

# Econometrics - sVAR and Co-integration

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# 1 Exercise 1 - sIRF

## 1.1 Structural VAR and structural Impulse Response Function

In the initial stage, as we are dealing with time series data and we have to estimate a sVAR model, it is essential to assess whether the series are stationary, checking for the presence or absence of a unit root. This examination can be performed through both visual inspection and a more precise test, such as the Augmented Dickey Fuller test. In both instances, as there is no evidence of volatility clustering or trend in the graphs(Figure 1), and the Augmented Dickey Fuller test (Figure 2) consistently rejects the null hypothesis, we can reasonably assume that the three series are stationary.

Having established stationarity, the subsequent step involves estimating the Vector Autoregression model. To accomplish this, it is crucial to identify the optimal number of lags by evaluating which number minimizes information criteria (AIC, BIC and HQC), ensuring the correct specification of the model. As depicted in the graph below (Figure 3), the optimal number of lags is one. Therefore, we can proceed with the estimation process.

To assess the validity of the model, I conducted tests to ensure that the residuals of the VAR(1) model exhibit no autocorrelation. In each lag, none of the p-values exceeded conventional alpha levels, indicating a failure to reject the null hypothesis of no autocorrelation. Furthermore, I examined the presence of heteroskedasticity using an ARCH test, and once again, the null hypothesis was not rejected, confirming homoskedasticity. The final test performed aimed to verify if the residuals followed a normal distribution, and the results aligned with previous findings, providing further confirmation of the model's correctness(Figure 4).

Furthermore, I examined whether all the equations estimated by the VAR(1) are statistically significant by assessing if, jointly, all the regressors in each equation significantly differ from zero. The result (Figure 5) indicates that, for each equation, the p-value of the F-Test rejects the null hypothesis, yielding a desirable outcome. Given that the model is correctly specified, the next step is to compute the structural impulse response function. It is worth mentioning that the VAR model has limitations in providing a clear understanding of cause-and-effect relationships between variables, as it primarily focuses on describing how variables move together over time without explicitly addressing the underlying economic mechanisms. To address this limitation, it is recommended to use Structural VARs, which aim to enhance the interpretability of relationships among variables by incorporating economic theories and assumptions.

Once the Structural VAR is estimated and the impulse-response functions are computed, we can delve into the propagation mechanism of oil price shocks (Figure 6, first column) on itself, inflation, and GDP. Regarding the persistence of an oil price shock, which can be considered as the "memory" of the shock, it is observed that before dissipating, the shock persists in the financial time series for approximately seven months. This persistence follows an exponential decay after an immediate peak. Concerning the impact of an oil shock on inflation, the peak of the impact is reached after two months, followed by a gradual decay lasting up to nine months. This aligns with economic theory, as the computation of inflation heavily weights the energy component. Therefore, an oil shock is expected to significantly influence inflation, as seen in recent events following the war in Ukraine.

In contrast, even though at the very beginning it seems to have a positive influence, the impact of an oil shock on GDP is negative. The negative impact reaches its peak after 5 months before gradually diminishing, with reabsorption occurring after the tenth month. This is plausible given the extensive use of oil in industry, being a fundamental component in almost every daily

consumable product. Thus, a shock in oil prices could lead to an economic slowdown.

## 1.2 ARMA Model vs VAR

Regarding the second part, I decided to analyze the oil time series and compute an ARMA model. An ARMA(p,q) model is a statistical model that combines autoregressive (AR) and moving average (MA) components. The AR part models the relationship between an observation and its lagged values, while the MA part models the relationship between an observation and a linear combination of past error terms.

To estimate our model, similar to what I did previously for the VAR, we need to determine the number of lags, p and q. This can be achieved in two ways: graphically by inspecting the ACF and the PACF (Figure 7) or by checking which model minimizes the information criteria (Figure 8). A gradually geometrically declining ACF and a PACF that is significant for only a few lags indicate an AR process. In this case, the PACF has only one significant lag (the ninth lag goes outside the bands, but we are hypothetizing an alpha at 5%, so we can neglect it) followed by a drop in PACF values, becoming insignificant. Thus, the correlogram suggests that the series follows an AR(1) process, where the lags are determined by the number of significant lags in the PACF.

The result is confirmed even when looking at the models that minimize the information criteria. Although, following this method, the model that minimizes the AIC criteria is an ARMA(3,3), once estimating that model, some of the coefficients are not significant, and, in addition, the model is not so parsimonious. For these reasons and since the model that minimizes the BIC and HQC is an AR(1), this is the model I estimated (Figure 9). The coefficient associated with the lagged value of oil is significant, the residuals of the model are not autocorrelated, and they follow a normal distribution (Figure 10). Moreover, we also do not reject the null hypothesis of the presence of an ARCH effect, so the model is overall correctly specified.

In a comparison with the equation in the VAR model, some differences can be spotted. The  $R^2$  in the VAR equation is higher, but this could be due to the obvious presence of more regressors. However, even looking at the  $R^2$  adjusted, the final conclusion is the same, so the VAR equation explains more variance of the dependent variable than the AR(1) model. As mentioned, the VAR equation has more regressors. Indeed, it seems that the lagged value of the growth has an explanatory power with respect to oil, which is something that the ARMA model, working only on the dependent variable, can't capture. Moreover, all the information criteria of the VAR equation are lower. Another difference, from a theoretical point of view, is that the VAR model considers contemporaneous relationships among all variables, while the autoregressive model does not.

