



**POLITECNICO**  
MILANO 1863

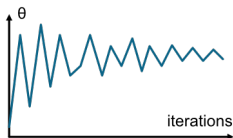
# Coupled Markov chains with applications to Approximate Bayesian Computation for model based clustering

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# Introduction

## A complex problem

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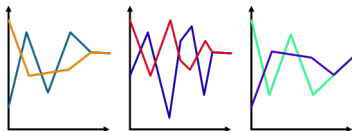


likelihood

intractable



Unbiased Markov chain  
Monte Carlo methods with  
couplings



Approximate Bayesian  
Computation



# Approximate Bayesian Computation

*Inputs:*

- a target posterior density  $\pi(\theta|y_{obs}) \propto p(y_{obs}|\theta)\pi(\theta)$ , consisting of a prior distribution  $\pi(\theta)$  and a procedure of generating data under the model  $p(y_{obs}|\theta)$ ;
- a Markov proposal density  $g(\theta, \theta')=g(\theta'|\theta)$ ;
- an integer  $N > 0$ ;
- a kernel function  $K_h(u)$  and a scale parameter  $h > 0$ ;
- a low dimensional vector of summary statistics  $s = S(y)$ .

*Initialise:*

repeat:

- ① choose an initial parameter vector  $\theta^{(0)}$  from the support of  $\pi(\theta)$ ;
- ② generate  $y^{(0)} \sim p(y|\theta^{(0)})$  from the model and compute summary statistics  $s^{(0)} = S(y^{(0)})$ , until  $K_h(\|s^{(0)} - s_{obs}\|) > 0$ .

*Sampling* for  $i = 1, \dots, N$ :

- ① generate candidate vector  $\theta' \sim g(\theta^{(i-1)}, \theta)$  from the proposal density  $g$ ;
- ② generate  $y' \sim p(y|\theta')$  from the model and compute summary statistics  $s' = S(y')$ ;
- ③ with probability

$$\min\left\{1, \frac{K_h(\|s' - s_{obs}\|)\pi(\theta')g(\theta', \theta^{(i-1)})}{K_h(\|s^{(i-1)} - s_{obs}\|)\pi(\theta^{(i-1)})g(\theta^{(i-1)}, \theta')}\right\}$$

set  $(\theta^{(i)}, s^{(i)}) = (\theta', s')$ . Otherwise set  $(\theta^{(i)}, s^{(i)}) = (\theta^{(i-1)}, s^{(i-1)})$ .

*Output*:

- a set of correlated parameter vectors  $\theta^{(1)}, \dots, \theta^{(N)}$  from a Markov chain with stationary distribution  $\pi_{ABC}(\theta|S_{obs})$ .

**Summary statistic:**

Sample mean

**Distance:**

2-norm of the difference.

**Kernel:**

$$K(u) = \frac{1}{\sqrt{2\pi}} e^{-\frac{1}{2}u^2}, \quad K_h(u) = \frac{K(\frac{u}{h})}{h}.$$

## Model

$$Y_i | \mu \stackrel{iid}{\sim} \mathcal{N}(\mu, \sigma_{obs}^2)$$

$$\mu \sim \mathcal{N}(\mu_0, \sigma_0^2)$$

$$\mu_0 = 8, \quad \sigma_0^2 = 4$$

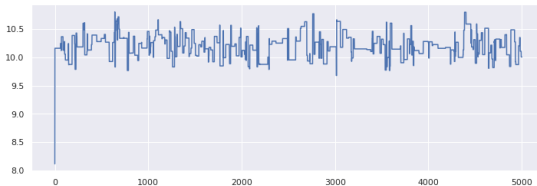
## Dataset

100 samples generated from a Gaussian distribution:

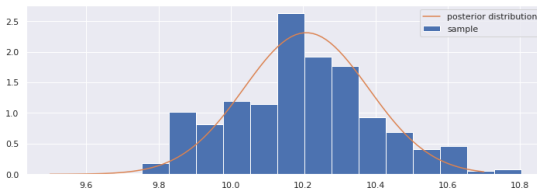
$$Y_{obs} \sim \mathcal{N}(\mu_{obs}, \sigma_{obs}^2)$$

$$\mu_{obs} = 10, \quad \sigma_{obs}^2 = 3$$

### Sampling



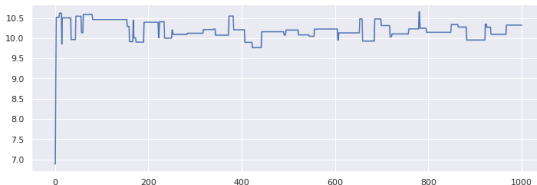
### Sampling histogram with real distribution



