《Probability and Statistics》

2022-2023 (2) 期末考试卷参考答案及评分标准

Some possible useful critical values:

 $z_{0.025} = 1.96$; $z_{0.05} = 1.645$;

 $\begin{array}{l} t_{0.025,8}=2.306~;~t_{0.025,9}=2.262~;~t_{0.05,8}=1.86~;~t_{0.05,9}=1.833~;\\ \chi^2_{0.025,8}=17.534;~\chi^2_{0.025,9}=19.022;~\chi^2_{0.975,8}=2.18;~\chi^2_{0.975,9}=2.70; \end{array}$

- I. (20 points) In a game, A, B and C take turns flipping a fair coin. The first one to get a head wins. Assume that A flips first, then B, then C, then A, and so on. Let X be the number of flips in total when the game ends.
 - Find the probability that B wins in his/her fifth flip.
 - b) Find the probability that C wins.
 - c) Calculate $\mathbb{P}(X \leq 2023)$.
 - d) Find the expected value of X.

a) The probability that B win in his/her fifth flip is

b) Let
$$C_n$$
 be the event that C wins in his/her n th flip. Then
$$(\frac{1}{2} \cdot \frac{1}{2} \cdot \frac{1}{2})^4 \cdot \frac{1}{2} \cdot \frac{1}{2} = \frac{1}{2^{14}}.$$

$$\mathbb{P}(C_n) = \left(\frac{1}{2} \cdot \frac{1}{2} \cdot \frac{1}{2}\right)^n = \frac{1}{8^n}$$

and

$$\mathbb{P}(C \text{ wins}) = \sum_{n=1}^{+\infty} \mathbb{P}(C_n) = \sum_{n=1}^{+\infty} \frac{1}{8^n} = \frac{1}{7}.$$
 10 $\frac{1}{2}$

c) Since $\mathbb{P}(X = n) = \frac{1}{2^{n}}$ we have

$$\mathbb{P}(X \le 2023) = \sum_{i=1}^{2023} \mathbb{P}(X = i) = \sum_{i=1}^{2023} \frac{1}{2^i} = 1 - 2^{-2023}.$$
15 \(\frac{1}{2}\)

d) From c), we know that X is a geometric random variable with parameter $\frac{1}{2}$. Thus

$$\mathbb{E}(X) = \frac{1}{1/2} = 2.$$

- II. (10 points) You ask your roommate to water a sickly plant while you are on vacation. Without water, it will die with probability 0.8; with water, it will die with probability 0.15. You are 90 percent certain that your roommate will remember to water the plant.
 - **a.** What is the probability that the plant will be alive when you return?
 - **b.** If the plant is dead upon your return, what is the probability that your roommate forgot to water it?

Solution.

- a) $\mathbb{P}(\text{alive}) = \mathbb{P}(\text{alive} \mid \text{remember})\mathbb{P}(\text{remember}) + \mathbb{P}(\text{alive} \mid \text{forget})\mathbb{P}(\text{forget}) =$

- III. (10 points) Balls numbered 1 through 2023 are in an urn. Suppose that n, $n \le 2023$, of them are randomly selected without replacement. Let Y denote the largest number selected.
 - **a.** Find the probability mass function of Y.
 - **b.** Calculate $\mathbb{E}(Y)$.

Solution.

a) If $n \le k \le 2023$, then

$$\mathbb{P}(Y=k) = \frac{\binom{k-1}{n-1}}{\binom{2023}{n}}$$

otherwise $\mathbb{P}(Y = k) = 0$.

b) From a), we see that

$$1 = \sum_{k=n}^{2023} \mathbb{P}(Y = k) = \sum_{k=n}^{2023} \frac{\binom{k-1}{n-1}}{\binom{2023}{n}} \Rightarrow \sum_{k=n}^{2023} \binom{k-1}{n-1} = \binom{2023}{n}.$$

