

18.650. Fundamentals of Statistics Spring 2026. Problem Set 1

Due Wednesday, Feb 18

Problem 1 (Probability Review). *Please use the accompanying bubble sheet for submitting your solutions for this problem.*

In what follows, Φ is the CDF of the standard Gaussian (Normal) distribution.

Let X be a random variable taking values between 0 and π , with pdf given by

$$f(x) = c \sin x, \quad x \in [0, \pi].$$

1. What is the value of c ? (a) π (b) **$1/2$** (c) 2 (d) $1/\pi$
2. What is $\mathbb{E}[X]$? (a) **$\pi/2$** (b) π (c) 1 (d) $1/2$

Let X be a Gaussian random variable with mean $\mu > 0$ and variance σ^2 .

3. What is $\mathbb{E}[X]$? (a) 0 (b) σ^2 (c) $\mu^2 + \sigma^2$ (d) **μ**
4. What is $\mathbb{V}[X]$? (a) **σ^2** (b) $\mu^2 + \sigma^2$ (c) μ^2 (d) σ
5. What is $\mathbb{E}[X^2]$? (a) μ^2 (b) $(\mu + \sigma)^2$ (c) **$\mu^2 + \sigma^2$** (d) σ^2
6. What is $\mathbb{E}[X^3]$? (a) $\mu^3 + 3\mu\sigma$ (b) **$\mu^3 + 3\mu\sigma^2$** (c) μ^3 (d) $3\mu\sigma^2$
7. What is $\mathbb{V}[X^2]$? (a) $4\mu^2\sigma^2$ (b) $(\mu^2 + \sigma^2)^2$ (c) $2\sigma^4$ (d) **$4\mu^2\sigma^2 + 2\sigma^4$**
8. What is $\mathbb{P}(X > 0)$ in terms of Φ ?
(a) **$\Phi(\mu/\sigma)$** (b) $\Phi(\mu/\sigma^2)$ (c) $1 - \Phi(\mu/\sigma)$ (d) $\Phi(\mu)$

Let $X \sim \text{Lognormal}(\mu, \sigma^2)$, i.e., $\log X \sim \mathcal{N}(\mu, \sigma^2)$ with $\sigma > 0$.

9. What is $\mathbb{E}[X]$? (a) e^μ (b) **$e^{\mu+\sigma^2/2}$** (c) $e^{\mu+\sigma^2}$ (d) $\mu + \sigma^2/2$
10. What is the median of X ? (a) $e^{\mu+\sigma^2/2}$ (b) $e^{\mu-\sigma^2}$ (c) **e^μ** (d) μ
11. What is $\mathbb{P}(X > 1)$ in terms of Φ ?
(a) $1 - \Phi(\mu/\sigma)$ (b) $\Phi(\mu)$ (c) $\Phi(\mu/\sigma^2)$ (d) **$\Phi(\mu/\sigma)$**

Let $X, Y \sim \text{Lognormal}(0, 1)$ be independent.

12. What is $\mathbb{P}(XY > 1)$? (a) $1/4$ (b) **$1/2$** (c) $1/\sqrt{2\pi}$ (d) $1/e$

Let X be a random variable such that

$$X = \begin{cases} 1 & \text{with probability } p \\ -1 & \text{with probability } 1-p \end{cases}$$

for some $p \in [0, 1]$.

13. What is $\mathbb{E}[X]$?
 (a) $-p$ (b) p (c) $1 - 2p$ (d) **$2p - 1$**

14. What is $\mathbb{V}[X]$?
 (a) $p(1 - p)$ (b) $4p - p^2$ (c) **$4p(1 - p)$** (d) $4p^2(1 - p)$

15. For what p is $\mathbb{V}[X]$ maximized?
 (a) 1 (b) 0 (c) **0.5** (d) $1/\sqrt{2}$

16. What is $\mathbb{E}[X^k]$?
 (a) p^k (b) $p^k - (1 - p)^k$ (c) $p(-1)^k + (1 - p)$ (d) **$p + (1 - p)(-1)^k$**

Let X, Y be two independent standard Gaussian random variables.