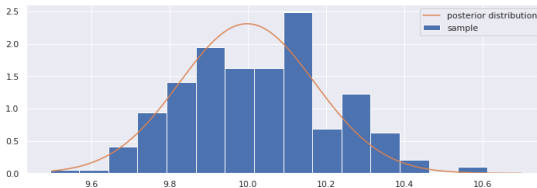



The same model using as summary statistic a vector of 10 quantiles:

**Sampling**

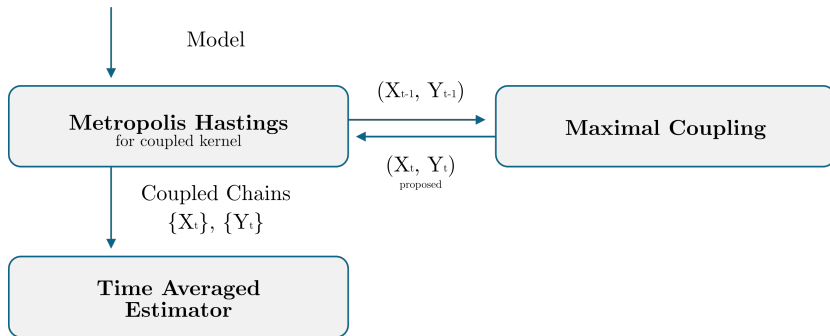


**Sampling histogram with real distribution**





# Unbiased Markov chain Monte Carlo methods with couplings



## Time-averaged estimator

- ① draw  $X_0$  and  $Y_0$  from an initial distribution  $\pi_0$  and draw  $X_1 \sim P(X_0, \cdot)$ ;
- ② set  $t = 1$ : while  $t < \max\{m, \tau\}$  and:
  - a draw  $(X_{t+1}, Y_t) \sim \bar{P}\{(X_t, Y_{t-1}), \cdot\}$ ;
  - b set  $t \leftarrow t + 1$ ;
- ③ compute the time-averaged estimator:

$$H_{k:m}(X, Y) = \frac{1}{m - k + 1} \sum_{l=k}^m h(X_l) + \sum_{l=k+1}^{\tau-1} \min(1, \frac{l - k}{m - k + 1}) \{h(X_l) - h(Y_{l-1})\}.$$

- ① sample  $(X^*, Y^*) | (X_t, Y_{t-1})$  from a maximal coupling of  $q(X_t, \cdot)$  and  $q(Y_{t-1}, \cdot)$ ;
- ② sample  $U \sim \mathcal{U}([0, 1])$ ;
- ③ if

$$U \leq \min \left\{ 1, \frac{\pi(X^*)q(X^*, X_t)}{\pi(X_t)q(X_t, X^*)} \right\}$$

then  $X_{t+1} = X^*$ ; otherwise  $X_t = X_{t-1}$ ;

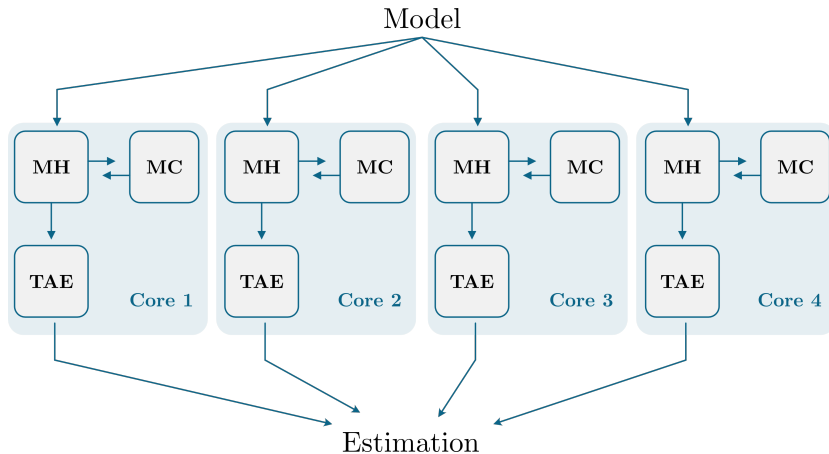
- ④ if

$$U \leq \min \left\{ 1, \frac{\pi(Y^*)q(Y^*, Y_t)}{\pi(Y_t)q(Y_t, Y^*)} \right\}$$

then  $Y_{t+1} = Y^*$ ; otherwise  $Y_t = Y_{t-1}$ .

