

CHANGE DETECTION ALGORITHMS

ADVANCED

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LECTURE SUMMARY

1. Advanced probabilistic tests

2. Conclusions

ADVANCED PROBABILISTIC TESTS

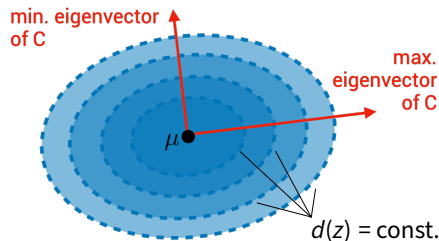
ADVANCED PROBABILISTIC TESTS


MULTIVARIATE CASE: MAHALANOBIS MEETS CHEBYCHEV

- ▶ $\mathcal{H}_0 =: z \in \mathbb{R}^n$ is a **random vector** with **known mean** μ and **covariance** C
- ▶ The **Mahalanobis distance (MD)** $d(z)$ tells us **how far** a sample z is from its nominal distribution

Definition (Mahalanobis distance)

$$d(z) \triangleq \sqrt{(z - \mu)^\top C^{-1} (z - \mu)}$$



 Xinjia Chen. "A new generalization of Chebyshev inequality for random vectors". In: *arXiv:0707.0805* (2007)

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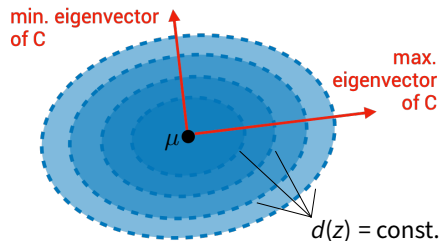
MULTIVARIATE CASE: MAHALANOBIS MEETS CHEBYCHEV

- ▶ The Chebychev inequality allows to test \mathcal{H}_0 with a desired FPR α , whatever is the distribution that generated z

Definition (Chebychev inequality)

$$\mathbb{P}[d(z) \geq \bar{d}] \leq \alpha,$$

$$\text{with } \bar{d} = \sqrt{\frac{n}{\alpha}} \rightarrow \begin{array}{l} n. \text{ of dimensions} \\ \alpha \rightarrow \text{desired FPR} \end{array}$$



📖 Xinjia Chen. “A new generalization of Chebyshev inequality for random vectors”. In: *arXiv:0707.0805* (2007)

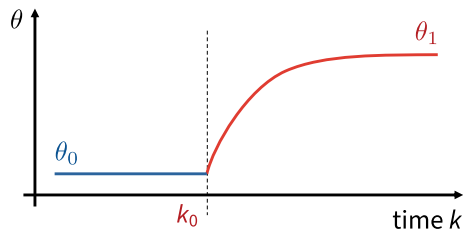
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STATISTICAL HYPOTHESIS TESTING À-LA NEYMAN-PEARSON

- **Two** hypotheses are considered

null $\mathcal{H}_0 : z \sim p_{\theta_0}$

alternate $\mathcal{H}_1 : z \sim p_{\theta_1}$



Problem

Design a decision function

$$d(z(k)) : \mathbb{R} \mapsto \{\mathcal{H}_0, \mathcal{H}_1\}$$

with a given FPR α and the highest possible TPR

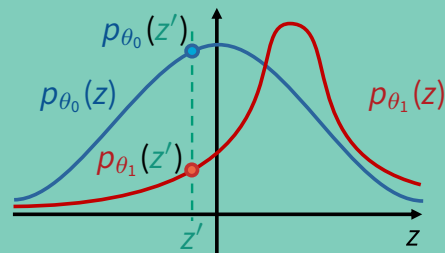
$\Rightarrow d$ is said to be **Most Powerful** (MP)

■ Michele Basseville, Igor V Nikiforov, et al. *Detection of abrupt changes: theory and application*. Vol. 104. prentice Hall Englewood Cliffs, 1993, Ch. 4

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STATISTICAL HYPOTHESIS TESTING À-LA NEYMAN-PEARSON

Intuition: likelihood matters



$$\mathbb{E}[p_{\theta_i} | \mathcal{H}_i] \geq \mathbb{E}[p_{\theta_i} | \mathcal{H}_j]$$


Neyman-Pearson Lemma (simplified)

- Let the decision function g be defined as

$$d(z) = \frac{p_{\theta_1}(z)}{p_{\theta_0}(z)} \underset{\mathcal{H}_0}{\overset{\mathcal{H}_1}{\geq}} \lambda$$

where λ is chosen such that $\text{FPR} = \alpha$.

- Then d is the MP detector with $\text{FPR} \leq \alpha$.

