Imperial College London

Coursework 1

IMPERIAL COLLEGE LONDON

DEPARTMENT OF MATHEMATICS

Stochastic Simulation

Author: 01844579

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1 Question 1

1.1 Question 1.1

To compute $M_{\lambda} = \sup_{x} p_{\nu}(x)/q_{\lambda}(x)$ we will first optimise over x and find x^* and then substitute it into our expression. First we let,

$$R(x) = \frac{p_{\nu}(x)}{q_{\lambda}(x)} = \frac{1}{2^{\frac{\nu}{2}}\Gamma(\frac{\nu}{2})} \cdot \frac{1}{\lambda} \cdot x^{\frac{\nu}{2}-1} \cdot e^{-\frac{x}{2}} \cdot e^{\lambda x} = \frac{1}{2^{\frac{\nu}{2}}} \cdot \frac{1}{\Gamma(\frac{\nu}{2})} \cdot \frac{1}{\lambda} \cdot x^{\frac{\nu}{2}-1} \cdot e^{x(\lambda-\frac{1}{2})}.$$

To find the turning points of R(x) it is easier to first take the logarithm,

$$\log(R(x)) = \log\left(\frac{1}{2^{\frac{\nu}{2}}} \cdot \frac{1}{\Gamma(\frac{\nu}{2})} \cdot \frac{1}{\lambda}\right) + \left(\frac{\nu}{2} - 1\right)\log(x) + x\left(\lambda - \frac{x}{2}\right),$$

and then differentiate

$$\frac{d\log(R(x))}{dx} = \frac{1}{x}\left(\frac{\nu}{2} - 1\right) + \lambda - \frac{1}{2}.$$

Setting the derivative equal to zero,

$$\frac{1}{x}\left(\frac{\nu}{2}-1\right)+\lambda-\frac{1}{2}=0 \implies \frac{1}{x}\left(\nu-2\right)=1-2\lambda.$$

We find that $x^* = \frac{v-2}{1-2\lambda}$. To check x^* is the optimal value of x we find the second derivative of $\log(R(x))$.

$$\frac{d^2 \log(R(x))}{dx^2} = -\frac{1}{x^2} \left(\frac{\nu}{2} - 1 \right)$$

Upon substituting $x^* = \frac{v-2}{1-2\lambda}$ we get

$$\frac{d^2\log(R(x^*))}{dx^2} = -\left(\frac{1-2\lambda}{\nu-2}\right)^2\left(\frac{\nu}{2}-1\right)$$

Noticing that our first term is strictly positive since $0 < \lambda < 1/2$ and that we have that $\nu/2 - 1 > 0$ since we assumed that $\nu > 2$. So our second derivative is strictly smaller than 0 and therefore a maximum. Now that we have found x^* and proved it is indeed the maximum, we substitute it into M_{λ} . So our expression reads

$$M_{\lambda} = \sup_{x} \frac{p_{\nu}(x)}{q_{\lambda}(x)} = \frac{p_{\nu}(x^*)}{q_{\lambda}(x^*)} = \frac{1}{2^{\frac{\nu}{2}} \Gamma(\frac{\nu}{2}) \lambda} \cdot \left(\frac{\nu-2}{1-2\lambda}\right)^{\frac{\nu}{2}-1} \cdot e^{\left(\lambda - \frac{1}{2}\right)\left(\frac{\nu-2}{1-2\lambda}\right)}.$$

1.2 Question 1.2

After finding x^* and an expression for M_{λ} we want to now find the optimal λ^* . To do so we first take the logarithm of M_{λ}

$$\begin{split} \log(M_{\lambda}) &= \log\left(\frac{1}{2^{\frac{\nu}{2}}\Gamma(\frac{\nu}{2})\lambda}\right) + \left(\frac{\nu}{2} - 1\right)\log\left(\frac{\nu - 2}{1 - 2\lambda}\right) + \left(\lambda - \frac{1}{2}\right)\left(\frac{\nu - 2}{1 - 2\lambda}\right) \\ &= \log\left(\frac{1}{2^{\frac{\nu}{2}}\Gamma(\frac{\nu}{2})}\right) - \log(\lambda) + \left(\frac{\nu}{2} - 1\right)\log(\nu - 2) - \left(\frac{\nu}{2} - 1\right)\log(1 - 2\lambda) + (\nu - 2)\left(\frac{\lambda - \frac{1}{2}}{1 - 2\lambda}\right) \end{split}$$

