

Advanced Regression Analysis

5. Inference and Interpretation of Linear Regression

GOVT 6029 - Spring 2020

Cornell University

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First, we'll talk about what different specifications mean

Later, we'll talk about how to choose a specification (fitting the model)

Let's imagine we know that the *true* model for some data Y is

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- “methods”: includes the method of estimation; e.g. least squares
- “model”: includes the choice of specification; e.g., which controls, transformations, and interactions

Omitted variable bias

Because Y_i was constructed by adding together

$\beta_0^{\text{true}} + \beta_1^{\text{true}} X_i + \beta_2^{\text{true}} Z_i$ + a Normal disturbance,

estimation by least squares will not only be unbiased,
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We won't get any estimate for β_2 (because we assumed it was zero by omitting it)

Moreover, it will usually be the case that $\hat{\beta}_1^*$ is a *biased*
estimate of β_1

Omitted variable bias

The source of this bias can be shown formally.

We estimated this

$$Y_i = \beta_0^* + \beta_1^* X_i \varepsilon_i^*$$

leaving out Z_i . Suppose we ran an auxiliary regression of Z_i on X_i :

$$Z_i = \gamma_0 + \gamma_1 X_i + \nu_i$$

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The estimate we get of β_1 is:

$$\begin{aligned} E(\hat{\beta}_1) &= \beta_1^{\text{true}} + \beta_2^{\text{true}}\gamma_1 \\ &= \beta_1^{\text{true}} + \beta_2^{\text{true}} \left(\frac{\sum_i (X_i - \bar{X})(Z_i - \bar{Z})}{\sum_i (X_i - \bar{X})^2} \right) \end{aligned}$$

Which is unbiased only if $\beta_2^{\text{true}} = 0$ or $\text{corr}(X_i, Z_i) = 0$.

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Specification is arguably *the* major concern in most observational research (along with selection & endogeneity)

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Yes, but it is smaller. We lose efficiency in two ways:

- Lost degrees of freedom
- Lost variance in relevant covariates after conditioning on irrelevant ones

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Kevin Clarke emphasizes that you only *solve* OVB when you have the “right” specification”. It is not usually clear whether getting “closer” to the right specification reduces bias.

My view: robustness to specification choice is a nice standard. Try to break your findings, and report how easily they are broken.

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Omitted variable bias

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We can validly omit variables W which only affect Y through included X 's if we want β to absorb the impact of W through X .

What makes linear regression “linear”?

As long as this remains a valid statement:

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We can make any algebraic substitutions we want.

To see this, replace the big X 's of the above equation with functions of small x 's.

Transformations of covariates

Suppose $X_1 = x_1^2$ and $X_2 = x_1$. Then, by algebraic substitution:

$$Y = \beta_0 + \beta_1 x_1^2 + \beta_2 x_1 + \dots + \varepsilon$$

Sample output (note we're leaving out x_1 to make a point; normally we need it, too):

Call:

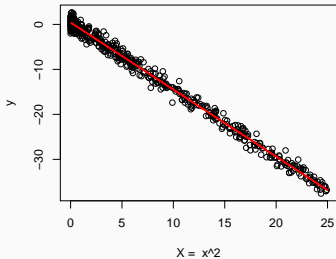
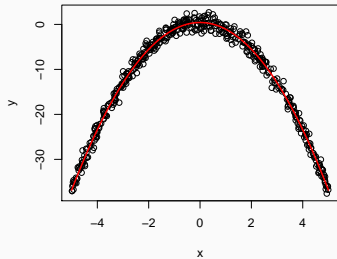
```
lm(formula = y ~ I(x^2))
```

Residuals:

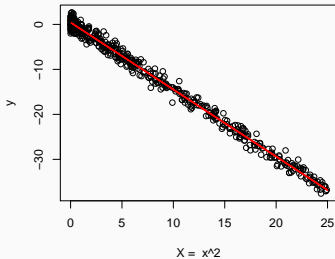
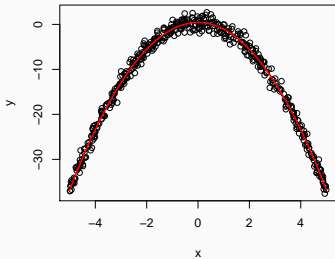
Min	1Q	Median	3Q	Max
-4.3876	-0.1713	0.2031	0.3971	0.7495

