

$$f(x) = \frac{1}{\sqrt{2\pi}} e^{-\frac{x^2}{2}}$$
$$\int_0^x f(t) dt = F(x)$$

The Acceptance-Rejection Method

Chapter 2 (2)

STATS 102C: Introduction to Monte Carlo Methods

UCLA



Introduction

- ▶ This indirect method will allow us to simulate virtually any distribution, and only require us to know the form of the *target density* f up to a multiplicative constant.
- ▶ The main idea is that we generate a candidate random variable from a simpler density g (called the instrumental or candidate density) and only accept it subject to passing a test. The only constraints we impose on this candidate density g are that

(i) $f(x) > 0$ and $g(x) > 0$

(ii) There is a constant M with $\frac{f(x)}{g(x)} \leq M$ for all x .

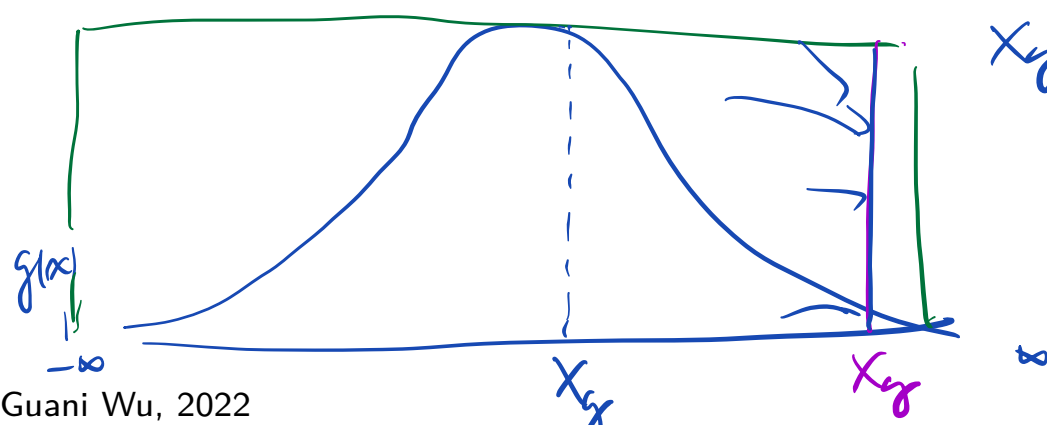
$$\underline{u} \sim \text{Unif}(0,1)$$

$$\underline{x}_g \sim g(x|\theta_g)$$

$$0 \leq \left| \frac{f(x)}{g(x)} \right| \leq 1$$

$$X_g \sim \text{Unif}(-\infty, \infty)$$

$$u \sim \text{Unit}(0,1)$$



$u > \frac{f(x_g)}{M \cdot g(x_g)}$ reject x_g

Rejection Method Algorithm

(1) Generate $X_g \sim g(x)$, $U \sim \text{Uniform}(0, 1)$;

(2) Accept X_g if $U < \frac{f(x_g)}{Mg(x_g)}$;

(3) Return to (1) otherwise.

$$u \sim \text{Unif}(0, 1)$$

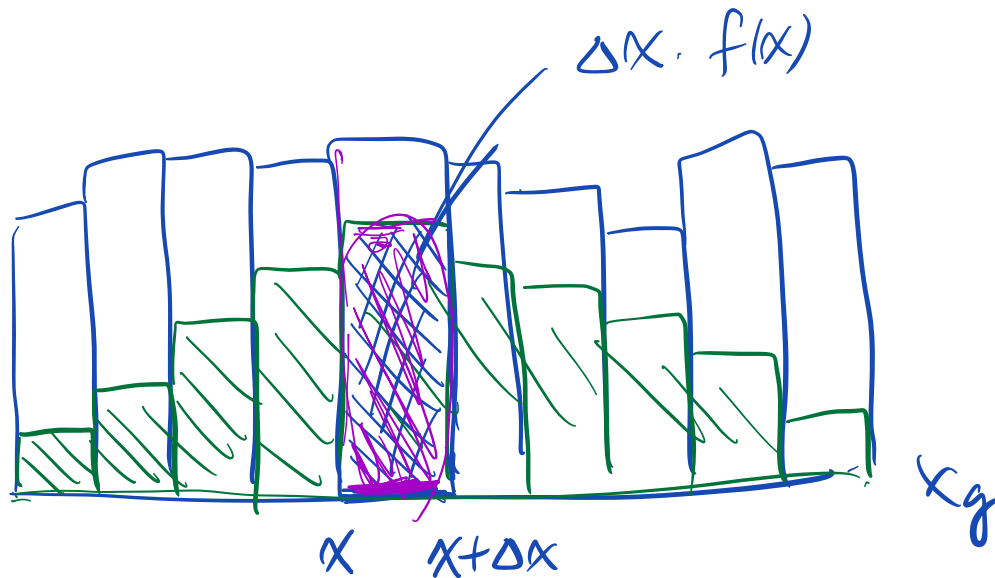
$$u < \frac{f(x_g)}{Mg(x_g)} \quad \text{accept } x_g$$