Now differentiating our expression with respect to λ

$$\frac{d\log(M_{\lambda})}{d\lambda} = -\frac{1}{\lambda} + \left(\frac{\nu}{2} - 1\right) \frac{2}{1 - 2\lambda} + (\nu - 2) \left(\frac{(1 - 2\lambda) + 2\left(\lambda - \frac{1}{2}\right)}{(1 - 2\lambda)^2}\right).$$

Noticing that the last term is equal to 0 and setting the expression equal to 0 we obtain

$$-\frac{1}{\lambda} + \frac{\nu - 2}{1 - 2\lambda} = 0 \implies \lambda^* = \frac{1}{\nu}.$$

To check λ^* indeed minimises M_{λ} we find that the second derivative of $\log(M_{\lambda})$ is

$$\frac{d^2\log(M_\lambda)}{d\lambda^2} = \frac{1}{\lambda^2} + \frac{2(\nu-2)}{(1-2\lambda)^2}$$

Plugging in our value of $\lambda^* = 1/\nu$ we obtain

$$\frac{d^2 \log(M_{\lambda^*})}{d\lambda^2} = \nu^2 + \frac{2\nu^2(\nu-2)}{\nu^2-4\nu+4} = \frac{\nu^2\big(\nu^2-4\nu+4\big)+2\nu^2(\nu-2)}{\nu^2-4\nu+4} = \frac{\nu^3(\nu-2)}{(\nu-2)^2} = \frac{\nu^3}{\nu-2}.$$

Note that our final result is strictly greater than 0 since we have that $\nu > 2$, therefore $\lambda^* = 1/\nu$ minimises M_{λ} .

1.3 Question 1.3

Figure 1 shows a histogram of our rejection sampler for $\nu = 4$ with the corresponding optimal λ . The theoretical acceptance rate is $\hat{a} = 0.6796$ and the calculated acceptance rate for Figure 1 is a = 0.6816. As we can see the acceptance rates obtained are very close values to one another, demonstrated by the histogram and density plot.

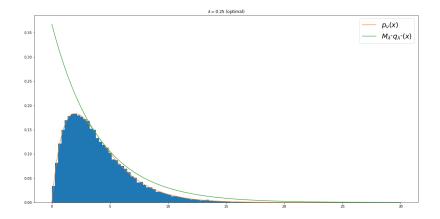


Figure 1: Rejection sampling procedure with optimal $\lambda = 1/\nu = 0.25$ for n = 100000.

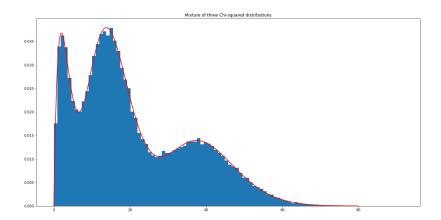


Figure 2: Histogram and density plot of mixture of Chi-squared distributions

2 Question 2

Figure 2 shows sampling from discrete mixture of Chi-squared distributions. First, indices are sampled from a discrete distribution using the inversion method. Then modifying our sampling previous rejection sampler we can take samples of respective Chi-squared distributions.