## 2 Exercise 2 - Purchasing Power Parity

### 2.1 Stationarity and Co-Integration

The Purchasing Power Parity (PPP) is an economic concept suggesting that national price levels, once converted to a common currency, should equalize. To investigate the potential relationship of cointegration between the log prices of the product in the first country (denoted as  $p_a$ ), the log prices of the product in the second country (denoted as  $p_b$ ), and the log nominal exchange rate (denoted as  $e_{AB}$ ), we first need to check whether these variables are stationary. This can be done visually or with the ADF test. As illustrated in the appendix, the three time series appear to exhibit a common upward trend (Figure 11), suggesting the presence of a unit root. In a statistical context, the ADF tests (Figure 12, Figure 13) consistently fail to reject the null hypothesis ( $H_0$ ) of non-stationarity, indicating that our series possess a unit root and are integrated of order 1 ( $I(1)$ ).

With the non-stationarity of our series established, the objective is to identify evidence of a linear combination among these  $I(1)$  variables capable of generating  $I(0)$  (or stationary) residuals. Given the aim to verify the reaction of the three variables to disequilibria and the intention to estimate a Vector Error Correction Model (VECM), the Johansen Test is employed to check for cointegration and identify the possible number of cointegration equations.

The use of the Johansen Test is preferred as it avoids the issue of choosing a dependent variable and is more suitable for multivariate analysis. It can detect multiple cointegrating vectors and treats every test variable as endogenous variables, making it more appropriate than the Engle-Granger test for multivariate analysis.

The Johansen Co-Integration test is run with the number of lags chosen based on what would be used if estimating a Vector Autoregression (VAR), which, in this case, is one (Figure 14). The trace-statistic in the Johansen table (Figure 15), for the rank 0, rejects the null hypothesis of no cointegration at any alpha level, indicating the presence of cointegration. Further examination for a higher rank of the matrix reveals that the null hypothesis of the presence of more than one cointegration equation is not rejected, leading to the conclusion that the rank of the matrix is one. Thus, only one cointegration relationship is identified.

In the context of a VECM, variables are in their first difference form. Therefore, the number of lags when estimating the VECM is one less than the number of lags in the corresponding VAR. In this case, the number of lags is zero, resulting in each equation in the VECM consisting of a constant term and the Error Correction Term (ECT), with no lagged variables included. This setup implies that the short-run equilibrium cannot be analyzed directly.

The equations derived from the VECM (Figure 16), reveal that only two out of the three variables react to a disequilibrium, namely  $p_a$  (ECT1 coefficient  $\approx -0,304$ ) and  $p_b$  (ECT1 coefficient  $\approx -0,02$ ), although the latter not at a 1% significance level. This implies that if the long-run equilibrium is disturbed,  $P_a$  and  $P_b$  will adjust, leading to a restoration of the equilibrium. Even in this case, the residuals of the model are not autocorrelated and they're normal. Moreover, there's no evidence of any ARCH effect (Figure 17).

To assess whether the PPP exists in its strong form, linear restrictions are imposed on the coefficients in the cointegrating vector related to  $p_b$  and  $e_{AB}$ . However, running a restricted VECM (Figure 18) by imposing restrictions such as  $b[1] = 1$ ,  $b[2] = -1$ , and  $b[3] = -1$  leads to the rejection of the null hypothesis that these linear restrictions are respected. In conclusion, the PPP exists, but not in a strong form.

