Model Identification and Data Analysis

Matteo Secco

March 9, 2021

Contents

I Prediction **4**

1 Probability Recall **4**
1.1 Random Vectors 4
1.2 Random processes 4
1.3 Important process classes 4

2 Spectral Analysis **6**
2.1 Foundamentals 6
2.2 Fundamental theorem of Spectral Analysis 6
2.3 Canonical representation of a Stationary Process 7

3 Moving Average Processes **8**
3.1 MA(1): 8
3.2 MA(n) 9
3.3 MA(∞) 9
3.4 Well definition of an MA(∞) 10

4 Auto Regressive Processes **11**
4.1 AR(1) 11
4.2 AR(n) 12

5 ARMA Processes **14**

6 Prediction problem **15**
6.1 Fake problem 15
6.2 True Problem 15
6.3 Prediction with eXogenous variables 16
6.3.1 ARX model 16
6.3.2 ARMAX model 16

II Identification **17**

7 Prediction Error Minimization **18**
7.1 Least Square Error 18
7.2 Estimation of mean, variance, spectrum 18
7.2.1 Mean 19
7.2.2 Covariance (a) 19
7.2.3 Covariance (b) 19
7.2.4 Spectrum 19

III Cheatsheet **20**

A Probability Recall **20**

B Spectral analysis **21**

C Moving Average MA(n) **22**
C.1 MA(∞) 22

D Auto Regressive AR(n) **23**

E Known predictors **24**

Part I
Prediction

1 Probability Recall

1.1 Random Vectors

Variance $Var[v] = E[(v - E[v])^2]$

Cross-Variance $Var[v, u] = E[(v - E[v])(u - E[u])]$

Variance Matrix $\begin{vmatrix} Var[v_1] & . & . & Var[v_1, v_k] \\ . & . & & . \\ . & & . & . \\ Var[v_k, v_1] & . & . & Var[v_k] \end{vmatrix}$

Covariance coefficient $\delta[i, j] = \frac{Var[i,j]}{\sqrt{Var[i]}\sqrt{Var[j]}}$
$\delta[i, j] = 0 \implies$ i, j uncorrelated

$abs\delta[i, j] = 1 \implies i = \alpha j$

1.2 Random processes

$v(t, s)|$ t time instant, s expetiment outcome (generally given)

Mean $m(t) = E[v(t, s)]$

Variance $\lambda^2(t) = Var[v(t)]$

Covariance function $\gamma(t_1, t_2) = E[(v(t_1) - m(t_1))(v(t_2) - m(t_2))] = \gamma(t_2, t_1)$

Normalized Covariance Function $\rho(\tau) = \frac{\gamma(\tau)}{\gamma(0)}$
$\forall$ stationary processes: $|\rho(\tau)| \leq 1 \quad \forall\tau$

1.3 Important process classes

Stationary process

- $m(t) = m$ constant
- $\lambda^2(t) = \lambda^2$ constant
- $\gamma(t_1, t_2) = f(t_2 - t_1) = \gamma(\tau)$ covariance depends only on time difference τ

$|\gamma(\tau)| \leq \gamma(0) \quad \forall\tau$

White noise $\eta(t) \sim WN(m, \lambda^2)$

- Stationary process
- $\gamma(\tau) = 0 \quad \forall \tau \neq 0$

$v(t) = \alpha\eta(t) + \beta \quad \eta(t) \sim WN(0, \lambda^2) \qquad \Longrightarrow \qquad v(t) \sim WN(\beta, \alpha^2\lambda^2)$

2 Spectral Analysis

2.1 Foundamentals

Spectrum

$$\Gamma(\omega) = \overbrace{F(\gamma(\tau))}^{\text{Fourier transform}} = \sum_{\tau=-\infty}^{+\infty} \gamma(\tau) \cdot e^{-j\omega\tau}$$

Euler formula $\Gamma(\omega) = \gamma(0) + 2cos(\omega)\gamma(1) + 2cos(2\omega)\gamma(2) + \dots$

