

## 18.650. Fundamentals of Statistics Spring 2026. Problem Set 1

Due Wednesday, Feb 18

**Problem 1** (Probability Review). *Please use the accompanying bubble sheet for submitting your solutions for this problem.*

In what follows,  $\Phi$  is the CDF of the standard Gaussian (Normal) distribution.

Let  $X$  be a random variable taking values between 0 and  $\pi$ , with pdf given by

$$f(x) = c \sin x, \quad x \in [0, \pi].$$

1. What is the value of  $c$ ? (a)  $\pi$  (b)  **$1/2$**  (c) 2 (d)  $1/\pi$
2. What is  $\mathbb{E}[X]$ ? (a)  **$\pi/2$**  (b)  $\pi$  (c) 1 (d)  $1/2$

Let  $X$  be a Gaussian random variable with mean  $\mu > 0$  and variance  $\sigma^2$ .

3. What is  $\mathbb{E}[X]$ ? (a) 0 (b)  $\sigma^2$  (c)  $\mu^2 + \sigma^2$  (d)  **$\mu$**
4. What is  $\mathbb{V}[X]$ ? (a)  **$\sigma^2$**  (b)  $\mu^2 + \sigma^2$  (c)  $\mu^2$  (d)  $\sigma$
5. What is  $\mathbb{E}[X^2]$ ? (a)  $\mu^2$  (b)  $(\mu + \sigma)^2$  (c)  **$\mu^2 + \sigma^2$**  (d)  $\sigma^2$
6. What is  $\mathbb{E}[X^3]$ ? (a)  $\mu^3 + 3\mu\sigma$  (b)  **$\mu^3 + 3\mu\sigma^2$**  (c)  $\mu^3$  (d)  $3\mu\sigma^2$
7. What is  $\mathbb{V}[X^2]$ ? (a)  $4\mu^2\sigma^2$  (b)  $(\mu^2 + \sigma^2)^2$  (c)  $2\sigma^4$  (d)  **$4\mu^2\sigma^2 + 2\sigma^4$**
8. What is  $\mathbb{P}(X > 0)$  in terms of  $\Phi$ ?  
(a)  **$\Phi(\mu/\sigma)$**  (b)  $\Phi(\mu/\sigma^2)$  (c)  $1 - \Phi(\mu/\sigma)$  (d)  $\Phi(\mu)$

Let  $X \sim \text{Lognormal}(\mu, \sigma^2)$ , i.e.,  $\log X \sim \mathcal{N}(\mu, \sigma^2)$  with  $\sigma > 0$ .

9. What is  $\mathbb{E}[X]$ ? (a)  $e^\mu$  (b)  **$e^{\mu+\sigma^2/2}$**  (c)  $e^{\mu+\sigma^2}$  (d)  $\mu + \sigma^2/2$
10. What is the median of  $X$ ? (a)  $e^{\mu+\sigma^2/2}$  (b)  $e^{\mu-\sigma^2}$  (c)  **$e^\mu$**  (d)  $\mu$
11. What is  $\mathbb{P}(X > 1)$  in terms of  $\Phi$ ?  
(a)  $1 - \Phi(\mu/\sigma)$  (b)  $\Phi(\mu)$  (c)  $\Phi(\mu/\sigma^2)$  (d)  **$\Phi(\mu/\sigma)$**

Let  $X, Y \sim \text{Lognormal}(0, 1)$  be independent.

12. What is  $\mathbb{P}(XY > 1)$ ? (a)  $1/4$  (b)  **$1/2$**  (c)  $1/\sqrt{2\pi}$  (d)  $1/e$

Let  $X$  be a random variable such that

$$X = \begin{cases} 1 & \text{with probability } p \\ -1 & \text{with probability } 1 - p \end{cases}$$

for some  $p \in [0, 1]$ .

13. What is  $\mathbb{E}[X]$ ?

- (a)  $-p$  (b)  $p$  (c)  $1 - 2p$  (d)  **$2p - 1$**

14. What is  $\mathbb{V}[X]$ ?

- (a)  $p(1 - p)$  (b)  $4p - p^2$  (c)  **$4p(1 - p)$**  (d)  $4p^2(1 - p)$

15. For what  $p$  is  $\mathbb{V}[X]$  maximized?

- (a)  $1$  (b)  $0$  (c)  **$0.5$**  (d)  $1/\sqrt{2}$

16. What is  $\mathbb{E}[X^k]$ ?

- (a)  $p^k$  (b)  $p^k - (1 - p)^k$  (c)  $p(-1)^k + (1 - p)$  (d)  **$p + (1 - p)(-1)^k$**

Let  $X, Y$  be two independent standard Gaussian random variables.