Coefficients:

	Estimate	Std. Error	t value	Pr(> t)
(Intercept)	0.9008332	0.0310244	29.036	<2e-16 ***
I(x^2)	-0.0006920	0.0007175	-0.965	0.335



The same regression; two views



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Left is the regression in the original, untransformed scale of x

Right is the regression as R sees it, in the transformed scale $x^2 = X$

Linear regression is always linear in the transformed covariate

It may be very curvilinear in the scale we care about

Suppose $X_1 = x_1^3$, $X_2 = x_1^2$, $X_3 = x_1$.

This is a cubic polynomial specification:

$$Y = \beta_0 + \beta_1 x_1^3 + \beta_2 x_1^2 + \beta_3 x_1 + \dots + \varepsilon$$

Or even add a quartic term, $X_4 = x_1^4$. Then,

$$Y = \beta_0 + \beta_1 x_1^3 + \beta_2 x_1^2 + \beta_3 x_1 + \beta_4 x_1^4 + \dots + \varepsilon$$

Transformations in R

Sample output for a cubic (3rd order) polynomial:

Call:

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lm(formula = y ~ I(x^3) + I(x^2) + I(x))
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Residuals:

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-4.3835	-0.1706	0.2006	0.3983	0.7536

Coefficients:

	Estimate	Std. Error	t value	Pr(> t)
(Intercept)	0.8840553	0.0830528	10.644	<2e-16 ***
I(x^3)	0.0002337	0.0011134	0.210	0.834
I(x^2)	-0.0044455	0.0168552	-0.264	0.792
I(x)	0.0166821	0.0721171	0.231	0.817

None of the coefficients are significant, but they collectively explain a lot of variance.
(How is this possible?)

Transformations in R

Sample output for a cubic (3rd order) polynomial:

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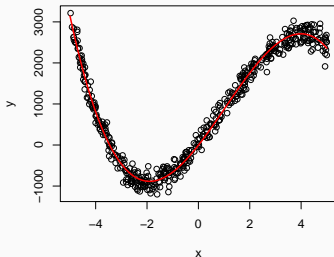
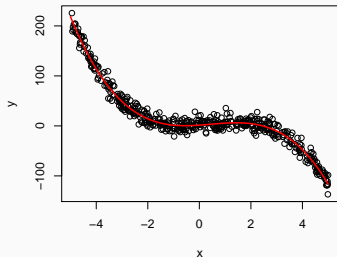
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With polynomials (and with other interactions)

t -tests of individual parameters are less interesting than CIs around \hat{Y}

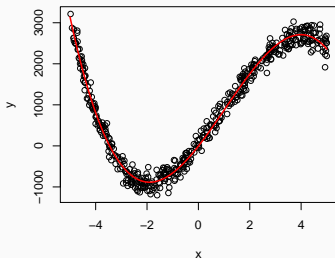
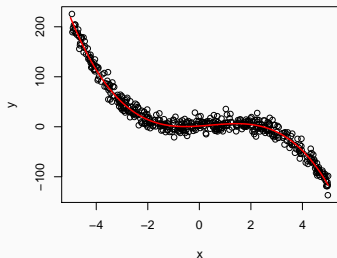


These are two different regressions.