### 3 Appendix - First Exercise

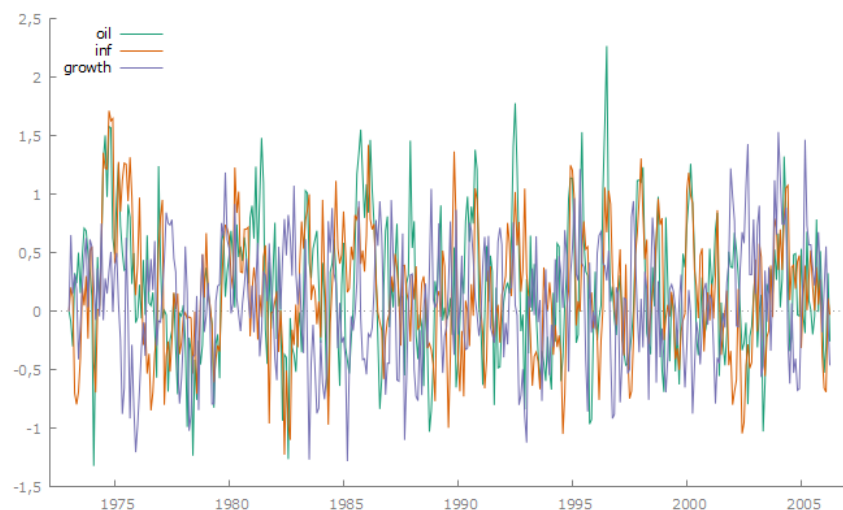


Figure 1: Time Series Plot

```

Test Dickey-Fuller aumentato per oil
test all'indietro da 16 ritardi, criterio AIC
Amplezza campionaria 399
Ipotesi nulla di radice unitaria: a = 1

Test con costante
inclusi 0 ritardi di (1-L)oil
Modello: (1-L)y = b0 + (a-1)*y(-1) + e
Valore stimato di (a - 1): -0,426515
Statistica test: tau_c(1) = -10,364
p-value asintotico 2,233e-20
Coefficiente di autocorrelazione del prim'ordine per e: 0,011

Test Dickey-Fuller aumentato per inf
test all'indietro da 16 ritardi, criterio AIC
Amplezza campionaria 399
Ipotesi nulla di radice unitaria: a = 1

Test con costante
inclusi 0 ritardi di (1-L)inf
Modello: (1-L)y = b0 + (a-1)*y(-1) + e
Valore stimato di (a - 1): -0,347793
Statistica test: tau_c(1) = -9,14088
p-value asintotico 1,742e-16
Coefficiente di autocorrelazione del prim'ordine per e: 0,028

Test Dickey-Fuller aumentato per growth
test all'indietro da 16 ritardi, criterio AIC
Amplezza campionaria 399
Ipotesi nulla di radice unitaria: a = 1

Test con costante
inclusi 0 ritardi di (1-L)growth
Modello: (1-L)y = b0 + (a-1)*y(-1) + e
Valore stimato di (a - 1): -0,521735
Statistica test: tau_c(1) = -11,817
p-value asintotico 4,126e-25
Coefficiente di autocorrelazione del prim'ordine per e: 0,021

```

Figure 2: ADF Test for each variable

The VAR Lag Selection of Gretl suggests that the number of lags that minimize all the information criteria is one.

Sistema VAR, ordine massimo ritardi 12					
Gli asterischi indicano i valori migliori (ossia minimizzati) dei rispettivi criteri di informazione, AIC = criterio di Akaike, BIC = criterio bayesiano di Schwartz e HQC = criterio di Hannan-Quinn.					
ritardi	logver	p(LR)	AIC	BIC	HQC
1	-653,64006		3,431134*	3,553640*	3,479706*
2	-650,55264	0,72230	3,461612	3,675996	3,546612
3	-647,11085	0,64924	3,490262	3,796525	3,611691
4	-641,59519	0,27357	3,508223	3,906365	3,666080
5	-637,24020	0,46446	3,532166	4,022187	3,726452
6	-630,84614	0,17243	3,545599	4,127499	3,776313
7	-626,37609	0,44282	3,568949	4,242728	3,836092
8	-621,40895	0,35583	3,589737	4,355395	3,893309
9	-612,19478	0,03052	3,588633	4,446170	3,928633
10	-608,58294	0,61384	3,616407	4,565823	3,992836
11	-603,48424	0,33474	3,636517	4,677812	4,049374
12	-594,42107	0,03374	3,636191	4,769365	4,085477

Figure 3: VAR Lag Selection

There are the tests to check whether the model is correctly specified or not.

Test per l'autocorrelazione fino all'ordine 12			
	Rao F	Approx dist.	p-value
lag 1	0,653	F(9, 949)	0,7515
lag 2	0,665	F(18, 1095)	0,8472
lag 3	0,777	F(27, 1122)	0,7851
lag 4	0,784	F(36, 1126)	0,8169
lag 5	0,862	F(45, 1123)	0,7289
lag 6	0,814	F(54, 1118)	0,8289
lag 7	0,916	F(63, 1111)	0,6618
lag 8	1,079	F(72, 1103)	0,3092
lag 9	1,072	F(81, 1095)	0,3168
lag 10	1,036	F(90, 1087)	0,3929
lag 11	1,143	F(99, 1078)	0,1693
lag 12	1,111	F(108, 1070)	0,2155
Test per ARCH di ordine 12			
	LM	df	p-value
lag 1	28,992	36	0,7900
lag 2	77,360	72	0,3116
lag 3	117,663	108	0,2470
lag 4	139,703	144	0,5857
lag 5	179,638	180	0,4936
lag 6	205,669	216	0,6819
lag 7	237,248	252	0,7391
lag 8	272,305	288	0,7384
lag 9	302,408	324	0,8000
lag 10	340,858	360	0,7585
lag 11	377,099	396	0,7450
lag 12	404,741	432	0,8225
Matrice di correlazione dei residui, C (3 x 3)			
	1,0000	0,28441	-0,016664
	0,28441	1,0000	-0,045188
	-0,016664	-0,045188	1,0000
Autovalori di C			
	0,71415		
	0,994851		
	1,291		
Test di Doornik-Hansen			
Chi-quadro(6) = 0,676364 [0,9950]			