Spectrum properties

- $\Gamma : \mathbb{R} \to \mathbb{R}$
- Γ is periodic with $T = 2\pi$
- Γ is even $[\Gamma(-\omega) = \Gamma(\omega)]$
- $\Gamma(\omega) \geq 0 \quad \forall\omega$

$\eta(t) \sim WN(0, \lambda^2) \quad \implies \quad \Gamma_\eta(\omega) = \gamma(0) = Var[\eta(t)] = \lambda^2$

Anti-Transform

$$\gamma(\tau) = \frac{1}{2\pi} \int_{-\pi}^{+\pi} \Gamma(\omega) e^{k\omega\tau} \, dw$$

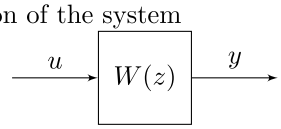
Complex spectrum

$$\phi(z) = \sum_{\tau=-\infty}^{+\infty} \omega(\tau) z^{-\tau}$$

$$\Gamma(\omega) = \Phi(e^{j\omega})$$

2.2 Fundamental theorem of Spectral Analysis

Fundamental theorem of Spectral Analysis allows to derive the (real and/or complex) spectrum of the output from the input and the transfer function of the system



$$\Gamma_{yy}(\omega) = |W(e^{j\omega})|^2 \quad \cdot \Gamma_{uu}(\omega)$$
$$\Phi_{yy}(z) = W(z)W(z^{-1}) \cdot \Phi_{uu}(z)$$

2.3 Canonical representation of a Stationary Process

A stationary process can be represented by an infinite number of transfer functions. The canonical representation is the transfer function $W(z)$ such that:

- Numerator and denominator have <u>same degree</u>
- Numerator and denominator are <u>monic</u> (highest grade coefficient is 1)
- Numerator and denominator are <u>coprime</u> ($W(z)$ cannot be simplified)
- numerator and denominator are <u>stable polynomials</u> (all poles and zeros of $W(z)$ are inside the unit disk)

3 Moving Average Processes

Given $\eta(t) \sim WN(0, \lambda^2)$

3.1 MA(1):

Model

$$v(t) = c_0\eta(t) + c_1\eta(t-1)$$

Mean

$$\begin{aligned} E[v(t)] &= c_0 \cdot E[\eta(t)] + c_1 \cdot E[\eta(t)] \\ &= c_0 \cdot 0 \qquad + c_1 \cdot 0 \end{aligned}$$

$$\boxed{E[v(t)] = 0}$$

Variance

$$\begin{aligned} Var[v(t)] &= E[(v(t) \underbrace{-E[v(t)])^2}_{0}] \\ &= E[(v(t))^2] \\ &= E[(c_0 \cdot \eta(t)^2 \qquad +c_1 \cdot \eta(t-1))^2] \\ &= c_0^2 \cdot E[\eta(t)^2] \qquad +c_1^2 \cdot E[\eta(t-1)^2] + \underbrace{2c_0c_1 \cdot E[\eta(t)\eta(t-1)]}_{0} \\ &= c_0^2 \cdot E[\eta(t)^2] \qquad +c_1^2 \cdot E[\eta(t-1)^2] \\ &= c_0^2\lambda^2 \qquad +c_1^2\lambda^2 \end{aligned}$$

$$\boxed{Var[v(t)] = (c_0^2 + c_1^2)\lambda^2}$$

Covariance

$$\begin{aligned} \gamma(t_1, t_2) &= E[(v(t_1) - E[v(t_1)]) \qquad \cdot(v(t_2) - E[v(t_2)])] \\ &= E[(c_0\eta(t_1) + c_1\eta(t_1 - 1)) \ \cdot(c_0\eta(t_2) + c_1\eta(t_2 - 1))] \\ &= c_0^2E[\eta(t_1)\eta(t_2)] \qquad +c_1^2E[\eta(t_1 - 1)\eta(t_2 - 1) \\ &\qquad +c_0c_1E[\eta(t_1)\eta(t_2 - 1)] + c_0c_1E[\eta(t_1 - 1)\eta(t_2)] \end{aligned}$$

$$\boxed{\gamma(\tau) = \begin{cases} c_0^2\lambda^2 + c_1^2\lambda^2 & \text{if } \tau = 0 \\ c_0c_1\lambda^2 & \text{if } \tau = \pm 1 \\ 0 & \text{otherwise} \end{cases}}$$

