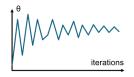


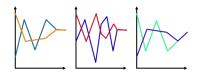
Coupled Markov chains with applications to Approximate Bayesian Computation for model based clustering

E. Bertoni, M. Caldarini, F. Di Filippo, G. Gabrielli, E. Musiari 10 January 2022





Unbiased Markov chain Monte Carlo methods with couplings



Approximate Bayesian
Computation

Unbiased Markov chain Monte Carlo methods with couplings

- **1** draw X_0 and Y_0 from an initial distribution π_0 and draw $X_1 \sim P(X_0, \cdot)$;
- 2 set t = 1: while $t < \max\{m, \tau\}$ and:
 - a draw $(X_{t+1}, Y_t) \sim \bar{P}\{(X_t, Y_{t-1}), \cdot\};$
 - b set $t \leftarrow t + 1$;
- 3 compute the time-averaged estimator:

$$H_{k:m}(X,Y) = \frac{1}{m-k+1} \sum_{l=k}^{m} h(X_l) + \sum_{l=k+1}^{\tau-1} \min(1, \frac{l-k}{m-k+1}) \{h(X_l) - h(Y_{l-1})\}.$$

Metropolis–Hasting algorithm allow us to calculate the coupled kernel $\bar{P}\{(X_t, Y_{t-1}), \cdot\}$:

- sample $(X^*, Y^*)|(X_t, Y_{t-1})$ from a maximal coupling of $q(X_t, \cdot)$ and $q(Y_{t-1}, \cdot)$;
- 2 sample $U \sim \mathcal{U}([0,1])$;
- 3 if

$$U \leq \min \left\{1, \frac{\pi(X^*)q(X^*, X_t)}{\pi(X_t)q(X_t, X^*)}\right\}$$

then $X_{t+1} = X^*$; otherwise $X_t = X_{t-1}$;

4 if

$$U \leq \min \left\{ 1, \frac{\pi(\mathsf{Y}^{\star})q(\mathsf{Y}^{\star}, \mathsf{Y}_{t})}{\pi(\mathsf{Y}_{t})q(\mathsf{Y}_{t}, \mathsf{Y}^{\star})} \right\}$$

then $Y_{t+1} = Y^*$; otherwise $Y_t = Y_{t-1}$.

The algorithm:

Set
$$p = \mathcal{N}(X_{t-1}, 1)$$
 and $q = \mathcal{N}(Y_{t-1}, 1)$,

- **1** sample $X_t \sim p$;
- 2 sample $W|X_t \sim \mathcal{U}\{[0, p(X_t)]\};$
- 3 if $W \le q(X_t)$ then output (X_t, X_t) , otherwise:
 - 1 sample $Y_t \sim q$;
 - 2 sample $W^*|Y_t \sim \mathcal{U}\{[0, q(Y_t)]\}$ until $W^* > p(Y_t)$ and output (X_t, Y_t) .

Model

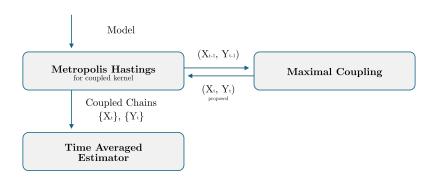
$$Y_i | \mu \stackrel{\text{iid}}{\sim} \mathcal{N}(\mu, \sigma_{\text{obs}}^2)$$
$$\mu \sim \mathcal{N}(\mu_0, \sigma_0^2)$$
$$\mu_0 = 38, \quad \sigma_0^2 = 4$$

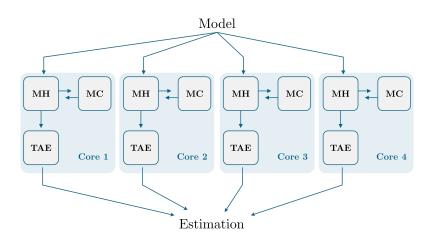
Dataset

1000 samples generated from a Gaussian distribution:

$$Y_{\text{obs}} \sim \mathcal{N}(\mu_{\text{obs}}, \sigma_{\text{obs}}^2)$$

 $\mu_{\text{obs}} = 43, \quad \sigma_{\text{obs}}^2 = 5$

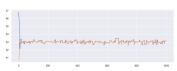




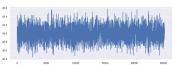
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$$\mathcal{N}(\mu_{\mathrm{n}},\sigma_{\mathrm{n}}^2), \quad \mu_{\mathrm{n}} = \frac{1}{\frac{1}{\sigma_{\mathrm{o}}^2} + \frac{n}{\sigma_{\mathrm{obs}}^2}} \cdot \left(\frac{\mu_0}{\sigma_0^2} + \frac{\sum \mathrm{y}_{\mathrm{obs}}}{\sigma_{\mathrm{obs}}^2}\right), \quad \sigma_{\mathrm{n}}^2 = \frac{1}{\frac{1}{\sigma_0^2} + \frac{n}{\sigma_{\mathrm{obs}}^2}}$$

