

Stochastic Process Homework 02

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0 Reference

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1 Optimal Coupling

Proof. Since

$$D_{\text{TV}}(\mu, \nu) = \frac{1}{2} \sum_{i \in \Omega} |\mu(i) - \nu(i)| = \sum_{i \in \Omega, \nu(i) \geq \mu(i)} (\nu(i) - \mu(i)), \quad (1)$$

it is natural to maximize $\omega(i, i)$, i.e. set $\omega(i, i) = \min \{\mu(i), \nu(i)\}$.

Considering $\mu(i) - \min \{\mu(i), \nu(i)\} = \max \{0, \mu(i) - \nu(i)\}$, $\nu(j) - \min \{\mu(j), \nu(j)\} = \max \{0, \nu(j) - \mu(j)\}$, we construct a feasible coupling ω as follows.

$$\omega(i, j) = \begin{cases} \min \{\mu(i), \nu(i)\}, & i = j \\ C \max \{0, \mu(i) - \nu(i)\} \max \{0, \nu(j) - \mu(j)\}, & i \neq j \end{cases}$$

Now we solve C . We know

$$\begin{aligned} \mu(i) &= \sum_{j \in \Omega} \omega(i, j) \\ &= \sum_{j \neq i \in \Omega} C \max \{0, \mu(i) - \nu(i)\} \max \{0, \nu(j) - \mu(j)\} + \min \{\mu(i), \nu(i)\} \\ &= \min \{\mu(i), \nu(i)\} + C \max \{0, \mu(i) - \nu(i)\} \sum_{j \neq i \in \Omega} \max \{0, \nu(j) - \mu(j)\} \end{aligned}$$

CASE 01. When $\min(\mu(i), \nu(i)) = \mu(i)$, obvious the equation above holds.

CASE 02. When $\min(\mu(i), \nu(i)) = \nu(i)$,

by (1), we have

$$\begin{aligned}
\mu(i) &= \nu(i) + C(\mu(i) - \nu(i)) \sum_{j \neq i \in \Omega, \nu(j) \geq \mu(j)} (\nu(j) - \mu(j)) \\
&= \nu(i) + C(\mu(i) - \nu(i)) D_{\text{TV}}\{\mu, \nu\}
\end{aligned}$$

Thus,

$$C = \frac{1}{D_{\text{TV}}(\mu, \nu)}$$

i.e.

$$\omega(i, j) = \begin{cases} \min\{\mu(i), \nu(i)\}, & i = j \\ \frac{1}{D_{\text{TV}}(\mu, \nu)} \max\{0, \mu(i) - \nu(i)\} \max\{0, \nu(j) - \mu(j)\}, & i \neq j \end{cases}$$

By the symmetry of μ and ν , i and j in the equation, we know

$$\nu(j) = \sum_{i \in \Omega} \omega(i, j)$$

also holds. Thus, ω is a valid coupling.

Now we prove $\mathbf{Pr}_{(X,Y) \sim \omega} [X \neq Y] = D_{\text{TV}}(\mu, \nu)$.

$$\begin{aligned}
\mathbf{Pr}_{(X,Y) \sim \omega} [X = Y] &= \sum_{i \in \Omega} \omega(i, i) \\
&= \sum_{i \in \Omega} \min\{\mu(i), \nu(i)\} \\
&= \sum_{i \in \Omega, \mu(i) \geq \nu(i)} \nu(i) + \sum_{i \in \Omega, \mu(i) < \nu(i)} \mu(i) \\
&= \sum_{i \in \Omega} \nu(i) + \sum_{i \in \Omega, \nu(i) \geq \mu(i)} (\mu(i) - \nu(i)) \\
&= 1 - \sum_{i \in \Omega, \nu(i) \geq \mu(i)} (\nu(i) - \mu(i)) \\
&= 1 - D_{\text{TV}}(\mu, \nu) \quad (\text{by (1)})
\end{aligned}$$

i.e.,

$$\mathbf{Pr}_{(X,Y) \sim \omega} [X \neq Y] = 1 - \mathbf{Pr}_{(X,Y) \sim \omega} [X = Y] = D_{\text{TV}}(\mu, \nu).$$

Thus, there exists a coupling

$$\omega(i, j) = \begin{cases} \min\{\mu(i), \nu(i)\}, & i = j \\ \frac{1}{D_{\text{TV}}(\mu, \nu)} \max\{0, \mu(i) - \nu(i)\} \max\{0, \nu(j) - \mu(j)\}, & i \neq j \end{cases}$$

s.t. $\mathbf{Pr}_{(X,Y) \sim \omega} [X \neq Y] = D_{\text{TV}}(\mu, \nu)$. ■

2 Stochastic Dominance

2.1 Binomial Distribution Case

Proof. First we prove the sufficiency.

