

Note 14 - Mar 05

Review

Period is a class property : $i \leftrightarrow \Rightarrow d_i = d_j$. Irreducible \Rightarrow period of the MC

Basic Limit Theorem

$\{X_n\}_{n=0,1,\dots}$ irreducible, aperiodic, positive recurrent DTMC; then a unique stationary distribution $\underline{\pi} = (\pi_0, \pi_1, \dots)$ exists. Moreover,

$$(*) \quad \underbrace{\lim_{n \rightarrow \infty} P_{ij}^{(n)}}_{\substack{\text{limiting distribution} \\ \text{(does not depend on the initial state i)}}} = \lim_{n \rightarrow \infty} \underbrace{\frac{\sum_{k=1}^n \mathbb{1}_{\{X_k=j\}}}{n}}_{\text{long-run fraction of time spent in j}} = \frac{1}{\underbrace{\mathbb{E}(T_j | X_0 = j)}_{\substack{T_j = \min\{n > 0 : X_n = j\} \\ \text{expected revisit time}}}}} = \pi_j, \quad i, j \in S$$

Periodic extension:

$$\frac{\lim_{n \rightarrow \infty} P_{jj}^{(nd)}}{d} = \lim_{n \rightarrow \infty} \frac{\sum_{k=1}^n \mathbb{1}_{\{X_k=j\}}}{n} = \frac{1}{\mathbb{E}(T_j | X_0 = j)} = \pi_j$$

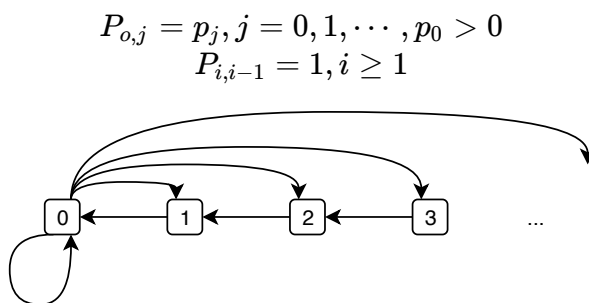
Examples show:

- Irreducibility is related to the uniqueness of the stationary distribution;
- Aperiodicity is related to the existence of the limiting distribution.

4. Stochastic Processes (cont'd)

4.5 Limiting Distribution (cont'd)

Example 4.5.3



$$\begin{aligned}\Rightarrow \mathbb{E}(T_0|X_0 = 0) &= \sum_{n=0}^{\infty} (n+1)p_n \\ &= 1 + \sum_{n=0}^{\infty} np_n\end{aligned}$$

We can construct p_n such that $\sum_{n=0}^{\infty} np_n = \infty$. (For example, $p_0 = \frac{1}{2}, p_2 = \frac{1}{4}, p_4 = \frac{1}{4}, \dots$)

In this case, the chain is **null recurrent**. It is irreducible and aperiodic ($P_{00} = p_0 > 0$)

A stationary distribution does not exist. Reason:

$$p = \begin{pmatrix} p_0 & p_1 & p_2 & \cdots & p_i & \cdots \\ 1 & & 0 & & & \\ & 1 & & & & \\ \cdots & & & & & \\ & & & & 1 & \end{pmatrix}$$

$$\underline{\pi} \cdot P = \underline{\pi} \Rightarrow$$

$$p_0 \pi_0 + \pi_1 = \pi_0$$

$$p_1 \pi_0 + \pi_2 = \pi_1$$

$$\vdots$$

$$p_{i-1} \pi_0 + \pi_i = \pi_{i-1}$$

$$p_i \pi_0 + \pi_{i+1} = \pi_i$$

Add the first i equations:

$$(p_0 + \cdots + p_{i-1})\pi_0 + (\cancel{\pi_1} + \cancel{\pi_2} + \cdots + \pi_i) = \pi_0 + \cancel{\pi_1} + \cancel{\pi_2} + \cdots + \pi_i$$

$$(p_0 + \cdots + p_{i-1})\pi_0 + \pi_i = \pi_0$$

$$\Rightarrow \pi_i = (1 - (p_0 + \cdots + p_{i-1}))\pi_0 = \sum_{k=i}^{\infty} p_k \pi_0$$

Try to normalize:

$$\begin{aligned}1 &= \sum_{i=1}^{\infty} \pi_i \\ &= \sum_{i=0}^{\infty} \sum_{k_i}^{\infty} p_k \pi_0 \\ &= \sum_{k_i}^{\infty} \sum_{i=0}^{\infty} p_k \pi_0 \\ &= \sum_{k_i}^{\infty} p_k \sum_{i=0}^{\infty} \pi_0 \\ &= \underbrace{\left(\sum_{k_i}^{\infty} (k+1)p_k \right)}_{\infty} \pi_0 \\ \Rightarrow \pi_0 &= 0, \quad p_i \pi_i = 0 \quad \forall i\end{aligned}$$

This is not a distribution. Thus, a stationary distribution does not exist.

positive recurrence is related to the existence of the stationary distribution

Example 4.5.4 : Electron

$$P = \begin{pmatrix} 1-\alpha & \alpha \\ \beta & 1-\beta \end{pmatrix} \quad \alpha, \beta \in (0, 1)$$

Irreducible, aperiodic, positive recurrence.