3.2 MA(n)

Model

$$\begin{aligned} v(t) &= c_0\eta(t) + c_1\eta(t-1) + \cdots + c_n\eta(t-n) \\ &= (c_0 + c_1 z^{-1} + \cdots + c_n z^{-n})\eta(t) \end{aligned}$$

Transfer function

$$W(z) = c_0 + c_1 z^{-1} + \cdots + c_n z^{-n} = \frac{c_0 z^n + c_1 z^{n-1} + \cdots + c_n}{z^n}$$

All poles are in the complex origin

Mean

$$E[v(t)] = (c_0 + c_1 + \cdots + c_n)\underbrace{E[\eta(t)]}_{0}$$

$$\boxed{E[v(t) = 0}$$

Covariance function

$$\boxed{\gamma(\tau) = \begin{cases} \lambda^2 \cdot \sum_{i=0}^{n-\tau} c_i c_{i-\tau} & |\tau| \leq n \\ 0 & \text{otherwise} \end{cases}}$$

example

$$\begin{aligned} &\gamma(0) = (c_0^2 + c_1^2 + \cdots + c_n^2)\lambda^2 \\ &\gamma(1) = (c_0c_1 + c_1c_2 + \cdots + c_{n-1}c_n)\lambda^2 \\ &\gamma(2) = (c_0c_2 + c_1c_3 + \cdots + c_{n-2}c_n)\lambda^2 \\ &\qquad \cdots \\ &\lambda(n) = (c_0c_n)\lambda^2 \\ &\lambda(k) = 0\, \forall k > n \end{aligned}$$

3.3 MA(∞)

Model

$$v(t) = c_0\eta(t) + c_1\eta(t-1) + \cdots + c_k\eta(t-k) + \cdots = \sum_{i=0}^{\infty} c_i\eta(t-i)$$

Variance

$$\gamma(0) = (c_0^2 + c_1^2 + \cdots + c_k^2 + \ldots)\lambda^2 = \lambda^2 \sum_{i=0}^{\infty} c_i^2$$

3.4 Well definition of an MA(∞)

We need to have $|\gamma(\tau)| \leq \gamma(0)$, so we must require that

$$\boxed{\gamma(0) = \lambda^2 \sum_{i=0}^{\infty} c_i^2 \text{ is finite}}$$

4 Auto Regressive Processes

4.1 AR(1)

Model

$$v(t) = av(t-1) + \eta(t)$$

Mean

$$\begin{aligned}
E[v(t)] &= E[av(t-1)] + \overbrace{E[\eta(t)]}^{0} \\
&= aE[v(t-1)] \\
&= aE[v(t)] \\
(1-a)E[v(t)] &= 0 \\
&\boxed{E[v(t)] = 0}
\end{aligned}$$

Covariance

MA(∞) method Observe as an AR(1) can be axpressed as an MA(∞)

$$\begin{aligned}
v(t) &= av(t-1) && +\eta(t) \\
&= a[av(t-2) + \eta(t-1)] && +\eta(t) \\
&= a^2 v(t-2) && +a\eta(t-1) + \eta(t) \\
&= a^2[v(t-3) + \eta(t-2)] && +a\eta(t-1) + \eta(t) \\
&= \underbrace{a^n v(t-n)}_{\to 0} + \underbrace{\sum_{i=0}^{\infty} a^i \eta(t-i)}_{\text{MA}(\infty)}
\end{aligned}$$