Set  $p = \mathcal{N}(X_{t-1}, 1)$  and  $q = \mathcal{N}(Y_{t-1}, 1)$ , then:

- ① sample  $X_t \sim p$ ;
- ② sample  $W|X_t \sim \mathcal{U}\{[0, p(X_t)]\}$ ;
- ③ if  $W \leq q(X_t)$  then output  $(X_t, X_t)$ , otherwise:
  - ① sample  $Y_t \sim q$ ;
  - ② sample  $W^*|Y_t \sim \mathcal{U}\{[0, q(Y_t)]\}$  until  $W^* > p(Y_t)$  and output  $(X_t, Y_t)$ .



## Model

$$Y_i | \mu \stackrel{iid}{\sim} \mathcal{N}(\mu, \sigma_{obs}^2)$$

$$\mu \sim \mathcal{N}(\mu_0, \sigma_0^2)$$

$$\mu_0 = 8, \quad \sigma_0^2 = 4$$

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100 samples generated from a Gaussian distribution:

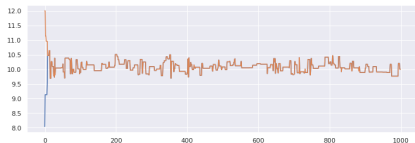
$$Y_{obs} \sim \mathcal{N}(\mu_{obs}, \sigma_{obs}^2)$$

$$\mu_{obs} = 10, \quad \sigma_{obs}^2 = 3$$

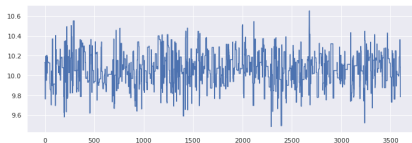


$$\mathcal{N}(\mu_n, \sigma_n^2), \quad \mu_n = \frac{1}{\frac{1}{\sigma_0^2} + \frac{n}{\sigma_{obs}^2}} \cdot \left( \frac{\mu_0}{\sigma_0^2} + \frac{\sum y_{obs}}{\sigma_{obs}^2} \right) \simeq 42.99, \quad \sigma_n^2 = \frac{1}{\frac{1}{\sigma_0^2} + \frac{n}{\sigma_{obs}^2}} \simeq 0.025$$

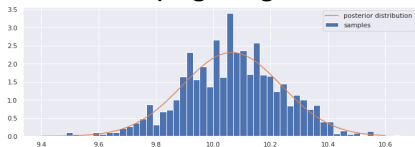
Coupled chains



Complete sampling




Sampling histogram



**Time Averaged Estimators mean:**

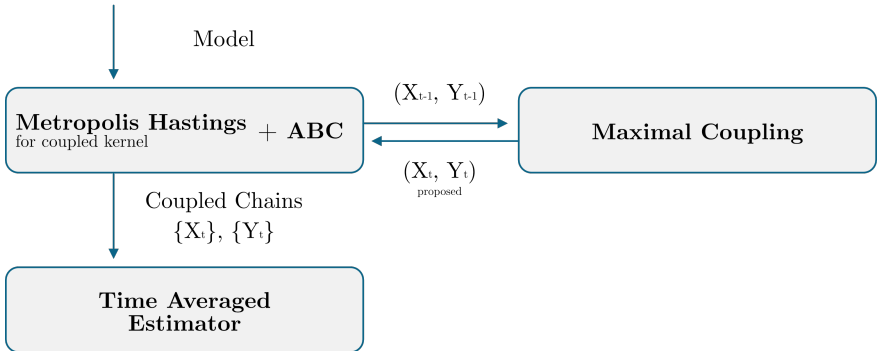
$$\mathbb{E}[H_{k:m}(X, Y)] = 42.9498$$



# The complete method: MCMC + Couplings + ABC

# The complete method: MCMC + Couplings + ABC Implementation

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## Metropolis Hastings with couplings and ABC