Figure 4: Validation of the VAR Model

This is the VAR model with one lag. In all the equations the F-Test rejects the null, so at least one coefficient is different from zero.

```
Sistema VAR, ordine ritardi 1
Stime OLS usando le osservazioni 1973:02-2006:04 (T = 399)
Log-verosimiglianza = -669,05555
Determinante della matrice di covarianza = 0,0057418759
AIC = 3,4138
BIC = 3,5338
HQC = 3,4613
Test portmanteau: LB(48) = 496,792, df = 423 [0,0076]

Equazione 1: oil
-----
coefficiente  errore std.  rapporto t  p-value
-----
const        0,0650306  0,0251229   2,589      0,0100 ***
oil_l        0,536946   0,0473278  11,35     5,15e-026 ***
inf_l        0,0730337  0,0510778   1,430     0,1535
growth_l     0,144265   0,0460109   3,135     0,0018 ***

Media var. dipendente  0,187046  SQM var. dipendente  0,576148
Somma quadr. residui  86,39327  E.S. della regressione  0,467645
R-quadro              0,346149  R-quadro corretto    0,341183
F(3, 395)             69,70441  P-value(F)          3,39e-36
rho                   0,019763  Durbin-Watson       1,957488
Note: SQM = scarto quadratico medio; E.S. = errore standard

Test F per zero vincoli:

Tutti i ritardi di oil      F(1, 395) = 128,71 [0,0000]
Tutti i ritardi di inf     F(1, 395) = 2,0445 [0,1535]
Tutti i ritardi di growth  F(1, 395) = 9,8310 [0,0018]
Tutte le variabili, ritardo 1 F(3, 395) = 69,704 [0,0000]

Equazione 2: inf
-----
coefficiente  errore std.  rapporto t  p-value
-----
const        0,0223693  0,0205834   1,087     0,2778
oil_l        0,286222   0,0387761   7,381     9,35e-013 ***
inf_l        0,501980   0,0418484  12,00     1,76e-028 ***
growth_l     0,0279317  0,0376971   0,7410    0,4592

Media var. dipendente  0,156672  SQM var. dipendente  0,538433
Somma quadr. residui  57,98613  E.S. della regressione  0,383145
R-quadro              0,497452  R-quadro corretto    0,493635
F(3, 395)             130,3316  P-value(F)          1,08e-58
rho                   0,049019  Durbin-Watson       1,900789
Note: SQM = scarto quadratico medio; E.S. = errore standard

Test F per zero vincoli:

Tutti i ritardi di oil      F(1, 395) = 54,485 [0,0000]
Tutti i ritardi di inf     F(1, 395) = 143,88 [0,0000]
Tutti i ritardi di growth  F(1, 395) = 0,54901 [0,4592]
Tutte le variabili, ritardo 1 F(3, 395) = 130,33 [0,0000]

Equazione 3: growth
-----
coefficiente  errore std.  rapporto t  p-value
-----
const        0,0547284  0,0240839   2,272     0,0236 **
oil_l        0,0504742  0,0453706   1,112     0,2666
inf_l        -0,179873   0,0489655  -3,673    0,0003 ***
growth_l     0,451900   0,0441081  10,25     5,32e-022 ***

Media var. dipendente  0,066655  SQM var. dipendente  0,517453
Somma quadr. residui  79,38628  E.S. della regressione  0,448306
R-quadro              0,255061  R-quadro corretto    0,249403
F(3, 395)             45,08150  P-value(F)          4,48e-25
rho                   0,021554  Durbin-Watson       1,949737
Note: SQM = scarto quadratico medio; E.S. = errore standard

Test F per zero vincoli:

Tutti i ritardi di oil      F(1, 395) = 1,2376 [0,2666]
Tutti i ritardi di inf     F(1, 395) = 13,494 [0,0003]
Tutti i ritardi di growth  F(1, 395) = 104,97 [0,0000]
Tutte le variabili, ritardo 1 F(3, 395) = 45,082 [0,0000]
```

Figure 5: Vector Autoregressive Model



Below, the structural Impulse Response Functions. We must focus on the oil shocks, so the first column.

