Hashing: Construction of A Hash Family With Smaller Description Length

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Abstract

TO BE UPDATED: The following is a placeholder for abstract. (Gavin

After reading through some related research papers, we decide to look into one of them, titled Ball and bins: smaller hash families and faster evaluation. This paper introduces two new constructions that we could guarantee $O(\log n/\log\log n)$ maximum load when throw n balls into n with either a smaller description length or a faster calculation time. It is well-known that, with high probability, a $O(\log n / \log \log n)$ -wise independent hash family would guarantee max load of $O(\log n/\log\log n)$. Such $O(\log n/\log\log n)$ -wise independent function can be described by $O(\log^2 n/\log\log n)$ bits, which already yields a dramatic improvement over a truly random function. This paper aims to find an even smaller description length or a faster calculation time of the function while maintaining the guarantee of $O(\log n/\log\log n)$ max load. The special part of this research, specifically about the first construction, is that it constructs an innovative structure of a multi-layer random graph. With such construction of multi-layer graph, each layer is considered as a different hash process with different input and output sizes, which is the number of bins in each layer. Thus, we want to look deep in the first construction of this research and elaborate more details.

1 - Introduction

A traditional analysis of randomized algorithm to map m balls into n bins independently and uniformly guarantees that each bin contains at most $O(\log n/\log\log n)$ balls with high probability, as known as the maximum load of the balls and bins problem. For a truly random hash function $h(x): M \to N$, it would take O(mlogn) space to store it. The traditional analysis with the use of truly random hash functions is impractical in various real-world applications because of the space to store the hash functions. Hence, a weaker notion of randomness, kwise independence, is introduced to solve this issue. It is specifically well-studied in the case of mapping n balls into n bins that any $O(\log n/\log\log n)$ -wise independent hash families can guarantee the maximum load of $O(\log n/\log\log n)$ with high probability.

This paper will continue to study the problems of mapping n balls into n bins with a construction of hash functions that require a smaller description length given the inspiration from the paper, titled Ball and bins: smaller hash families and faster evaluation. By using a $O(\log n/\log\log n)$ -wise independent hash families, the hash functions can be described by $O(\log^2 n/\log\log n)$ bits, which itself yields a dramatic improvement over the description length of a truly random functions. We would like to provide an explicit family of hash functions to guarantee the same maximum load of $O(\log n/\log\log n)$ with high probability, and each hash function can be strictly described by $o(\log^2 n/\log\log n)$ bits. We provide an overview of the construction below.

1.1 - Construction: COPIED AND PASTED(Gavin 8.9)

Our construction is based on concatenating the outputs of O(loglogn) functions which are gradually more independent: each function f in our construction is described using d functions

$$f(x) = h_1(x) \circ \dots \circ h_d(x),$$

where we view the output of each h_i as a binary string, and a \circ denotes the concatenation operator on binary strings. The first function h_1 is only O(1)-wise independent, and the level of independence gradually increases to O(logn/loglogn)-wise independence for the last function h_d . As we increase the level of independence, we decrease the output length of the functions from $\Omega(logn)$ bits for h_1 to O(loglogn) bits for h_d . We instantiate these O(loglogn) functions using ϵ -biased distributions. The trade-off between the level of independence and the output length implies that each of these functions can be described using only O(logn) bits and evaluated in time O(logn).

1.2 - Contribution:

We note that the above construction is from the paper, titled *Ball and bins: smaller hash families and faster evaluation*. This paper was written to fully understand the construction presented in Section 1.1. This construction was the study results of L. Elisa Celis, Omer Reingold, Gil Segev, and Udi Wieder. This paper will present a full and complete tour guide in understanding that this construction indeed guaranteed the same maximum load with a smaller description length. In the original paper, proofs of many lemmas and theorems are neglected. We have expanded many of lemmas and theorems given in the original paper with elaboration and proofs.