Left is a cubic (3rd order) polynomial specification. It has 2 bends

Right is a quartic (4th order) polynomial. It has 3 bends

Each polynomial order we add puts another bend in the line



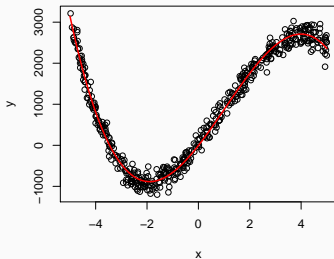
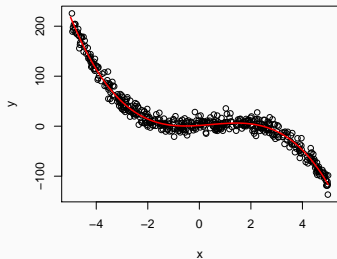
If we include n terms, there will be a bend for every observation

Called “curve-fitting”: a perfect (& perfectly useless) model

Always have a theoretical reason to include polynomial terms

Seldom is more than a quadratic justified by theory

If you include polynomial terms, you need to interpret the result graphically

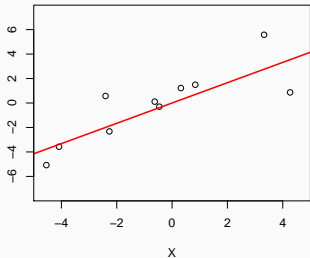


Warnings about polynomial fits:

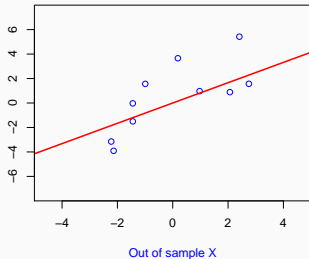
- High order polynomials will *always* fit the sample well, but seldom fit the population well (curve-fitting)
- Extrapolation from polynomial or interactive specifications is dangerous

These functional forms behave wildly outside the known data

Number of Obs: 10. Order of polynomial: 1.
se(regression): 1.651. R-Squared: 0.679.



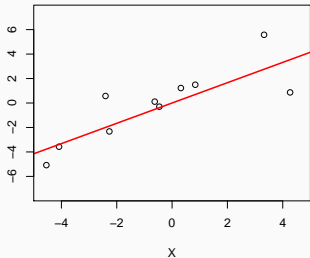
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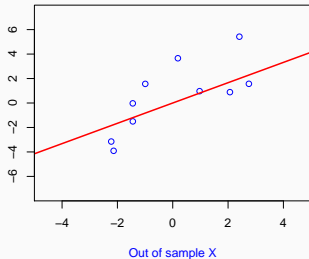
Polynomial overfitting experiment: generate 10 obs from the “true” model:

$$Y = x + \varepsilon, \quad \varepsilon \sim N(0, 3)$$

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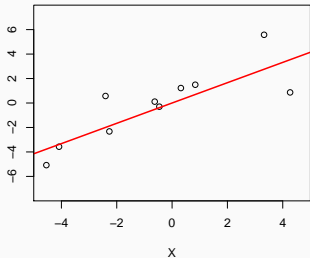
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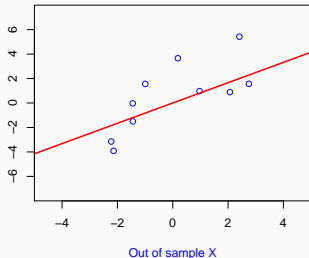
and fit these data using different polynomials of x .

We will show the fit of the model to the original data on the left

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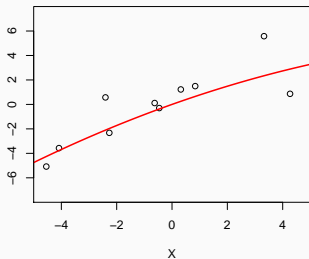
We'll draw new “out of sample data” from the same true model:

$$Y_{\text{OOS}} = x_{\text{OOS}} + \varepsilon_{\text{OOS}}, \quad \varepsilon \sim N(0, 3)$$

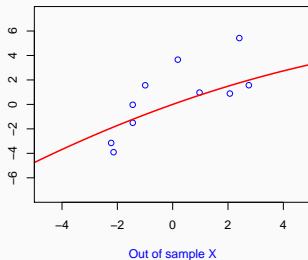
use the model as fitted on the orig data to predict out-of-sample cases

Then show the fit of the old model to the out-of-sample data on the right

Number of Obs: 10. Order of polynomial: 2.
se(regression): 1.551. R-Squared: 0.6916.