In order to find of P^n ; we use the diagonalization technique.

$$P = Q\Lambda Q^{-1} \quad \text{where } \Lambda \text{ is diagonal}$$

$$\Lambda = \begin{pmatrix} 1 & 0 \\ 0 & 1-\alpha-\beta \end{pmatrix} \quad Q = \begin{pmatrix} 1 & \alpha \\ 1 & 1-\beta \end{pmatrix} \quad Q^{-1} = \frac{1}{\alpha+\beta} \begin{pmatrix} \beta & \alpha \\ 1 & -1 \end{pmatrix}$$

Then

$$\begin{aligned} P^n &= (Q\Lambda Q^{-1})(Q\Lambda Q^{-1}) \cdots (Q\Lambda Q^{-1}) \\ &= Q\Lambda^n Q^{-1} \\ &= \begin{pmatrix} 1 & \alpha \\ 1 & -\beta \end{pmatrix} \begin{pmatrix} 1 & \\ & (1-\alpha-\beta)^n \end{pmatrix} \frac{1}{\alpha+\beta} \begin{pmatrix} \beta & \alpha \\ 1 & -1 \end{pmatrix} \\ &= \frac{1}{\alpha+\beta} \begin{pmatrix} \beta + \alpha(1-\alpha-\beta)^n & \alpha - \alpha(1-\alpha-\beta)^n \\ \beta - \beta(1-\alpha-\beta)^n & \alpha + \beta(1-\alpha-\beta)^n \end{pmatrix} \\ &\Rightarrow \lim_{n \rightarrow \infty} P^n = \frac{1}{\alpha+\beta} \begin{pmatrix} \beta & \alpha \\ \beta & \alpha \end{pmatrix} = \begin{pmatrix} \frac{\beta}{\alpha+\beta} & \frac{\alpha}{\alpha+\beta} \\ \frac{\beta}{\alpha+\beta} & \frac{\alpha}{\alpha+\beta} \end{pmatrix} \end{aligned}$$

Note that $\lim_{n \rightarrow \infty} P^n$ has identical rows. This corresponds to the result that $\lim_{n \rightarrow \infty} P_{ij}^{(n)}$ does not depend on i

We saw that the stationary distribution $\underline{\pi} = (\frac{\beta}{\alpha+\beta}, \frac{\alpha}{\alpha+\beta})$. So we verify that $\pi_j = \lim_{n \rightarrow \infty} P_{ij}^{(n)}$

Also, given $X_0 = 0$, $\mathbb{P}(T_0 = 1 | X_0 = 0) = 1 - \alpha$.

For $k = 2, 3, \dots$

$$\begin{aligned}
\mathbb{P}(T_0 = k | X_0 = 0) &= \mathbb{P}(X_k = 0, X_{k-1} = 1, \dots, X_1 = 1 | X_0 = 0) \\
&= \alpha(1 - \beta)^{k-2} \beta \\
&\Rightarrow \mathbb{E}(T_0 | X_0 = 0) \\
&= 1 \cdot (1 - \alpha) + \sum_{k=2}^{\infty} \alpha(1 - \beta)^{k-2} \beta k \\
&= 1 - \alpha + \sum_{k=1}^{\infty} \underbrace{\alpha(1 - \beta)^{k-2} \beta(k-1)}_{\mathbb{E}(\text{Geo}(\beta))} + \sum_{k=2}^{\infty} \underbrace{\alpha(1 - \beta)^{k-2} \beta}_{\text{pmf of Geo}(\beta)} \\
&= 1 - \alpha + \alpha \sum_{k=1}^{\infty} (1 - \beta)^{k-2} \beta(k-1) + \sum_{k=2}^{\infty} \alpha(1 - \beta)^{k-2} \beta \\
&= 1\alpha + \alpha \cdot \frac{1}{\beta} + \alpha \cdot 1 \\
&= 1 - \alpha + \frac{\alpha}{\beta} + \alpha \\
&= \frac{\alpha + \beta}{\beta}
\end{aligned}$$

Hence we verify that $\mathbb{E}(T_0 | X_0 = 0) = \frac{1}{\pi_0}$

4.6 Generating function and branching processes

Definition 4.6.1

Let $\underline{p} = (p_0, p_1, \dots)$ be a distribution on $\{0, 1, 2, \dots\}$. Let ξ be a r.v. following distribution \underline{p} . That is $\mathbb{P}(\xi = i) = p_i$. Then the generating function of ξ , or of \underline{p} , is defined by

$$\begin{aligned}
\psi(s) &= \mathbb{E}(s^\xi) \\
&= \sum_{k=0}^{\infty} p_k s^k \quad \text{for } 0 \leq s \leq 1
\end{aligned}$$

Properties of generating function

1. $\psi(0) = p_0, \quad \psi(1) = \sum_{k=0}^{\infty} p_k = 1$
2. Generating function determines the distribution

$$p_k = \frac{1}{k!} \frac{d^k \psi(s)}{ds^k} \Big|_{s=0}$$

Reason:

$$\psi(s) = p_0 + p_1 s^1 + \dots + p_{k-1} s^{k-1} + p_k s^k + p_{k+1} s^{k+1} + \dots$$

$$\frac{d^k \psi(s)}{ds^k} = k! p_k + (\dots)s + (\dots)s^2 + \dots$$

$$\frac{d^k \psi(s)}{ds^k} \Big|_{s=0} = k! p_k$$

In particular, $p_1 \geq 0 \Rightarrow \psi(s)$ is increasing. $p_2 \geq 0 \Rightarrow \psi(s)$ is climax

