

Comparing distributions: ℓ_1 geometry improves kernel two-sample testing

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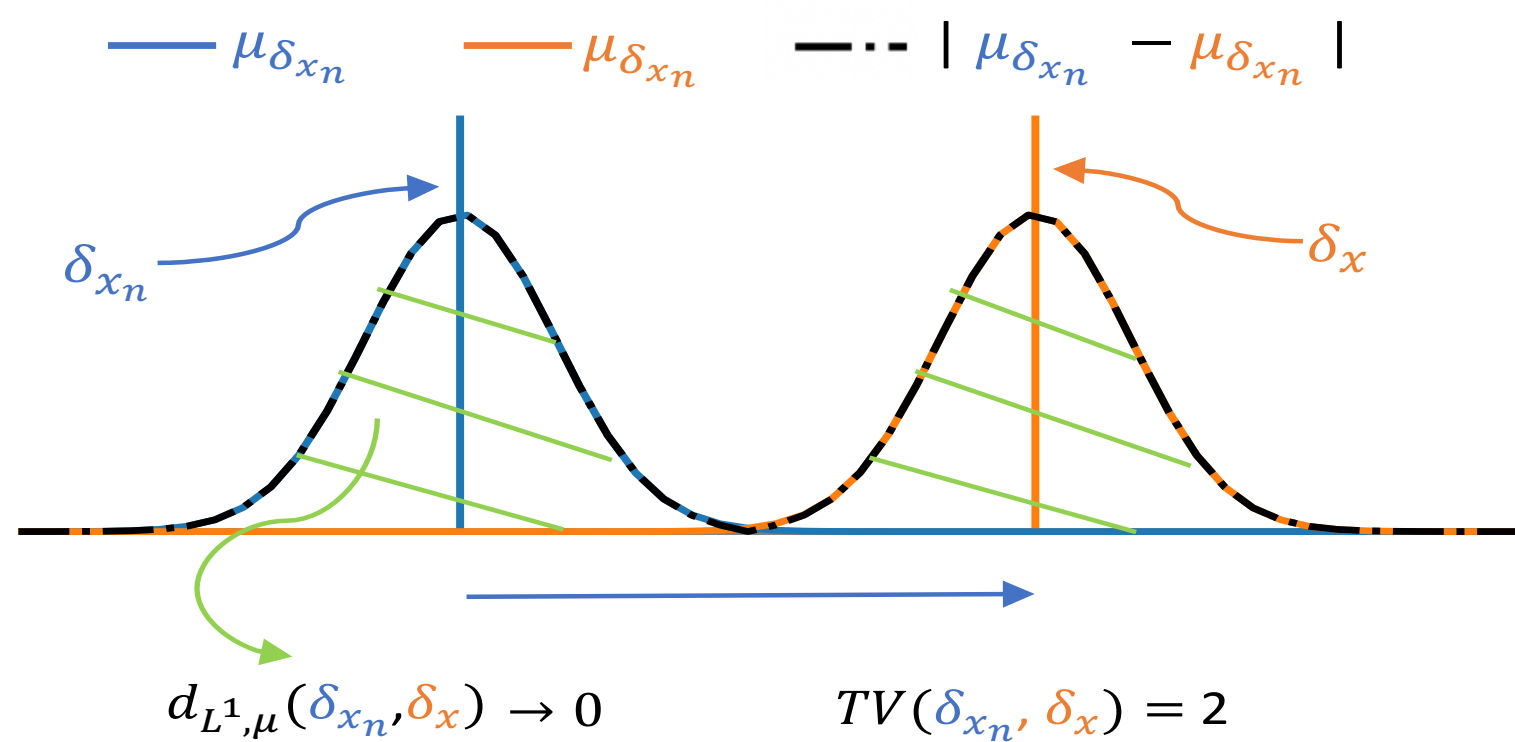
Overview

Problem: Are two sets of observations drawn from the same distribution?

