

Note 15 - Mar 07

Review

Positive recurrence is related to the existence of the stationary distribution.

generating function:

$$\begin{aligned}\psi(s) &= \mathbb{E}(s^\xi) \\ &= \sum_{k=0}^{\infty} \underbrace{P_k}_{\mathbb{P}(\xi=k), k=0,1,\dots} s^k \quad \text{for } 0 \leq s \leq 1\end{aligned}$$

Properties

1. $\psi(0) = p_0, \quad \psi(1) = \sum_{k=0}^{\infty} p_k = 1$
2. Generating function determines the distribution

$$p_k = \frac{1}{k!} \left. \frac{d^k \psi(s)}{ds^k} \right|_{s=0}$$

generating function determines the distribution.

$$\frac{d^k \psi(s)}{ds^k} = k! P_k + (\dots)s + (\dots)s^2 + \dots$$

Since $P_k \geq 0$ for all $k = 0, 1, \dots$, $\frac{d^k \psi(s)}{ds^k} \geq 0$ for all $k = 1, 2, \dots$, $s \in [0, 1]$.

In particular, $\psi(s)$ is increasing and convex between 0 and 1

4. Stochastic Processes (cont'd)

4.6 Generating function and branching processes

Properties of generating function

1. $\psi(0) = p_0, \quad \psi(1) = \sum_{k=0}^{\infty} p_k = 1$
2. Generating function determines the distribution

$$p_k = \frac{1}{k!} \left. \frac{d^k \psi(s)}{ds^k} \right|_{s=0}$$

Reason:

$$\psi(s) = p_0 + p_1 s^1 + \dots + p_{k-1} s^{k-1} + p_k s^k + p_{k+1} s^{k+1} + \dots$$

$$\frac{d^k \psi(s)}{ds^k} = k! p_k + (\dots)s + (\dots)s^2 + \dots$$

$$\left. \frac{d^k \psi(s)}{ds^k} \right|_{s=0} = k! p_k$$

In particular, $p_1 \geq 0 \Rightarrow \psi(s)$ is increasing. $p_2 \geq 0 \Rightarrow \psi(s)$ is concave

3. Let ξ_1, \dots, ξ_n be independent r.v. with generating function ψ_1, \dots, ψ_n ,

$$X = \xi_1 + \dots + \xi_n \Rightarrow \psi_X(s) = \psi_1(s) \psi_2(s) \dots \psi_n(s)$$

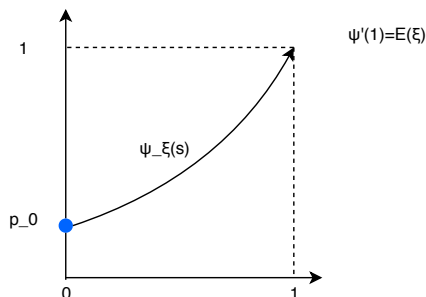
Proof:

$$\begin{aligned}
 \psi_X(s) &= s^{\mathbb{X}} \\
 (\text{independent}) &= \mathbb{E}(s^{\xi_1} s^{\xi_2} \dots s^{\xi_n}) \\
 &= \mathbb{E}(s^{\xi_1}) \dots \mathbb{E}(s^{\xi_n}) \\
 &= \psi_1(s) \dots \psi_n(s)
 \end{aligned}$$

$$4. \quad \left. \frac{d^k \psi(s)}{ds^k} \right|_{s=1} = \left. \frac{d^k \mathbb{E}(s^\xi)}{ds^k} \right|_{s=1} = \mathbb{E} \left(\left. \frac{d^k s^\xi}{ds^k} \right|_{s=1} \right) = \mathbb{E}(\xi(\xi-1)(\xi-2)\dots(\xi-k+1)s^{\xi-k}) \Big|_{s=1} = \mathbb{E}(\xi(\xi-1)\dots(\xi-k+1))$$

$$\text{In particular, } \mathbb{E}(\xi) = \psi'(1) \text{ and } Var(\xi) = \mathbb{E}(\xi^2) - (\mathbb{E}(\xi))^2 = \mathbb{E}(\xi^2 - \xi) + \mathbb{E}(\xi) - (\mathbb{E}(\xi))^2 = \psi''(1) + \psi(1) - (\psi'(1))^2$$