In particular, the result depends on an $MA(\infty)$ having $\sum_{i=0}^{\infty} c_i = \sum_{i=0}^{\infty} a^i$. To check if the variance is finite we check $\gamma(0) = \lambda^2 \sum_{i=0}^{\infty} a^{2i} < \infty$. The given is a geometric series, convergent for $|a| < 1$. Under this hypothesis its value is

$$\gamma(0) = \lambda^2 \sum_{i=0}^{\infty} a^{2i} = \frac{\lambda^2}{1-a^2}$$

Applying the formula of the variance of MA processes we get

$$\gamma(1) = (c_0c_1 + c_1c_2 + \dots)\lambda^2 = (a + aa^2 + \dots)\lambda^2 = a(1 + a^2 + a^4 + \dots)\lambda^2 \quad = a\lambda^2 \sum_{i=0}^{\infty} a^{2i} = a\frac{\lambda^2}{1-a^2} \quad = a\gamma(0)$$

$$\gamma(2) = (c_0c_2 + c_1c_3 + \dots)\lambda^2 = (a^2 + aa^3 + ...)\lambda^2 = a^2(1 + a^2 + a^4 + \dots)\lambda^2 = a^2\lambda^2 \sum_{i=0}^{\infty} a^{2i} = a^2\frac{\lambda^2}{1-a^2} = a^2\gamma(0)$$

$$\boxed{\gamma(\tau) = a^{|\tau|}\frac{\lambda^2}{1-a^2}}$$

Yule-Walkler Equations

$$\begin{aligned} Var[v(t)] &= E[v(t)^2] \\ &= E[(av(t) + \eta(t))^2] \\ &= a^2 \underbrace{E[v(t-1)^2]}_{\substack{=Var[v(t-1)] \\ =Var[v(t)] \\ =\gamma(0)}} + \underbrace{E[\eta(t)^2]}_{=\lambda^2} + 2a \quad \underbrace{E[v(t-1)\eta(t)]}_{\substack{v(t-1) \text{ depends on } \eta(t-2) \\ \eta(t) \text{ independent of } \eta(t-2) \\ \Longrightarrow \\ E[v(t-1)\eta(t)]=0}} \end{aligned}$$

$$\gamma(0) = a^2\gamma(0) + \lambda^2$$

$$\boxed{\gamma(0) = \frac{\lambda^2}{1-a^2}}$$

To find $\gamma(\tau)$, we start from the model $v(t) = av(t-1) + \eta(t)$.

$$\begin{aligned} v(t) &= av(t-1) & &+\eta(t) \\ v(t)v(t-\tau) &= av(t-1)v(t-\tau) & &+\eta(t)v(t-\tau) \\ \underbrace{E[v(t)v(t-\tau)]}_{\gamma(\tau)} &= a\underbrace{E[v(t-1)v(t-\tau)]}_{\gamma(\tau-1)} & &+ \underbrace{E[\eta(t)v(t-\tau)]}_{0} \end{aligned}$$

$$\boxed{\gamma(\tau) = a\gamma(\tau-1)}$$

We can join the two by inductive reasoning, obtaining

$$\boxed{\gamma(\tau) = a^{|\tau|}\frac{\lambda^2}{1-a^2}}$$

Long Division Leads to same result, but is boring

4.2 AR(n)

Model

$$v(t) = a_1v(t-1) + a_2v(t-2) + \dots + a_nv(t-n) + \eta(t)$$

Transfer function

$$W(z) = \frac{z^n}{z^n - a_1 z_{n-1} - \cdots - a_n}$$

Mean

$$E[v(t)] = a_1 E[v(t-1])+a_2 E[v(t-2)]+\ldots+a_n E[v(t-n)]+\underbrace{E[\eta(t)]}_{0}$$

$$m = a_1 m \qquad +a_2 m \qquad +\ldots+a_n m$$

$$(1 - a_1 - a_2 - \cdots - a_n)m = 0$$

$$\boxed{E[v(t)] = 0}$$

5 ARMA Processes

Model

$$v(t) = a_1 v(t-1) + \cdots + a_{n_a} v(t-n_a) + c_0 \eta(t) + \cdots + c_{n_c} v(t-n_c)$$

Can also be espressed as $V(t) = \frac{C(z)}{A(z)}\eta(t)$, where

$$C(z) = c_0 + c_1 z^{-1} + \cdots + c_{n_c} z^{-n_c}$$
$$A(z) = 1 - a_1 z^{-1} - \cdots - a_{n_a} z^{-n_a}$$

Such process is stationary if all the poles of $W(z)$ are inside the unit disk.