Assume $p \geq q$. Let $X \sim \text{Binom}(n, p), Y \sim \text{Binom}(n, q)$.

For any $a \in \Omega$,

CASE 01. When $a \geq n + 1$, $\Pr[X \geq a] = \Pr[Y \geq a] = 0$.

CASE 02. When $a \leq -1$, $\Pr[X \geq a] = \Pr[Y \geq a] = 1$.

CASE 03. When $1 \leq a \leq n$, we have

$$\begin{aligned}\Pr[X \geq a] &= \sum_{i=a}^n \Pr[X = i] = \sum_{i=a}^n \binom{n}{i} p^i (1-p)^{n-i} \\ \Pr[Y \geq a] &= \sum_{i=a}^n \Pr[Y = i] = \sum_{i=a}^n \binom{n}{i} q^i (1-q)^{n-i}\end{aligned}$$

Consider $f(x) = \sum_{i=a}^n \binom{n}{i} x^i (1-x)^{n-i}$. When $x \in [0, 1]$,

$$\begin{aligned}\frac{df}{dx} &= nx^{n-1} + \sum_{i=a}^{n-1} \binom{n}{i} (ix^{i-1}(1-x)^{n-i} - (n-i)x^i(1-x)^{n-i-1}) \\ &= nx^{n-1} + \sum_{i=a}^{n-1} \binom{n}{i} (i-nx)x^{i-1}(1-x)^{n-i-1} \geq 0\end{aligned}$$

i.e. $f(x)$ is monotonously increasing on $[0, 1]$.

Since $p, q \in [0, 1]$, $p \geq q$, we know $f(p) \geq f(q)$, i.e.

$$\Pr[X \geq a] \geq \Pr[Y \geq a].$$

Thus, for any $a \in \Omega$, $\Pr[X \geq a] \geq \Pr[Y \geq a]$, i.e. $\text{Binom}(n, p) \succeq \text{Binom}(n, q)$.

Now we prove the necessity.

Assume $\text{Binom}(n, p) \succeq \text{Binom}(n, q)$. Let $X \sim \text{Binom}(n, p), Y \sim \text{Binom}(n, q)$.

Then we have

$$\begin{aligned}\Pr[X \geq n] \geq \Pr[Y \geq n] &\iff \Pr[X = n] \geq \Pr[Y = n] \\ &\iff p^n \geq q^n \\ &\iff p \geq q.\end{aligned}$$

In conclusion, for any $p, q \in [0, 1]$, $\text{Binom}(n, p) \succeq \text{Binom}(n, q)$ **iff.** $p \geq q$. ■

2.2 Monotone Coupling

Proof. First we prove the sufficiency.

Assume exists a monotone coupling ω of μ and ν . Then we know

$$1 = \mathbf{Pr}_{(X,Y) \sim \omega} [X \geq Y] = \sum_{i \in \Omega} \sum_{j \leq i} \omega(i, j) = 1 = \sum_{i \in \Omega} \sum_{j \in \Omega} \omega(i, j)$$

Thus, when $i < j$, $\omega(i, j) = 0$.

$$\begin{aligned} \forall a \in \Omega, \quad \mathbf{Pr}_{X \sim \mu} [X \geq a] &= \mathbf{Pr}_{(X,Y) \sim \omega} [X \geq a] = \sum_{i \in \Omega, i \geq a} \mathbf{Pr}_{(X,Y) \sim \omega} [X = i] \\ &= \sum_{i \in \Omega, i \geq a} \sum_{j \in \Omega} \omega(i, j) = \sum_{j \in \Omega} \sum_{i \in \Omega, i \geq a} \omega(i, j) \\ &\geq \sum_{j \in \Omega, j \geq a} \sum_{i \in \Omega, i \geq j} \omega(i, j) = \sum_{j \in \Omega, j \geq a} \sum_{i \in \Omega, i \geq j} \omega(i, j) + 0 \\ &= \sum_{j \in \Omega, j \geq a} \sum_{i \in \Omega, i \geq j} \omega(i, j) + \sum_{j \in \Omega, j \geq a} \sum_{i \in \Omega, i < j} \omega(i, j) \\ &= \sum_{j \in \Omega, j \geq a} \sum_{i \in \Omega} \omega(i, j) = \sum_{j \in \Omega, j \geq a} \mathbf{Pr}_{(X,Y) \sim \omega} [Y = j] \\ &= \mathbf{Pr}_{(X,Y) \sim \omega} [Y \geq a] \\ &= \mathbf{Pr}_{Y \sim \nu} [Y \geq a]. \end{aligned}$$

i.e. $\mu \succeq \nu$.

