

LOUISIANA STATE UNIVERITY

College of Agriculture School of Plant, Environmental, and Soil Sciences AGRO 7075 Prediction-based Breeding



Mating designs

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Importance of mating designs

- Estimate components of variance
- Understand the genetic control
- Identify:
- the best parents,
- population structure (heterotic groups),
- testers, and
- the best combinations (hybrids)
- Support decisions populations and breeding schemes

Principle and ideal scenario

Principle: generate various levels of relatedness among the relative progenies by effecting a series of crosses among individuals of a random mating population that caused different statistical components of variation from which genetic components of variations were estimated.

Ideal: Complete matings of all possible, desirable crosses, but generally not practical and not economically or biologically necessary

Balanced mating designs are preferred to estimate general combining ability (GCA) and specific combining ability (SCA)

- Therefore, breeders have moved from full diallel to top cross over the years due to practical issues.
- E.g., 49 lines, divided into two groups (34 and 15)

	L_1	L_2	L_3	L_4
L_1	L ₁	HS ₁₂	HS ₁₃	HS 14
L ₂	HS ₂₁	L2	HS 23	HS 24
L_3	HS 31	HS 32	L3	HS 3 4
L_4	HS ₄₁	HS 42	HS 43	L4

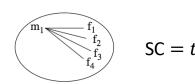
$$SC = n(n-1)/2$$

$$SC = \frac{49(48)}{2} = 1,176$$
 $SC = na \times n_b$

	P1	P2
P3	F1(1,3)	F1(2,3)
P4	F1(1,4)	F1(2,4)

$$SC = na x n_b$$

$$SC = 34 \times 15 = 510$$



$$SC = 1x34 + 1x15 = 49$$

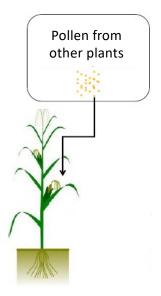
Mating Designs: A Survey

- Two schools have contributed the most: North Carolina State University, USA, and the University of Birmingham, England (UK)
- The former concentrated on the analyses of random mating populations. They are extensions of the one-way, two-way, and nested anova proposed by Fisher during early 20th century
- The latter worked on a series of mating designs that can be easily selfed or crossed
- Two important issues should be considered before choosing any mating design:
- <u>The kind of relatives that will be available for analysis:</u> certain types of relatives are observed more readily in some species than others, and some types of phenotypic covariance between relatives are more likely to approximate the desired quantities than others.
- <u>The degree of precision:</u> is a function of the number of individuals measured and how effort is allocated to the number of families versus the number of individuals within families an issue of optimum experimental design

Sib Analysis – progeny test

- An attractive alternative to parent-offspring regression in estimating genetic parameters
- There are three types of sib analysis: full-sib families, half-sib families, and; combinations of the half-sib and full-sib families
- A comparison between families and within families' variances provides an estimate of VA (additive genetic variance), in some cases, dominance (VD)
- The common environmental effects are the main drawbacks of any sib analysis.

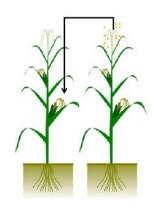
Progeny types



Half-sibs (HS)

Open pollination

$$\sigma_g^2 = \frac{1}{4}\sigma_A^2$$



Full-sibs (FS)

Manual pollination

$$\sigma_g^2 = \frac{1}{2}\sigma_A^2 + \frac{1}{4}\sigma_D^2$$



Self-sibs (Sn)

Natural or manual pollination

$$\sigma_g^2 = \frac{1}{2}\sigma_A^2$$

Sib Analysis – progeny test

Table 1. ANOVA table of polycross design with many replicated.

Source	df	MS	EMS	Variance components
Progenies	g-1	M_1	$\sigma_e^2 + r \sigma_{prog}^2$	$\sigma^2 prog = Cov(HS) = \frac{1+F}{4}\sigma^2 A$
Blocks	r-1	M_2	-	-
error	(g-1)(r-1)	M_3	σ_e^2	$\sigma^2_{\ e} = \sigma^2$

Source: Wricke and Weber, 1986

$$RS = i. \frac{\frac{1}{4}\sigma_a^2}{\sqrt{\sigma_G^2 + \sigma_D^2 + \sigma_E^2/r}}$$

North Carolina Designs

Comstock and Robinson (1952) proposed three mating designs, I, II, and III

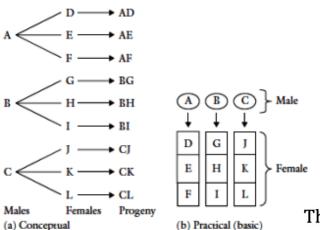


Table 2. Skeleton of general ANOVA for North Carolina design I.