1.3 - Outline:

In Section 2, we will introduce a few pieces of terminology, definitions, lemmas, and theorems that we will be using in latter sections. Section 3 is the essential part of this research paper. It will contain the formal introduction of our construction and formal proof of why this construction works. It will first give a formal description of the construction, and we will be analyzing the construction in which we will be essentially explaining the interpretation of such construction. Finally, we will prove step-by-step that the construction guarantees the description length of hashing functions. Finally, Section 4 will introduce some extensional use of this construction. It will also analyze this construction with correspondence to the trade-off it makes to obtain the smaller description length of the function.

2 - Preliminary Tools

In this section, we present the relevant definitions, lemmas, theorems that we will use to prove the construction.

2.1 - Definitions:

GIVE EXAMPLE to differentiate from original paper like item 6(Gavin 8.9) FINISHed basic definition, need example (Eric 8.9)

We use the unit RAM model throughout the paper. In the RAM model, we assume that we can access an arbitrary position of an array in O(1) time. We also assume that word size is large enough such that it takes O(1) word operation. For the balls and bins problem, we are considering the case where there are exactly n balls and n bins. We want to achieve a maximum load of $O(\log n/\log\log n)$

with smaller than usual description length under this condition. The following are some definitions and terms that we will be using frequently throughout the paper.

- 1. We set log to be base-2 as default if no base is noted.
- 2. For a natural number u, we define the set of integers $\{1, 2, ..., n\}$ as [u].
- 3. We represent a uniform distribution over the set $\{0,1\}^n$ by U_n .
- 4. The term $x \in X$ represents a sample x from a random variable X.
- 5. SD(X,Y) represents the statistical distance between two random variables over finite domain

$$SD(X,Y) = \frac{1}{2} \sum_{w \in \Omega} |Pr(X \in \omega) - Pr(Y \in \omega)|$$

- 6. The term $x \in S$ means that we draw a random sample x uniformly from a finite set S.
- 7. For two bit-string x and y, $x \circ y$ represents the concatenation of x and y bit-string. For example, for x = 1001 and y = 0111,

$$x \circ y = 1001 \circ 0111 = 10010111$$

2.2 - More Definitions:

We present a few more definitions. These definitions are more complicated but essential to understanding the proof the construction.

k-wise δ -dependent: For a family of function $f:[u] \to [v]$, it is k-wise δ -dependent iff

$$SD(X,Y) < \delta$$

for X =distribution $(f(x_1), f(x_2), ..., f(x_k))$ for any distinct $x_1, x_2, ..., x_k \in [u]$ and Y = uniform distribution over $[v]^k$. It will take $O(k \max\{\log u, \log v\})$ bits to describe f and O(k) times to calculate f in the unit RAM model.

 ϵ -biased distribution: A sequence of random variables $X_1, X_2, ..., X_n$ over $\{0, 1\}$ are ϵ -biased if, for any non-empty set $S \subseteq [n]$,

$$|Pr(\bigoplus_{i \in S} X_i = 1) - Pr(\bigoplus_{i \in S} X_i = 0)| \le \epsilon$$

DELETE DEFINITION OF min-entropy , k-source, and any block source since we are not using them (Eric $8.9)\,$

Following definitions will be used to prove the existence of an ϵ -biased distirbution used in the proof of our construction. This ϵ -biased distirbution is based on the work did by Alon et al.included in the paper titled *Simple Constructions*

of Almost k-wise Independent Random Variables (Alon et al., 2002)

Almost k-wise independence: Let S_n be a sample space and $X = x_1...x_n$ be selected uniformly from S_n

1. S_n is (ϵ, k) -independent (in max form) if for any k positions $i_1 < i_2 < ... < i_k$ and any k-bit string α , we have

$$|Pr[x_{i_1}x_{i_2}...x_{i_k}] - 2^{-k}| \le \epsilon$$

2. S_n is ϵ -away (in L_1 form) for k-wise independence if for any k positions $i_1 < i_2 < ... < i_k$, we have

$$\sum_{\alpha \in \{0,1\}^k} |Pr[x_{i_1} x_{i_2} ... x_{i_k}] - 2^{-k}| \le \epsilon$$

