## Multiple birth-death model – conditioning probability

The idea is to keep track of the number of species  $m_1$  and  $m_2$  descending from each of the crown species. We construct the Q-equation for the components  $Q_{m_1,m_2}$ . Multiple-birth events can affect all species  $(2 + m_1 + m_2)$  in total). Their contribution to the Q-equation is

$$\begin{aligned} \text{contrib} &= \nu \sum_{a_1=0}^{\lfloor \frac{1+m_1}{2} \rfloor} \sum_{a_2=0}^{\lfloor \frac{1+m_2}{2} \rfloor} \sum_{j_1=0}^{\min(1,a_1)} \sum_{j_2=0}^{\min(1,a_2)} 2^{j_1+j_2} \binom{1}{j_1} \binom{1}{j_2} \binom{m_1-a_1}{a_1-j_1} \binom{m_2-a_2}{a_2-j_2} \\ & q^{a_1+a_2} (1-q)^{2+m_1+m_2-2a_1-2a_2} Q_{m_1-a_1,m_2-a_2} \\ &= \nu \sum_{a_1=0}^{\lfloor \frac{1+m_1}{2} \rfloor} \sum_{a_2=0}^{\lfloor \frac{1+m_2}{2} \rfloor} \frac{m_1+1}{m_1-2a_1+1} \binom{m_1-a_1}{a_1} \frac{m_2+1}{m_2-2a_2+1} \binom{m_2-a_2}{a_2} \\ & = \nu \sum_{a_1=0}^{\lfloor \frac{1+m_1}{2} \rfloor} \sum_{a_2=0}^{\lfloor \frac{1+m_2}{2} \rfloor} \frac{m_1+1}{m_1-2a_1+1} \binom{m_1-a_1}{a_1} \frac{m_2+1}{m_2-2a_2+1} \binom{m_2-a_2}{a_2} \\ & = \nu \sum_{a_1=0}^{\lfloor \frac{1+m_1}{2} \rfloor} \sum_{a_2=0}^{\lfloor \frac{1+m_2}{2} \rfloor} \frac{(m_1+1)(m_1-a_1)!}{(m_1-2a_1+1)!} \frac{(m_2+1)(m_2-a_2)!}{(m_2-2a_2+1)!} \\ & = \nu \sum_{a_1=0}^{\lfloor \frac{1+m_1}{2} \rfloor} \sum_{a_2=0}^{\lfloor \frac{1+m_2}{2} \rfloor} \frac{(m_1+1)(m_1-a_1)!}{(m_1-2a_1+1)!} \frac{(m_2+1)(m_2-a_2)!}{(m_2-2a_2+1)!} \\ & = \nu \sum_{a_1=0}^{\lfloor \frac{1+m_1}{2} \rfloor} \sum_{a_2=0}^{\lfloor \frac{1+m_2}{2} \rfloor} \frac{(m_1+1)(m_1-a_1)!}{(m_1-2a_1+1)!} \frac{(m_2-a_2)!}{(m_2-2a_2+1)!} \\ & = \nu \sum_{a_1=0}^{\lfloor \frac{1+m_2}{2} \rfloor} \sum_{a_2=0}^{\lfloor \frac{1+m_2}{2} \rfloor} \frac{(m_1+1)(m_1-a_1)!}{(m_1-2a_1+1)!} \frac{(m_2-a_2)!}{(m_2-2a_2+1)!} \\ & = \nu \sum_{a_1=0}^{\lfloor \frac{1+m_2}{2} \rfloor} \sum_{a_2=0}^{\lfloor \frac{1+m_2}{2} \rfloor} \frac{(m_1+1)(m_1-a_1)!}{(m_1-2a_1+1)!} \frac{(m_2-a_2)!}{(m_2-2a_2+1)!} \\ & = \nu \sum_{a_1=0}^{\lfloor \frac{1+m_2}{2} \rfloor} \sum_{a_2=0}^{\lfloor \frac{1+m_2}{2} \rfloor} \frac{(m_1+1)(m_1-a_1)!}{(m_1-2a_1+1)!} \frac{(m_2-a_2)!}{(m_2-2a_2+1)!} \\ & = \nu \sum_{a_1=0}^{\lfloor \frac{1+m_2}{2} \rfloor} \sum_{a_2=0}^{\lfloor \frac{1+m_2}{2} \rfloor} \frac{(m_1+1)(m_1-a_1)!}{(m_1-2a_1+1)!} \frac{(m_2+1)(m_2-a_2)!}{(m_2-2a_2+1)!} \\ & = \nu \sum_{a_1=0}^{\lfloor \frac{1+m_2}{2} \rfloor} \sum_{a_1=0}^{\lfloor \frac{1+m_2}{2} \rfloor} \frac{(m_1+1)(m_1-a_1)!}{(m_1-2a_1+1)!} \frac{(m_2+1)(m_2-a_2)!}{(m_2-2a_2+1)!} \\ & = \nu \sum_{a_1=0}^{\lfloor \frac{1+m_2}{2} \rfloor} \frac{(m_1+1)(m_1-a_1)!}{(m_1-2a_1+1)!} \frac{(m_1+1)(m_1-a_1)!}{(m_2-2a_2+1)!} \\ & = \nu \sum_{a_1=0}^{\lfloor \frac{1+m_2}{2} \rfloor} \frac{(m_1+1)(m_1-a_1)!}{(m_1-2a_1+1)!} \frac{(m_2+1)(m_2-a_2)!}{(m_2-2a_2+1)!} \\ & = \nu \sum_{a_1=0}^{\lfloor \frac{1+m_2}{2} \rfloor} \frac{(m_1+1)(m_1-a_1)!}{(m_1-2a_1+1)!} \\ & = \nu \sum_{a_1=0}^{\lfloor \frac{1+m_2}{2} \rfloor} \frac{(m_1+1)(m_1-a_1)!}{(m_1-2a_1+1)!} \frac{(m_2+1)(m_2-a_2)!}{(m_2-2a_2+1)!} \\ & = \nu \sum_{a_1=0}^{\lfloor \frac{1+m_2}{2} \rfloor} \frac{(m_1+1)(m_1-a_1)!}{(m_$$

