

CSE – 426 Homework 2

Griffin Kent

1) Verify that $\sigma(-z) = 1 - \sigma(z)$.

(Proof): By the definition of the sigmoid function, we know that

$$\sigma(z) = \frac{1}{1 + e^{-z}}$$

Thus, we have

$$\begin{aligned} 1 - \sigma(z) &= 1 - \frac{1}{1 + e^{-z}} = \frac{1 + e^{-z}}{1 + e^{-z}} - \frac{1}{1 + e^{-z}} \\ &= \frac{1 + e^{-z} - 1}{1 + e^{-z}} = \frac{e^{-z}}{1 + e^{-z}} = \frac{e^z e^{-z}}{e^z(1 + e^{-z})} \\ &= \frac{1}{e^z + 1} = \frac{1}{1 + e^z} = \sigma(-z). \end{aligned}$$

Therefore, completing the proof. ■

2) Suppose the probability of $y = 1$ given z is $\sigma(z) = \mathbb{P}(y = 1|z)$. Using the fact that $y \in \{0,1\}$, verify that $f(y) = \sigma(z)^y(1 - \sigma(z))^{1-y}$ is indeed the conditional probability of the random variable y given z .

(Proof): Since $y \in \{0,1\}$, it can only take on one of two values and thus is a Bernoulli distributed random variable. The standard form of the probability density function of a Bernoulli random variable is given by

$$f(Y = y) = \mathbb{P}(y)^y(1 - \mathbb{P}(y))^{1-y}$$

Where $\mathbb{P}(y)$ is defined as the probability of y taking on the value of 1. Since in this case $\mathbb{P}(y = 1|z) = \sigma(z)$, we obtain the following conditional probability density function for y given z by substituting in the defined probability:

$$f(y|z) = \sigma(z)^y(1 - \sigma(z))^{1-y}$$

Thus, completing the proof. ■

3) Find $\frac{d}{dz} \log(1 - \sigma(z))$ and $\frac{d}{dz} \log(\sigma(z))$.

(Solution): Solving for $\frac{d}{dz} \log(1 - \sigma(z))$ first, we have

$$\begin{aligned} \frac{d}{dz} \log(1 - \sigma(z)) &= \frac{d}{dz} \log(\sigma(-z)) = \frac{d}{dz} \log\left(\frac{1}{1 + e^z}\right) \\ &= (1 + e^z) \frac{d}{dz} \left(\frac{1}{1 + e^z}\right) \end{aligned}$$

$$\begin{aligned}
&= (1 + e^z) \left(\frac{(1 + e^z) \frac{d}{dz} (1) - (1) \frac{d}{dz} (1 + e^z)}{(1 + e^z)^2} \right) \\
&= \frac{0 - e^z \frac{d}{dz} (z)}{(1 + e^z)} = -\frac{e^z}{1 + e^z} = -\frac{e^{-z} e^z}{e^{-z} (1 + e^z)} \\
&= -\frac{1}{1 + e^{-z}} = -\sigma(z).
\end{aligned}$$

Now, solving for $\frac{d}{dz} \log(\sigma(z))$, we have the following

$$\begin{aligned}
\frac{d}{dz} \log(\sigma(z)) &= \frac{d}{dz} \log \left(\frac{1}{1 + e^{-z}} \right) = (1 + e^{-z}) \frac{d}{dz} \left(\frac{1}{1 + e^{-z}} \right) \\
&= (1 + e^{-z}) \left(\frac{(1 + e^{-z}) \frac{d}{dz} (1) - (1) \frac{d}{dz} (1 + e^{-z})}{(1 + e^{-z})^2} \right) \\
&= \frac{0 - e^{-z} \frac{d}{dz} (-z)}{(1 + e^{-z})} = \frac{e^{-z}}{1 + e^{-z}}
\end{aligned}$$

Then, from the properties we proved in question (1) above, we have

$$\frac{d}{dz} \log(\sigma(z)) = \frac{e^{-z}}{1 + e^{-z}} = 1 - \sigma(z).$$

■

4) Find the partial derivatives of the first entry of the output of the softmax function

$$\text{softmax}([z_1, z_2, \dots, z_k]) = \left[\frac{e^{z_1}}{\sum_j e^{z_j}}, \frac{e^{z_2}}{\sum_j e^{z_j}}, \dots, \frac{e^{z_k}}{\sum_j e^{z_j}} \right] = [\phi_1, \phi_2, \dots, \phi_k]$$

With respect to z_1 and z_2 .