Clearly, we could transfer from these two definitions of almost k-wise independence:

1. S_n is (ϵ, k) -independent (in max form). Then, since there are at most 2^k different α , we could state this equations:

$$|Pr[x_{i_1}x_{i_2}...x_{i_k}] - 2^{-k}| \le \epsilon \Rightarrow \sum_{\alpha \in \{0,1\}^k} |Pr[x_{i_1}x_{i_2}...x_{i_k}] - 2^{-k}| \le 2^k \epsilon$$

Thus, S_n is at most $2^k \epsilon$ -away from k-wise independence.

2. S_n is ϵ -away (in L_1 form) for k-wise independence. Then, we could simply state that:

$$\sum_{\alpha \in \{0,1\}^k} |Pr[x_{i_1} x_{i_2} ... x_{i_k}] - 2^{-k}| \le \epsilon \Rightarrow |Pr[x_{i_1} x_{i_2} ... x_{i_k}] - 2^{-k}| \le \epsilon$$

Thus, S_n is also (ϵ, k) -independent

Linear Test: Linear test refers to "linear Boolean tests" or test which take the **exclusive-or** of the bits in some fixed location in the strings. In particular, Linear test of size k means to perform linear boolean tests on k different select positions.

Definition Set: This set includes various definitions used to prove the existence of an ϵ -biased distirbution:

1. Let $(\alpha, \beta)_2$ denote the **inner-product-mod-2** of the binary string $\alpha = \alpha_1 \alpha_2 ... \alpha_n$ and $\beta = \beta_1 \beta_2 ... \beta_n$.

$$(\alpha, \beta)_2 = (\alpha_1 \alpha_2 ... \alpha_n, \beta_1 \beta_2 ... \beta_n)_2 = (\sum_{i=1}^n \alpha_i \beta_i) \mod 2$$

2. A 0-1 random variable X is ϵ -biased if

$$|Pr([x=0]) - Pr([x=1])| \le \epsilon$$

- 3. Let S_n be a sample space and $X = x_1x_2...x_n$ be selected uniformly randomly from S_n . The sample space S_n is said to be ϵ -biased with respect to linear tests if, for every $\alpha = \alpha_1\alpha_2...\alpha_n \in \{0,1\}^n \{0\}^n$, the random variable $(\alpha, X)_2$ is ϵ -biased.
- 4. The sample space S_n is said to be ϵ -biased with respect to linear tests of size at most \mathbf{k} if, for every $\alpha = \alpha_1 \alpha_2 ... \alpha_n \in \{0,1\}^n \{0\}^n$ such that at most \mathbf{k} of the α_i are 1, the random variable $(\alpha, X)_2$ is ϵ -biased.

Notice that the definition of ϵ -biased distribution and ϵ -biased sample space are actually the same with respect to linear tests using XOR operation since we could treat:

- 1. $X_1, X_2, ..., X_n$ in the definition of ϵ -biased distribution as $x_1, x_2, ..., x_n$ which is bits of random variable X in the definition of ϵ -biased sample
- 2. S in the definition of ϵ -biased distribution as α in the definition of ϵ -biased sample space
- 3. \oplus XOR operation in the definition of ϵ -biased distribution as the inner-product-mod-2 operation in the definition of ϵ -biased sample space

2.3 - Theorems:

In this section, we can present the key lemmas, corollaries, and theorems that we use to prove the construction. Refer to Section 2.1 and Section 2.2 if encountering unknown definitions.