Contributions:

- We exhibit a family of L^p -based metrics which **metrize the weak convergence**.
- We derive linear-time, nonparametric, a.s - consistent **L^1 -based two sample tests**.
- We show L^1 geometry provides **better power** than its L^2 counterpart.
- We maximize a **lower bound** on the test power and learn distinguishing features between distributions.

Weak Convergence



Theorem: Let k a characteric and bounded kernel. For all $p \geq 1$,

$$d_{L^p, \mu}(P, Q) := \left(\int_t |\mu_P(t) - \mu_Q(t)|^p d\Gamma(t) \right)^{1/p}$$

where $\mu_P(t) := \int_t k(x, t) dP(x)$ is a metric which metrize the weak convergence.

Sketch of proof:

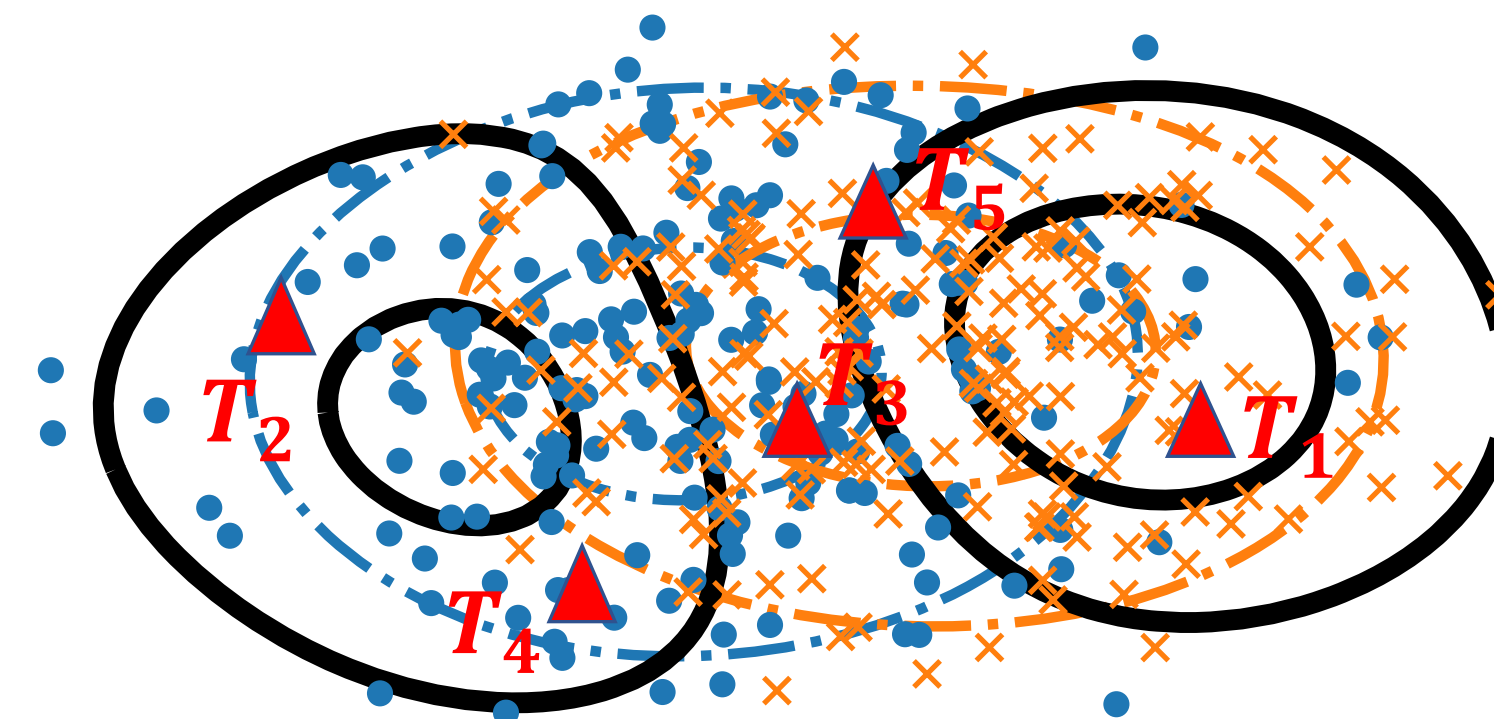
- Integral Operator: $T_k: f \in L_2^{d\Gamma}(\mathbb{R}^d) \rightarrow \int_{t \in \mathbb{R}^d} k(t, \cdot) f(t) d\Gamma$
- Unit Ball of $L_\infty^{d\Gamma}(\mathbb{R}^d)$: $B_\infty^{d\Gamma} := \{f: \sup |f(x)| \leq 1 \Gamma\text{-a.s}\}$
- IPM formulation: $d_{L^1, \mu}(P, Q) = \sup_{f \in T_k(B_\infty^{d\Gamma})} \{E_P(f(X)) - E_Q(f(Y))\}$