Now we prove the necessity, i.e. to construct a coupling ω s.t. $\mathbf{Pr}_{(X,Y) \sim \omega} [X \geq Y] = 1$.

Assume $\mu \succeq \nu$. We construct ω as follows.

First we set $\omega(i, j) = 0$ for $i < j$.

We have $\mu(n) = \mathbf{Pr}_{X \sim \mu} [X \geq n] \geq \mathbf{Pr}_{Y \sim \nu} [Y \geq n] = \nu(n)$. Set $\omega(n, n) = \nu(n)$.

For the remaining part, we determine $\omega(i, j)$ in the following order

$$\begin{aligned} &\omega(n, n-1), \omega(n-1, n-1), \\ &\omega(n, n-2), \omega(n-1, n-2), \omega(n-2, n-2), \\ &\dots, \\ &\omega(n, 1), \omega(n-1, 1), \dots, \omega(1, 1) \end{aligned}$$

by the following method.

$$\omega(i, j) = \min \left\{ \mu(i) - \sum_{k=j+1}^n \omega(i, k), \quad \nu(j) - \sum_{k=i+1}^n \omega(k, j) \right\}$$

We prove ω is a valid coupling as follows. Obvious $\mu(i) = \sum_{j \in \Omega} \omega(i, j)$, $\nu(j) = \sum_{i \in \Omega} \omega(i, j)$, which is maintained and guaranteed by how we determined $\omega(i, j)$.

Meanwhile, $\omega(i, j) \geq 0$, (since $\sum_{k=j+1}^n \omega(i, k) \leq \mu(i)$, $\sum_{k=i+1}^n \omega(k, j) \leq \nu(j)$).

Thus, ω is a valid coupling s.t. $\Pr_{(X,Y) \sim \omega} [X \geq Y] = 1$.

In conclusion, $\mu \succeq \nu$ **iff.** exists a coupling ω s.t. $\Pr_{(X,Y) \sim \omega} [X \geq Y] = 1$. ■

2.3 Erdős–Rényi Model Random Graph

- The completion of this problem is based on the notes of **AI2613** last year.

Proof. We can generate $G \sim \mathcal{G}(n, p)$ and $H \sim \mathcal{G}(n, q)$ simultaneously, where $p, q \in [0, 1]$, $p \geq q$.

For each pair of vertices (u, v) , we independently pick $r_{\{u,v\}} \sim \text{Uniform}(0, 1)$. We determine whether graph G and H has edge $\{u, v\}$ as follows.

$$\begin{cases} \text{both } G \text{ and } H \text{ have edge } \{u, v\}, & r_{\{u,v\}} \in [0, q] \\ \text{only } G \text{ has edge } \{u, v\}, & r_{\{u,v\}} \in (q, p] \\ \text{neither has edge } \{u, v\}, & r_{\{u,v\}} \in (p, 1] \end{cases}$$

Obvious H is always a subgraph of G . If H is connected, G is for sure connected.

Thus, $\Pr[G \text{ is connected}] \geq \Pr[H \text{ is connected}]$.

Therefore, for any $p, q \in [0, 1]$ s.t. $p \geq q$,

$$\Pr_{G \sim \mathcal{G}(n,p)} [G \text{ is connected}] \geq \Pr_{H \sim \mathcal{G}(n,q)} [H \text{ is connected}].$$

Qed. ■

3 Total Variation Distance is Non-Increasing

Proof. By **Coupling Lemma**, we know

$$\Delta(t) \leq \mathbf{Pr}_{(X,Y) \sim \omega_t} [X \neq Y],$$

where ω_t is a coupling of μ_t and π .

We construct a coupling ω_t as follows, where $(X_t, Y_t) \sim \omega_t, (X_{t+1}, Y_{t+1}) \sim \omega_{t+1}$.

$$\begin{cases} X_{t+1} = X_t = Y_t = Y_{t+1}, & \text{if } X_t = Y_t \\ X_{t+1} \sim \mu_t, Y_{t+1} \sim \pi, & \text{if } X_t \neq Y_t \end{cases}$$

Then we have

$$\begin{aligned} \Delta(t+1) &\leq \mathbf{Pr}_{(X,Y) \sim \omega_{t+1}} [X \neq Y] \\ &= \mathbf{Pr} [X_{t+1} \neq Y_{t+1}] \\ &\leq \mathbf{Pr} [X_t \neq Y_t] \\ &= \Delta(t). \end{aligned}$$

Qed. ■