(Solution): To start, we will take the partial derivative of the first entry of the output of the softmax function with respect to z_1 .

$$\begin{aligned}
\frac{\partial \phi_1}{\partial z_1} &= \frac{\partial}{\partial z_1} \left(\frac{e^{z_1}}{\sum_j e^{z_j}} \right) = \frac{(\sum_j e^{z_j}) \frac{\partial}{\partial z_1} (e^{z_1}) - (e^{z_1}) \frac{\partial}{\partial z_1} (\sum_j e^{z_j})}{(\sum_j e^{z_j})^2} \\
&= \frac{(\sum_j e^{z_j})(e^{z_1})}{(\sum_j e^{z_j})^2} - \frac{(e^{z_1}) \frac{\partial}{\partial z_1} (e^{z_1} + e^{z_2} + \dots + e^{z_k})}{(\sum_j e^{z_j})^2} \\
&= \frac{(e^{z_1})}{\sum_j e^{z_j}} - \frac{(e^{z_1})(e^{z_1})}{(\sum_j e^{z_j})^2} = \phi_1 - \phi_1^2.
\end{aligned}$$

Now, taking the partial derivative of the first entry of the output of the softmax function with respect to z_2 , we have

$$\begin{aligned}\frac{\partial \phi_1}{\partial z_2} &= \frac{\partial}{\partial z_2} \left(\frac{e^{z_1}}{\sum_j e^{z_j}} \right) = \frac{(\sum_j e^{z_j}) \frac{\partial}{\partial z_2} (e^{z_1}) - (e^{z_1}) \frac{\partial}{\partial z_2} (\sum_j e^{z_j})}{(\sum_j e^{z_j})^2} \\ &= \frac{0 - (e^{z_1})(e^{z_2})}{(\sum_j e^{z_j})^2} = 0 - \frac{(e^{z_1})(e^{z_2})}{(\sum_j e^{z_j})^2} = -\phi_1 \phi_2.\end{aligned}$$

■

5) (Graduate Only) Prove that the Hessian matrix of the negative log-likelihood function for logistic regression is positive semi-definite.

(Proof): The negative log-likelihood function (cross-entropy loss function) for logistic regression is defined as the following

$$-\ell(\boldsymbol{\theta}) = -\log L(\boldsymbol{\theta}) = -\sum_{i=1}^m \{y^{(i)} \log \sigma(z^{(i)}) + (1 - y^{(i)}) \log \sigma(-z^{(i)})\}$$

Where m is the number of training examples $\{(\mathbf{x}^{(i)}, y^{(i)})\}_{i=1}^m$. Likewise, the Hessian matrix of the negative log-likelihood function is defined as

$$\mathbf{H} = \sum_{i=1}^m \mathbf{x}^{(i)} (\mathbf{x}^{(i)})^T \sigma(\boldsymbol{\theta}^T \mathbf{x}^{(i)}) (1 - \sigma(\boldsymbol{\theta}^T \mathbf{x}^{(i)}))$$

$$\mathbf{H} = \mathbf{X} \boldsymbol{\Sigma} \mathbf{X}^T$$

Where $\boldsymbol{\Sigma} = \text{diag}(\sigma(\boldsymbol{\theta}^T \mathbf{x}^{(i)}) (1 - \sigma(\boldsymbol{\theta}^T \mathbf{x}^{(i)})))$ is a diagonal matrix with the i -th diagonal element being $\sigma(\boldsymbol{\theta}^T \mathbf{x}^{(i)}) (1 - \sigma(\boldsymbol{\theta}^T \mathbf{x}^{(i)}))$ and $\mathbf{X}^T = [\mathbf{x}^{(1)}, \mathbf{x}^{(2)}, \dots, \mathbf{x}^{(m)}]$, where each $\mathbf{x}^{(i)}$ is a column vector. (*Logistic Regression Lecture Notes*).

By definition, a matrix will be positive semi-definite if all its eigenvalues $\lambda_i \geq 0$. Since \mathbf{H} is a symmetric matrix, (Using the properties defined in Section 3.13 of the *Linear Algebra Review and Reference*) we know that \mathbf{X} is an orthogonal matrix and $\boldsymbol{\Sigma}$ is the diagonal matrix whose entries are the eigenvalues of \mathbf{H} . Thus, we have

$$\lambda_i = \sigma(\boldsymbol{\theta}^T \mathbf{x}^{(i)}) (1 - \sigma(\boldsymbol{\theta}^T \mathbf{x}^{(i)}))$$

Since the sigmoid function maps any real number to the range of (0,1), we know that

$$\lambda_i = \sigma(\boldsymbol{\theta}^T \mathbf{x}^{(i)}) (1 - \sigma(\boldsymbol{\theta}^T \mathbf{x}^{(i)})) \geq 0, \forall i = 1, 2, \dots, m.$$

Therefore, since we have shown for the hessian matrix \mathbf{H} that the eigenvalues $\lambda_i \geq 0$, we know that the matrix \mathbf{H} must be positive semi-definite.

■