Graph of a g.f.:



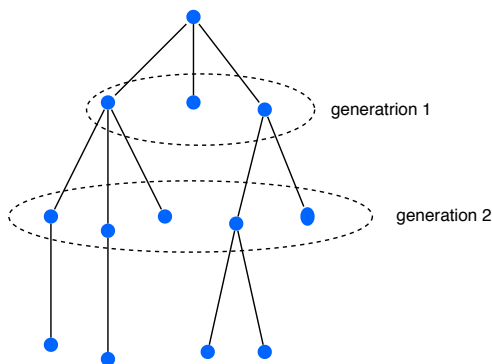
4.6.1 Branching Process

Each organism, at the end of its life, produces a random number Y of offsprings.

$$\mathbb{P}(Y = k) = P_k, \quad k = 0, 1, 2, \dots, \quad P_k \geq 0, \quad \sum_{k=0}^{\infty} P_k = 1$$

The number of offsprings of different individuals are independent.

Start from one ancestor $X_0 = 1$, X_n : # of individuals (population in n -th generation)



Then $X_{n+1} = Y_1^{(n)} + Y_2^{(n)} + \dots + Y_{X_n}^{(n)}$, where $Y_1^{(n)}, \dots, Y_{X_n}^{(n)}$ are independent copies of Y , $Y_i^{(n)}$ is the number of offsprings of the i -th individual in the n -th generation

4.6.1.2 Mean and Variance

Mean: $\mathbb{E}(X_n)$ and Variance: $Var(X_n)$

Assume, $\mathbb{E}(Y) = \mu$, $Var(Y) = \sigma^2$.

$$\begin{aligned}
 \mathbb{E}(X_{n+1}) &= \mathbb{E}(Y_1^{(n)} + \dots + Y_{X_n}^{(n)}) \\
 &= \mathbb{E}(\mathbb{E}(Y_1^{(n)} + \dots + Y_{X_n}^{(n)} | X_n)) \\
 &= \mathbb{E}(X_n \mu)
 \end{aligned}$$

$$\text{Wald's identity (tutorial 3)} \quad = \mu \mathbb{E}(X_n)$$

$$\begin{aligned}
\Rightarrow \mathbb{X}_n &= \mu \mathbb{E}(X_{n-1}) \\
&= \mu^2 \mathbb{X}_{n-2} \\
&\vdots \\
&= \mu^n \mathbb{E}(X_0) = \mu^n, \quad n = 0, 1, \dots
\end{aligned}$$

$$\text{Var}(X_{n+1}) = \mathbb{E}(\text{Var}(X_{n+1}|X_n) + \text{Var}(\mathbb{E}X_{n+1}|X_n))$$

$$\begin{aligned}
\text{Var}(\mathbb{E}(X_{n+1}|X_n)) &= \text{Var}(\mu X_n) \\
&= \mu^2 \text{Var}(X_n)
\end{aligned}$$

$$\Rightarrow \text{Var}(X_{n+1}) = \sigma^2 \mu^n + \mu^2 \text{Var}(X_n)$$

$$\text{Var}(X_1) = \sigma^2$$

$$\text{Var}(X_2) = \sigma^2 \mu + \mu^2 \sigma^2 = \sigma^2 (\mu^1 + \mu^2)$$

$$\text{Var}(X_3) = \sigma^2 \mu^2 + \mu^2 (\sigma^2 (\mu^1 + \mu^2)) = \sigma^2 (\mu^2 + \mu^3 + \mu^4)$$

$$\vdots$$

In general, (can be proved by induction)

$$\text{Var}(X_n) = \sigma^2 (\mu^{n-1} + \dots + \mu^{2n-2})$$

$$= \begin{cases} \sigma^2 \mu^{n-1} \frac{1 - \mu^n}{1 - \mu} & \mu \neq 1 \\ \sigma^2 n & \mu = 1 \end{cases}$$

$$\begin{aligned}
\mathbb{E}(\text{Var}(X_{n+1}|X_n)) &= \mathbb{E}(\text{Var}(Y_1^{(n)} + \dots + Y_{X_N}^{(N)} | X_N)) \\
&= \mathbb{E}(X_n \cdot \sigma^2) \\
&= \sigma^2 \mu^n
\end{aligned}$$