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se(regression): 1.979. R-Squared: 0.468.

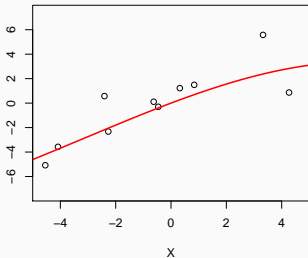


Above is the fit from a quadratic specification of x , ie, we estimated:

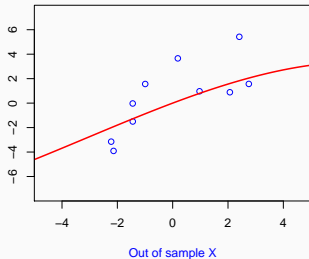
$$Y = \hat{\beta}_1 x + \hat{\beta}_2 x^2 + \hat{\varepsilon}$$

Note that we omitted the constant for didactic reasons

Number of Obs: 10. Order of polynomial: 3.
se(regression): 1.546. R-Squared: 0.6919.



Number of Obs: 10. Order of polynomial: 3.
se(regression): 1.960. R-Squared: 0.4758.



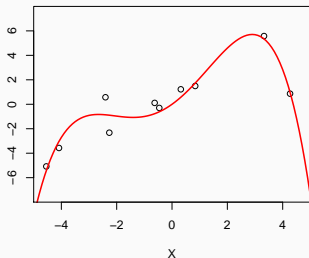
Above is the fit from a cubic specification of x , ie, we estimated:

$$Y = \hat{\beta}_1 x + \hat{\beta}_2 x^2 + \hat{\beta}_3 x^3 + \hat{\varepsilon}$$

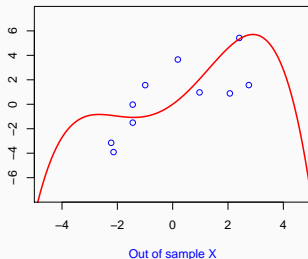
How many polynomials can we add and still find $\hat{\beta}$?

What will happen to the fit in and out of sample as we add polynomials?

Number of Obs: 10. Order of polynomial: 4.
se(regression): 0.796. R-Squared: 0.9255.



Number of Obs: 10. Order of polynomial: 4.
se(regression): 2.577. R-Squared: 0.1113.

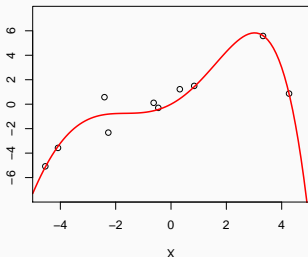


Above is the fit from a quartic specification of x , ie, we estimated:

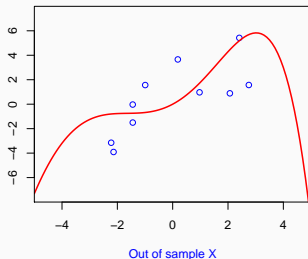
$$Y = \hat{\beta}_1 x + \hat{\beta}_2 x^2 + \hat{\beta}_3 x^3 + \hat{\beta}_4 x^4 + \hat{\varepsilon}$$

On the left, we see the model finds a curious non-linearity, by which low and high x suppress y , but middle values of x increase y . Do you trust this finding?

Number of Obs: 10. Order of polynomial: 5.
se(regression): 0.7585. R-Squared: 0.9324.



Number of Obs: 10. Order of polynomial: 5.
se(regression): 2.513. R-Squared: 0.1427.

