

# STATS100B – Introduction to Mathematical Statistics

## Homework 7

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### Question a

Solution:

$$\begin{aligned} L &= \frac{1}{(2\pi)^{\frac{n}{2}}} |\Sigma|^{-\frac{1}{2}} e^{\frac{1}{2}(\mathbf{Y}-\mu\mathbf{1})'\Sigma^{-1}(\mathbf{Y}-\mu\mathbf{1})} \\ \ln L &= -\frac{n}{2}\ln(2\pi\sigma^2) - \frac{1}{2}\ln|\mathbf{V}| - \frac{1}{2}\sigma^2(\mathbf{Y}-\mu)\mathbf{V}^{-1}(\mathbf{Y}-\mu) \\ \frac{\partial \ln L}{\partial \mu} &= -\frac{1}{2\sigma^2} [-\mathbf{Y}'\mathbf{V}^{-1}\mathbf{1} - \mathbf{1}'\mathbf{V}^{-1}\mathbf{Y} + 2\mu\mathbf{1}'\mathbf{V}^{-1}\mathbf{1}] = 0 \\ \therefore \hat{\mu} &= \frac{\mathbf{1}'\mathbf{V}^{-1}\mathbf{Y}}{\mathbf{1}'\mathbf{V}^{-1}\mathbf{1}} \\ \frac{\partial \ln L}{\partial \sigma^2} &= -\frac{n}{2\sigma^2} + \frac{1}{\sigma^4} (\mathbf{Y}-\mu\mathbf{1})'\mathbf{V}^{-1}(\mathbf{Y}-\mu\mathbf{1}) = 0 \\ \therefore \hat{\sigma}^2 &= \frac{(\mathbf{Y}-\hat{\mu}\mathbf{1})'\mathbf{V}^{-1}(\mathbf{Y}-\hat{\mu}\mathbf{1})}{n} \end{aligned}$$

### Question b

Solution:

$$E(\hat{\mu}) = E\left(\frac{\mathbf{1}'\mathbf{V}^{-1}\mathbf{Y}}{\mathbf{1}'\mathbf{V}^{-1}\mathbf{1}}\right) = \frac{\mathbf{1}'\mathbf{V}^{-1}E(\mathbf{Y})}{\mathbf{1}'\mathbf{V}^{-1}\mathbf{1}} = \frac{\mathbf{1}'\mathbf{V}^{-1}\mu\mathbf{1}}{\mathbf{1}'\mathbf{V}^{-1}\mathbf{1}} = \mu$$

### Question c

Solution:

$$\begin{aligned} \mathbf{I}(\boldsymbol{\theta}) &= -E\begin{pmatrix} \frac{\partial^2 \ln L}{\partial \mu^2} & \frac{\partial^2 \ln L}{\partial \mu \partial \sigma^2} \\ \frac{\partial^2 \ln L}{\partial \sigma^2 \partial \mu} & \frac{\partial^2 \ln L}{\partial \sigma^2 \partial \sigma^2} \end{pmatrix} \\ \frac{\partial^2 \ln L}{\partial \mu^2} &= -\frac{\mathbf{1}'\mathbf{V}^{-1}\mathbf{1}}{\sigma^2} \\ \frac{\partial^2 \ln L}{\partial \mu \partial \sigma^2} &= \frac{1}{\sigma^4} [\mu\mathbf{1}'\mathbf{V}^{-1}\mathbf{1} - \mathbf{1}'\mathbf{V}^{-1}\mathbf{Y}] \end{aligned}$$

$$\begin{aligned}\frac{\partial^2 \ln L}{\partial \sigma^2 \partial \mu} &= \frac{1}{\sigma^4} [\mu 1' V^{-1} 1 - 1' V^{-1} Y] \\ \frac{\partial \ln L}{\partial \sigma^2} &= \frac{n}{2\sigma^4} - \frac{1}{\sigma^6} (Y - \mu 1)' V^{-1} (Y - \mu 1) \\ \therefore I(\theta) &= \left( -\frac{1' V^{-1} 1}{\sigma^2} \right)\end{aligned}$$