17. What is $\mathbb{E}[X^2Y]$?
 (a) **0** (b) 1 (c) 2 (d) 3

18. What is $\mathbb{V}(X + Y)$?
 (a) 0 (b) 1 (c) **2** (d) 3

19. What is $\mathbb{V}(XY)$?
 (a) 0 (b) **1** (c) 2 (d) 3

20. What is $\text{Cov}(X, X + Y)$?
 (a) 0 (b) **1** (c) 2 (d) 3

21. What is $\text{Cov}(X, XY)$?
 (a) **0** (b) 1 (c) 2 (d) 3

Let X be an exponential random variable with parameter $1/2$ that models the lifetime (in years) of a lightbulb.¹

22. What is (approximately) the probability that the lightbulb will last at least 2 years?
 (a) 0.002 (b) **0.018** (c) 0.180 (d) 0.810

¹ We use the convention from AoS for the parameter of an exponential distribution.

23. Given that the lightbulb has already lasted for at least 3 years, what is (approximately) the probability that it will last for at least two more years?

- (a) 0.002 (b) **0.018** (c) 0.180 (d) 0.810

Let X_1, \dots, X_n be i.i.d with mean μ and variance σ^2 .

24. What is $\mathbb{E}[\sum_{i=1}^n X_i]$? (a) μ (b) $n\sigma$ (c) **$n\mu$** (d) σ

25. What is $\mathbb{V}[\sum_{i=1}^n X_i]$? (a) $n^2\sigma^2$ (b) **$n\sigma^2$** (c) $n\sigma^2 + n^2\mu^2$ (d) $n\mu$

26. What is $\mathbb{E}[(\sum_{i=1}^n X_i)^2]$? (a) $n^2\mu^2$ (b) $n\sigma^2$ (c) $n\mu$ (d) **$n\sigma^2 + n^2\mu^2$**

27. What is $\mathbb{E}[\frac{1}{n} \sum_{i=1}^n X_i]$? (a) σ (b) $n\sigma^2$ (c) $n\mu$ (d) **μ**

28. What is $\mathbb{V}[\frac{1}{n} \sum_{i=1}^n X_i]$? (a) μ (b) **σ^2/n** (c) σ^2 (d) $n\mu$

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The following problems are of “show-your-work” type. You get complete points for a solution if you show your complete work. That includes all computations that lead to your answer with appropriate reasoning. Please upload your work to Gradescope.

Problem 2. Let $X_n \sim \text{Unif}(-\frac{1}{n}, \frac{1}{n})$ and let X be a random variable such that $\mathbb{P}(X = 0) = 1$.

1. Compute and draw the CDF $F_n(x)$ and $F(x)$ of X_n and X respectively.

Solution. $F_n(x) = 0$ for $x \leq -1/n$ and $F_n(x) = 1$ for $x \geq 1/n$. For $x \in (-1/n, 1/n)$, $F_n(x)$ is linear, from the value zero at $x = -1/n$ to the value 1 at $x = 1/n$.

For F , we have $F(x) = 0$ when $x < 0$, and $F(x) = 1$ when $x \geq 0$. (In particular, F is right continuous.)

2. Does $X_n \xrightarrow{\mathbb{P}} X$? (prove or disprove)

Solution. For any $\epsilon > 0$ we have $\mathbb{P}(|X_n - X| > \epsilon) = \mathbb{P}(|X_n| > \epsilon) = 0$ for all $n \geq 1/\epsilon$. Therefore, $\mathbb{P}(|X_n - X| > \epsilon) \rightarrow 0$ as $n \rightarrow \infty$, so X_n converges to X in probability.

3. Does $X_n \rightsquigarrow X$? (prove or disprove)

Solution. X_n converges to X in distribution because it converges in probability.

