

STATS100B – Introduction to Mathematical Statistics

Homework 7

Feng Shiwei UID:305256428

Question a

Solution:

$$E(\bar{X}) = \mu = \lambda, E(S^2) = \sigma^2 = \lambda$$

$$f(x) = \frac{\lambda^x e^{-\lambda}}{x!}$$

$$\ln f(x) = x \ln \lambda - \lambda - \ln x!$$

$$\frac{\partial \ln f(x)}{\partial \lambda} = \frac{x}{\lambda} - 1$$

$$\frac{\partial^2 \ln f(x)}{\partial \lambda^2} = -\frac{x}{\lambda^2}$$

Let's find the Cramer-Rao Lower Bound.

$$\text{var}(\hat{\lambda}) \geq -\frac{1}{nE\left(-\frac{x}{\lambda^2}\right)} = \frac{\lambda^2}{n\lambda} = \frac{\lambda}{n}$$

$$\therefore \text{var}(\bar{X}) = \frac{\sigma^2}{n} = \frac{\lambda}{n}$$

$$\therefore \text{var}(\sigma^2) \geq \text{var}(\bar{X})$$

So \bar{X} is a better estimator.

Question b

Solution:

$$E(\alpha \bar{X} + (1 - \alpha) cS) = \alpha E(\bar{X}) + (1 - \alpha) E(cS) = \alpha \theta + (1 - \alpha) \theta = \theta$$

$$\text{var}(\alpha \bar{X} + (1 - \alpha) cS) = \alpha \cdot \text{var}(\bar{X}) + (1 - \alpha)^2 \cdot \text{var}(cS)$$

$$\frac{d \text{var}(\alpha \bar{X} + (1 - \alpha) cS)}{d\alpha} = 2\alpha \cdot \text{var}(\bar{X}) + 2(\alpha - 1) \cdot \text{var}(cS) = 0$$

$$\alpha = \frac{\text{var}(cS)}{\text{var}(cS) + \text{var}(\bar{X})}$$

$$\therefore \text{var}(\bar{X}) = \frac{\theta^2}{n}, \text{var}(cS) = E(c^2 S^2) - (E(cS))^2 = c^2 \theta^2 - \theta^2 = (c^2 - 1) \theta^2$$

$$\therefore \alpha = \frac{n(c^2 - 1)}{n(c^2 - 1) + 1}$$

Question c

Solution:

$$\begin{aligned}
 X &\sim (\alpha, \beta), \bar{X} \sim (n\alpha, \frac{\beta}{n}) \\
 E\left(\frac{1}{\bar{X}}\right) &= \frac{\Gamma(n\alpha - 1) \left(\frac{\beta}{n}\right)^{-1}}{\Gamma(n\alpha)} = \frac{n}{(n\alpha - 1)\beta} \\
 \therefore \hat{\theta} &= \frac{n\alpha - 1}{n\bar{X}}, E(\hat{\theta}) = \frac{n\alpha - 1}{n} E\left(\frac{1}{\bar{X}}\right) = \frac{1}{\beta}
 \end{aligned}$$

Question d

Solution:

$$\begin{aligned}
 \sum_{i=1}^4 (X_i - \bar{X})^2 &= \sum_{i=1}^4 X_i^2 - 4\bar{X}^2 \\
 &= \sum_{i=1}^4 X_i^2 - 4 \left[\frac{1}{4} (X_1 + X_2 + X_3 + X_4) \right]^2 \\
 &= \sum_{i=1}^4 X_i^2 - \frac{1}{2} \left(\sum_{i=1}^4 X_i^2 + 2 \sum_{1 \leq i < j \leq 4} X_i X_j \right) \\
 &= \frac{3}{4} \sum_{i=1}^4 X_i^2 - \frac{1}{2} \sum_{1 \leq i < j \leq 4} X_i X_j
 \end{aligned}$$

$$\begin{aligned}
 RHS &= \left(\frac{1}{2} X_1^2 - X_1 X_2 + \frac{1}{2} X_2^2 \right) + \left(\frac{2}{3} X_3^2 + \frac{1}{6} (X_1 + X_2)^2 - \frac{2}{3} X_3 (X_1 + X_2) \right) + \frac{3}{4} \left(X_4 - \frac{1}{3} (X_1 + X_2 + X_3) \right)^2 \\
 &= \dots\dots\dots (simple but tedious simplifications) \\
 &= \frac{3}{4} \sum_{i=1}^4 X_i^2 - \frac{1}{2} \sum_{1 \leq i < j \leq 4} X_i X_j
 \end{aligned}$$

$$\therefore \sum_{i=1}^4 (X_i - \bar{X})^2 = \frac{(X_1 - X_2)^2}{2} + \frac{\left[X_3 - \frac{(X_1 + X_2)}{2} \right]^2}{\frac{3}{2}} + \frac{\left[X_4 - \frac{(X_1 + X_2 + X_3)}{3} \right]^2}{\frac{4}{3}}$$

$$\therefore X_1 - X_2 \sim N(0, \sqrt{2})$$

$$\therefore \frac{(X_1 - X_2)^2}{2} \sim \chi_1^2$$

$$\therefore X_3 - \frac{X_1 + X_2}{2} \sim N\left(0, \sqrt{\frac{3}{2}}\right)$$

$$\therefore \frac{(X_3 - \frac{X_1 + X_2}{2})^2}{\frac{3}{2}} \sim \chi_1^2$$

$$\therefore X_4 - \frac{X_1 + X_2 + X_3}{3} \sim N\left(0, \sqrt{\frac{4}{3}}\right)$$

$$\therefore \frac{(X_4 - \frac{X_1 + X_2 + X_3}{3})^2}{\frac{4}{3}} \sim \chi_1^2$$

$$\mathbf{X} = \begin{pmatrix} X_1 & X_2 & X_3 & X_4 \end{pmatrix}'$$

$$\begin{pmatrix} X_1 - X_2 \\ X_3 - \frac{X_1 + X_2}{2} \\ X_4 - \frac{X_1 + X_2 + X_3}{3} \end{pmatrix} = \begin{pmatrix} 1 & -1 & 0 & 0 \\ -\frac{1}{2} & -\frac{1}{2} & 1 & 0 \\ -\frac{1}{3} & -\frac{1}{3} & -\frac{1}{3} & 1 \end{pmatrix} \begin{pmatrix} X_1 \\ X_2 \\ X_3 \\ X_4 \end{pmatrix} = \mathbf{A}\mathbf{X}$$

$$var(\mathbf{A}\mathbf{X}) = \mathbf{A}var(\mathbf{X})\mathbf{A}' = \begin{pmatrix} 1 & -1 & 0 & 0 \\ -\frac{1}{2} & -\frac{1}{2} & 1 & 0 \\ -\frac{1}{3} & -\frac{1}{3} & -\frac{1}{3} & 1 \end{pmatrix} \begin{pmatrix} 1 & 0 & 0 & 0 \\ 0 & 1 & 0 & 0 \\ 0 & 0 & 1 & 0 \\ 0 & 0 & 0 & 1 \end{pmatrix} \begin{pmatrix} 1 & -\frac{1}{2} & -\frac{1}{3} \\ -1 & -\frac{1}{2} & -\frac{1}{3} \\ 0 & 1 & -\frac{1}{3} \\ 0 & 0 & 1 \end{pmatrix} = \begin{pmatrix} 2 & 0 & 0 \\ 0 & \frac{3}{2} & 0 \\ 0 & 0 & \frac{4}{3} \end{pmatrix}$$

So the three terms in the RHS are independent each with a χ_1^2 distribution.