Source of variation	df	MS	Expected mean squares
Males	(m-1)	M_{1}	$\sigma^2 w + r \sigma^2_{mf} + r f \sigma^2_{m}$
Females	m(f-1)	\pmb{M}_2	$\sigma^2 w + r \sigma^2_{mf}$
Within plots	mf(r-1)	$M_{\scriptscriptstyle 1}$	$\sigma^2 w$
Total	rmf-1	1	

Source: Acquaah, 2012.

The parameter $\sigma^2 w$ refers to the average variance within the full sib families and is given as;

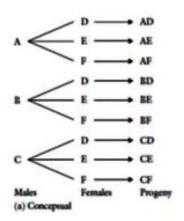
$$\sigma^{2}w = M_{1} = \frac{1}{2}V_{A} + \frac{3}{4}V_{D} + V_{E}$$

$$\sigma^{2}_{m} = (M_{1} - M_{2})/rf = \frac{1}{4}V_{A}$$

$$r\sigma^{2}_{mf} = (M_{2} - M_{3})/r = \frac{1}{4}V_{A} + \frac{1}{4}V_{D}$$

$$RS. \, males = i. \frac{\frac{1}{4}\sigma_a^2}{\sqrt{\sigma_G^2 + \sigma_D^2 + \sigma_E^2/r}}$$

North Carolina II



	P1	P2
P3	F1(1,3)	F1(2,3)
P4	F1(1,4)	F1(2,4)

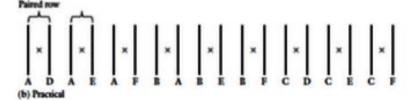


Table 5. Skeleton of general ANOVA of NC II.

Source of variation	df	MS	Expected mean squares
Replications	r-1		
Males	m-1	$M_{_1}$	$\sigma^2 w + r \sigma^2_{mf} + r f \sigma^2_{m}$
Females	f-1	M_2	$\sigma^2 w + r \sigma^2_{mf} + r m \sigma^2_{f}$
Males x females	(m-1)(f-1)	M_3	$\sigma^2 w + r \sigma^2_{mf}$
Within progenies	mf(r-1)	$M_{\scriptscriptstyle 4}$	$\sigma^2 w$
Error	(r-1)(mf-1)	M_{5}	σ^2
Total	rmf-1		

Source: Kearsey and Pooni, 1996

$$RS. males = i. \frac{\frac{1}{4}\sigma_a^2}{\sqrt{\sigma_G^2 + \sigma_D^2 + \sigma_E^2/r}}$$

$$RS. males = i. \frac{\frac{1}{4}\sigma_a^2}{\sqrt{\sigma_c^2 + \sigma_D^2 + \sigma_E^2/r}} \qquad RS. females = i. \frac{\frac{1}{4}\sigma_a^2}{\sqrt{\sigma_G^2 + \sigma_D^2 + \sigma_E^2/r}}$$

Where: $\sigma^2 w$ is the within progenies genetic and environmental variances. In the absence of epistasis and common environmental effects, σ^2_{mf} is a function of dominance variance V_D only (Kearsey and Pooni, 1996). If there is environmental variation between FS families, this could be due to general and specific maternal effects. The general maternal effects $(V_{\scriptscriptstyle EM})$ will appear in $\sigma^2_f = \frac{1}{4}V_A + V_{EM}$, while the specific maternal effects $V_{EC} - V_{EM}$ will appear in $\sigma^2_{mf} = \frac{1}{4}V_D + (V_{EC} - V_{EM})$ and will be confounded with V_D (Kearsey and Pooni, 1996). If the number of males and females is the same,

 $n_1 = n_2 = n$, we can have a test of maternal effects by

comparing M_1/M_2 as a variance ratio.