Corollary 2.1: There are many variations of passages of Lorem Ipsum available, but the majority have suffered alteration in some form, by injected humour, or randomised words which don't look even slightly believable. If you are going to use a passage of Lorem Ipsum, you need to be sure there isn't anything embarrassing hidden in the middle of text. All the Lorem Ipsum generators on the Internet tend to repeat predefined chunks as necessary, making this the first true generator on the Internet. It uses a dictionary of over 200 Latin words, combined with a handful of model sentence structures, to generate Lorem Ipsum which looks reasonable. The generated Lorem Ipsum is therefore always free from repetition, injected humour, or non-characteristic words etc.

Corollary 2.2: There are many variations of passages of Lorem Ipsum available, but the majority have suffered alteration in some form, by injected humour, or randomised words which don't look even slightly believable. If you are going to use a passage of Lorem Ipsum, you need to be sure there isn't anything embarrassing hidden in the middle of text. All the Lorem Ipsum generators on the

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Lemma 2.3: There are many variations of passages of Lorem Ipsum available, but the majority have suffered alteration in some form, by injected humour, or randomised words which don't look even slightly believable. If you are going to use a passage of Lorem Ipsum, you need to be sure there isn't anything embarrassing hidden in the middle of text. All the Lorem Ipsum generators on the Internet tend to repeat predefined chunks as necessary, making this the first true generator on the Internet. It uses a dictionary of over 200 Latin words, combined with a handful of model sentence structures, to generate Lorem Ipsum which looks reasonable. The generated Lorem Ipsum is therefore always free from repetition, injected humour, or non-characteristic words etc.

Corollary 2.4: There are many variations of passages of Lorem Ipsum available, but the majority have suffered alteration in some form, by injected humour, or randomised words which don't look even slightly believable. If you are going to use a passage of Lorem Ipsum, you need to be sure there isn't anything embarrassing hidden in the middle of text. All the Lorem Ipsum generators on the Internet tend to repeat predefined chunks as necessary, making this the first true generator on the Internet. It uses a dictionary of over 200 Latin words, combined with a handful of model sentence structures, to generate Lorem Ipsum which looks reasonable. The generated Lorem Ipsum is therefore always free from repetition, injected humour, or non-characteristic words etc.

3 - Construction

In this section, we will first present the construction, mentioned in Section 1.1, based on the gradually increasing independence. This construction will guarantee a maximum load of $O(\log n/\log\log n)$ using $O(\log n/\log\log n)$ with a smaller description length. This constructions allows us to prove the following theorem:

Theorem 3.1: For any constant c > 0, integers n and u = poly(n), there exists a family \mathscr{F} of $f : [u] \to [v]$ such that:

- 1. could be described using $O(\log n \log \log n)$ bits
- 2. f(x) can be computed in $O(\log n \log \log n)$ using unit cost RAM model for any $f \in \mathscr{F}$ and $x \in [u]$

3. $\exists \gamma > 0$ such that, for any $S \subseteq [u]$, |S| = n and an error parameter c,

$$Pr_{f \leftarrow \mathscr{F}} \left[\max_{i \in [n]} |f^{-1}(i) \cap S| \le \frac{\gamma \log n}{\log \log n} \right] > 1 - \frac{1}{n^c}$$

Under our construction, statement 3 of Theorem 3.1 could be interpreted in following way:

For any bin in the last layer,

the maximum load of each bin is $O(\frac{\log n}{\log \log n})$ with high probability.

In what follows we provide a more formal description of our construction (see Section 3.1), and then prove Theorem 3.1 (see Section 3.2).

3.1 - Formal Description of Construction

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Add modification from note (Eric 8.9).

There are two cases about n: whether n is a power of 2 or not. We only consider the case that $n=2^k$ since we could set the number of bins to be $m=2^{\lfloor \log_2 n \rfloor}$ and the influence on maximum load will only be at most a factor of two. Let $d=O(\log\log n)$, and for every $i\in [d]$, let \mathscr{H}_i be a family of k_i -wise δ -dependent functions $[u]\to 0, 1^{l_i}$, where:

1.
$$n_0 = n, n_i = \frac{n^{i-1}}{2^{l_i}} \ \forall i \in [d]$$

2.
$$l_i = \lfloor \frac{\log n_{i-1}}{4} \rfloor \ \forall i \in [d-1], \ l_d = \log n - \sum_{i=1}^{d-1} l_i$$

3.
$$k_i l_i = \Theta(\log n) \ \forall i \in [d-1], \ k_d = \Theta(\frac{\log n}{\log \log n})$$

4.
$$\delta = \text{poly}(\frac{1}{n})$$

We randomly select h_i from corresponding function family \mathcal{H}_i for every $i \in [d]$, and generate our overall hash function f as:

$$f(x) = h_1(x) \circ \dots \circ h_d(x),$$

which just simply concatenates all output of d k_i -wise δ -dependent functions.

We could visualize this construction as a reversed tree of d+1 layers. Given the sets of balls $S \subseteq [n]$, we will have:

layer 0: 1 bin with $n_0 = n$

layer 1: 2^{l_1} bins with $n_1 = \frac{n_0}{2^{l_1}}$

layer 2:
$$2^{l_1+l_2}$$
 bins with $n_2 = \frac{n_1}{2^{l_1}}$

.

layer i:
$$2^{\sum_{j=1}^{i} l_j}$$
 bins with $n_i = \frac{n_{i-1}}{2^{l_i}}$

Now, we start in layer 0 where we only have 1 bin containing all n balls. Then, we use function h_1 to hash all n elements into 2^{l_1} bins in the next layer. Now, if h_1 is an uniformly independent hash function, the expected number of balls in one bin in layer 1 is $\frac{n}{2^{l_1}=n_1}=\frac{n_0}{2^{l_1}}$. Then, for each bin in layer 1, we run h_2 on all balls inside this bin and hashed them to a unique group of 2^{l_2} bins from $2^{l_1+l_2}$ bins in layer 2. We repeat this step for all 2^{l_1} bins in layer 1. Since we expect each bin in layer 1 to have n_1 balls and we hash these balls to 2^{l_2} bins in layer 2, the expected number of balls in each bin in layer 2 is $\frac{n_1}{2^{l_1}}=n_2$ if function h_2 is uniformly independent. In general, level i will have $2^{\sum_{j=1}^{i} l_j}$ bins, and we will hash all balls in each bin in layer i-1 to 2^{l_i} bins in layer i. The expected load of each bin in layer i is n_i .

3.2 - Step-by-Step Proof of Theorem 3.1

To proof Theorem 3.1, we first need support from following Lemma.

Lemma 3.2: For any $i=0,...,d-2,\alpha=\Omega(\frac{1}{\log\log n}),0<\alpha_i<1,$ and set $S_i\subseteq [u]$ of size at most $(1+\alpha_i)n_i$,

$$Pr_{h_{i+1} \leftarrow \mathscr{H}_{i+1}} \left[\max_{y \in 0, 1^{l_{i+1}}} |h_{i+1}^{-1}(y) \cap S_i| \le (1+\alpha)(1+\alpha_i)n_{i+1} \right] > 1 - \frac{1}{n^{c+1}}$$

Proof for Lemma 3.2:

We first fix $y \in \{0,1\}^{l_{i+1}}$, let $X = |h_{i+1}^{-1}(y) \cap S|$. Assume without loss of generosity, $|S_i| \ge \lfloor (1+\alpha_i)n_i \rfloor$. If $|S_i| < \lfloor (1+\alpha_i)n_i \rfloor$, we could enlarge S_i by adding dummy elements. Then, we introduce indicator random variables X_i such that:

$$X_i = \begin{cases} 1 & \text{if element } j \in S_i \text{ is hashed into bin y by } h_{i+1} \\ 0 & \text{Otherwise} \end{cases}$$

Thus, we could rewrite X as the sum of X_i and, if h_{i+1} is uniformly independent, the expectation of X will become:

$$E[X] = \sum_{j=1}^{|S_i|} E[X_j] = \sum_{j=1}^{|S_i|} \frac{1}{2^{l_{i+1}}} = \frac{|S_i|}{2^{l_{i+1}}}$$