- ① Compute  $s_{obs} = S(y_{obs})$ ;
- ② generate  $\theta_x^{(0)} \sim \pi(\mu)$  and  $\theta_y^{(0)} \sim \pi(\mu)$  from prior density;
- ③ generate with a maximal coupling two samples of N observations such that  $y_{1i} \sim \mathcal{N}(\theta_x^{(0)}, \sigma_{obs}^2)$  and  $y_{2j} \sim \mathcal{N}(\theta_y^{(0)}, \sigma_{obs}^2)$ ;
- ④ compute  $s_x^{(0)} = S(y_1)$  and  $s_y^{(0)} = S(y_2)$ ;
- ⑤ until  $Kh(||s_x^{(0)} - s_{obs}||) > 0$ :
  - ▶ generate  $\theta_x^{(0)} \sim \pi(\mu)$  from prior density;
  - ▶ generate a sample of N observations such that  $y_{1i} \sim \mathcal{N}(\theta_x^{(0)}, \sigma_{obs}^2)$ ;
  - ▶ compute  $s_x^{(0)} = S(y_1)$ ;
- ⑥ until  $Kh(||s_y^{(0)} - s_{obs}||) > 0$ :
  - ▶ generate  $\theta_y^{(0)} \sim \pi(\mu)$  from prior density;
  - ▶ generate a sample of N observations such that  $y_{2j} \sim \mathcal{N}(\theta_y^{(0)}, \sigma_{obs}^2)$ ;
  - ▶ compute  $s_y^{(0)} = S(y_2)$ ;

⑧ for  $i = 1, \dots, N$ :

- ▶ generate  $[\theta_x^{(i)}, \theta_y^{(i)}]$  from a maximal coupling given  $[\theta_x^{(i-1)}, \theta_y^{(i-1)}]$ ;
- ▶ generate from a maximal coupling two samples of  $N$  observations  $y_1 \sim p(y|\theta_x^{(i)})$  and  $y_2 \sim p(y|\theta_y^{(i)})$ ;
- ▶ compute  $s_x^{(i)} = S(y_1)$  and  $s_y^{(i)} = S(y_2)$ ;
- ▶ accept  $\theta_x^{(i)}$  with probability

$$\frac{Kh(\|s_x^{(i)} - s_{obs}\|)\pi(\theta_x^{(i)})}{Kh(\|s_x^{(i-1)} - s_{obs}\|)\pi(\theta_x^{(i-1)})}$$

and accept  $\theta_y^{(i)}$  with probability

$$\frac{Kh(\|s_y^{(i)} - s_{obs}\|)\pi(\theta_y^{(i)})}{Kh(\|s_y^{(i-1)} - s_{obs}\|)\pi(\theta_y^{(i-1)})}.$$

As output we get two sets of parameter vectors:

$$\theta_x^{(1)}, \dots, \theta_x^{(N)} \sim \pi_{ABC}(\theta|y_{obs});$$

$$\theta_y^{(1)}, \dots, \theta_y^{(N)} \sim \pi_{ABC}(\theta|y_{obs}).$$

**Summary statistic:**

Sample mean

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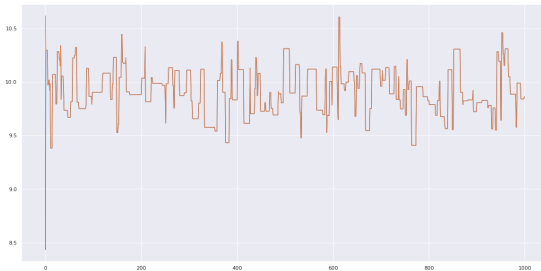
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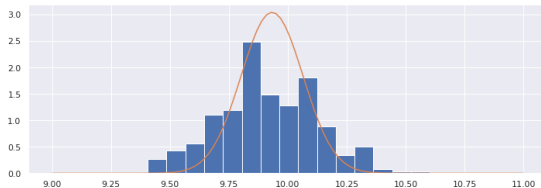
$$\mu_{obs} = 10, \quad \sigma_{obs}^2 = 3$$



### Coupled chains



### Sampling histogram with real distribution





# Conclusions

The next step will be the conclusion of the **multivariate implementation** the MCMC with couplings and approximate bayesian computation.

Further steps will be testing on more complex data.

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