4.6.1.2 Extinction Probability

Q: What is the probability that the population size is eventually reduced to 0

Note that for a branching process, $X_n = 0 \Rightarrow X_k = 0$ for all $k \geq n$. Thus, state 0 is absorbing. ($P_{00} = 1$). Let N be the time that extinction happens.

$$N = \min\{n : X_n = 0\}$$

Define

$$U_n = \mathbb{P}(\underbrace{N \leq n}_{\substack{\text{extinction happens} \\ \text{before or at } n}}) = \mathbb{P}(X_n = 0)$$

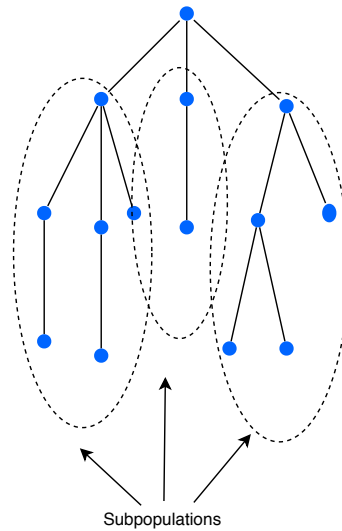
Then U_n is increasing in n , and

$$\begin{aligned}
u &= \lim_{n \rightarrow \infty} U_n = \mathbb{P}(N < \infty) \\
&= P(\text{the extinction eventually happens}) \\
&= \text{extinction probability}
\end{aligned}$$

Our goal : find u

We have the following relation between U_n and U_{n-1} :

$$U_n = \sum_{k=0}^{\infty} P_k(U_{n-1})^k = \underbrace{\psi}_{\text{gf of } Y}(U_{n-1})$$



Each subpopulation has the same distribution as the whole population.

Total population dies out in n steps if and only if each subpopulation dies out in $n - 1$ steps

$$\begin{aligned}
 U_n &= \mathbb{P}(N \leq n) \\
 &= \sum_k \mathbb{P}(N \leq n | X_1 = k) \underbrace{\mathbb{P}(X_1 = k)}_{=P_k} \\
 &= \sum_k \mathbb{P}(\underbrace{N_1 \leq n-1}_{\substack{\text{\# of steps for} \\ \text{subpopulation 1 to die out}}}, \dots, N_k \leq n-1 | X_1 = k) \cdot P_k \\
 &= \sum_k P_k \cdot U_{n-1}^k \\
 &= \mathbb{E}(U_{n-1}^Y) \\
 &= \psi(U_{n-1})
 \end{aligned}$$

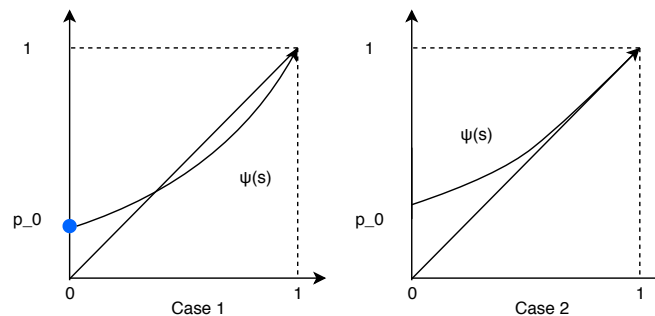
Thus, the question is :

With initial value $U_0 = 0$ (since $X_0 = 1$), relation $U_n = \psi(U_{n-1})$. What is $\lim_{n \rightarrow \infty} U_n = u$?

Recall that we have

1. $\psi(0) = P_0 \geq 0$
2. $\psi(1) = 1$
3. $\psi(s)$ is increasing
4. $\psi(s)$ is convex

Draw $\psi(s)$ and function $f(s) = s$ between 0 and 1, we have two cases:



The extinction probability u will be the smallest intersection of $\psi(s)$ and $f(s)$. Equivalently, it is the smallest solution of the equation $\psi(s) = s$ between 0 and 1.