6 Prediction problem

We want to predict $v(t+r)$ from $v(t), v(t-1), \dots$, where r is called prediction horizon, of the following stationary process:

$$\xrightarrow{\eta}\boxed{W(z)}\xrightarrow{v}$$

6.1 Fake problem

Having a process with transfer function $W(z)$, we can compute it in polynomial form using the long division algorithm

$$W(z) = w_0 + w_1 z^{-1} + w_2 z^{-2} + \dots$$

We can calculate

$$v(t+r) = W(z)\eta(t+r) = \underbrace{w_0\eta(t+r) + w_1\eta(t+r-1) + \dots + w_{r-1}\eta(t+1)}_{\alpha(t)\ \textbf{unpredictable:}\ \text{future of } \eta \text{ involved}} + \underbrace{w_r\eta(t) + w_{r+1}\eta(t-1) + \dots}_{\beta(t)\ \textbf{predictable}}$$

The optimal fake predictor is then

$$\boxed{v(t+r|t) = w_r\eta(t) + w_{r+1}\eta(t-1) + \dots} = \beta(t)$$

And the prediction error is

$$\begin{aligned}\epsilon(t) &= v(t+r) & &-\hat{v}(t+r|t)\\ &= \alpha(t) + \beta(t) & &-\beta(t)\\ &= \alpha(t)\end{aligned}$$

$$\boxed{\epsilon(t) = w_0\eta(t+r) + w_1\eta(t+r-1) + \dots + w_{r-1}\eta(t+1)}$$

$$\boxed{Var[\epsilon(t)] = (w_0^2 + w_1^2 + \dots + w_{r-1}^2)\lambda^2}$$

6.2 True Problem

We want to estimate $v(t+r)$ form $v(t)$, having transfer function $W(z)$ and $\hat{W}_r(z)$ the solution to the fake problem. We can calculate the transfer function of the real predictor from the process as

$$\boxed{W_r(z) = W(z)^{-1} \cdot \hat{W}_r(z)}$$

For ARMA processes a shortcut exists:

$$\hat{v}_{\text{ARMA}}(t|t-1) = \frac{C(z)A(z)}{C(z)} \qquad \text{having } W(z) = \frac{C(z)}{A(z)}$$

6.3 Prediction with eXogenous variables

An exogenous variable is a deterministic input variable in the system

6.3.1 ARX model

$$v(t) = a_1 v(t-1) + \cdots + a_{n_a} v(t-n_a) + b_1 u(t-1) + \cdots + b_{n_b} u(t-n_b) + \eta(t) A(z)v(t) \quad = B(z)u(t-1)$$

Transfer functions from u and η

$$W_u(z) = \frac{B(z)}{A(z)} \qquad\qquad W_\eta(z) = \frac{1}{A(z)}$$

6.3.2 ARMAX model

$$\begin{aligned} A(z)v(t) &= C(z)\eta(t) + B(z)u(t-1) \\ y(t) &= W(z)\eta(t) + G(z)u(t) \end{aligned}$$

Predictor

$$\hat{y}(t|t-1) = \frac{C(z) - A(z)}{C(z)} y(t) + \frac{B(z)}{C(z)} u(t-1)$$

Part II
Identification

Consists of estimating a model from data.