A Code for Question 1

```
n = 100000 # Number of samples to draw
2 nu = np.array([4, 16, 40]) # Array of nu values (will only use first
     element for Q1)
3 lam = 1/nu # Lambda values (will only use first element for Q1)
4 count = 0 # Counter to calculate accepted samples
s w = [0.2, 0.5, 0.3] # Weights of discrete distribution
_{6} s = [0, 1, 2] # Support of discrete distribution
8 x_samples = [] # Empty list to append samples from exponential
     distribution
9 samples = [] # Empty list to append samples from a rejection sampling
      procedure
# Defining functions
12 def p(x, nu):
     return x ** (nu / 2 - 1) * np.exp(-x / 2) / (2 ** (nu / 2) * np.
     math.factorial(int(nu / 2) - 1))
 def q(x, lam):
     return lam*np.exp(-lam * x)
16
17
18 def m(nu):
     return nu * nu ** (nu / 2 - 1) * np.exp(1 - nu / 2) / (2 ** (nu /
      2) * np.math.factorial(int(nu / 2) - 1))
```

```
20
21 m = np.array([m(nu[0]), m(nu[1]), m(nu[2])]) # M values for
     acceptance probabilities (derived)
22
23 for i in range(n): # Code to sample from exponential distribution
     using inversion
24
      a = np.random.uniform(0,1)
25
      b = -(1/lam[0]) * np.log(1-a)
26
      x_samples = np.append(x_samples, b)
28
 for i in range(n): # Rejection algorithm
29
30
      x = x_samples[i] # Proposal
31
      u = np.random.uniform(0,1) # Uniform
32
      acceptance = p(x, nu[0])/(m[0]*q(x, lam[0])) # Acceptance
     probability
      if u < acceptance: # Accept/reject</pre>
35
          samples = np.append(samples, x) # Store accepted values
36
          count += 1 # Increase count for accepted samples (for
37
     acceptance rate)
38
39 fig = plt.figure(figsize=(20,10)) # Code for plotting histogram and
     densities
x = np.linspace(0, 30, 1000)
41 plt.hist(samples,bins=100, density =True)
42 plt.plot(x, p(x, nu[0]), label = r'$p_{\nu}(x)$')
43 plt.plot(x, m[0]*q(x,lam[0]), label = r'$M_{\lambda}^{1}ambda^{*}} q_{\lambda}^{1}ambda
      ^{*}}(x)$')
plt.legend(fontsize=20)
45 plt.title(r'$\lambda = 0.25$ (optimal)')
46 plt.show()
48 # Calculations for acceptance rates
49 print(f"Acceptance Rate: {count/n}")
print(f"Theoretical Acceptance Rate: {1/m[0]}")
fig.savefig('q1.png')
```

B Code for Question 2

```
def rejection_sampler(nu, lam, m): # Modified previous rejection
    sampler to obtain a single sample

sample = []

while len(sample) < 1:

a = np.random.uniform(0,1)
    x = -(1/lam) * np.log(1-a)
    u = np.random.uniform(0,1)
    acceptance = p(x, nu)/(m*q(x, lam))</pre>
```

```
if u < acceptance:</pre>
12
              sample = np.append(sample, x)
13
      return sample
14
15
16 def discrete_sampler(s, w): # Samples from discrete distribution
     defined on s with probabilities w (for indicies)
17
      cdf = np.cumsum(w)
18
      sample = []
19
      u = np.random.uniform(0,1)
20
21
      for k in range(len(cdf)):
22
          if cdf[k] > u:
23
              sample = s[k]
              break
25
      return sample
26
28 samples = [] # Empty list to append accepted samples
30 for i in range(n): # Samples from mixture of distributions
31
      u = discrete_sampler(s,w) # Sample from discrete distribution
32
      samples = np.append(samples, rejection_sampler(nu[u],lam[u], m[u
33
     ])) # Samples from appropriate Chi-squared
def mixture_density(x, w, nu): # Code for plotting density
     return w[0]*p(x,nu[0]) + w[1]*p(x,nu[1]) + w[2]*p(x,nu[2])
38 figg = plt.figure(figsize=(20,10)) # Code for plotting histogram and
     density
xx = np.linspace(0, 80, 1000)
40 plt.hist(samples, bins =100, density=True)
plt.plot(xx, mixture_density(xx, w, nu), color='r', linewidth=2)
42 plt.title('Mixture of three Chi-squared distributions')
43 plt.show()
44 figg.savefig('q2.png')
```