The full Q-equation is

$$\frac{dQ_{m_1,m_2}}{dt} = \lambda \left( (m_1+1)Q_{m_1-1,m_2} + (m_2+1)Q_{m_1,m_2-1} \right)$$

$$+ \mu \left( (m_1+1)Q_{m_1+1,m_2} + (m_2+1)Q_{m_1,m_2+1} \right)$$

$$- (\lambda + \mu)(m_1 + m_2 + 2)Q_{m_1,m_2}$$

$$+ (contribution of multiple-birth events)$$

The solution of the Q-equation has the following interpretation: the probability that the diversification process has  $m_1 + 1$  descendant species of the first crown species and  $m_2 + 1$  descendant species of the second crown species is given by

$$P_{m_1,m_2}^c = \frac{Q_{m_1,m_2}}{(m_1+1)(m_2+1)}$$

Hence, the conditional probability  $P_c$  is given by

$$P^{c} = \sum_{m_{1}, m_{2}} P^{c}_{m_{1}, m_{2}} = \sum_{m_{1}, m_{2}} \frac{Q_{m_{1}, m_{2}}}{(m_{1} + 1)(m_{2} + 1)}$$

My claim is that

- in the model without multiple births ( $\nu = 0$ ), the terms in this sum for fixed  $m_1 + m_2$  are equal, i.e., all combinations ( $m_1, m_2$ ) for fixed  $m_1 + m_2$  are equally probable. In the PRSB paper we took advantage of this property to efficiently compute the conditioning probability.
- in the model with multiple births ( $\nu > 0$ ), this symmetry property does not hold. As a result, we have to compute the conditioning probability explicitly.

Here is a (quick – it can be improved) implementation of this approach for your example (this is Matlab code):

```
la=0.0204942104;
mu=0.0001333249;
nu=1.5728643216:
pq=0.0787076385;
tt=10; % time between crown age and present
lq=100; % maximal number of missing species
% construct auxiliary matrix
cc=probcond2_matr(pq,lq);
% integrate equations
[~,X4]=ode45(@(t,x) probcond2_rhs(x,la,mu,0,zeros(100),1),[0 tt],x0);
Q4=X4(end,:);
Q4=reshape(Q4,lq,lq);
% compute conditioning probability
m1=ones(lq,1)*(0:(lq-1));
Pc4=Q4./((m1+1).*(m2+1));
sum(Pc4(:))
```

The code for the right-hand side of the differential equation:

And the code for the auxiliary matrix used in the differential equation:

```
function [cc,cp]=probcond2_matr(pq,lq)

cc=zeros(lq);
cp=zeros(lq);
for m1=0:(lq-1)
    for a1=0:floor((1+m1)/2)
        aux=log(m1+1)+gammaln(m1-a1+1)-gammaln(m1-2*a1+2)-gammaln(a1+1);
        aux=exp(aux);
        cp(m1+1,m1-a1+1)=aux;
        aux=aux*pq^a1*(1-pq)^(m1+1-2*a1);
        cc(m1+1,m1-a1+1)=aux;
    end
end
```

The conditioning probability I obtain is 0.9985889671, smaller than one as it should be. I checked that the symmetry property is verified for  $\nu=0$  and that it is not verified for  $\nu>0$ . I also checked that for  $\nu=0$  I get the same conditioning probability as previously.