### Question d

Solution:

$$\begin{aligned}\sum_{i=1}^4 (X_i - \bar{X})^2 &= \sum_{i=1}^4 X_i^2 - 4\bar{X}^2 \\ &= \sum_{i=1}^4 X_i^2 - 4 \left[ \frac{1}{4} (X_1 + X_2 + X_3 + X_4) \right]^2 \\ &= \sum_{i=1}^4 X_i^2 - \frac{1}{2} \left( \sum_{i=1}^4 X_i^2 + 2 \sum_{1 \leq i < j \leq 4} X_i X_j \right) \\ &= \frac{3}{4} \sum_{i=1}^4 X_i^2 - \frac{1}{2} \sum_{1 \leq i < j \leq 4} X_i X_j\end{aligned}$$

$$\begin{aligned}RHS &= \left( \frac{1}{2} X_1^2 - X_1 X_2 + \frac{1}{2} X_2^2 \right) + \left( \frac{2}{3} X_3^2 + \frac{1}{6} (X_1 + X_2)^2 - \frac{2}{3} X_3 (X_1 + X_2) \right) + \frac{3}{4} \left( X_4 - \frac{1}{3} (X_1 + X_2 + X_3) \right)^2 \\ &= \dots\dots\dots (simple but tedious simplifications)\end{aligned}$$

$$= \frac{3}{4} \sum_{i=1}^4 X_i^2 - \frac{1}{2} \sum_{1 \leq i < j \leq 4} X_i X_j$$

$$\therefore \sum_{i=1}^4 (X_i - \bar{X})^2 = \frac{(X_1 - X_2)^2}{2} + \frac{\left[ X_3 - \frac{(X_1 + X_2)}{2} \right]^2}{\frac{3}{2}} + \frac{\left[ X_4 - \frac{(X_1 + X_2 + X_3)}{3} \right]^2}{\frac{4}{3}}$$

$$\therefore X_1 - X_2 \sim N(0, \sqrt{2})$$

$$\therefore \frac{(X_1 - X_2)^2}{2} \sim \chi_1^2$$

$$\therefore X_3 - \frac{X_1 + X_2}{2} \sim N\left(0, \sqrt{\frac{3}{2}}\right)$$

$$\therefore \frac{(X_3 - \frac{X_1 + X_2}{2})^2}{\frac{3}{2}} \sim \chi_1^2$$

$$\therefore X_4 - \frac{X_1 + X_2 + X_3}{3} \sim N\left(0, \sqrt{\frac{4}{3}}\right)$$

$$\therefore \frac{(X_4 - \frac{X_1 + X_2 + X_3}{3})^2}{\frac{4}{3}} \sim \chi_1^2$$

$$\mathbf{X} = \begin{pmatrix} X_1 & X_2 & X_3 & X_4 \end{pmatrix}'$$

$$\begin{pmatrix} X_1 - X_2 \\ X_3 - \frac{X_1 + X_2}{2} \\ X_4 - \frac{X_1 + X_2 + X_3}{3} \end{pmatrix} = \begin{pmatrix} 1 & -1 & 0 & 0 \\ -\frac{1}{2} & -\frac{1}{2} & 1 & 0 \\ -\frac{1}{3} & -\frac{1}{3} & -\frac{1}{3} & 1 \end{pmatrix} \begin{pmatrix} X_1 \\ X_2 \\ X_3 \\ X_4 \end{pmatrix} = \mathbf{A}\mathbf{X}$$

$$\text{var}(\mathbf{A}\mathbf{X}) = \mathbf{A}\text{var}(\mathbf{X})\mathbf{A}' = \begin{pmatrix} 1 & -1 & 0 & 0 \\ -\frac{1}{2} & -\frac{1}{2} & 1 & 0 \\ -\frac{1}{3} & -\frac{1}{3} & -\frac{1}{3} & 1 \end{pmatrix} \begin{pmatrix} 1 & 0 & 0 & 0 \\ 0 & 1 & 0 & 0 \\ 0 & 0 & 1 & 0 \\ 0 & 0 & 0 & 1 \end{pmatrix} \begin{pmatrix} 1 & -\frac{1}{2} & -\frac{1}{3} \\ -1 & -\frac{1}{2} & -\frac{1}{3} \\ 0 & 1 & -\frac{1}{3} \\ 0 & 0 & 1 \end{pmatrix} = \begin{pmatrix} 2 & 0 & 0 \\ 0 & \frac{3}{2} & 0 \\ 0 & 0 & \frac{4}{3} \end{pmatrix}$$

So the three terms in the RHS are independent each with a  $\chi_1^2$  distribution.