Mean Embedding test

- Test $H_0: P = Q$ vs $H_1: P \neq Q$:
- Samples: $X := \{x_i\}_{i=1}^n \sim P$ and $Y := \{y_i\}_{i=1}^n \sim Q$
- Empirical ME: $\mu_X(T) := \frac{1}{n} \sum_{i=1}^n k_\sigma(x_i, T)$
- k_σ the Gaussian kernel of width σ
- Test locations: $\{T_i\}_{i=1}^J \sim \Gamma$
- Test statistic:

$$\left(\hat{d}_{\ell_p, \mu}(X, Y) \right)^p := n^{\frac{p}{2}} \sum_{i=1}^J |\mu_X(T_i) - \mu_Y(T_i)|^p$$

$$\text{--- } |\mu_P - \mu_Q| \quad \text{--- } \mu_P \quad \text{--- } \mu_Q \quad \text{--- } \hat{\mu} \quad \times$$



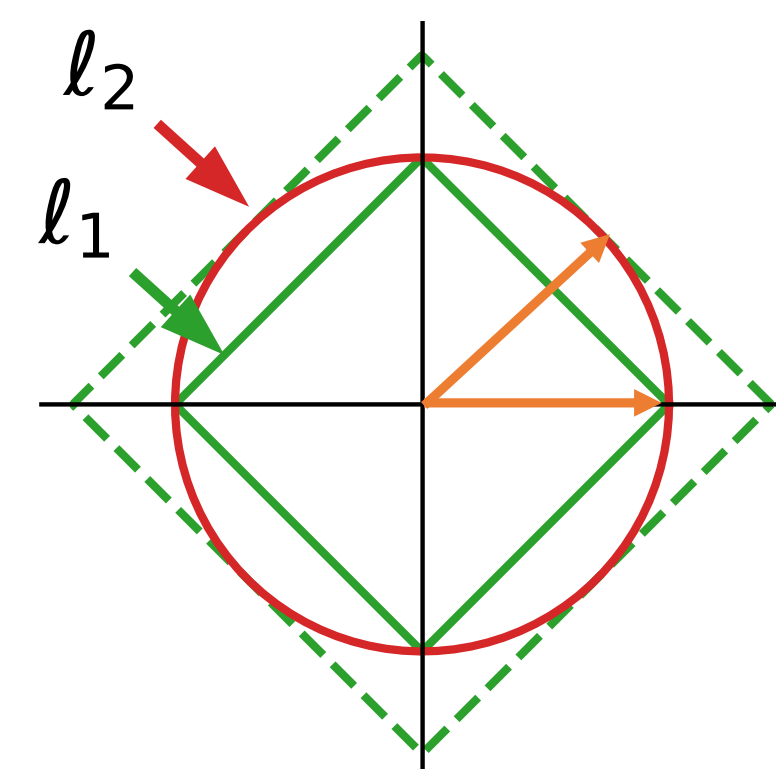
Test of level α : Compute $\left(\hat{d}_{\ell_p, \mu}(X, Y) \right)^p$ and reject H_0 if $\left(\hat{d}_{\ell_p, \mu}(X, Y) \right)^p > \zeta_{\alpha, p} = 1 - \alpha$ quantile of the null distribution.