Problem 3. Let $X \sim \mathcal{N}(2, 1.44)$. Compute the following probabilities:

1. $\mathbb{P}(2X - 1 < 0) = \Phi(-5/4)$

2. $\mathbb{P}\left(\frac{7}{5} \leq X \leq \frac{16}{5}\right) = \Phi(1) - \Phi(-1/2)$

$$3. \mathbb{P}(X > \frac{16}{5} \mid X > \frac{7}{5}) = \frac{1 - \Phi(1)}{\Phi(1/2)}$$

$$4. \mathbb{P}(X \leq \frac{4}{5} \text{ or } X \geq \frac{16}{5}) = 2(1 - \Phi(1))$$

Problem 4. Let

$$\begin{pmatrix} X \\ Y \end{pmatrix} \sim \mathcal{N}\left(\begin{pmatrix} 1 \\ 0 \end{pmatrix}, \begin{pmatrix} 1 & 1 \\ 1 & 2 \end{pmatrix}\right)$$

Compute the following quantities (show your work):

1. $\mathbb{V}[X] = 1.$
2. $\mathbb{E}[Y^2 + X] = \mathbb{V}(Y) + (\mathbb{E}Y)^2 + \mathbb{E}X = 2 + 0^2 + 1 = 3.$
3. $\mathbb{E}[(X - Y)^2] = \mathbb{E}X^2 + \mathbb{E}Y^2 - 2\mathbb{E}XY = \mathbb{V}(X) + (\mathbb{E}X)^2 + \mathbb{V}(Y) + (\mathbb{E}Y)^2 - 2\text{Cov}(X, Y) - 2(\mathbb{E}X)(\mathbb{E}Y) = 1 + 1^2 + 2 + 0^2 - 2 - 0 = 2.$
4. $\mathbb{V}[X + 2Y] = \mathbb{V}(X) + 4\mathbb{V}(Y) + 4\text{Cov}(X, Y) = 1 + 8 + 4 = 13.$
5. Find $\alpha > 0$ such that $\alpha X = Y$ with probability 1 or prove that no such α exists.

Solution. Notice that $\mathbb{E}[\alpha X] = \alpha$ while $\mathbb{E}Y = 0$. Since $A = B$ with probability 1 implies $\mathbb{E}A = \mathbb{E}B$, no $\alpha > 0$ can exist with the desired property.

Problem 5. We are testing n lightbulbs. Each bulb independently *passes* some quality check with probability $p \in (0, 1)$ and *fails* with probability $1 - p$. Let $X_i \in \{0, 1\}$ indicate the outcome, where $X_i = 1$ means the bulb passes.

Conditioned on X_i , the lifetime Y_i of bulb i is exponentially distributed:

$$Y_i \mid (X_i = 1) \sim \text{Exp}(\lambda_1), \quad Y_i \mid (X_i = 0) \sim \text{Exp}(\lambda_0),$$

where $\lambda_0, \lambda_1 > 0$.² Assume the pairs (X_i, Y_i) are i.i.d. across i .

1. Let (X, Y) be a copy of (X_1, Y_1) and let $T := XY$. Compute the following quantities in terms of p, λ_0, λ_1 :
 $\mathbb{E}[Y], \mathbb{V}(Y), \text{Cov}(X, Y), \mathbb{E}[T], \mathbb{V}(T), \text{ and Cov}(X, T).$

Solution. We have $\mathbb{E}[X] = p$ and $\text{Var}(X) = p(1 - p)$.

Let $m := \mathbb{E}[Y]$. Then

$$m = \mathbb{E}[\mathbb{E}[Y \mid X]] = p\lambda_1 + (1 - p)\lambda_0.$$

By the law of total variance,

$$\begin{aligned} \text{Var}(Y) &= \mathbb{E}[\text{Var}(Y \mid X)] + \text{Var}(\mathbb{E}[Y \mid X]) \\ &= p\lambda_1^2 + (1 - p)\lambda_0^2 + p(1 - p)(\lambda_1 - \lambda_0)^2. \end{aligned}$$

²See 1

Next,

$$\begin{aligned}\text{Cov}(X, Y) &= \mathbb{E}[XY] - \mathbb{E}[X]\mathbb{E}[Y] = \mathbb{E}[T] - pm, \\ \mathbb{E}[T] &= \mathbb{E}[\mathbb{E}[XY | X]] = p\mathbb{E}[Y | X = 1] = p\lambda_1,\end{aligned}$$

so $\text{Cov}(X, Y) = p(1-p)(\lambda_1 - \lambda_0)$.