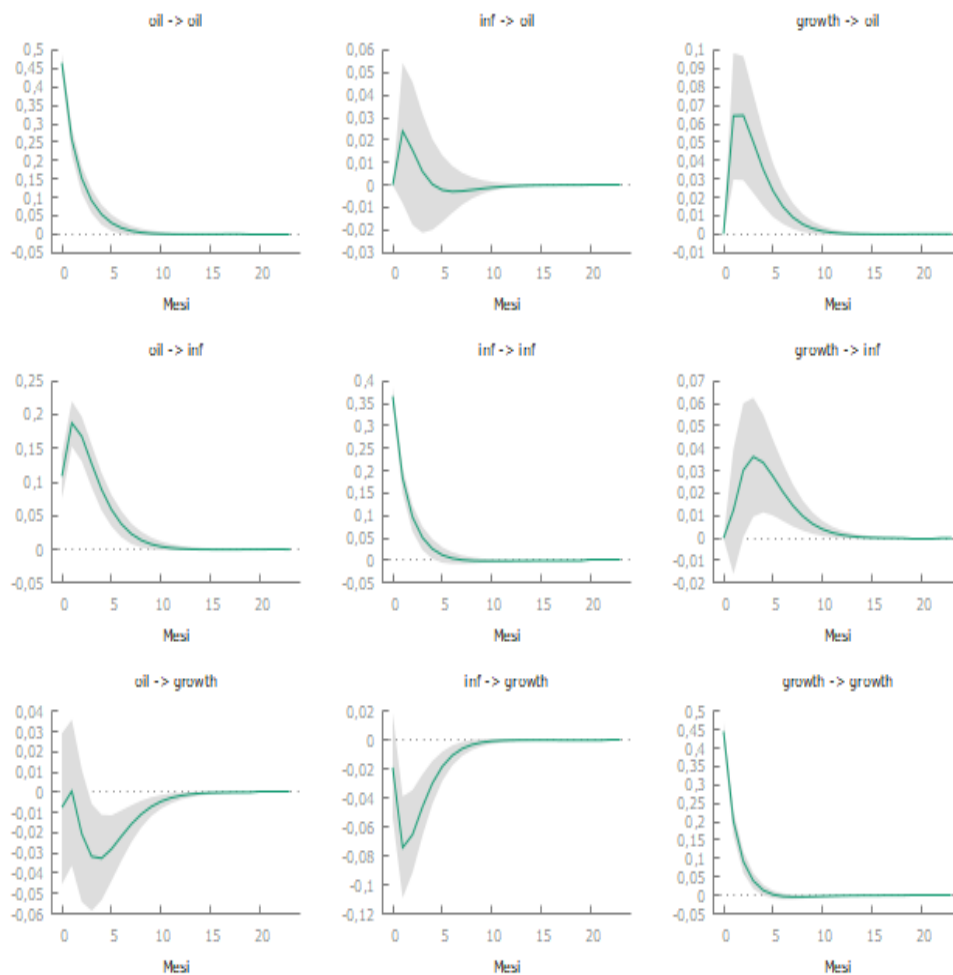


Figure 6: Impulse Response Function

### 3.1 ARMA Model - Oil prices

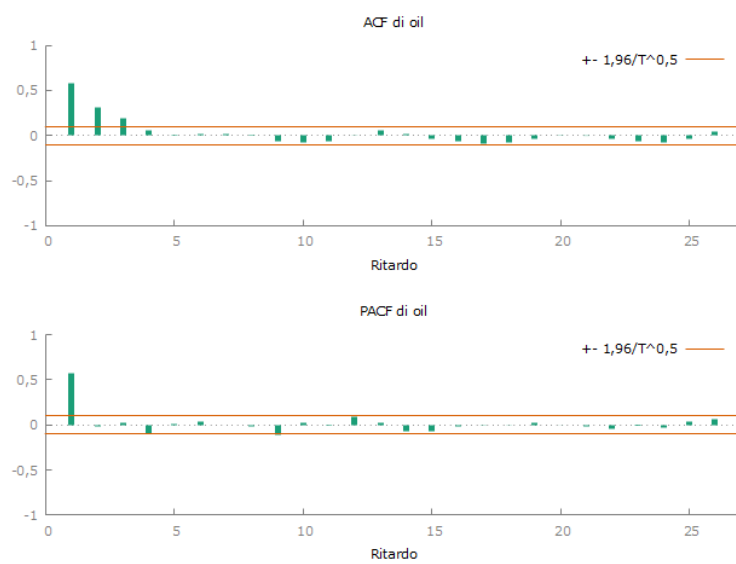


Figure 7: ACF - PACF

Stimato usando AS 197 (MV esatta)  
Dependent variable oil, T = 400  
Criteria for ARMA(p, q) specifications

p, q	AIC	BIC	HQC	loglik
0, 0	696,1386	704,1215	699,3000	-346,0693
0, 1	569,3741	581,3485	574,1161	-281,6870
0, 2	554,5525	570,5184	560,8752	-273,2763
0, 3	541,9689	561,9262	549,8723	-265,9845
1, 0	539,2531	551,2275*	543,9951*	-266,6265
1, 1	541,0723	557,0381	547,3950	-266,5361
1, 2	542,6668	562,6241	550,5701	-266,3334
1, 3	541,2283	565,1770	550,7123	-264,6141
2, 0	541,0931	557,0589	547,4158	-266,5465
2, 1	542,1191	562,0765	550,0225	-266,0596
2, 2	543,6708	567,6196	553,1548	-265,8354
2, 3	541,5614	569,5016	552,6261	-263,7807
3, 0	542,7394	562,6968	550,6428	-266,3697
3, 1	543,4716	567,4204	552,9556	-265,7358
3, 2	539,3918	567,3321	550,4565	-262,6959
3, 3	538,3418*	570,2735	550,9871	-261,1709

Figure 8: ARMA Lag Selection

Modello 7: ARMA, usando le osservazioni 1973:01-2006:04 (T = 400)  
Stimato usando AS 197 (MV esatta)  
Variabile dipendente: oil  
Errori standard basati sull'Hessiana

	coefficiente	errore std.	z	p-value	
const	0,184476	0,0548699	3,362	0,0008	***
phi_1	0,572210	0,0409091	13,99	1,86e-044	***

Media var. dipendente	0,186578	SQM var. dipendente	0,575501
Media innovazioni	0,000347	SQM innovazioni	0,471014
R-quadro	0,328478	R-quadro corretto	0,328478
Log-verosimiglianza	-266,6265	Criterio di Akaike	539,2531
Criterio di Schwarz	551,2275	Hannan-Quinn	543,9951

Note: SQM = scarto quadratico medio; E.S. = errore standard

Figure 9: AR(1) Model

Below, the test for the normality of the residual and the test for the presence of autocorrelation.