7 Prediction Error Minimization

Aims to minimize $\epsilon(t) = v(t) - \hat{v}(t|t-r)$
Steps:

1. **Data collection:** collect $\vec{u}$ and $\vec{y}$
2. **Family selection:** choose a family of models $M(\theta)$

   MA(1) $\theta = [a]$

   MA(n) $\theta = [a_1, \dots, a_n]$

   ARMA(n_a, n_c) $\theta = [a_1, \dots, a_{n_a}, c_1, \dots, c_{n_c}]$

   ...

3. **Select an optimization criterion**

   Mean Squared error $J(\theta) = \frac{1}{N}\sum_{t=1}^{N} \epsilon_\theta(t)^2$

   Mean absolute error $J(\theta) = \frac{1}{N}\sum_{t=1}^{N} |\epsilon_\theta(t)|$

   ...

4. **Optimization** find $\hat{\theta} = argmin\, J(\theta) \implies \frac{dJ(\theta)}{d\theta} = 0$
5. **Validation** verify if the result satisfies the requirements

7.1 Least Square Error

Consider the ARX model:

$$y(t) = a_1(t-1) + \dots + a_{n_a} y(t-n_a) + b_1 u(t-1) + \dots + b_{n_b} u(t-n_b) + \xi(t) = \theta^T \phi(t) + \eta(t)$$

where $\theta = (a_1, \dots, a_{n_a}, b_1, \dots, b_{n_b})$ and $\phi(t) = y(t-1), \dots, y(t-n_a), u(t-1), \dots, u(t-n_b)$ We then use the MSE as optimization criterion and impose its detivative to 0, obtaining the normal equations:

$$\sum_{t=1}^{N} \phi(t)\phi^T(t)\theta = \sum_{t=1}^{N} y(t)\phi^T(t) \implies \hat{\theta} = \left[\sum_{t=1}^{N} \phi(t)\phi^T(t)\right]^{-1} \cdot \sum_{t=1}^{N} y(t)\phi^T(t)$$

Toepliz matrix $R_{uu} = \begin{vmatrix} \gamma(0) & \gamma(1) & \gamma(2) & \dots \\ \gamma(1) & \gamma(2) & \dots & \\ \gamma(2) & \dots & & \\ \dots & & & \end{vmatrix}$ If R_{uu} is invertible, then the normal equations will have a unique solution.

7.2 Estimation of mean, variance, spectrum

Given the data $y_1, \dots, y_N$, we want to calculate an estimator $\hat{\bullet}_N$ (or $\hat{\bullet}$) from the data

7.2.1 Mean

$$\hat{m}_n = \frac{1}{N}\sum_{i=1}^{N} y_i$$

Correct $E[\hat{m}_N] = m \quad \forall N$

Consistent $N \to \infty \implies Var[\hat{m}_s] \to 0$

7.2.2 Covariance (a)

$$\hat{\gamma}_N^a(\tau) = \frac{1}{N}\sum_{i=1}^{N-|\tau|} y_i y_{i+|\tau|}$$

Has positive semi-definite Toeplitz matrix

Asymptotically correct $N \to \infty \implies E[\hat{\gamma}_N^a(\tau)] \to \gamma(\tau)$

7.2.3 Covariance (b)

$$\hat{\gamma}_N^a(\tau) = \frac{1}{N-|\tau|}\sum_{i=1}^{N-|\tau|} y_i y_{i+|\tau|}$$

Correct $E[\hat{\gamma}_N^b] = m \quad \forall N$

7.2.4 Spectrum

$$\hat{\Gamma}(\omega_k) = \sum_{\tau=1-N}^{N-1} \hat{\gamma}_N(\tau) e^{-j\omega_k \tau}$$

Asymptotically correct $N \to \infty \implies E[\hat{\Gamma}(\omega)] \to \Gamma(\omega)$

Not correct $N \to \infty \implies Var[\hat{\Gamma}(\omega)] \to \Gamma(\omega)^2$