Since $T = XY$ takes the value Y when $X = 1$ and 0 when $X = 0$, we have

$$\mathbb{E}[T^2] = \mathbb{E}[XY^2] = p\mathbb{E}[Y^2 | X = 1] = p(\text{Var}(Y | X = 1) + (\mathbb{E}[Y | X = 1])^2) = 2p\lambda_1^2,$$

hence

$$\text{Var}(T) = \mathbb{E}[T^2] - (\mathbb{E}[T])^2 = 2p\lambda_1^2 - (p\lambda_1)^2 = p(2-p)\lambda_1^2.$$

Finally, $XT = X^2Y = XY = T$, so

$$\text{Cov}(X, T) = \mathbb{E}[XT] - \mathbb{E}[X]\mathbb{E}[T] = \mathbb{E}[T] - p(p\lambda_1) = p(1-p)\lambda_1.$$

Define

$$\bar{X}_n = \frac{1}{n} \sum_{i=1}^n X_i, \quad \bar{Y}_n = \frac{1}{n} \sum_{i=1}^n Y_i, \quad \text{and} \quad \bar{T}_n = \frac{1}{n} \sum_{i=1}^n X_i Y_i.$$

Write a central limit theorem for each of the following quantities in the form

$$\sqrt{n}(Z_n - \mu) \rightsquigarrow \mathcal{N}(0, \sigma^2) \quad \text{or} \quad \sqrt{n}(Z_n - \mu) \rightsquigarrow \mathcal{N}(0, \Sigma),$$

depending on whether Z_n is a random variable or a random vector.

$$2. \ Z_n = \begin{pmatrix} \bar{X}_n \\ \bar{T}_n \end{pmatrix}.$$

Solution. Since (\bar{X}_n, \bar{T}_n) is the sample mean of the i.i.d. vectors $(X_i, X_i Y_i)$, the CLT applies. The mean is $\mu = (p, p\lambda_1)$ and the covariance matrix is

$$\Sigma_{X,XY} = \begin{pmatrix} p(1-p) & p(1-p)\lambda_1 \\ p(1-p)\lambda_1 & p(2-p)\lambda_1^2 \end{pmatrix}.$$

Therefore,

$$\sqrt{n} \left(\begin{pmatrix} \bar{X}_n \\ \bar{T}_n \end{pmatrix} - \begin{pmatrix} p \\ p\lambda_1 \end{pmatrix} \right) \rightsquigarrow \mathcal{N}(0, \Sigma_{X,XY}).$$

$$3. \ Z_n = \log(\bar{Y}_n).$$

Solution. By the CLT, $\sqrt{n}(\bar{Y}_n - m) \rightsquigarrow \mathcal{N}(0, \text{Var}(Y))$.

Let $g(y) = \log y$. Then $g(m) = \log m$ and $g'(m) = 1/m$. By the delta method,

$$\sqrt{n}(\log \bar{Y}_n - \log m) \rightsquigarrow \mathcal{N}\left(0, \frac{\text{Var}(Y)}{m^2}\right).$$

Define the average lifetime among passed bulbs

$$\hat{\lambda}_{1,n} = \begin{cases} \frac{\sum_{i=1}^n X_i Y_i}{\sum_{i=1}^n X_i} = \frac{\bar{T}_n}{\bar{X}_n}, & \text{if } \sum_{i=1}^n X_i > 0, \\ 0, & \text{if } \sum_{i=1}^n X_i = 0, \end{cases}$$

and the corresponding rate estimator

$$\hat{\rho}_{1,n} = \begin{cases} \frac{1}{\hat{\lambda}_{1,n}} = \frac{\bar{X}_n}{\bar{T}_n}, & \text{if } \sum_{i=1}^n X_i > 0, \\ 0, & \text{if } \sum_{i=1}^n X_i = 0. \end{cases}$$

(Note that $\mathbb{P}(\sum_{i=1}^n X_i = 0) = (1-p)^n \rightarrow 0$, so this convention does not affect any CLT/delta-method limits.)