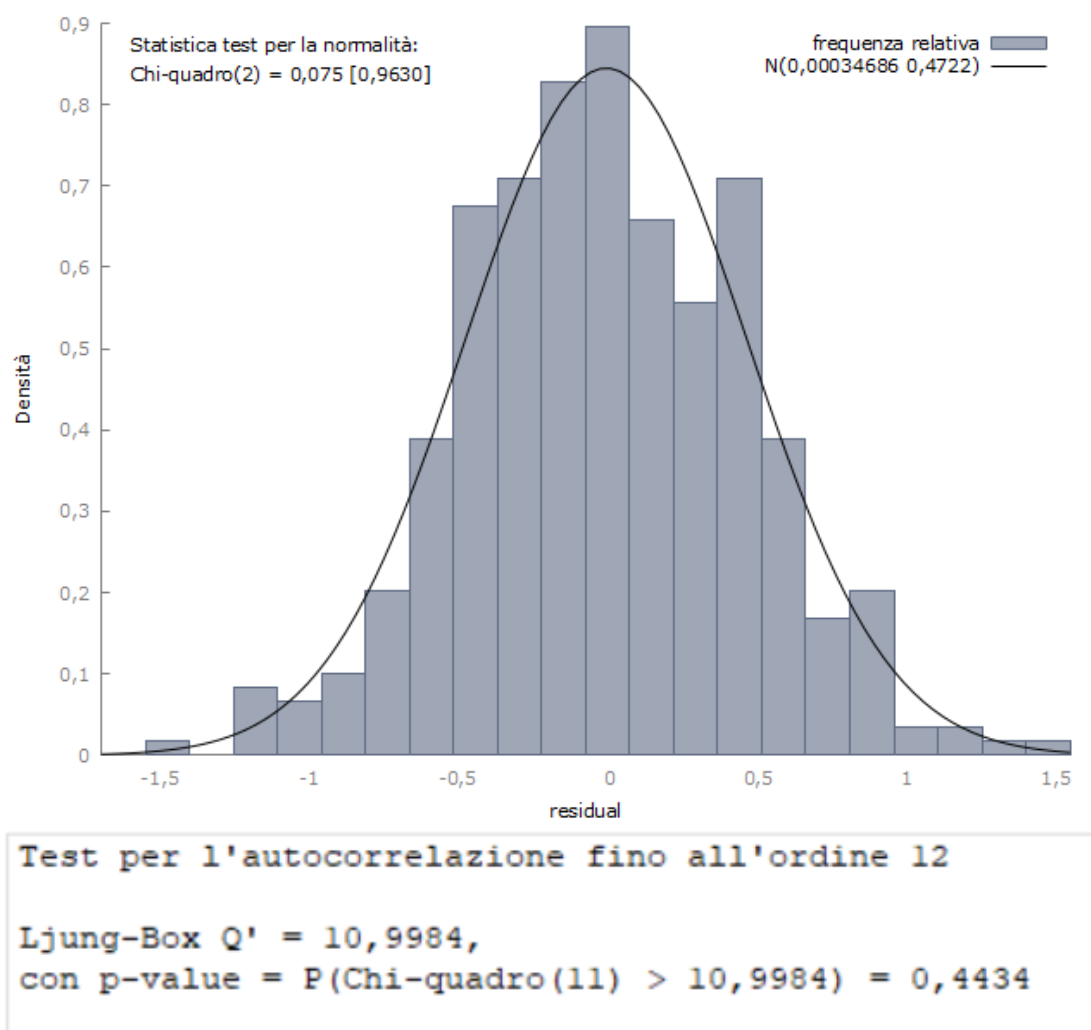


Figure 10: Residuals distribution and Autocorrelation Test

## 4 Appendix - Second Exercise

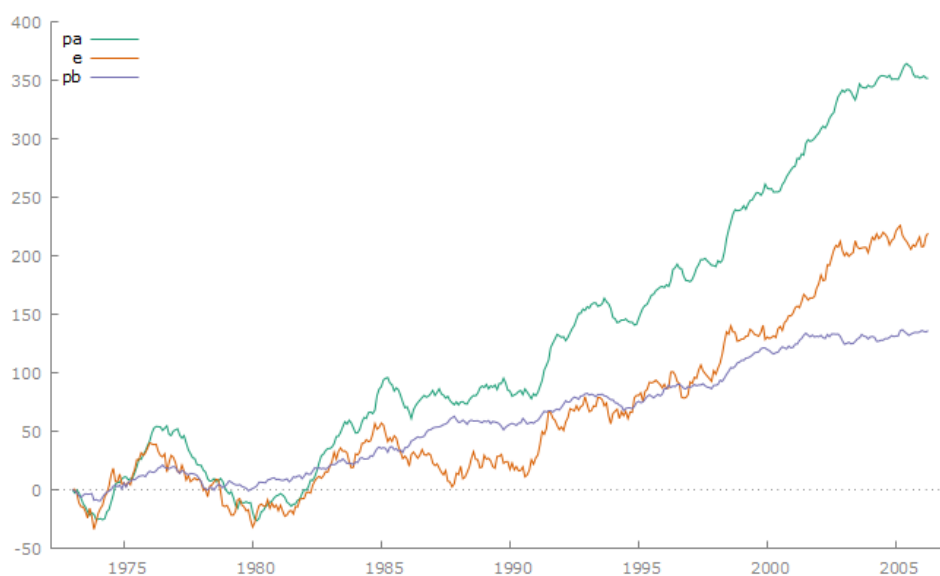


Figure 11: Time Series Plot

```
Test Dickey-Fuller aumentato per pa
test all'indietro da 16 ritardi, criterio AIC
Ampiezza campionaria 391
Ipotesi nulla di radice unitaria: a = 1

Test con costante
inclusi 8 ritardi di (1-L)pa
Modello: (1-L)y = b0 + (a-1)*y(-1) + ... + e
Valore stimato di (a - 1): 0,000260122
Statistica test: tau_c(1) = 0,184954
p-value asintotico 0,9717
Coefficiente di autocorrelazione del prim'ordine per e: -0,002
differenze ritardate: F(8, 381) = 10,935 [0,0000]

Con costante e trend
inclusi 8 ritardi di (1-L)pa
Modello: (1-L)y = b0 + b1*t + (a-1)*y(-1) + ... + e
Valore stimato di (a - 1): -0,00768343
Statistica test: tau_ct(1) = -1,75197
p-value asintotico 0,728
Coefficiente di autocorrelazione del prim'ordine per e: -0,003
differenze ritardate: F(8, 380) = 11,115 [0,0000]
```

Figure 12: ADF Pa

```

Test Dickey-Fuller aumentato per e
test all'indietro da 16 ritardi, criterio AIC
Amplezza campionaria 399
Ipotesi nulla di radice unitaria: a = 1

Test con costante
inclusi 0 ritardi di (1-L)e
Modello: (1-L)y = b0 + (a-1)*y(-1) + e
Valore stimato di (a - 1): 0,00218475
Statistica test: tau_c(1) = 0,615708
p-value asintotico 0,9902
Coefficiente di autocorrelazione del prim'ordine per e: -0,002

Con costante e trend
inclusi 0 ritardi di (1-L)e
Modello: (1-L)y = b0 + b1*t + (a-1)*y(-1) + e
Valore stimato di (a - 1): -0,0138472
Statistica test: tau_ct(1) = -1,6944
p-value asintotico 0,7541
Coefficiente di autocorrelazione del prim'ordine per e: 0,002