North Carolina III

- A random sample of F2 plants is backcrossed to the two inbred lines from which the F2 was descended triple testcross
- Capable of testing non-allelic (epistatic) interactions

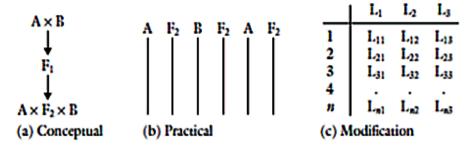
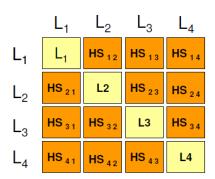


Table 7. Skeleton of NC III ANOVA.

Source of variation	df	MS	Expected mean squares
Testers, p	1	M_4	$\sigma^2 + r\sigma^2_{mp} + rmK^2_p$
Males (F ₂), m	m-1	M_3	$\sigma^2 + 2r\sigma^2_m$
Testers x parents	m-1	M_2	$\sigma^2 + r\sigma^2_{mp}$
Within FS families/error	(r-1)(2m-1)	$M_{\scriptscriptstyle 1}$	$oldsymbol{\sigma}^2$
Total	2mr-1		

Source: Hallauer et al., 2010

Griffing's diallel



- A set of genotypes are crossed in all possible combinations
- Lead to the testing and estimation of the general (GCA) and specific (SCA) combining abilities
- Four diallel mating methods depend on whether parents and reciprocals are retained or excluded from a particular design:
- Method I or full diallel: parents, one set of F1s, and reciprocal F1s. n² genotypes
- Method II: parents and one set of F1s without reciprocals F1s. p(p+1)/2 genotypes
- Method III: one set of F1s and the reciprocals are included. p(p-1) genotypes
- Method IV: only one set of F1s is included. It is the most common of the diallel crossing systems. p(p-1)/2 genotypes

Diallel

The table 8 presents the estimates of variances for variance components both fixed and random model.

Table 8. Skeleton of ANOVA for method I diallel design.

	16		140	Expected mean squares	
Source	df	SS	MS	Model I	Model II
GCA	p-1	S_g	M_{g}	$\sigma^2 + 2p(\frac{1}{p-1})\sum g^2{}_i$	$\sigma^2 + \frac{2(p-1)}{p}\sigma^2_g + 2p\sigma^2_g$
SCA	p(p-1)/2	Ss	M_s	$\sigma^2 + \frac{2}{p(p-1)} \sum \sum S_{ij}^2$	$\sigma^2 + \frac{2(p^2 - p + 1)}{p^2} \sigma_s^2$
Reciprocal eff.	p(p-1)/2	S_r	M_r	$\sigma^2 + 2(\frac{2}{p(p-1)}) \sum_{i < j} \sum r_{ij}^2$	$\sigma^2 + 2\sigma^2_r$
Error	m	S_e	M_{e}		

Source: Griffing (1956b)

Table 10. Skeleton of ANOVA of Diallel method III.

Cannac	df	SS	MS	Expected mean squares		
Source df	aı	33		Model I	Model II	
GCA	p-1	S_g	M_{g}	$\sigma^2 + 2p(p-2)(\frac{1}{p-1})\sum g^2$	$\sigma^2 + 2\sigma_s^2 + 2(p-2)\sigma_g^2$	
SCA	p(p-3)/2	Ss	M_s	$\sigma^2 + \frac{2}{p(p-3)} \sum_{i \in j} \sum_{i \in j} S_{ij}^2$	$\sigma^2 + 2\sigma_s^2$	
Reciprocal eff.	p(p-1)/2			$\sigma^2 + 2(\frac{2}{p(p-1)}) \sum_{i < j} \sum r_{ij}^2$	$\sigma^2 + 2\sigma_r^2$	
Error	m	S_e		σ^2	σ^2	

Source: Griffing (1956b)

Table 9. The analysis of variance for method II.