However, since our function h_{i+1} is k_{i+1} -wise δ -dependent, $X = \text{sum of } |S_i|$ k_{i+1} -wise δ -dependent random variables. Then, we could directly apply Lemma

2.2 with
$$k = \frac{k_{i+1}}{2}$$
 and $\mu = E[X] = \frac{|S_i|}{2^{l_{i+1}}}$,

$$Pr[X > (1+\alpha)\mu] \le 2\left(\frac{|S_i|k_{i+1}}{(\alpha\mu)^2}\right)^{\frac{k_{i+1}}{2}} + \delta\left(\frac{|S_i|}{\alpha\mu}\right)^{\frac{k_{i+1}}{2}}$$

$$= 2\left(\frac{|S_i|k_{i+1}}{\alpha^2 \frac{|S_i|^2}{2^{2l_{i+1}}}}\right)^{\frac{k_{i+1}}{2}} + \delta\left(\frac{|S_i|}{\alpha \frac{|S_i|}{2^{l_{i+1}}}}\right)^{\frac{k_{i+1}}{2}} \quad \text{since } \mu = \frac{|S_i|}{2^{l_{i+1}}}$$

$$= 2\left(\frac{2^{2l_{i+1}}k_{i+1}}{\alpha^2|S_i|}\right)^{\frac{k_{i+1}}{2}} + \delta\left(\frac{2^{2l_{i+1}}}{\alpha}\right)^{\frac{k_{i+1}}{2}}$$

Now, we will try to upper bound each item in this expression.

1. Notice that, in our construction,

(a)
$$l_{i+1} = \lfloor \frac{\log n_i}{4} \rfloor \le \frac{\log n_i}{4}$$

(b)
$$|S| \ge (1 + \alpha_i)n_i - 1 \ge n_i$$
, and

(c)
$$\alpha = \Omega(\frac{1}{\log \log n})$$

Then, for the first item, we could derive:

$$2\left(\frac{2^{2l_{i+1}}k_{i+1}}{\alpha^{2}|S_{i}|}\right)^{\frac{k_{i+1}}{2}} \le 2\left(\frac{2^{\frac{\log n_{i}}{2}}k_{i+1}}{\alpha^{2}|S_{i}|}\right)^{\frac{k_{i+1}}{2}} \quad \text{since (a)}$$

$$= 2\left(\frac{\sqrt{n_{i}}k_{i+1}}{\alpha^{2}|S_{i}|}\right)^{\frac{k_{i+1}}{2}} \quad \text{since } 2^{\frac{\log n_{i}}{2}} = \sqrt{n_{i}}$$

$$\le 2\left(\frac{\sqrt{n_{i}}k_{i+1}}{\alpha^{2}n_{i}}\right)^{\frac{k_{i+1}}{2}} \quad \text{since (b)} \to \frac{1}{S_{i}} \le \frac{1}{n_{i}}$$

$$= 2\left(\frac{k_{i+1}}{\alpha^{2}\sqrt{n_{i}}}\right)^{\frac{k_{i+1}}{2}}$$

$$= 2\left(\frac{k_{i+1}}{\alpha^{2}2^{2l_{i+1}}}\right)^{\frac{k_{i+1}}{2}}$$
Note that $2^{2l_{i+1}} \le 2^{\frac{\log n_{i}}{2}} = \sqrt{n_{i}} \to \frac{1}{\sqrt{n_{i}}} \le \frac{1}{2^{2l_{i+1}}}$, then
$$\le 2\left(\frac{k_{i+1}}{\alpha^{2}2^{2l_{i+1}}}\right)^{\frac{k_{i+1}}{2}}$$

$$= 2\left(\frac{k_{i+1}^{k_{i+1}/2}}{\alpha^{k_{i+1}}2^{k_{i+1}l_{i+1}}}\right)$$

by construction, we set $k_{i+1}l_{i+1} = \log n^c \in \Theta(\log n)$, then

$$= 2 \left(\frac{k_{i+1}^{k_{i+1}/2}}{\alpha^{k_{i+1}} 2^{\log n^c}} \right)$$
$$= 2 \left(\frac{k_{i+1}}{\alpha^2} \right)^{k_{i+1}/2} \times \frac{1}{n^c}$$