Test Dickey-Fuller aumentato per pb
test all'indietro da 16 ritardi, criterio AIC
Amplezza campionaria 399
Ipotesi nulla di radice unitaria: a = 1

Test con costante
inclusi 0 ritardi di (1-L)pb
Modello: (1-L)y = b0 + (a-1)*y(-1) + e
Valore stimato di (a - 1): 0,000534501
Statistica test: tau_c(1) = 0,281572
p-value asintotico 0,9774
Coefficiente di autocorrelazione del prim'ordine per e: 0,051

Con costante e trend
inclusi 0 ritardi di (1-L)pb
Modello: (1-L)y = b0 + b1*t + (a-1)*y(-1) + e
Valore stimato di (a - 1): -0,0223825
Statistica test: tau_ct(1) = -2,36121
p-value asintotico 0,4001
Coefficiente di autocorrelazione del prim'ordine per e: 0,059

```

Figure 13: ADF Test for E and Pb

Sistema VAR, ordine massimo ritardi 6					
Gli asterischi indicano i valori migliori (ossia minimizzati) dei rispettivi criteri di informazione, AIC = criterio di Akaike, BIC = criterio bayesiano di Schwartz e HQC = criterio di Hannan-Quinn.					
ritardi	logver	p(LR)	AIC	BIC	HQC
1	-2774,27176		14,143511*	14,264619*	14,191500*
2	-2772,02898	0,87665	14,177812	14,389750	14,261792
3	-2769,56934	0,84129	14,211012	14,513780	14,330983
4	-2765,57607	0,53550	14,236427	14,630025	14,392389
5	-2761,00653	0,42454	14,258916	14,743345	14,450870
6	-2753,25806	0,07816	14,265269	14,840528	14,493214

Figure 14: Lag Selection for Johansen test and VECM

Caso 3: costante non vincolata

Log-verosimiglianza = -1675,55 (termine costante incluso: -2807,86)

Rango	Autovalore	Test	traccia	p-value	Test Lmax	p-value
0	0,67348	450,90	[0,0000]		446,59	[0,0000]
1	0,010719	4,3111	[0,8717]		4,3000	[0,8218]
22,7941e-005		0,011149	[0,9159]		0,011149	[0,9159]

Corretto per ampiezza campionaria (df = 395)