Therefore,

$$\mathbb{E}(Y) = \sum_{k=n}^{2023} k \cdot \mathbb{P}(Y = k) = \sum_{k=n}^{2023} \frac{k \cdot \binom{k-1}{n-1}}{\binom{2023}{n}} = \frac{n}{\binom{2023}{n}} \sum_{k=n}^{2023} \binom{k}{n}$$

$$\text{set } j = k+1 = \frac{n}{\binom{2023}{n}} \sum_{j=n+1}^{2024} \binom{j-1}{(n+1)-1} = \frac{n}{\binom{2023}{n}} \binom{2024}{n+1}$$

$$= \frac{2024 \cdot n}{n+1}.$$

IV. (10 points) Let W have the exponential distribution with mean 1. Let $0 < w_1 < w_2 < \infty$. Set g(w) = 0 for $0 \le w < w_1, g(w) = 1$ for $w_1 \le w < w_2$, and g(w) = 2 for $w \ge w_2$. If W can be used to construct a random variable Y = g(W) such that $P(Y = 0,1,2) = \frac{1}{3}$. Find w_1 and w_2

Solution.

The probability function of Y = g(W) is given by

$$P(Y = 0) = P(0 \le W < w_1) = 1 - e^{-w_1},$$

 $P(Y = 1) = P(w_1 \le W < w_2) = e^{-w_1} - e^{-w_2},$
 $P(Y = 2) = P(W \ge w_2) = e^{-w_2}.$

Thus *Y* is uniformly distributed on $\{0,1,2\}$ provided that $1 - e^{-w_1} = 1/3$ and $e^{-w_2} = 1/3$; that is, $w_1 = \log(3/2)$ and $w_2 = \log 3$.

V. (10 points) Let U and V be independent random variables, each uniformly distributed on [0,1]. Determine the mean and variance of the random variable $Y = 3U^2 - 2V$

Solution.

Now $E(U^2) = \int_0^1 u^2 du = 1/3$ and $E(U^4) = \int_0^1 u^4 du = 1/5$, so $var(U^2) = E(U^4) - [E(U^2)]^2 = 1/5 - 1/9 = 4/45$. Also, E(V) = 1/2 and $E(V^2) = 1/3$, so $var(V) = E(V^2) - [E(V)]^2 = 1/3 - 1/4 = 1/12$. So var(V) = 1/3. Consequently, $E(Y) = E(3U^2 - 2V) = 3E(U^2) - 2E(V) = 1 - 1 = 0$ and $var(Y) = var(3U^2 - 2V) = 9var(U^2) + 4var(V) = 9(4/45) + 4(1/12) = 4/5 + 1/3 = 17/15$.

VI. (10 points) The number of parking tickets issued in a certain city on any given

weekday has a Poisson distribution with parameter $\mu = 50$. What is the approximate probability that between 30 and 70 tickets are given out on a particular day (Use the CLT)? $(\Phi(2.90) = 0.9981, \Phi(2.83) = 0.9977)$

Solution.

With Y = # of tickets, Y has approximately a normal distribution with $\mu = 50$ and $\sigma = \sqrt{\mu} = 7.071.$

So, using a continuity correction from [30,70] to [29.5,70.5],

$$P(30 \le Y \le 70) \approx P\left(\frac{29.5 - 50}{7.071} \le Z \le \frac{70.5 - 50}{7.071}\right) = P(-2.90 \le Z \le 2.90)$$

$$= 2\Phi(2.90) - 1 = 0.9962.$$

VII. (10 points) Consider a random sample $X_1, X_2, ..., X_n$ from the pdf $f(x; k) = \begin{cases} kx^{k-1} & 0 \le x \le 1\\ 0 & \text{otherwise} \end{cases}$

$$f(x;k) = \begin{cases} kx^{k-1} & 0 \le x \le 1\\ 0 & \text{otherwise} \end{cases}$$

where k > 0.