17. What is  $\mathbb{E}[X^2Y]$ ?

- (a)  **$0$**  (b)  $1$  (c)  $2$  (d)  $3$

18. What is  $\mathbb{V}(X + Y)$ ?

- (a)  $0$  (b)  $1$  (c)  **$2$**  (d)  $3$

19. What is  $\mathbb{V}(XY)$ ?

- (a)  $0$  (b)  **$1$**  (c)  $2$  (d)  $3$

20. What is  $\text{Cov}(X, X + Y)$ ?

- (a)  $0$  (b)  **$1$**  (c)  $2$  (d)  $3$

21. What is  $\text{Cov}(X, XY)$ ?

- (a)  **$0$**  (b)  $1$  (c)  $2$  (d)  $3$

Let  $X$  be an exponential random variable with parameter  $1/2$  that models the lifetime (in years) of a lightbulb.<sup>1</sup>

22. What is (approximately) the probability that the lightbulb will last at least 2 years?

- (a)  $0.002$  (b)  **$0.018$**  (c)  $0.180$  (d)  $0.810$

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<sup>1</sup> We use the convention from AoS for the parameter of an exponential distribution.

23. Given that the lightbulb has already lasted for at least 3 years, what is (approximately) the probability that it will last for at least two more years?

(a) 0.002      (b) **0.018**      (c) 0.180      (d) 0.810

Let  $X_1, \dots, X_n$  be i.i.d with mean  $\mu$  and variance  $\sigma^2$ .

24. What is  $\mathbb{E}[\sum_{i=1}^n X_i]$ ?    (a)  $\mu$     (b)  $n\sigma$     (c)  **$n\mu$**     (d)  $\sigma$
25. What is  $\mathbb{V}[\sum_{i=1}^n X_i]$ ?    (a)  $n^2\sigma^2$     (b)  **$n\sigma^2$**     (c)  $n\sigma^2 + n^2\mu^2$     (d)  $n\mu$
26. What is  $\mathbb{E}[(\sum_{i=1}^n X_i)^2]$ ?    (a)  $n^2\mu^2$     (b)  $n\sigma^2$     (c)  $n\mu$     (d)  **$n\sigma^2 + n^2\mu^2$**
27. What is  $\mathbb{E}[\frac{1}{n} \sum_{i=1}^n X_i]$ ?    (a)  $\sigma$     (b)  $n\sigma^2$     (c)  $n\mu$     (d)  **$\mu$**
28. What is  $\mathbb{V}[\frac{1}{n} \sum_{i=1}^n X_i]$ ?    (a)  $\mu$     (b)  **$\sigma^2/n$**     (c)  $\sigma^2$     (d)  $n\mu$

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The following problems are of “show-your-work” type. You get complete points for a solution if you show your complete work. That includes all computations that lead to your answer with appropriate reasoning. Please upload your work to Gradescope.

**Problem 2.** Let  $X_n \sim \text{Unif}(-\frac{1}{n}, \frac{1}{n})$  and let  $X$  be a random variable such that  $\mathbb{P}(X = 0) = 1$ .

1. Compute and draw the CDF  $F_n(x)$  and  $F(x)$  of  $X_n$  and  $X$  respectively.

**Solution.**  $F_n(x) = 0$  for  $x \leq -1/n$  and  $F_n(x) = 1$  for  $x \geq 1/n$ . For  $x \in (-1/n, 1/n)$ ,  $F_n(x)$  is linear, from the value zero at  $x = -1/n$  to the value 1 at  $x = 1/n$ .

For  $F$ , we have  $F(x) = 0$  when  $x < 0$ , and  $F(x) = 1$  when  $x \geq 0$ . (In particular,  $F$  is right continuous.)

2. Does  $X_n \xrightarrow{\mathbb{P}} X$ ? (prove or disprove)

**Solution.** For any  $\epsilon > 0$  we have  $\mathbb{P}(|X_n - X| > \epsilon) = \mathbb{P}(|X_n| > \epsilon) = 0$  for all  $n \geq 1/\epsilon$ . Therefore,  $\mathbb{P}(|X_n - X| > \epsilon) \rightarrow 0$  as  $n \rightarrow \infty$ , so  $X_n$  converges to  $X$  in probability.