Rango	Test	traccia	p-value
0	450,90	[0,0000]	
1	4,3111	[0,8727]	
2	0,011149	[0,9165]	

Figure 15: Johansen Test

Sistema VECM, ordine ritardi 1  
 Stime Massima verosimiglianza usando le osservazioni 1973:02-2006:04 (T = 399)  
 Rango di cointegrazione = 1  
 Caso 3: costante non vincolata

beta (vettori di cointegrazione, errori standard tra parentesi)

pa 1,0000  
 (0,00000)  
 e -0,99082  
 (0,012463)  
 pb -1,0691  
 (0,019550)

alpha (vettori di aggiustamento)

pa -0,30454  
 e -0,014125  
 pb -0,019888

Log-verosimiglianza = -2810,0181  
 Determinante della matrice di covarianza = 262,87335  
 AIC = 14,1455  
 BIC = 14,2654  
 HQC = 14,1930

Equazione 1: d\_pa

	coefficiente	errore std.	rapporto t	p-value	
const	0,262149	0,100961	2,597	0,0098	***
EC1	-0,304539	0,0107420	-28,35	1,90e-097	***

Media var. dipendente	0,881044	SQM var. dipendente	3,419977
Somma quadr. residui	1539,117	E.S. della regressione	1,968977
R-quadro	0,669370	R-quadro corretto	0,668537
rho	0,028894	Durbin-Watson	1,932139

Note: SQM = scarto quadratico medio; E.S. = errore standard

Equazione 2: d\_e

	coefficiente	errore std.	rapporto t	p-value	
const	0,520731	0,259683	2,005	0,0456	**
EC1	-0,0141254	0,0276298	-0,5112	0,6095	

Media var. dipendente	0,549437	SQM var. dipendente	5,059747
Somma quadr. residui	10182,51	E.S. della regressione	5,064448
R-quadro	0,000658	R-quadro corretto	-0,001859
rho	-0,015444	Durbin-Watson	2,030526

Note: SQM = scarto quadratico medio; E.S. = errore standard

Equazione 3: d\_pb

	coefficiente	errore std.	rapporto t	p-value	
const	0,300510	0,0880630	3,412	0,0007	***
EC1	-0,0198877	0,0093697416	-2,123	0,0344	**

Media var. dipendente	0,340926	SQM var. dipendente	1,724990
Somma quadr. residui	1170,996	E.S. della regressione	1,717444
R-quadro	0,011221	R-quadro corretto	0,008730
rho	0,006996	Durbin-Watson	1,984176

Note: SQM = scarto quadratico medio; E.S. = errore standard

Figure 16: VECM Cointegration Vector and Pa equation



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Test per l'autocorrelazione fino all'ordine 12

	Rao F	Approx dist.	p-value
lag 1	0,486	F(9, 956)	0,8845
lag 2	0,525	F(18, 1103)	0,9474
lag 3	0,651	F(27, 1130)	0,9143
lag 4	0,729	F(36, 1135)	0,8814
lag 5	0,901	F(45, 1132)	0,6587
lag 6	0,972	F(54, 1127)	0,5342
lag 7	0,897	F(63, 1120)	0,7020
lag 8	0,884	F(72, 1112)	0,7426
lag 9	0,824	F(81, 1104)	0,8655
lag 10	0,828	F(90, 1096)	0,8736
lag 11	0,775	F(99, 1087)	0,9468
lag 12	0,866	F(108, 1079)	0,8296

---

Test per ARCH di ordine 12

	LM	df	p-value
lag 1	40,266	36	0,2870
lag 2	73,748	72	0,4208
lag 3	120,657	108	0,1909
lag 4	154,864	144	0,2534
lag 5	194,946	180	0,2113
lag 6	223,229	216	0,3535
lag 7	255,763	252	0,4221
lag 8	307,563	288	0,2047
lag 9	339,773	324	0,2624
lag 10	382,136	360	0,2023
lag 11	419,043	396	0,2041
lag 12	445,301	432	0,3190

---

Matrice di correlazione dei residui, C (3 x 3)

1,0000	-0,077449	-0,046608
-0,077449	1,0000	0,28990
-0,046608	0,28990	1,0000

Autovalori di C

0,708388  
0,977175  
1,31444

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Test di Doornik-Hansen

Chi-quadro(6) = 3,83493 [0,6990]

Figure 17: ARCH Effects

```

Insieme di vincoli
  1: b[1] = 1
  2: b[2] = -1
  3: b[3] = -1
Rango dello Jacobiano = 5, numero di parametri liberi = 3
Il modello è pienamente identificato
Basato sullo Jacobiano, df = 0
Log-verosimiglianza non vincolata (lu) = -2810,0181
Log-verosimiglianza vincolata (lr) = -2838,8329
2 * (lu - lr) = 57,6297
P(Chi-quadro(2) > 57,6297) = 3,06103e-013

Vettori di cointegrazione

pa      1,0000
e       -1,0000
pb       -1,0000

Alfa (vettori di aggiustamento) (errori standard tra parentesi)

pa      -0,29446
        (0,011580)
e       -0,0034977
        (0,027777)
pb       -0,019391
        (0,0094199)

```

Figure 18: Restricted VECM