.$$

now we fix f_1 and then

$$\log P\{G = \mathcal{G}_k\} - \frac{1}{K-1} f_k = \log P\{G = \mathcal{G}_1\} - \frac{1}{K-1} f_1,$$

hence

$$f_k = f_1 + (K - 1) \log \frac{P\{G = \mathcal{G}_k\}}{P\{G = \mathcal{G}_1\}},$$

since $1^T f = 0$ we have

$$Kf_1 + (K-1)\sum_{k=2}^{K} \log \frac{P\{G = \mathcal{G}_k\}}{P\{G = \mathcal{G}_1\}} = 0$$

which yields that

$$f_1 = \frac{K-1}{K} \left[(K-1) \log P\{G = \mathcal{G}_1\} - \sum_{k=2}^K \log P\{G = \mathcal{G}_k\} \right] = (K-1) \log P\{G = \mathcal{G}_1\} - \frac{K-1}{K} \sum_{k=1}^K \log P\{G = \mathcal{G}_k\},$$

hence

$$f_k = (K-1)\log P\{G = \mathcal{G}_k\} - \frac{K-1}{K} \sum_{k=1}^K \log P\{G = \mathcal{G}_k\}.$$

Finally we get

$$f_k^* = (K-1)\log P\{G = \mathcal{G}_k\} - \frac{K-1}{K} \sum_{k=1}^K \log P\{G = \mathcal{G}_k\}.$$

5.2 b

Now the optimization problem is

$$(\beta_m, G_m) = \arg\min_{\beta, G} \sum_i L[y_i, f_{m-1}(x_i) + \beta G(x_i)]$$

which can be written as

$$(\beta_m, G_m) = \arg\min_{\beta, G} \sum_i w_i^{(m)} \exp\left[-\frac{\beta}{K} y_i^T G(x_i)\right]$$

here $w_i^{(m)} = \exp\left[-y_i^T f_{m-1}(x_i)\right]$. Now we take the derivative of β and set it to zero we have

$$\sum_{i} w_i^{(m)} y_i^T G(x_i) \exp\left[-\frac{\beta}{K} y_i^T G(x_i)\right] = 0,$$

since when $G(x_i) = y_i$ we have

$$y_i^T G(x_i) = \frac{K-1}{(K-1)^2} + 1 = \frac{K}{K-1},$$

when $G(x_i) \neq y_i$ we have

$$y_i^T G(x_i) = \frac{K-2}{(K-1)^2} - \frac{2}{K-1} = \frac{-K}{(K-1)^2},$$

hence

$$\sum_{y_i = G(x_i)} w_i^{(m)} y_i^T G(x_i) \exp\left[-\frac{1}{K-1}\beta\right] + \sum_{y_i \neq G(x_i)} w_i^{(m)} y_i^T G(x_i) \exp\left[\frac{1}{(K-1)^2}\beta\right] = 0$$

which yields

$$\exp\left[\frac{K}{(K-1)^2}\beta\right] = (K-1)\frac{\sum_{y_i = G(x_i)} w_i^{(m)}}{\sum_{y_i \neq G(x_i)} w_i^{(m)}},$$

hence

$$\beta_m = \frac{(k-1)^2}{K} \left[\log \left(\frac{1 - \operatorname{err}_m}{\operatorname{err}_m} \right) + \log(K - 1) \right].$$

6 ESL 10.8

6.1 a

First we have

$$p_k(x) = \frac{e^{f_k(x)}}{\sum_j e^{f_j(x)}},$$

the log-likelihood can be written as

$$\begin{split} l &= \sum_{x_i \in R} \sum_{k=1}^K y_{ik} \log p_k(x_i) = \sum_{x_i \in R} \sum_{k=1}^K y_{ik} \left[\log e^{f_k(x_i) + \gamma_k} - \log \left(\sum_j e^{f_j(x_i) + \gamma_j} \right) \right] \\ &= \sum_{x_i \in R} \sum_{k=1}^K y_{ik} \left[f_k(x_i) + \gamma_k \right] - \sum_{x_i \in R} \sum_{k=1}^K y_{i,k} \log \left(\sum_j e^{f_j(x_i) + \gamma_j} \right) \\ &= \sum_{x_i \in R} \sum_{k=1}^K y_{ik} \left[f_k(x_i) + \gamma_k \right] - \sum_{x_i \in R} \log \left(\sum_j e^{f_j(x_i) + \gamma_j} \right). \end{split}$$