Write a CLT for each of the following choices of Z_n :

$$4. Z_n = \hat{\lambda}_{1,n}.$$

Solution. We use the vector CLT from the previous part and apply the delta method to $g(\bar{X}_n, \bar{T}_n)$, with $\mu = (p, p\lambda_1)$ and covariance $\Sigma_{X,XY}$.

Let $g(s, t) = t/s$. Then $g(\mu) = \lambda_1$ and $\nabla g(s, t) = (-t/s^2, 1/s)$, so $\nabla g(\mu) = (-\lambda_1/p, 1/p)$. Therefore,

$$\sqrt{n}(\hat{\lambda}_{1,n} - \lambda_1) \rightsquigarrow \mathcal{N}\left(0, \frac{\lambda_1^2}{p}\right).$$

$$5. Z_n = \hat{\rho}_{1,n}.$$

Solution. Let $g(s, t) = s/t$. Then $g(\mu) = 1/\lambda_1$ and $\nabla g(s, t) = (1/t, -s/t^2)$, so $\nabla g(\mu) = (1/(p\lambda_1), -1/(p\lambda_1^2))$. Therefore,

$$\sqrt{n}\left(\hat{\rho}_{1,n} - \frac{1}{\lambda_1}\right) \rightsquigarrow \mathcal{N}\left(0, \frac{1}{p\lambda_1^2}\right).$$

Extra probability practice (not graded)

Problem 6. Let X be a random variable with pmf given by

$$\mathbb{P}(X = k) = \frac{c\lambda^k}{k!}, k = 0, 1, 2, \dots$$

for some $\lambda > 0$.

1. What is the value of c ? (a) 1 (b) λ (c) $e^{-\lambda}$ (d) e^λ
2. What is $\mathbb{E}[X]$? (a) 1 (b) λ (c) $e^{-\lambda}$ (d) e^λ
3. What is $\mathbb{V}[X]$? (a) 1 (b) λ (c) $e^{-\lambda}$ (d) e^λ

Let X be a uniform random variable in the interval $[2, 8]$.

4. What is $\mathbb{E}[X]$?
(a) 2 (b) 3 (c) **5** (d) 8
5. What is $\mathbb{V}[X]$?
(a) 2 (b) **3** (c) 5 (d) 8
6. What³ is $\mathbb{P}[\log(X) \leq 1]$ approximately?
(a) **.12** (b) 0.8 (c) -.1 (d) 0

Let X be an exponential random variable with parameter 3 and Y be a Poisson random variable with parameter 2.⁴ Assume that X and Y are independent.

7. What is $\mathbb{E}[X^2 + Y^2]$?
(a) 12 (b) 23 (c) **24** (d) 36
8. What is $\mathbb{E}[X^2 Y]$?
(a) 12 (b) 23 (c) 24 (d) **36**
9. What is $\mathbb{V}(2X + 3Y)$?
(a) 24 (b) 34 (c) 44 (d) **54**

Let $X \geq 0$ be a positive random variable such that $\mathbb{E}[X] = \lambda$.

10. Which is correct?
(a) $\mathbb{E}[1/X] = 1/\lambda$ (b) **$\mathbb{E}[1/X] \geq 1/\lambda$** (c) $\mathbb{E}[1/X] \leq 1/\lambda$

Let X and Y be two random variables such that X is a Bernoulli random variable with parameter $p \in (0, 1)$, and $Y^2 + 2XY = 3X^2$ almost surely.

³all logs are natural (base e) unless specified otherwise

⁴See 1

11. What is $\mathbb{E}[Y]$?

- (a) 0 (b) $-3p$ (c) X (d) $-3X$ (e) **Some number in $[-3, 1]$**

Let X and Y be two independent, identically distributed random variables.

12. Compute the conditional expectation $\mathbb{E}[X|X + Y = x]$.

- (a) **$x/2$** (b) x (c) $-x$ (d) 0