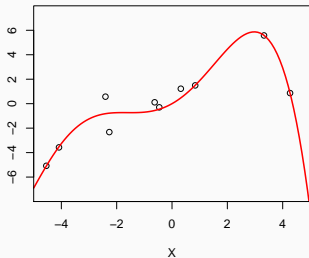


In small samples, outliers can easily create the illusion of complex curves relating x and y

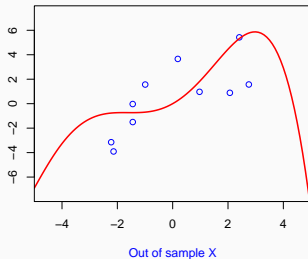
We need *a lot* of data to discern if such curves are more than spurious

(And so we probably need a strong theory, too, to justify the data collection)

Number of Obs: 10. Order of polynomial: 6.
se(regression): 0.7591. R-Squared: 0.9326.

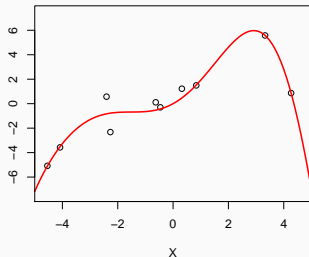


Number of Obs: 10. Order of polynomial: 6.
se(regression): 2.543. R-Squared: 0.1113.

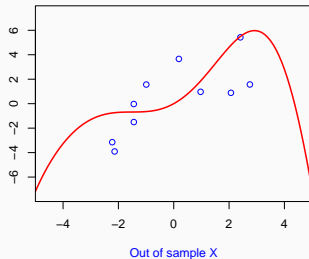


What happens as we approach a tenth-order polynomial?

Number of Obs: 10. Order of polynomial: 7.
se(regression): 0.7602. R-Squared: 0.9326.

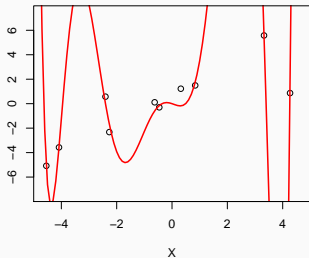


Number of Obs: 10. Order of polynomial: 7.
se(regression): 2.571. R-Squared: 0.07751.

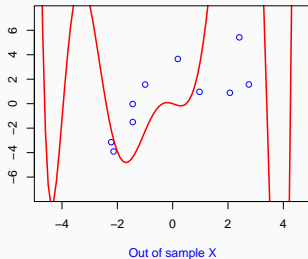


What happens as we approach a tenth-order polynomial?

Number of Obs: 10. Order of polynomial: 8.
se(regression): 0.5832. R-Squared: 0.9584.

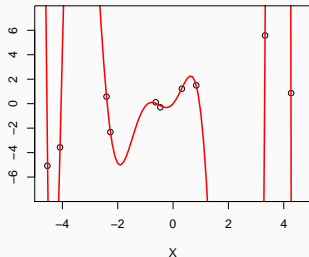


Number of Obs: 10. Order of polynomial: 8.
se(regression): 15.12. R-Squared: -35.

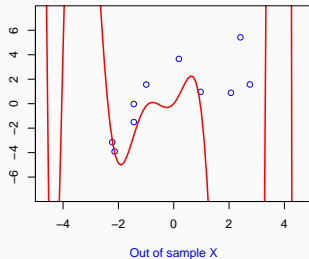


What happens as we approach a tenth-order polynomial?

Number of Obs: 10. Order of polynomial: 9.
se(regression): 0.05322. R-Squared: 0.9997.

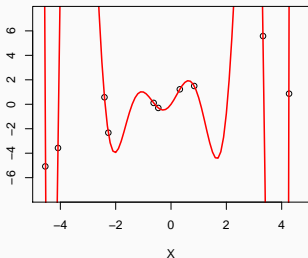


Number of Obs: 10. Order of polynomial: 9.
se(regression): 46.64. R-Squared: -384.0.

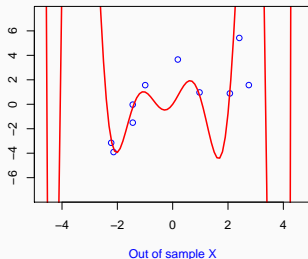


What happens as we approach a tenth-order polynomial?

Number of Obs: 10. Order of polynomial: 10.
se(regression): 0. R-Squared: 1.