Why $\ell_1 \gg \ell_2$?

Definition: (Analytic kernel) A positive definite kernel k is analytic if for all $x \in \mathbb{R}^d$, the feature map $x \rightarrow k(x, \cdot)$ is an analytic function on \mathbb{R}^d .

Proposition: Let $\delta > 0$. Under the alternative hypothesis H_1 , almost surely there exists $N \geq 1$, such that for all $n \geq N$, with a probability of $1 - \delta$:

$$\left(\hat{d}_{\ell_2, \mu}(X, Y) \right)^2 > \zeta_{\alpha, 2} \Rightarrow \hat{d}_{\ell_1, \mu}(X, Y) > \zeta_{\alpha, 1}$$



Normalized Tests

Remark: Under H_0 , $\hat{d}_{\ell_1, \mu}(X, Y)$ converge to a sum of correlated Nakagami variables.

Normalized Mean Embedding (ME) Test:

$$L1\text{-ME}[X, Y] := \|\sqrt{n} \Sigma_n^{-\frac{1}{2}} S_n\|_1$$

- $S_n := \frac{1}{n} \sum_{i=1}^n Z_X^i - Z_Y^i$
- $\Sigma_n := \widehat{\text{cov}}(Z_X) + \widehat{\text{cov}}(Z_Y)$
- $Z_X^i := (k_\sigma(x_i, T_1), \dots, k_\sigma(x_i, T_J))$

Proposition: Under H_0 , $L1\text{-ME}[X, Y]$ is a.s asymptotically distributed as a sum of J i.i.d Nakagami variables of parameter $m = \frac{1}{2}$ and $\varpi = \frac{1}{2}$.

Normalized Smooth Characteristic Function (SCF) Test:

$$L1\text{-SCF}[X, Y] := \|\sqrt{n} \Sigma_n^{-\frac{1}{2}} S_n\|_1$$

- $Z_X^i := (\cos(x_i^T T_1) f(x_i), \sin(x_i^T T_1) f(x_i), \dots, \sin(x_i^T T_J) f(x_i))$
- f is the inverse Fourier transform of k_σ .

Optimization Procedure

Regularized Statistic:

$$L1\text{-ME}[X, Y] := \|\sqrt{n} (\Sigma_n + \gamma_n)^{-1/2} S_n\|_1$$

- $\gamma_n \rightarrow 0$

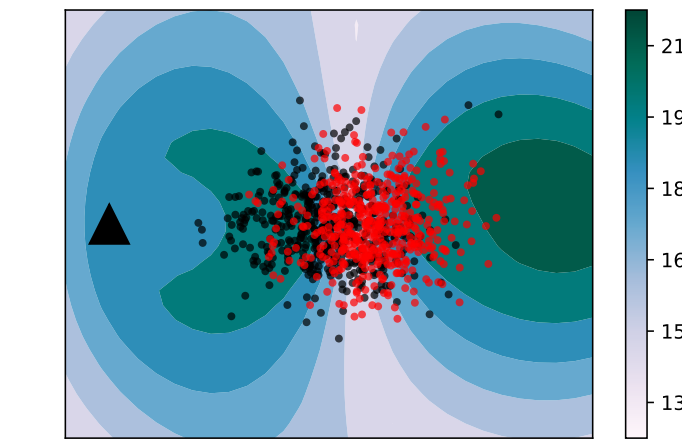
Proposition: The test power $P(L1\text{-ME}[X, Y] > \epsilon)$ of the the $L1\text{-ME}$ test satisfies $P(L1\text{-ME}[X, Y] > \epsilon) \geq L(\lambda_n)$ where $L(\lambda_n)$ is an increasing function of λ_n and goes to 1 when n goes to infinity.

