

Lecture 6:

pdf of $X \in \mathbb{R}^n$ is a func $f: \mathbb{R}^n \rightarrow \mathbb{R}_{\geq 0}$

$$\int_{\mathbb{R}^n} f(x) dx = 1$$

$$R \subseteq \mathbb{R}^n; \quad P(X \in R) = \int_R f(x) dx$$

$$R = [a_1, b_1] \times [a_2, b_2] \times \dots \times [a_n, b_n]$$

$$\{(x_1, \dots, x_n) : a_i \leq x_i \leq b_i \text{ for } i=1, \dots, n\}$$

$$\int_R f = \int_{a_n}^{b_n} \dots \int_{a_1}^{b_1} f(x_1, \dots, x_n) dx_1 \dots dx_n$$

Case 1: indep. case

$(x_1, \dots, x_n \text{ are indep})$

$$f(x_1, \dots, x_n) = f_1(x_1) \dots f_n(x_n) = \prod_{i=1}^n f_i(x_i)$$

f_i is the pdf of X_i

$$\text{esp if } X_i \text{ are i.i.d, } f(x_1, \dots, x_n) = \prod_{i=1}^n f_n(x_i)$$

$$\rightarrow f(x)$$

Conditional Density:

$X = (x_1, \dots, x_n)$ with pdf f
conditional pdf of X_n
given $x_1 = x_1, \dots, x_{n-1} = x_{n-1}$

$$\text{is } f(x_n | X_1 = x_1, \dots, X_{n-1} = x_{n-1})$$

$$= \frac{f(x_1, \dots, x_{n-1}, x_n)}{\int_{\mathbb{R}} f(x_1, \dots, x_{n-1}, x_n) dx_n}$$

↓
conditional density func

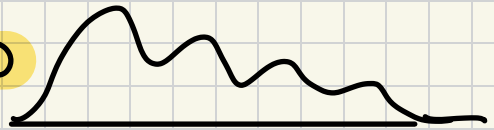
$(x_1, \dots, x_{n-1}) \mapsto \text{(cond) pdf}$
 $f(\cdot | X_1 = x_1, \dots, X_{n-1} = x_{n-1})$

Marginalization:

$$f_1(x_1) = \int_{\mathbb{R}} \dots \int_{\mathbb{R}} f(x_1, \dots, x_n) dx_2 dx_3 \dots dx_n$$

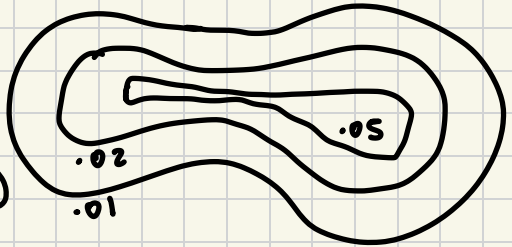
↓
Marginal density func

1-D
 (x_1)



2-D

(x_1, x_2)



3-D

(x_1, x_2, x_3)

	1	2	3
1	X		
2		X	
3			X

Case 2: Gaussian Case

Call $X \in \mathbb{R}^k$ a Gaussian if it's pdf is

$$f(x) = \frac{1}{(2\pi)^{k/2} \det(\Sigma)^{1/2}} \exp\left(-\frac{1}{2} (x-\mu)^T \Sigma^{-1} (x-\mu)\right)$$

here $\mu \in \mathbb{R}^k$ is mean vector

and $\Sigma \in \mathbb{R}^{k \times k}$ is p.d. matrix and
↓
covariance mtr

If A is a matrix, b is a vector

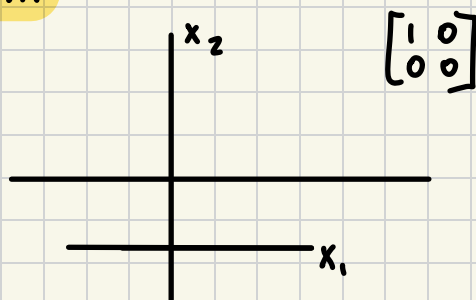
$AX + b$ is still a gaussian. → normal dist.

$$X \sim N(A\mu + b, A\Sigma A^T)$$

$\downarrow \quad \downarrow \quad \downarrow$
 $m \times k \quad k \times k \quad k \times m$

$X \in \mathbb{R}^k$, A is $m \times k$, $b \in \mathbb{R}^m$

$$AX + b \in \mathbb{R}^m$$



Standardization:

$$X \sim N(\mu, \Sigma)$$

$$X - \mu \sim N(0, \Sigma)$$

$$\Sigma^{1/2} = U \Lambda^{1/2} U^T$$

$$\Sigma^{1/2} \Sigma^{1/2} = \Sigma$$

↓

$$N(0, I_k)$$

Σ is p.d.

→ U is orthogonal

$$(UU^T = I_k)$$

Λ is diagonal

$$\Lambda = \begin{bmatrix} \lambda_1 & & 0 \\ & \ddots & \\ 0 & & \lambda_k \end{bmatrix}$$



another way:

$$\Sigma^{1/2} = U \Lambda^{1/2}$$

$$\Sigma^{1/2} (\Sigma^{1/2})^T = \Sigma$$

take $\Sigma^{-1/2} = (\Sigma^{-1})^{1/2} = (U \Lambda^{-1} U^T)^{1/2}$

$$\Sigma^{-1/2} (X - \mu) \sim N(0, \Sigma^{-1/2} \Sigma (\Sigma^{-1/2})^T)$$

$$\underbrace{U \Lambda^{-1/2} U^T U \Lambda U^T U \Lambda^{-1/2} U^T}_{I_k}$$

$$U I_k U^T = I_k$$

X_1, X_2, X_3, \dots iid w/ mean μ and cov. Σ .

each X_i is a rand vec. in \mathbb{R}^k

$$\bar{X}_n = \frac{1}{n} \sum_{i=1}^n X_i$$

$$\sqrt{n} (\bar{X}_n - \mu) \rightsquigarrow N(0, \Sigma)$$

as $n \rightarrow \infty$.