The first derivative is

$$\frac{\partial l}{\partial \gamma_k} = \sum_{x_i \in R} y_{i,k} - \sum_{x_i \in R} \frac{e^{f_k(x_i) + \gamma_k}}{\sum_j e^{f_j(x_i) + \gamma_j}}.$$

The second derivatives are

$$\frac{\partial^2 l}{\partial \gamma_k^2} = -\sum_{x_i \in R} \frac{e^{f_k(x_i) + \gamma_k} \sum_j e^{f_j(x_i) + \gamma_j} - e^{2f_k(x_i) + 2\gamma_k}}{\left(\sum_j e^{f_j(x_i) + \gamma_j}\right)^2}$$

and

$$\frac{\partial^2 l}{\partial \gamma_k \gamma_m} = \sum_{x_i \in R} \frac{e^{f_k(x_i) + \gamma_k} e^{f_m(x_i) + \gamma_m}}{\left(\sum_j e^{f_j(x_i) + \gamma_j}\right)^2}.$$

6.2 b

The Newton update can be written as

$$\gamma_{new} = \left(\frac{\partial^2 l}{\partial \gamma^2}\right)^{-1} \frac{\partial l}{\partial \gamma},$$

since we only use the diagonal of the Hessian matrix, we have

$$\gamma_k^1 = \frac{\partial l}{\partial \gamma_k} \cdot \left(\frac{\partial^2 l}{\partial \gamma_k^2} \right)^{-1} \bigg|_{\gamma_k = 0} = \frac{\sum_{x_i \in R} (y_{i,k} - p_k(x_i))}{\sum_{x_i \in R} p_k(x_i)(1 - p_k(x_i))}$$

6.3 c

We can verify that

$$\sum_{k=1}^{K} \hat{\gamma}_k = \frac{K-1}{K} \sum_{k=1}^{K} \gamma_k^1 - K \cdot \frac{K-1}{K} \sum_{k=1}^{K} \frac{1}{K} \gamma_k^1 = 0.$$

7 ESL 12.1

Consider (12.8), since C > 0, the optimal ξ_i should satisfy

$$\xi_i[y_i(x_i^T \beta + \beta_0) - 1 + \xi_i] = 0,$$

hence

$$\xi_i = 1 - y_i(x_i^T \beta + \beta_0) \quad \text{or} \quad \xi_i = 0$$

must hold, hence $\xi_i = [1 - y_i(x_i^T \beta + \beta_0)]_+$, hence (12.8) can be written as

$$\min_{\beta,\beta_0} \quad \frac{1}{2} ||\beta||^2 + C \sum_{i=1}^{N} [1 - y_i(x_i^T \beta + \beta_0)]_+,$$

with $\lambda = 1/C$ the optimization problem is also equivalent to

$$\min_{\beta,\beta_0} \quad \frac{\lambda}{2} ||\beta||^2 + \sum_{i=1}^{N} [1 - y_i(x_i^T \beta + \beta_0)]_+,$$

thereby we conclude the proof.

8 ESL 12.2

If we choose K such that

$$\langle h(x_j), h(x_i) \rangle = K(x_j, x_i),$$

then

$$f(x_j) = h(x_j)^T \beta + \beta_0 = \sum_{i=1}^{N} \alpha_i y_i \langle h(x_j), h(x_i) \rangle + \beta_0 = \sum_{i=1}^{N} \alpha_i y_j K(x_j, x_i) + \beta_0$$

for some α_i , where

$$\beta = \sum_{i=1}^{N} \alpha_i y_i \left\langle \cdot, h(x_i) \right\rangle,$$

let $\alpha_i' = \alpha_i y_i$ hence

$$\beta^T \beta = \sum_{i=1}^N \sum_{j=1}^N \alpha_i \alpha_j y_i y_j \langle h(x_j), h(x_i) \rangle = \sum_{i=1}^N \sum_{j=1}^N \alpha_i' \alpha_j' K(x_j, x_i) = (\alpha')^T K(\alpha'),$$

hence

$$h(x_j)^T \beta = \sum_{i=1}^N \alpha_i' K(x_j, x_i), \quad \beta^T \beta = (\alpha')^T K(\alpha'),$$

hence the solution to (12.29) is the same as the solution to (12.25) for a particular kernel $K(x_j, x_i) = \langle h(x_j), h(x_i) \rangle$, we have concluded the proof.