Course	df	CC	MS	Expecte	d mean squares
Source	uı	SS	MIS	Model I	Model II
GCA	p-1	S_g	M_{g}	$\sigma^2 + (2+p)(\frac{1}{p-1})\sum g_i^2$	$\sigma^2 + \sigma_s^2 + (p+2)\sigma_g^2$
SCA	p(p-1)/2	Ss	M_s	$\sigma^2 + \frac{2}{p(p-1)} \sum_i \sum_j s_{ij} ^2$	$\sigma^2 + \sigma_s^2$
Error	m	S_e	M_e '	σ^2	
					Source: Griffing (1956b)

Source	df	cc	MS	Expected mean squares		
Source	uı	SS	MIS	Model I	Model II	
GCA	p-1	S_g	M_{g}	$\sigma^2 + (p-2)(\frac{1}{p-1})\sum_i g_i^2$	$\sigma^2 + 2\sigma_s^2 + (p-2)\sigma_g^2$	
SCA	p(p-3)/2	Ss	M_s	$\sigma^2 + \frac{2}{p(p-3)} \sum_{i < j} \sum_{i < j} s_{ij}^2$	$\sigma^2 + \sigma_s^2$	
Error	m	S_{e}	M_e '	σ^2	$\sigma^{^2}$	

Source: Griffing (1956b)

Line x tester - topcross

- Involves hybridization between lines (f) and wide based testers in one-to-one fashion generating $f \times m = fm$ hybrids
- It provides both full-sibs and half-sibs simultaneously as opposed to topcross which provides only half-sibs
- Estimates SCA of each cross, but not provides GCA of lines only for testers

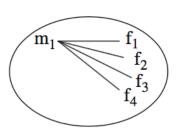


Table 3. Skeleton of ANOVA for Line X Tester Design

Source df		MS	Expected m	ean squares
Source	aı	MIS	Model I	Model II
Replication	r-1			
Lines	m-1	$M_{_1}$	$\sigma^2 + rf \frac{1}{m-1} + \sum_i g_i^2$	$\sigma^2 + v_{sca} + r f_{gca(m)}$
Testers	f-1	M_2	$\sigma^2 + rm(\frac{1}{f-1})\sum_i g_j^2$	$\sigma^2 + rv_{sca} + rm_{gca(f)}$
Line x tester	(m-1)(f-1)	M_3	$\sigma^2 + r \left[\frac{1}{(m-1)(f-1)} \right] \sum_{i} \sum_{j} s_{ij}$	$\sigma^2 + rv_{sca}$
Error	(r-1)(mf-1)	$M_{\scriptscriptstyle 4}$	σ^2	σ^2

Source: Sharma (2006)

Generation Mean Analysis - Mather (1949)

- It is an extensive analysis of family means, universally known as the "basic generation analysis"
- Six generations (P1, P2, F1), F2, F2xP1, and F2xP2
- Test and estimate the additive, dominance, and epistatic effects of genes at the means level
- Provided a rationale to estimate VA, VD, and VE, thus heritability
- Vp = F2
- Ve = V(P1 + P2 + F1) / 3
- Vg = Vp Ve

Mean analysis = m + a

• Two consecutive generations. E.g., F_1 and F_2

•
$$\mathbf{F_1} = m + a + d$$

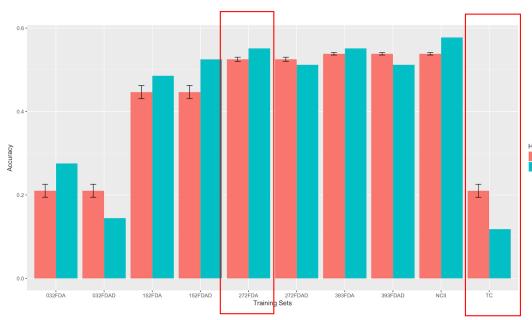
•
$$\mathbf{F}_2 = m + a + 1/2d$$

•
$$\mathbf{F_3} = m + a + 1/4d$$
 $m + a = 2(m + a + \frac{d}{2}) - (m + a + d)$

 $m + a = 2F_2 - F_1$

- ...
- F_{∞} homozygosity the lines performance will be m + a
- So, progenies with higher m + a are "superior"

Cross	F1 (Kg ha)	F2 (Kg ha)	2F2 – F1	m + a
L1 x L2	2,0	2,2	2 * 2,2 - 2,0	2,4
L3 x L4	2,5	2,5	2 * 2,5 - 2,5	2,5
L5 x L6	2,2	3,0	2 * 3,0 - 2,2	3,8



Theoretical and Applied Genetics (2018) 131:1153–1162 https://doi.org/10.1007/s00122-018-3068-8

ORIGINAL ARTICLE

Accuracy of genomic selection to predict maize single-crosses obtained through different mating designs

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