3. Does  $X_n \rightsquigarrow X$ ? (prove or disprove)

**Solution.**  $X_n$  converges to  $X$  in distribution because it converges in probability.

**Problem 3.** Let  $X \sim \mathcal{N}(2, 1.44)$ . Compute the following probabilities:

1.  $\mathbb{P}(2X - 1 < 0) = \Phi(-5/4)$
2.  $\mathbb{P}(\frac{7}{5} \leq X \leq \frac{16}{5}) = \Phi(1) - \Phi(-1/2)$

3.  $\mathbb{P}\left(X > \frac{16}{5} \mid X > \frac{7}{5}\right) = \frac{1-\Phi(1)}{\Phi(1/2)}$
4.  $\mathbb{P}\left(X \leq \frac{4}{5} \text{ or } X \geq \frac{16}{5}\right) = 2(1 - \Phi(1))$

**Problem 4.** Let

$$\begin{pmatrix} X \\ Y \end{pmatrix} \sim \mathcal{N}\left(\begin{pmatrix} 1 \\ 0 \end{pmatrix}, \begin{pmatrix} 1 & 1 \\ 1 & 2 \end{pmatrix}\right)$$

Compute the following quantities (show your work):

1.  $\mathbb{V}[X] = 1$ .
2.  $\mathbb{E}[Y^2 + X] = \mathbb{V}(Y) + (\mathbb{E}Y)^2 + \mathbb{E}X = 2 + 0^2 + 1 = 3$ .
3.  $\mathbb{E}[(X - Y)^2] = \mathbb{E}X^2 + \mathbb{E}Y^2 - 2\mathbb{E}XY = \mathbb{V}(X) + (\mathbb{E}X)^2 + \mathbb{V}(Y) + (\mathbb{E}Y)^2 - 2\text{Cov}(X, Y) - 2(\mathbb{E}X)(\mathbb{E}Y) = 1 + 1^2 + 2 + 0^2 - 2 - 0 = 2$ .
4.  $\mathbb{V}[X + 2Y] = \mathbb{V}(X) + 4\mathbb{V}(Y) + 4\text{Cov}(X, Y) = 1 + 8 + 4 = 13$ .
5. Find  $\alpha > 0$  such that  $\alpha X = Y$  with probability 1 or prove that no such  $\alpha$  exists.

**Solution.** Notice that  $\mathbb{E}[\alpha X] = \alpha$  while  $\mathbb{E}Y = 0$ . Since  $A = B$  with probability 1 implies  $\mathbb{E}A = \mathbb{E}B$ , no  $\alpha > 0$  can exist with the desired property.

**Problem 5.** We are testing  $n$  lightbulbs. Each bulb independently *passes* some quality check with probability  $p \in (0, 1)$  and *fails* with probability  $1 - p$ . Let  $X_i \in \{0, 1\}$  indicate the outcome, where  $X_i = 1$  means the bulb passes.

Conditioned on  $X_i$ , the lifetime  $Y_i$  of bulb  $i$  is exponentially distributed:

$$Y_i \mid (X_i = 1) \sim \text{Exp}(\lambda_1), \quad Y_i \mid (X_i = 0) \sim \text{Exp}(\lambda_0),$$

where  $\lambda_0, \lambda_1 > 0$ .<sup>2</sup> Assume the pairs  $(X_i, Y_i)$  are i.i.d. across  $i$ .

1. Let  $(X, Y)$  be a copy of  $(X_1, Y_1)$  and let  $T := XY$ . Compute the following quantities in terms of  $p, \lambda_0, \lambda_1$ :  
 $\mathbb{E}[Y]$ ,  $\mathbb{V}(Y)$ ,  $\text{Cov}(X, Y)$ ,  $\mathbb{E}[T]$ ,  $\mathbb{V}(T)$ , and  $\text{Cov}(X, T)$ .

**Solution.** We have  $\mathbb{E}[X] = p$  and  $\text{Var}(X) = p(1 - p)$ .

Let  $m := \mathbb{E}[Y]$ . Then

$$m = \mathbb{E}[\mathbb{E}[Y \mid X]] = p\lambda_1 + (1 - p)\lambda_0.$$

By the law of total variance,

$$\begin{aligned} \text{Var}(Y) &= \mathbb{E}[\text{Var}(Y \mid X)] + \text{Var}(\mathbb{E}[Y \mid X]) \\ &= p\lambda_1^2 + (1 - p)\lambda_0^2 + p(1 - p)(\lambda_1 - \lambda_0)^2. \end{aligned}$$

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<sup>2</sup>See 1

Next,

$$\begin{aligned}\text{Cov}(X, Y) &= \mathbb{E}[XY] - \mathbb{E}[X]\mathbb{E}[Y] = \mathbb{E}[T] - pm, \\ \mathbb{E}[T] &= \mathbb{E}[\mathbb{E}[XY \mid X]] = p \mathbb{E}[Y \mid X = 1] = p\lambda_1,\end{aligned}$$

so  $\text{Cov}(X, Y) = p(1 - p)(\lambda_1 - \lambda_0)$ .

