

CMSE 820 HW 11

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1 Question 1

Let Ω be a non-empty set and \mathcal{A} be a collection of events. Define

$$\mathcal{I}(\mathcal{A}) = \{\mathcal{F} | \mathcal{F} \text{ is a } \sigma\text{-algebra and } \mathcal{F} \supseteq \mathcal{A}\}.$$

and

$$\sigma(\mathcal{A}) = \bigcap_{\mathcal{F} \in \mathcal{I}(\mathcal{A})} \mathcal{F}.$$

Prove that $\sigma(\mathcal{A})$ is a σ -algebra. Note that $\sigma(\mathcal{A})$ is called the σ -algebra generated by \mathcal{A} .

Proof: Let $a \subseteq \mathcal{A}$ be any subset of \mathcal{A} . By the definition of $\sigma(\mathcal{A})$,

$$\forall \mathcal{F} \in \mathcal{I}, a \in \mathcal{F}.$$

Following from the facts that \mathcal{F} is a σ -algebra of \mathcal{A} , we have, for any sequence of subsets of \mathcal{A} , $\{a_i\}_{i=1}^{\infty}$,

$$[\forall \mathcal{F} \in \mathcal{I}, (a \in \mathcal{F} \implies \emptyset \in \mathcal{F})] \implies \emptyset \in \sigma(\mathcal{A}),$$

$$\forall \mathcal{F} \in \mathcal{I}, a_i^c \in \mathcal{F} \implies a_i^c \in \sigma(\mathcal{A}),$$

$$\forall \mathcal{F} \in \mathcal{I}, \bigcup_{i=1}^{\infty} a_i \in \mathcal{F} \implies \bigcup_{i=1}^{\infty} a_i \in \sigma(\mathcal{A}).$$

$\sigma(\mathcal{A})$ contains \emptyset , and is closed under taking complement and countable union. Thus, $\sigma(\mathcal{A})$ is a σ -algebra of \mathcal{A} .

2 Question 2

Let $\Omega = \mathbb{R}$. Define

$$\mathcal{A} = \{(a, b) | a, b \in \mathbb{R}, a \leq b\}.$$

Then $\sigma(\mathcal{A}) = \mathcal{B}(\mathbb{R})$ is the Borel σ -algebra on \mathbb{R} . Let

$$\mathcal{D} = \{(-\infty, a) : a \in \mathbb{R}\}.$$

Prove that $\sigma(\mathcal{D}) = \mathcal{B}(\mathbb{R})$.

Proof:

To show that $\sigma(\mathcal{D}) \supseteq \sigma(\mathcal{A})$, it suffices to show that for any $a, b \in \mathbb{R}$ such that $a \leq b$,

$$(a, b) \in \sigma(\mathcal{D}).$$

Note that σ -algebra is closed under countable intersection (which amounts to the complement of the countable union of the complements). So

$$\cap_{n=1}^{\infty} (-\infty, a + \frac{1}{n}) \in \sigma(\mathcal{D}).$$

In fact, we claim that $\cap_{n=1}^{\infty} (-\infty, a + \frac{1}{n}) = (-\infty, a]$, because

$$\forall x \in (-\infty, a], x \in (-\infty, a + \frac{1}{n}) \implies (-\infty, a] \subseteq \cap_{n=1}^{\infty} (-\infty, a + \frac{1}{n}),$$

and

$$\forall x > a, \exists n \text{ such that } x \notin (-\infty, a + \frac{1}{n}) \implies x \notin \cap_{n=1}^{\infty} (-\infty, a + \frac{1}{n}).$$

Now we can simply express (a, b) as

$$(a, b) = (a, \infty) \cap (-\infty, b) = \left(\cap_{n=1}^{\infty} (-\infty, a + \frac{1}{n}) \right)^c \cap (-\infty, b) \in \sigma(\mathcal{D}).$$

So $\sigma(\mathcal{A}) \subseteq \sigma(\mathcal{D})$.

Next, we will need to show that $\sigma(\mathcal{A}) \supseteq \sigma(\mathcal{D})$. That is, for any $a \in \mathbb{R}$, we need to show that $(-\infty, a) \in \sigma(\mathcal{A})$.

Consider a sequence of elements $\{A_n\}_{n=1}^{\infty}$ in $\sigma(\mathcal{A})$ defined by $A_n = (a - n, a)$. It is easy to see that

$$\cup_{n=1}^{\infty} A_n = \cup_{n=1}^{\infty} (a - n, a) = (-\infty, a) \subseteq \sigma(\mathcal{A}).$$

Thus, $\sigma(\mathcal{A}) = \sigma(\mathcal{D}) = \mathcal{B}(\mathbb{R})$.