Barlett method Given N data (with N large) divide it into r subsets of size $\hat{N} = \frac{N}{r}$ where $N >> r$. Compute the single spectra:

$$\hat{\Gamma}_{\hat{N}}^{(i)}(\omega) \forall i = 1, ..., r$$

and the average:

$$\overline{\hat{\Gamma}}(\omega) = \frac{1}{r}\sum_{i=1}^{r} \hat{\Gamma}_{\hat{N}}^{(i)}(\omega)$$

As $N >> r$, data from different series can be assumed uncorrelated. We can then consider

$$Var\left[\overline{\hat{\Gamma}}(\omega)\right] \approx \frac{1}{r}\Gamma^2(\omega) << \Gamma^2(\omega)$$

Part III
Cheatsheet

A Probability Recall

Cross-Variance $Var[v,u] = E[(v - E[v])(u - E[u])]$

Variance Matrix $\begin{vmatrix} Var[v_1] & . & . & Var[v_1, v_k] \\ . & . & & . \\ . & & . & . \\ Var[v_k, v_1] & . & . & Var[v_k] \end{vmatrix}$

Covariance coefficient $\delta[i,j] = \frac{Var[i,j]}{\sqrt{Var[i]}\sqrt{Var[j]}}$

Stationary process

- m constant
- λ^2 constant
- covariance $\gamma(\tau)$ depends only on time difference
- $|\gamma(\tau)| \leq \gamma(0) \quad \forall \tau$

White noise $\eta(t) \sim WN(m, \lambda^2)$

- Stationary process
- $\gamma(\tau) = 0 \quad \forall \tau \neq 0$
- $v(t) = \alpha\eta(t) + \beta \implies v(t) \sim WN(\beta, \alpha^2\lambda^2)$

Canonical representation

- Monic
- Same degree
- Coprime
- Poles and zeros in unit disk

B Spectral analysis

Spectrum

- $\Gamma(\omega) = \gamma(0) + 2cos(\omega)\gamma(1) + 2cos(2\omega)\gamma(2) + \dots$
- Periodic $T = 2\pi$
- Even
- $\Gamma_\eta(\omega) = \gamma_\eta(0) = \lambda^2$

Complex spectrum

- $\Phi(z) = \sum_{\tau=-\infty}^{+\infty} \omega(\tau) z^{-\tau}$
- $\Gamma(\omega) = \Phi(e^{j\omega})$

Fundamental theorem of spectral analysis

- $\Gamma_{\text{out}}(\omega) = |W(e^{j\omega})|^2 \cdot \Gamma_{\text{in}}(\omega)$
- $\Phi_{\text{out}}(z) = W(z)W(z^{-1}) \cdot \Phi_{\text{in}}(z)$

C Moving Average MA(n)

- $W(z) = \frac{c_0 z^n + c_1 z^{n-1} + \cdots + c_n}{z_n}$
- $m = 0$
- $\gamma(\tau) = \begin{cases} (c_0 c_\tau + c_1 c_{1+\tau} + \cdots + c_{n-\tau} c_\tau)\lambda^2 & |\tau| \leq n \\ 0 & \text{otherwise} \end{cases}$

C.1 MA(∞)

- $\gamma(0) = (c_0^2 + c_1^2 + \cdots + c_k^2 + \dots)\lambda^2$
- $\gamma(0)$ must converge to a finite value

D Auto Regressive AR(n)

- $m = 0$
- $W(z) = \frac{z^n}{z^n - a_1 z_{n-1} - \cdots - a_n}$
- Covariance calculated by its definition

E Known predictors

AR(1) $\hat{v}(t|t-r) = a^r v(t-r)$

MA(1) $\hat{v}(t|t-1) = v(t-1) - c\hat{v}(t-1|t-2)$

MA(n) $\hat{v}(t|t-\mathbf{k}) = 0 \quad \forall k > n$

ARMA(n_a, n_b**)** $\hat{v}(t|t-1) = \frac{C(z)-A(z)}{C(z)} v(t)$

ARMAX(n_a, n_b**)** $\hat{y}(t|t-1) = \frac{C(z)-A(z)}{C(z)} y(t) + \frac{B(z)}{C(z)} u(t-1)$