Since  $T = XY$  takes the value  $Y$  when  $X = 1$  and 0 when  $X = 0$ , we have

$$\mathbb{E}[T^2] = \mathbb{E}[XY^2] = p \mathbb{E}[Y^2 \mid X = 1] = p(\text{Var}(Y \mid X = 1) + (\mathbb{E}[Y \mid X = 1])^2) = 2p\lambda_1^2,$$

hence

$$\text{Var}(T) = \mathbb{E}[T^2] - (\mathbb{E}[T])^2 = 2p\lambda_1^2 - (p\lambda_1)^2 = p(2 - p)\lambda_1^2.$$

Finally,  $XT = X^2Y = XY = T$ , so

$$\text{Cov}(X, T) = \mathbb{E}[XT] - \mathbb{E}[X]\mathbb{E}[T] = \mathbb{E}[T] - p(p\lambda_1) = p(1 - p)\lambda_1.$$

Define

$$\bar{X}_n = \frac{1}{n} \sum_{i=1}^n X_i, \quad \bar{Y}_n = \frac{1}{n} \sum_{i=1}^n Y_i, \quad \text{and} \quad \bar{T}_n = \frac{1}{n} \sum_{i=1}^n X_i Y_i.$$

Write a central limit theorem for each of the following quantities in the form

$$\sqrt{n}(Z_n - \mu) \rightsquigarrow \mathcal{N}(0, \sigma^2) \quad \text{or} \quad \sqrt{n}(Z_n - \mu) \rightsquigarrow \mathcal{N}(0, \Sigma),$$

depending on whether  $Z_n$  is a random variable or a random vector.

$$2. \ Z_n = \begin{pmatrix} \bar{X}_n \\ \bar{T}_n \end{pmatrix}.$$

**Solution.** Since  $(\bar{X}_n, \bar{T}_n)$  is the sample mean of the i.i.d. vectors  $(X_i, X_i Y_i)$ , the CLT applies. The mean is  $\mu = (p, p\lambda_1)$  and the covariance matrix is

$$\Sigma_{X, XY} = \begin{pmatrix} p(1 - p) & p(1 - p)\lambda_1 \\ p(1 - p)\lambda_1 & p(2 - p)\lambda_1^2 \end{pmatrix}.$$

Therefore,

$$\sqrt{n} \left( \begin{pmatrix} \bar{X}_n \\ \bar{T}_n \end{pmatrix} - \begin{pmatrix} p \\ p\lambda_1 \end{pmatrix} \right) \rightsquigarrow \mathcal{N}(0, \Sigma_{X, XY}).$$

$$3. \ Z_n = \log(\bar{Y}_n).$$

**Solution.** By the CLT,  $\sqrt{n}(\bar{Y}_n - m) \rightsquigarrow \mathcal{N}(0, \text{Var}(Y))$ .

Let  $g(y) = \log y$ . Then  $g(m) = \log m$  and  $g'(m) = 1/m$ . By the delta method,

$$\sqrt{n}(\log \bar{Y}_n - \log m) \rightsquigarrow \mathcal{N}\left(0, \frac{\text{Var}(Y)}{m^2}\right).$$

Define the average lifetime among passed bulbs

$$\hat{\lambda}_{1,n} = \begin{cases} \frac{\sum_{i=1}^n X_i Y_i}{\sum_{i=1}^n X_i} = \frac{\bar{T}_n}{\bar{X}_n}, & \text{if } \sum_{i=1}^n X_i > 0, \\ 0, & \text{if } \sum_{i=1}^n X_i = 0, \end{cases}$$

and the corresponding rate estimator

$$\hat{\rho}_{1,n} = \begin{cases} \frac{1}{\hat{\lambda}_{1,n}} = \frac{\bar{X}_n}{\bar{T}_n}, & \text{if } \sum_{i=1}^n X_i > 0, \\ 0, & \text{if } \sum_{i=1}^n X_i = 0. \end{cases}$$

(Note that  $\mathbb{P}(\sum_{i=1}^n X_i = 0) = (1-p)^n \rightarrow 0$ , so this convention does not affect any CLT/delta-method limits.)