- $\lambda_n := \|\sqrt{n} \Sigma^{-\frac{1}{2}} S\|_1$ is the population counterpart of $L1\text{-ME}[X, Y]$.

Optimization Procedure:

- Optimize $\{T_i\}_{i=1}^J, \sigma = \argmax L(\lambda_n)$
- Estimation of λ_n on a separate training set.

Informative Features



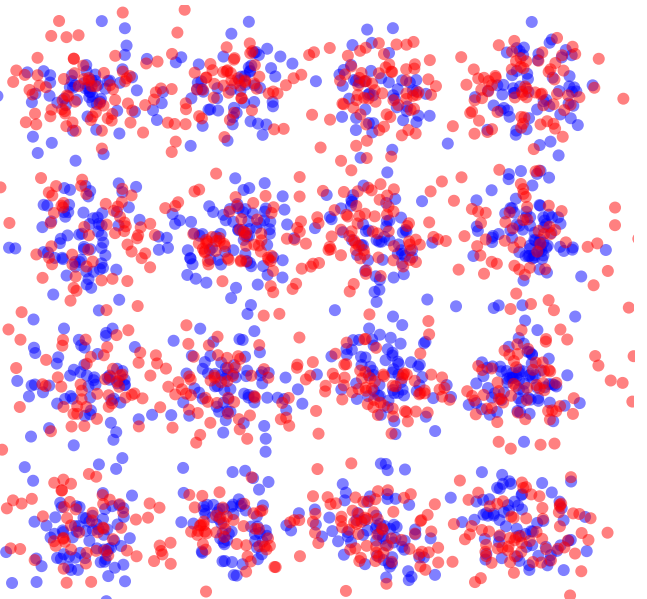
Contour plot of $L1\text{-ME}[X, Y]$ as a function of T_2 with $J = 2$ and T_1 fixed.

- $P \sim N([0, 0], I_2)$
- $Q \sim N([0, 1], I_2)$
- $L1\text{-ME}[X, Y]$ detects the differences.