Number of Obs: 10. Order of polynomial: 10.
se(regression): 9.433. R-Squared: -11.61.

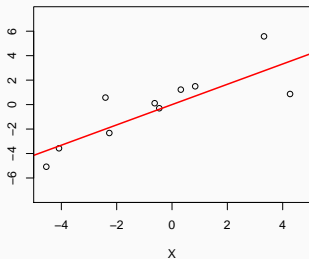


When the number of parameters in the model equals the number of observations, least squares is able to fit a line through every datapoint

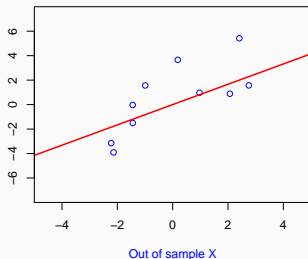
This means the fit will be “perfect”: no error.

And out of sample, it will be completely useless, and worse than guessing that y simply equals its sample mean in every case.

Number of Obs: 10. Order of polynomial: 1.
se(regression): 1.651. R-Squared: 0.679.



Number of Obs: 10. Order of polynomial: 1.
se(regression): 1.993. R-Squared: 0.4768.



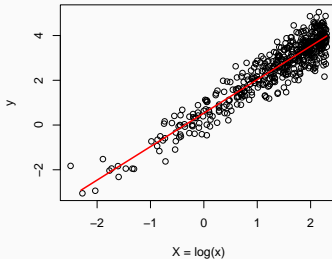
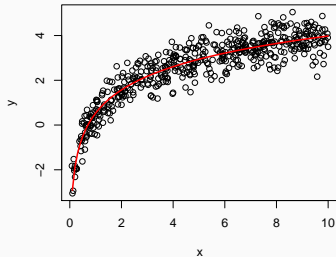
Two lessons:

Beware curve-fitting

Beware good fits in-sample unless they fit well out of sample, too

Suppose $X_1 = \log(x_1)$. Then,

$$Y = \beta_0 + \beta_1 \log(x_1) + \beta_2 X_2 + \dots + \varepsilon$$



These are the same regression

Left is on the original, untransformed scale

Right is on the transformed, log scale

Log transformations for effects that diminish in per unit potency as x increases

We could keep going, combining transformations and interactions to get very nonlinear models

Suppose $X_1 = x_1 \times x_2$ and $X_2 = x_2$. Then,

$$Y = \beta_0 + \beta_1 x_1 \times x_2 + \beta_2 x_2 + \dots + \varepsilon$$

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Suppose $X_1 = x_1 \times x_2$ and $X_2 = x_2$. Then,

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Suppose $X_1 = x_1 \times \log(x_2) \times \sqrt{x_3}$ and $X_2 = x_2 / (x_1 + x_2^2)$. Then,

$$Y = \beta_0 + \beta_1 (x_1 \times \log(x_2) \times \sqrt{x_3}) + \beta_2 x_2 / (x_1 + x_2^2) + \dots + \varepsilon$$

Without strong theoretical support, these models will be silly and produce unreliable results with little predictive power

We could even replace Y .

This is a good idea if you think Y is not a linear function of the regressors, but $g(y)$ is a linear function of them

Usually, this is the case for counts, e.g., or money.

Transformations of the response variable

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Usually, this is the case for counts, e.g., or money.

Raising a government budget from \$1 million to \$10 million may be “just as hard”

as raising it from \$10 to \$100 million

If X 's affect the order of magnitude of Y , you should log Y .

If $Y = \log(y)$, then,

$$\log(y) = \beta_0 + \beta_1 X_1 + \beta_2 X_2 + \dots + \varepsilon$$

Now all X 's have diminishing effect

If $Y = \log(y)$, then,

$$\log(y) = \beta_0 + \beta_1 X_1 + \beta_2 X_2 + \dots + \varepsilon$$

Now all X 's have diminishing effect

In fact, level changes in X yield *percentage* changes in Y

If you log both X and Y , then % changes in X cause % changes in Y

Transformations of the response variable

What if Y is bounded but continuous?