Set
$$k_{i+1} = \frac{k_{i+1}l_{i+1}}{l_{i+1}} = \frac{4c\log n}{\log n_{i+1}} \text{ and } \alpha = \frac{2n}{\sqrt{k_{i+1}}} = \Omega(\frac{1}{\log\log n})$$

$$= 2\left(\frac{k_{i+1}}{4n^2 \times \frac{1}{k_{i+1}}}\right)^{\frac{2c\log n}{\log n_{i+1}}} \times \frac{1}{n^c}$$

$$= 2\left(\frac{1}{4n^2}\right)^{\frac{2c\log n}{\log n_{i+1}}} \times \frac{1}{n^c}$$

$$\leq 2 \times \frac{1}{4n^2} \times \frac{1}{n^c} \quad \text{since } \frac{\log n}{\log n_i} \geq 1$$

$$= \frac{1}{2n^{c+2}}$$

2. Notice that $\delta = \text{poly}(\frac{1}{n})$. Then, for the second item, we could derive:

$$\delta\left(\frac{2^{2l_{i+1}}}{\alpha}\right)^{\frac{k_{i+1}}{2}} = \delta\left(\frac{2^{k_{i+1}l_{i+1}}}{\alpha_{i+1}^k}\right)$$

$$= \delta\left(\frac{2^{\log n^c}}{\alpha_{i+1}^k}\right) \quad \text{since } k_{i+1}l_{k+1} = \log n^c$$

$$\leq \delta n^c$$

$$\leq \delta n^c$$

$$\text{Set } \delta = \frac{1}{2n^{2c+2}}, \text{ then}$$

$$\leq \frac{1}{2n^{2c+2}} \times n^c$$

$$= \frac{1}{2n^{c+2}}$$

Thus, we will get:

$$Pr\left[X > (1+\alpha)(1+\alpha_i)n_{i+1}\right] = Pr\left[X > (1+\alpha)(1+\alpha_i)\frac{n_i}{2^{l_{i+1}}}\right] \quad \text{since } n_{i+1} = \frac{n_i}{2^{l_{i+1}}}$$

$$\leq Pr\left[X > (1+\alpha)\frac{|S_i|}{2^{l_{i+1}}}\right] \quad \text{since } |S_i| \leq (1+\alpha_i)n_i$$

$$= Pr\left[X > (1+\alpha)\mu\right] \quad \text{since } \mu = \frac{|S_i|}{2^{l_{i+1}}}$$

$$\leq \frac{1}{2n^{c+2}} + \frac{1}{2n^{c+2}} \quad \text{by upper bound derived above}$$

$$= \frac{1}{n^{c+2}}$$

We have at most $2^{l_{i+1}} \leq n$ different y since $2^{l_{i+1}} \leq 2^{\frac{\log n_{i-1}}{4}} \leq 2^{\log n} = n$. Then, we will apply an union bound on y. Then, we will get Lemma 2.2 by subtract the union bound result from 1.

4 - Extension and Appendix

This section presents some extensional use and appendix. There are many variations of passages of Lorem Ipsum available, but the majority have suffered alteration in some form, by injected humour, or randomised words which don't look even slightly believable. If you are going to use a passage of Lorem Ipsum, you need to be sure there isn't anything embarrassing hidden in the middle of text. There are many variations of passages of Lorem Ipsum available, but the majority have suffered alteration in some form, by injected humour, or randomised words which don't look even slightly believable. If you are going to use a passage of Lorem Ipsum, you need to be sure there isn't anything embarrassing hidden in the middle of text.

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