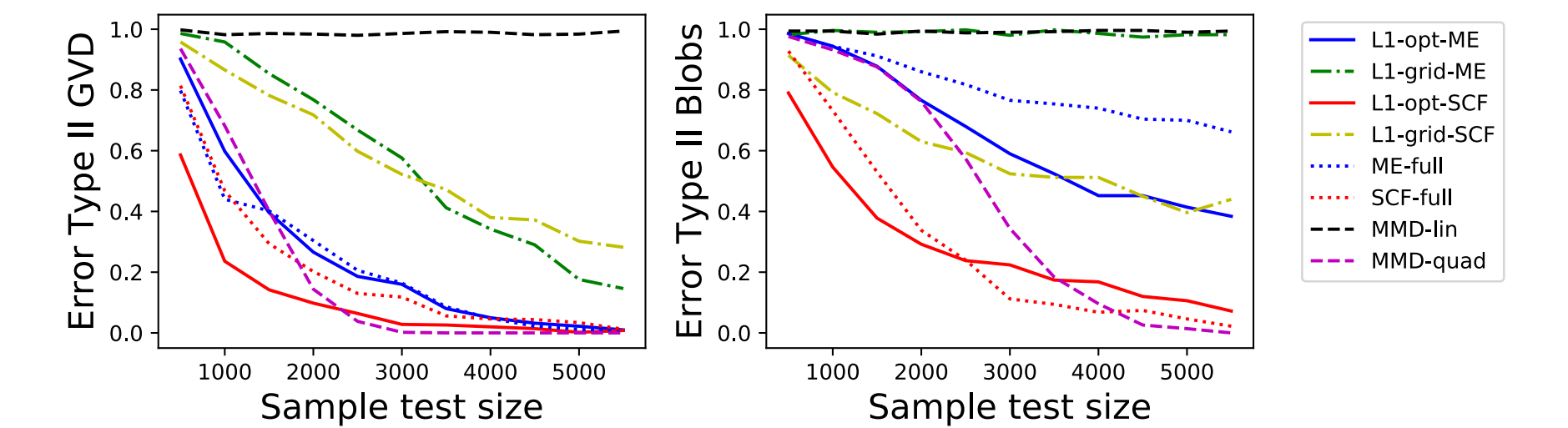
Test Power: Synthetic Problems

- Test Power vs. Sample Size

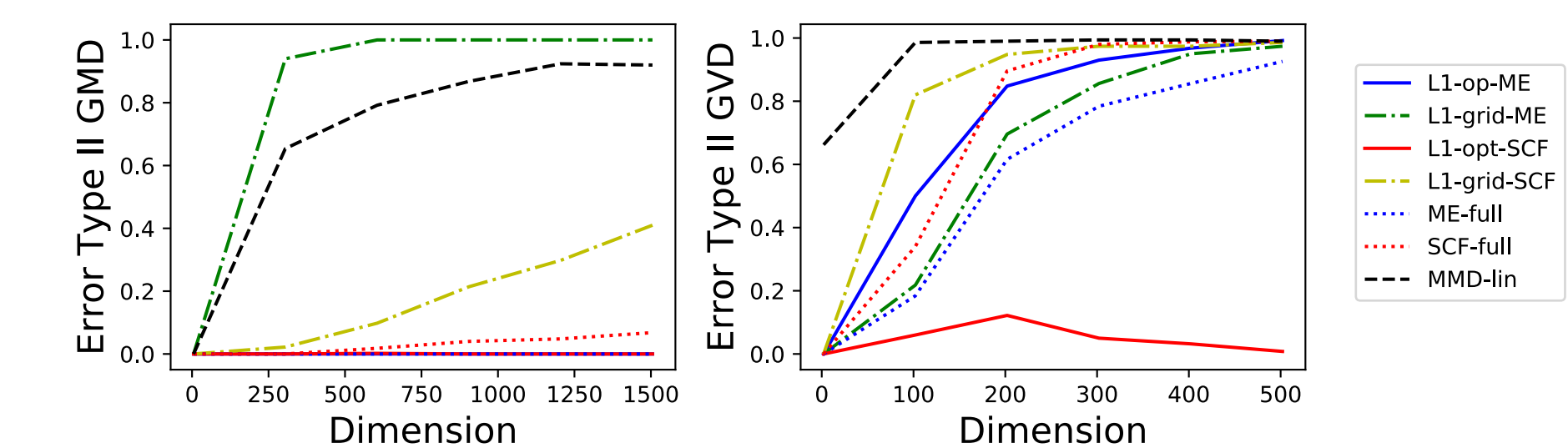
Data	P	Q
GMD	$\mathcal{N}(0, I_d)$	$\mathcal{N}((1, 0, \dots, 0)^T, I_d)$
GVD	$\mathcal{N}(0, I_d)$	$\mathcal{N}(0, \text{diag}(2, 1, \dots, 1))$
Blobs	Mixture of 16 Gaussians in \mathbb{R}^2	



- L1-opt-ME, L1-opt-SCF:** Proposed Methods
- L1-grid-ME, L1-grid-SCF:** Random settings
- ME-full, SCF-full: Optimized ℓ_2 -based methods
- MMD-quad, MMD-lin: Quadratic and linear-time MMD tests



- Test Power vs. Dimension



Higgs Dataset

Higgs Dataset: $d = 4$, $J = 3$. Plot of Type-II error for ℓ_1 and ℓ_2 based test.

- Optimized tests outperform their random versions.
- ℓ_1 norm provides better power.