Suppose it ranges between 0 and 1 (but doesn't include these values)?

Then we need to “stretch it out” to range from $-\infty$ to ∞

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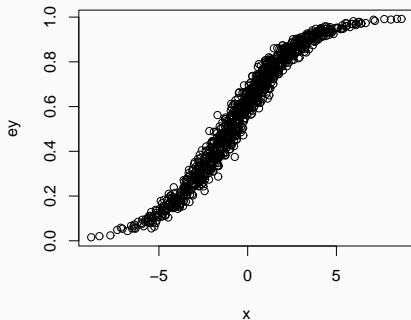
The logit transformation does this: $Y = \log(y/(1 - y))$.

$$\log\left(\frac{y}{1 - y}\right) = \beta_0 + \beta_1 X_1 + \beta_2 X_2 + \dots + \varepsilon$$

This model doesn't work if the original data include 0s or 1s

It is not “the logit model”, which is a non-linear model of 0s and 1s only

It is a linear regression with a logit transformed response variable



Note the curve is S-shaped.

We could rescale it to work for any bounds, not just (0, 1)

That is, just transform y to $y^* = \frac{y-a}{b-a}$, then run the regression of $\text{logit}(y^*)$ on your covariates

This works generally, with the caveat that for any bounds (a, b) , none of the data can be exactly a or b

All of the above models are examples of linear regression.

They are all linear in the parameters.

They can all be estimated by least squares.

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What specifications *can't* we estimate?

Transformations of the response variable

All of the above models are examples of linear regression.

They are all linear in the parameters.

They can all be estimated by least squares.

What specifications *can't* we estimate?

We can't use LS to estimate models that are *non-linear* in the parameters

Transformations of the response variable

Example of a model that is non-linear in the parameters:

$$Y = \beta_0 + \beta_1 X_1 + \beta_2 X_2 + \beta_1 \beta_2 X_3 + \dots + \varepsilon$$

No amount of algebra can turn the above into a linear model.

There are advanced methods to deal with this, e.g., non-linear least squares

Doesn't come up that often, because so many specifications *are* linear in the parameters

Lots of flexibility hidden in linear regression.

So just what is linear regression again?

We have expanded linear regression to encompass any model for unbounded continuous functions $f(\cdot)$ and arbitrary functions $g_k(\cdot)$:

$$y_i = \beta_0 + \beta_1 x_{1i} + \dots + \varepsilon_i$$

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$$f(y_i) = \beta_0 + \sum_{k=1}^K \beta_k g_k(x_{ki}) + \dots + \varepsilon_i$$

“Linearity” in linear regression just refers to the fact that the effect of a one unit change in $g(x_k)$ on $f(y)$ is β_k .

But the relationship between x_k and y themselves could be very non-linear, as a result of $f(\cdot)$ and $g_k(\cdot)$.

Various methods:

- For logged response, $\hat{\beta}$ is the % change in Y for a level change in X
- For logged covariate, $\hat{\beta}$ is the level change in Y for a % changes in X
- For both sides logged, $\hat{\beta}$ is the % change in Y for % changes in X, known as the elasticity of Y with respect to X
- For polynomial coefficients, make a plot of \hat{Y}
- For interactions, make a plot of \hat{Y}

So how do we get \hat{Y} for these models?

Could do it by hand fairly easily.

But what if we want confidence intervals around \hat{Y} too?

Use `predict()`.

Key is setting `newdata` input correctly

If `lm()` did the transformations and interactions

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Then `predict()` will construct interactions and polynomials from their base terms as needed

If not, you need to give `predict` properly constructed interactions and polynomials

What we've learned about linear regression so far

Now we know how to:

1. Specify a regression model
2. Estimate that model
3. Interpret our findings

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Now we know how to:

1. Specify a regression model
2. Estimate that model
3. Interpret our findings

But how do we know if our findings are any good?

That we used the right specification?

That our model explained the data well or poorly?

Need to learn one more skill:

4. Select models with good fit