3 Question 3

Pick $X \in [-1, 1]^d$ uniformly at random. Prove that $\|X\|$ is tightly concentrated around $\sqrt{d/3}$. Namely,

$$\Pr \left(\left| \|X\|^2 - \frac{d}{3} \right| > \epsilon d \right) \rightarrow 0,$$

as $d \rightarrow \infty$.

Proof: First of all, we compute a few moments of a uniform random variable $x \in [-1, 1]$,

$$\begin{aligned} \mathbb{E}[x] &= 0, \\ \mathbb{E}[x^2] &= \frac{1}{3}, \\ \mathbb{E}[x^4] &= \int_{-1}^1 \frac{1}{2} x^4 dx = \frac{1}{5}. \end{aligned}$$

Clearly, $\|X\|^2$ should be concentrated around $\frac{d}{3}$, because $\mathbb{E}[\|X\|^2] = \mathbb{E}[\sum_{i=1}^d x_i^2] = \sum_{i=1}^d \mathbb{E}[x_i^2] = \frac{d}{3}$. We apply the Chebyshev's inequality to $\|X\|^2$,

$$\Pr\left(\left|\|X\|^2 - \frac{d}{3}\right| > \epsilon d\right) \leq \frac{\mathbb{E}[(\|X\|^2 - d/3)^2]}{(\epsilon d)^2}.$$

We want to show that $\mathbb{E}[(\|X\|^2 - d/3)^2]$ is finite.

$$\begin{aligned} \mathbb{E}[(\|X\|^2 - d/3)^2] &= \mathbb{E}\left[\left(\sum_{i=1}^d (x_i^2 - \frac{1}{3})\right)^2\right] \\ &= \sum_{i=1}^d \mathbb{E}[(x_i^2 - 1/3)^2] + \sum_{i=1, j>i}^d \mathbb{E}[(x_i^2 - 1/3)(x_j^2 - 1/3)] \\ &= \sum_{i=1}^d \mathbb{E}[(x^2 - 1/3)^2] = \frac{4d}{45}. \end{aligned}$$

So the probability goes to 0 as d approaches ∞ at a rate of $4/45\epsilon^2 d$.

4 Question 4

4.1 a

A random variable X with mean $\mathbb{E}[X] = \mu$ is sub-Gaussian if there is a positive number σ such that, for all $\lambda \in \mathbb{R}$,

$$\mathbb{E}[e^{\lambda(X-\mu)}] \leq e^{\sigma^2 \lambda^2 / 2}.$$

If X is a sub-Gaussian random variable with σ as its sub-Gaussian parameter, prove that

$$\Pr(|X - \mu| \geq t) \leq 2e^{-\frac{t^2}{2\sigma^2}}, \text{ for all } t \in \mathbb{R}.$$

Proof: For any $t \in \mathbb{R}$, we apply the Cramer's inequality,

$$\begin{aligned}
\Pr(|X - \mu| \geq t) &= 2 \Pr[(X - \mu) \geq t] \leq 2 \inf_{\lambda > 0} e^{-\lambda t} \mathbf{E}[e^{\lambda(X - \mu)}] \\
&\leq 2 \inf_{\lambda > 0} \exp \left[\frac{\sigma^2}{2} \lambda^2 - t \lambda \right] \\
&\leq 2 \exp \left(- \frac{t^2}{2\sigma^2} \right).
\end{aligned}$$

4.2 b

Suppose that the random variables X_i , $i = 1, \dots, n$ are independent, and X_i has mean μ_i and sub-Gaussian parameter σ_i . Prove that for all $t \geq 0$ we have

$$\Pr \left[\sum_{i=1}^n (X_i - \mu_i) \geq t \right] \leq \exp \left\{ - \frac{t^2}{2 \sum_{i=1}^n \sigma_i^2} \right\}.$$

Proof: For any $t \in \mathbb{R}$, we again apply the Cramer's inequality,

$$\begin{aligned}
\Pr \left[\sum_{i=1}^n (X_i - \mu_i) \geq t \right] &\leq \inf_{\lambda > 0} e^{-\lambda t} \mathbf{E}[e^{\lambda \sum_{i=1}^n (X_i - \mu_i)}] \\
&= \inf_{\lambda > 0} \prod_{i=1}^n e^{-\lambda t/n} \mathbf{E}[e^{\lambda(X_i - \mu_i)}] \\
&\leq \inf_{\lambda > 0} \prod_{i=1}^n \exp \left[\frac{\sigma_i^2}{2} \lambda^2 - \frac{t}{n} \lambda \right] \\
&= \inf_{\lambda > 0} \exp \left[\frac{\sum_{i=1}^n \sigma_i^2}{2} \lambda^2 - t \lambda \right] \\
&= \exp \left(- \frac{t^2}{2 \sum_{i=1}^n \sigma_i^2} \right).
\end{aligned}$$