 Michele Basseville, Igor V Nikiforov, et al. *Detection of abrupt changes: theory and application*. Vol. 104. prentice Hall Englewood Cliffs, 1993, Ch. 4

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THE LOG-LIKELIHOOD RATIO


Definition (LLR)

$$s(z) \triangleq \ln \frac{p_{\theta_1}(z)}{p_{\theta_0}(z)}$$

- The expectation of $s(z)$ under **null** and **alternate** hypothesis has opposite sign!

$$\mathbb{E}[s|\mathcal{H}_0] = \int_{-\infty}^{\infty} s(z)p_{\theta_0}(z)dz < 0$$

$$\mathbb{E}[s|\mathcal{H}_1] = \int_{-\infty}^{\infty} s(z)p_{\theta_1}(z)dz > 0$$

 Mogens Blanke et al. *Diagnosis and fault-tolerant control*.
Vol. 2. Springer, 2006, Ch. 6

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THE CUSUM TEST

Intuition

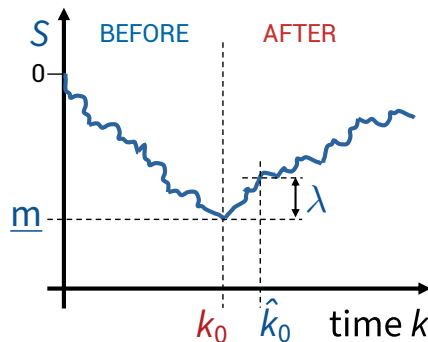
Approximate online the expectation $\mathbb{E}[s]$ with a **cumulative sum**, then look for **drift** changes


Definition (CUSUM test)

$$S(k) \triangleq \sum_{i=1}^k s(z(i)) = \sum_{i=1}^k \ln \frac{p_{\theta_1}(z(i))}{p_{\theta_0}(z(i))}$$

$$g(k) \triangleq S(k) - m(k) \quad m(k) \triangleq \min_{i=1:k} S(i)$$

$$d(k) \triangleq g(k) \underset{\mathcal{H}_0}{\overset{\mathcal{H}_1}{\geq}} \lambda$$



 Mogens Blanke et al. *Diagnosis and fault-tolerant control*. Vol. 2. Springer, 2006, Ch. 6

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THE CUSUM TEST – EXAMPLES

Example 6.12 *Change in the mean of a Gaussian sequence*

Remember that the Gaussian probability density function for a random variable with mean μ and variance σ is

$$p_{\mu}(z) = \frac{1}{\sqrt{2\pi}\sigma} \exp\left(-\frac{(z - \mu)^2}{2\sigma^2}\right). \quad (6.121)$$

The resulting likelihood ratio for detecting a change in the mean from μ_0 to μ_1 is

$$\frac{p_{\mu_1}(z)}{p_{\mu_0}(z)} = \exp\left(-\frac{(z - \mu_1)^2}{2\sigma^2} + \frac{(z - \mu_0)^2}{2\sigma^2}\right).$$

Hence straightforward computations yield the following expression for the log-likelihood ratio $s(z)$:

$$s(z) = \frac{2(\mu_1 - \mu_0)z + (\mu_0^2 - \mu_1^2)}{2\sigma^2} = \frac{\mu_1 - \mu_0}{\sigma^2} \left(z - \frac{\mu_0 + \mu_1}{2}\right). \quad (6.122)$$

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THE CUSUM TEST – EXAMPLES

Example 6.13 *Change in the mean and variance*

If both mean and variance change after a fault, the following relation

$$\frac{p_{\mu_1}(z)}{p_{\mu_0}(z)} = \frac{\sigma_0}{\sigma_1} \exp \left(-\frac{(z - \mu_1)^2}{2\sigma_1^2} + \frac{(z - \mu_0)^2}{2\sigma_0^2} \right)$$

holds and the log-likelihood ratio is

$$s(z) = \ln \frac{\sigma_0}{\sigma_1} + \frac{(z - \mu_0)^2}{2\sigma_0^2} - \frac{(z - \mu_1)^2}{2\sigma_1^2}.$$

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THE CUSUM TEST – PARAMETER TUNING

- ▶ Parameter θ_0 (i.e. $\theta_0 = (\mu_0, \sigma_0^2)$) is assumed to be known

Problem

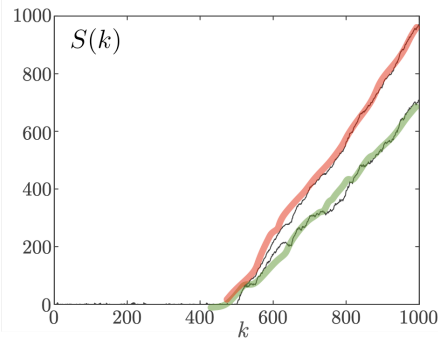
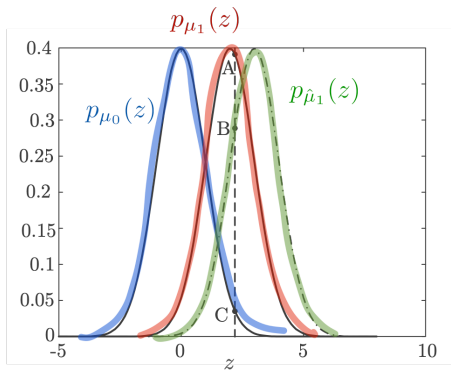
- ⊙ What values should we choose for θ_1 ?
- ⊙ What value should we choose for λ ?