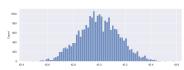
Coupled chains



Complete sampling



Sampling histogram



Time Averaged Estimators mean:

$$\mathbb{E}[H_{k:m}(X,Y)] = 42.9498$$

Approximate Bayesian Computation

Inputs:

- a target posterior density $\pi(\theta|y_{obs}) \propto p(y_{obs}|\theta)\pi(\theta)$, consisting of a prior distribution $\pi(\theta)$ and a procedure of generating data under the model $p(y_{obs}|\theta)$;
- a Markov proposal density $g(\theta, \theta') = g(\theta'|\theta)$;
- \blacksquare an integer N > 0;
- **a** a kernel function $K_h(u)$ and a scale parameter h > 0;
- lacksquare a low dimensional vector of summary statistics s = S(y).

Initialise:

repeat:

- **1** choose an initial parameter vector $\theta^{(0)}$ from the support of $\pi(\theta)$;
- 2 generate $\mathbf{y}^0 \sim p(\mathbf{y}|\theta^{(0)})$ from the model and compute summary statistics $\mathbf{s}^0 = \mathbf{S}(\mathbf{y}^{(0)})$, until $\mathbf{K}_h(\parallel \mathbf{s}^{(0)} \mathbf{s}_{obs} \parallel) > 0$.

ABC Metropolis Hastings

Sampling for i = 1, ..., N:

- **1** generate candidate vector $\theta' \sim g(\theta^{(i-1)}, \theta)$ from the proposal density g;
- 2 generate $y' \sim p(y|\theta')$ from the model and compute summary statistics s' = S(y');
- 3 with probability

$$\min\{1, \frac{\mathcal{K}_{h}(\parallel \mathbf{s}' - \mathbf{s}_{obs} \parallel) \pi(\theta') \mathbf{g}(\theta', \theta^{(i-1)})}{\mathcal{K}_{h}(\parallel \mathbf{s}^{(i-1)} - \mathbf{s}_{obs} \parallel) \pi(\theta^{(i-1)}) \mathbf{g}(\theta^{(i-1)}, \theta')}\}$$

set
$$(\theta^{(i)}, \mathbf{s}^{(i)}) = (\theta', \mathbf{s}')$$
. Otherwise set $(\theta^{(i)}, \mathbf{s}^{(i)}) = (\theta^{(i-1)}, \mathbf{s}^{(i-1)})$.

Output:

a set of correlated parameter vectors $\theta^{(1)},...,\theta^{(N)}$ from a Markov chain with stationary distribution $\pi_{ABC}(\theta|S_{obs})$.

Summary statistic:

Mean

Distance:

Norm (absolute value) of the difference.

Kernel:

$$\frac{1}{\sqrt{2\pi}}\mathbf{e}^{-\frac{1}{2}u^2}.$$

Same as previous:

Model

$$\begin{aligned} \mathbf{Y}_{\mathit{i}} | \mu &\stackrel{\mathit{iid}}{\sim} \mathcal{N}(\mu, \sigma_{\mathit{obs}}^2) \\ \mu &\sim \mathcal{N}(\mu_0, \sigma_0^2) \\ \mu_0 &= 38, \quad \sigma_0^2 = 4 \end{aligned}$$

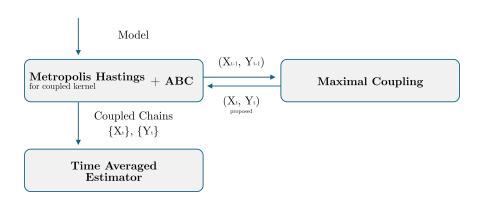
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The complete method: Monte Carlo Markov chain method with couplings and approximate bayesian computation



Multivariate study case

Conclusions

The next step will be the conclusion of the **separate multivariate implementation** of both solution to be tested on simulated data and the **parallelized multivariate implementation**.

Further steps will be testing on more complex data.

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