ST2131 AY21/22 SEM 2

github/jovyntls

01. COMBINATORIAL ANALYSIS

tricky - E18, E20-22, E23, E26

The Basic Principle of Counting

- combinatorial analysis → the mathematical theory of counting
- basic principle of counting → Suppose that two experiments are performed. If exp1 can result in any one of m possible outcomes and if, for each outcome of exp1, there are n possible outcomes of exp2, then together there are mn possible outcomes of the two experiments.
- generalized basic principle of counting \rightarrow If r experiments are performed such that the first one may result in any of n_1 possible outcomes and if for each of these n_1 possible outcomes, and if ..., then there is a total of $n_1 \cdot n_2 \cdot \cdots \cdot n_r$ possible outcomes of r experiments.

Permutations

factorials - 1! = 0! = 1

N1 - if we know how to count the number of different ways that an event can occur, we will know the probability of the event.

N2 - there are n! different arrangements for n objects.

N3 - there are $\frac{n!}{n_1! \, n_2! \dots n_r!}$ different arrangements of n objects, of which n_1 are alike, n_2 are alike, ..., n_r are alike.

Combinations

N4 - $\binom{n}{r} = \frac{n!}{(n-r)! \, r!}$ represents the number of different groups of size r that could be selected from a set of n objects when the order of selection is not considered

N4b -
$$\binom{n}{r} = \binom{n-1}{r-1} + \binom{n-1}{r}, \quad 1 \le r \le n$$

Proof: If object 1 is chosen $\Rightarrow \binom{n-1}{r-1}$ ways of choosing the remaining objects. If object 1 is not chosen $\Rightarrow \binom{n-1}{n}$ ways of choosing the remaining objects.

N5 - The Binomial Theorem -
$$(x+y)^n = \sum_{k=0}^n \binom{n}{k} x^k y^{n-k}$$

Proof: by mathematical induction: n=1 is true; expand; sub dummy variable; combine using N4b; combine back to final term

Multinomial Coefficients

 $\mathbf{N6} \cdot {n \choose n_1,n_2,\dots,n_r} = \frac{n!}{n_1!\,n_2!\dots n_r!} \text{ represents the number of possible divisions of } n_1!$ n distrinct objects into r distinct groups of respective sizes n_1, n_2, \ldots, n_3 , where $n_1 + n_2 + \cdots + n_r = n$

$$\begin{split} \textit{Proof:} & \text{ using basic counting principle,} \\ &= \binom{n}{n_1} \binom{n-n_1}{n_2} \binom{n-n_1-n_2}{n_3} \cdots \binom{n-n_1-n_2-n_{r-1}}{n_r} \\ &= \frac{n!}{(n-n_1)!} \sum_{\substack{n_1 \\ n_1 \mid n_2 \mid \dots \mid n_r \mid}} \times \frac{(n-n_1)!}{(n-n_1-n_2)!} \times \cdots \times \frac{(n-n_1-n_2-\cdots-n_{r-1})}{0! \, n_r!} \\ &= \frac{n!}{n_1! \, n_2! \, \dots \, n_r!} \end{split}$$

N7 - The Multinomial Theorem:
$$(x_1 + x_2 + \dots + x_r)^n = \sum_{\substack{(n_1,\dots,n_r): n_1+n_2+\dots+n_r=n}} \frac{n!}{n_1! \, n_2! \, \dots n_r!} x_1^{n_1} x_2^{n_2} \, \dots x_r^{n_r}$$

Number of Integer Solutions of Equations

N8 - there are $\binom{n-1}{r-1}$ distinct *positive* integer-valued vectors (x_1, x_2, \ldots, x_r) satisfying $x_1 + x_2 + \cdots + x_r = n$, $x_i > 0$, $i = 1, 2, \ldots, r$! cannot be directly applied to N8 as 0 value is not included

N9 - there are $\binom{n+r-1}{r-1}$ distinct *non-negative* integer-valued vectors

 (x_1, x_2, \dots, x_r) satisfying $x_1 + x_2 + \dots + x_r = n$

Proof: let $y_k = x_k + 1 \Rightarrow y_1 + y_2 + \cdots + y_r = n + r$

02. AXIOMS OF PROBABILITY

Sample Space and Events

- sample space → The set of all outcomes of an experiment (where outcomes are not predictable with certainty)
- event → Any subset of the sample space
- **union** of events E and $F \to E \cup F$ is the event that contains all outcomes that are either in E or F (or both).
- **intersection** of events E and $F \to E \cap F$ or EF is the event that contains all outcomes that are both in E and in F.
- **complement** of $E \to E^c$ is the event that contains all outcomes that are *not* in E.
- **subset** $\to E \subset F$ is all of the outcomes in E that are also in F.
 - $E \subset F \land F \subset E \Rightarrow E = F$

DeMorgan's Laws

$$(\bigcup_{i=1}^n \mathbf{E_i})^\mathbf{c} = \bigcap_{i=1}^n \mathbf{E_i^c}$$

 $\begin{array}{l} \textit{Proof} \colon \text{to show LHS} \subset \text{RHS} \colon \text{let } x \in (\bigcup_{i=1}^n E_i)^c \\ \quad \Rightarrow x \notin \bigcup_{i=1}^n E_i \Rightarrow x \notin E_1 \text{ and } x \notin E_2 \dots \text{ and } x \notin E_n \\ \quad \Rightarrow x \in E_1^c \text{ and } x \in E_2^c \dots \text{ and } x \in E_n^c \\ \quad \Rightarrow x \in \bigcap_{i=1}^n E_i^c \\ \quad \text{to show RHS} \subset \text{LHS} \colon \text{let } x \in \bigcap_{i=1}^n E_i^c \end{array}$

$$(\bigcap_{i=1}^n \mathbf{E_i})^\mathbf{c} = \bigcup_{i=1}^n \mathbf{E_i^c}$$

Proof: using the first law of DeMorgan, negate LHS to get RHS

Axioms of Probability

Defining Probability of an Event

definition 1: relative frequency

$$P(E) = \lim_{n \to \infty} \frac{n(E)}{n}$$

problems with this definition:

- 1. $\frac{n(E)}{n}$ may not converge when $n \to \infty$
- 2. $\frac{n(E)}{n}$ may not converge to the same value if the experiment is repeated

definition 2

Consider an experiment with sample space S. For each event E of the sample space S, we assume that a number P(E) is definned and satisfies the following 3 axioms:

- 1. 0 < P(E) < 1
- 2. P(S) = 1
- 3. For any sequence of mutually exclusive events E_1, E_2, \ldots (i.e., events for which $E_iE_j=\emptyset$ when $i\neq j$), $P(\bigcup_{i=1}^\infty E_i)=\sum_{i=1}^\infty P(E_i)$

P(E) is the probability of event E.

 $\begin{array}{lll} \textbf{commutative} & E \cup F = F \cup E & E \cap F = F \cap E \\ \textbf{associative} & (E \cup F) \cup G = E \cup (F \cup G) & (E \cap F) \cap G = E \cap (F \cap G) \\ \textbf{distributive} & (E \cup F) \cap G = (E \cap F) \cup (F \cap G) & (E \cap F) \cup G = (E \cup F) \cap (F \cup G) \\ \textbf{DeMorgan's} & (\bigcup_{i=1}^n E_i)^c = \bigcap_{i=1}^n E_i^c & (\bigcap_{i=1}^n E_i)^c = \bigcup_{i=1}^n E_i^c \\ \end{array}$