➡ some options next

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THE CUSUM TEST – TUNING μ_1

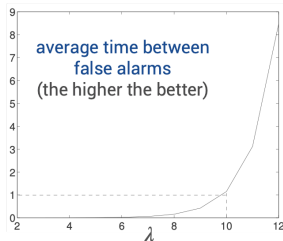
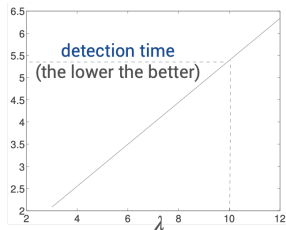
- 💡 **replace** by **minimum change** you want to detect (both $\pm \Rightarrow$ two sided CUSUM)
- 💡 **estimate** it from **data** (pro: ideal CUSUM conditions; con: detection delay!)



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THE CUSUM TEST – TUNING λ

- ▶ **exact methods** exist for some distributions (i.e. Gaussian)
 - ➔ book by Blanke, page 244-246
- ▶ otherwise, tuning of λ is **empirical**
 - 💡 $\lambda = \text{desired} - \text{detection} - \text{time} \times \text{slope} - \text{of} - S - \text{after} - a - \text{change}$
 - 💡 run a series of Monte Carlo simulations and compute detection time, FPR and TPR as a function of λ



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THE GENERALIZED LIKELIHOOD RATIO (GLR) TEST

Intuition

Estimate both the change time k_0 and the new parameter θ_1

Definition (GLR test)

$$S_j^k(\theta_1) \triangleq \sum_{i=j}^k s(z(i)) = \sum_{i=1}^k \ln \frac{p_{\theta_1}(z(i))}{p_{\theta_0}(z(i))}$$

$$(\hat{k}_0, \hat{\theta}_1) = \arg \left\{ \max_{1 \leq k_0 \leq k} \max_{\theta_1} S_j^k(\theta_1) \right\}$$

$$g(k) = \max_{1 \leq k_0 \leq k} \max_{\theta_1} S_j^k(\theta_1)$$

- ▶ Decision $d(k)$ as usual
- ▶ The solution is a **maximum likelihood** estimate
- ⚠ Solution involves a **double maximization**

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THE GLR TEST – EXAMPLE

Example 6.12 (cont.) *Change in the mean of a Gaussian sequence*

In this particular case, it is possible to find an explicit expression for $\hat{\mu}_1(k, j)$, the maximum likelihood estimate of μ_1 at the present time instant k , assuming that the fault occurred at time instant j . Indeed, from (6.122), $S_j^k(\mu_1)$ takes the following form:

$$S_j^k(\mu_1) = \frac{\mu_1 - \mu_0}{\sigma^2} \sum_{i=j}^k \left(z(i) - \frac{\mu_0 + \mu_1}{2} \right) \quad (6.140)$$

In order to maximise this expression with respect to μ_1 , one has to take the derivative of $S_j^k(\mu_1)$ with respect to μ_1 and equate that expression to zero:

$$\frac{\partial S_j^k(\mu_1)}{\partial \mu_1} = \frac{1}{\sigma^2} \sum_{i=j}^k \left(z(i) - \frac{\mu_0 + \mu_1}{2} \right) - \frac{k - j + 1}{2} \frac{(\mu_1 - \mu_0)}{\sigma^2} = 0. \quad (6.141)$$

Equation (6.141) yields:

$$\hat{\mu}_1(k, j) = \frac{1}{k - j + 1} \sum_{i=j}^k z(i). \quad (6.142)$$

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THE GLR TEST – EXAMPLE

Substituting this expression for μ_1 in (6.140) results, after straightforward computations, in:

$$S_j^k(\hat{\mu}_1(k, j)) = \frac{1}{2\sigma^2} \frac{1}{k-j+1} \left[\sum_{i=j}^k (z(i) - \mu_0) \right]^2. \quad (6.143)$$

Hence the GLR decision function can be written:

$$g(k) = \frac{1}{2\sigma^2} \max_{k-M+1 \leq j \leq k} \frac{1}{k-j+1} \left[\sum_{i=j}^k (z(i) - \mu_0) \right]^2. \quad (6.144)$$

If \mathcal{H}_1 is accepted in the above GLR algorithm, at the alarm time k_a , the estimated change occurrence time is given as:

$$\hat{k}_0 = \arg\left\{ \frac{1}{2\sigma^2} \max_{k_a-M+1 \leq j \leq k_a} \frac{1}{k_a-j+1} \left[\sum_{i=j}^{k_a} (z(i) - \mu_0) \right]^2 \right\}. \quad \square \quad (6.145)$$

CONCLUSIONS

CONCLUSIONS

IN THIS LECTURE WE COVERED

- ▶ Mahalanobis distance
- ▶ Neyman-Pearson approach to statistical hypothesis test
- ▶ CUSUM
- ▶ GLR

Next lecture: **Signal-based diagnosis**

CONCLUSIONS

THANK YOU FOR YOUR ATTENTION!

For further information:
Course page on **Brightspace**
or
our **MS Team**