Write a CLT for each of the following choices of  $Z_n$ :

4.  $Z_n = \hat{\lambda}_{1,n}$ .

**Solution.** We use the vector CLT from the previous part and apply the delta method to  $g(\bar{X}_n, \bar{T}_n)$ , with  $\mu = (p, p\lambda_1)$  and covariance  $\Sigma_{X,Y}$ .

Let  $g(s, t) = t/s$ . Then  $g(\mu) = \lambda_1$  and  $\nabla g(s, t) = (-t/s^2, 1/s)$ , so  $\nabla g(\mu) = (-\lambda_1/p, 1/p)$ . Therefore,

$$\sqrt{n}(\hat{\lambda}_{1,n} - \lambda_1) \rightsquigarrow \mathcal{N}\left(0, \frac{\lambda_1^2}{p}\right).$$

5.  $Z_n = \hat{\rho}_{1,n}$ .

**Solution.** Let  $g(s, t) = s/t$ . Then  $g(\mu) = 1/\lambda_1$  and  $\nabla g(s, t) = (1/t, -s/t^2)$ , so  $\nabla g(\mu) = (1/(p\lambda_1), -1/(p\lambda_1^2))$ . Therefore,

$$\sqrt{n}\left(\hat{\rho}_{1,n} - \frac{1}{\lambda_1}\right) \rightsquigarrow \mathcal{N}\left(0, \frac{1}{p\lambda_1^2}\right).$$

## Extra probability practice (not graded)

**Problem 6.** Let  $X$  be a random variable with pmf given by

$$\mathbb{P}(X = k) = \frac{c\lambda^k}{k!}, k = 0, 1, 2, \dots$$

for some  $\lambda > 0$ .

1. What is the value of  $c$ ? (a) 1 (b)  $\lambda$  (c)  $e^{-\lambda}$  (d)  $e^\lambda$
2. What is  $\mathbb{E}[X]$ ? (a) 1 (b)  $\lambda$  (c)  $e^{-\lambda}$  (d)  $e^\lambda$
3. What is  $\mathbb{V}[X]$ ? (a) 1 (b)  $\lambda$  (c)  $e^{-\lambda}$  (d)  $e^\lambda$

Let  $X$  be a uniform random variable in the interval  $[2, 8]$ .

4. What is  $\mathbb{E}[X]$ ?  
(a) 2 (b) 3 (c) **5** (d) 8
5. What is  $\mathbb{V}[X]$ ?  
(a) 2 (b) **3** (c) 5 (d) 8
6. What<sup>3</sup> is  $\mathbb{P}[\log(X) \leq 1]$  approximately?  
(a) **.12** (b) 0.8 (c)  $-1$  (d) 0

Let  $X$  be an exponential random variable with parameter 3 and  $Y$  be a Poisson random variable with parameter 2.<sup>4</sup> Assume that  $X$  and  $Y$  are independent.

7. What is  $\mathbb{E}[X^2 + Y^2]$ ?  
(a) 12 (b) 23 (c) **24** (d) 36
8. What is  $\mathbb{E}[X^2 Y]$ ?  
(a) 12 (b) 23 (c) 24 (d) **36**
9. What is  $\mathbb{V}(2X + 3Y)$ ?  
(a) 24 (b) 34 (c) 44 (d) **54**

Let  $X \geq 0$  be a positive random variable such that  $\mathbb{E}[X] = \lambda$ .

10. Which is correct?  
(a)  $\mathbb{E}[1/X] = 1/\lambda$  (b)  $\mathbb{E}[1/X] \geq 1/\lambda$  (c)  $\mathbb{E}[1/X] \leq 1/\lambda$

Let  $X$  and  $Y$  be two random variables such that  $X$  is a Bernoulli random variable with parameter  $p \in (0, 1)$ , and  $Y^2 + 2XY = 3X^2$  almost surely.

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<sup>3</sup>all logs are natural (base  $e$ ) unless specified otherwise

<sup>4</sup>See 1

11. What is  $\mathbb{E}[Y]$ ?

- (a) 0      (b)  $-3p$       (c)  $X$       (d)  $-3X$       (e) **Some number in  $[-3, 1]$**

Let  $X$  and  $Y$  be two independent, identically distributed random variables.

12. Compute the conditional expectation  $\mathbb{E}[X|X + Y = x]$ .

- (a)  **$x/2$**       (b)  $x$       (c)  $-x$       (d) 0