- a. Use the method of moments to obtain an estimator of k.
- b. Obtain the maximum likelihood estimator of k.

Solution.

a) $E(X) = \int_{-\infty}^{+\infty} x f(x; k) dx = \int_{0}^{1} k x^{k} dx = \frac{k}{k+1}$

Equating E(X) to \bar{X} gives $\frac{k}{k+1} = \bar{X}$. Solving it we get the moment estimator

$$\hat{k} = \frac{\bar{X}}{1 - \bar{X}}.$$

b) The likelihood function is

$$L(k) = k^n x_1^{k-1} x_2^{k-1} \cdots x_n^{k-1}, \quad 0 \le x_i \le 1 (i = 1, \dots, n).$$

$$ln L(k) = n ln k + (k-1) \sum ln x_i.$$

Let
$$\frac{\mathrm{d}L(k)}{\mathrm{d}k} = \frac{n}{k} + \sum \ln x_i = 0$$
 to obtain $\hat{k} = -\frac{n}{\sum \ln x_i}$.

And thus the mle is

$$\hat{k} = -\frac{n}{\sum \ln X_i}.$$

- VIII. (10 points) Suppose the weight distribution of packages sent in a certain manner is normal. A random sample of n = 9 packages yielded a sample mean weight of 10 and a sample standard deviation of 4.5.
 - Assume the true population standard deviation is unknown. Please compute a 95% CI for the true average weight.
 - b. Assume the true average weight is unknown. Please compute a 95% CI for the true population variance.
 - Assume the true population standard deviation is known, which happens to be 4.5. What sample size is necessary to ensure that the resulting 95% CI for the true average weight has a width of (at most) 2?

- a) A 95% CI for the true average weight is $\bar{x} \pm t_{0.025,8} \frac{4.5}{\sqrt{9}} = 10 \pm 2.306 \times \frac{4.5}{3} = 10 \pm 2.306 \times \frac{4.$ 10 ± 3.459 .
- b) A 95% CI for the true population variance, σ^2 , is given by $\left(\frac{(n-1)s^2}{\chi^2_{0.025,8}}, \frac{(n-1)s^2}{\chi^2_{0.975,8}}\right) = \left(\frac{8\times4.5^2}{17.534}, \frac{8\times4.5^2}{2.18}\right) = (9.2392, 74.3119).$
- c) A 95% CI has a width $2z_{0.025} \frac{4.5}{\sqrt{n}} = 9 \times 1.96 \times \frac{1}{\sqrt{n}} = \frac{17.64}{\sqrt{n}}$.

Solve
$$\frac{17.64}{\sqrt{n}} \le 2$$
 to obtain $n \ge \left(\frac{17.64}{2}\right)^2 = 77.7924$.
Since n must be an integer, n is at least 78.

IX. (10 points) Consider the following sample of diameters of n = 9 randomly selected ball bearings (滚珠):

Suppose the ball bearing diameter is normally distributed with $\sigma = 0.24$

- a. Does the data indicate conclusively that the true average diameter bigger than 0.5 at a significance level $\alpha = 0.05$?
- b. If the true average diameter is $\mu = 0.6316$ and a level $\alpha = 0.05$ test based on n = 9 is used, what is the probability of making a type II error?

Solution.

- a) Null hypothesis: H_0 : $\mu = 0.5$. Alternative hypothesis: H_α : $\mu > 0.5$. The testing statistic value $z = \frac{\bar{x} 0.5}{0.24/3}$. Rejection region: $z \ge z_{0.05} = 1.645$. Substituting $\bar{x} = 0.6$, z = 1.25. The computed value does not fall in the rejection region. So we cannot reject H_0 , which means that the sample data does not suggest that the true average diameter is bigger than 0.5 at a significance level $\alpha = 0.05$.
- b) $\beta(0.6316) = \Phi\left(z_{0.05} + \frac{0.5 0.6316}{0.24/3}\right) = \Phi(0) = 0.5.$