$$X_g^{(\text{accepted})} \sim f(x)$$

Can we prove that the accepted points $x \sim f(x)$ as desired? ✓

$$X_1, X_2, \dots, X_N \sim g(x|\theta)$$

$$u \sim \text{Unif}(0, 1)$$



$$= \Delta x \cdot f(x)$$

$$f(x) \approx \frac{\text{phob. of } x \text{ accepted in } (x, x+\Delta x)}{\Delta x} = \frac{P(X \in (x, x+\Delta x))}{\Delta x}$$

$$P(\text{event}) = \frac{\text{\# of event}}{\text{sample space}} = \frac{\frac{N}{M} \Delta x f(x)}{\frac{N}{M}} = \Delta x \cdot f(x)$$

$$\begin{aligned} \text{\# of data points accepted in } (x, x+\Delta x) &= N \cdot g(x) \Delta x \frac{f(x)}{M \cdot g(x)} \\ \text{total \# of points in } (x, x+\Delta x) &\times \frac{P(X \text{ being accepted})}{\frac{f(x)}{M \cdot g(x)}} = \frac{N}{M} \Delta x f(x) \end{aligned}$$

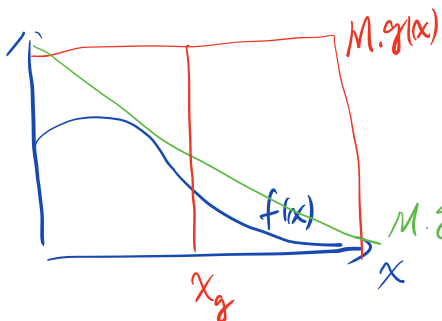
Across the all bins, the # of points accepted

$$\sum_{\text{bins}} \frac{N}{M} \Delta x f(x) = \frac{N}{M} \sum_{\text{bin}} \Delta x f(x) = \left(\frac{N}{M} \right)$$

$$\text{The acceptance rate} = \frac{\text{\# of data points accepted}}{\text{total data points}} = \frac{N/M}{N} = \frac{1}{M}$$

$\frac{1}{M} \uparrow \Rightarrow M \text{ must be small}$

$$0 \leq \frac{f(x)}{M \cdot g(x)} \leq 1 \Rightarrow \frac{f(x)}{g(x)} \leq M \quad \forall x$$



$$\max \frac{f(x)}{g(x)} \leq M$$

$$M = \max \left[\frac{f(x)}{g(x)} \right]$$

Toy example

X	X ₁	X ₂
P(X)	0.6	0.4

$$M = \max \left(\frac{P(X)}{g(X)} \right)$$

$$\begin{array}{c} \stackrel{=1.2}{\left(\frac{0.6}{0.5} \right)} \quad \frac{0.4}{0.5} \end{array}$$

① Generate 1000 random samples

X	X ₁	X ₂
g(X)	0.5	0.5

$$\begin{array}{c} 500 \quad X_1 \\ 500 \quad X_2 \end{array}$$

For our sample prob. function — For X₁, $500 \times \frac{P(X_1)}{M \cdot g(X_1)} = 500 \times \frac{0.6}{1.2(0.5)}$

$$P(X_1) = \frac{500}{500 + 333}$$

$$= 500$$

$$= \frac{500}{833} \approx 0.6$$

$$\text{For } X_2, \underline{500} \times \frac{P(X_2)}{M \cdot g(X_2)} = 500 \times \frac{0.4}{(1.2)(0.5)}$$

$$= 333$$

$$P(X_2) = \frac{333}{833} \approx 0.4$$

$$M = \max \left(\frac{f(x)}{g(x)} \right)$$

$$X_g \sim g(x) = \text{Unif}(0, \pi/2)$$

$$U_g \sim \text{Unif}(0, 1)$$

$$\underline{X_g} = \underline{U_g \frac{\pi}{2}}$$

$$M = \max \left(\frac{f(x_g)}{g(x_g)} \right)$$

$$u \sim \text{Unif}(0, 1)$$

$$\text{If } u \leq \frac{f(x_g)}{M g(x_g)} \quad \text{accept } x_g$$

Example 1: $f(x) = \sin x$

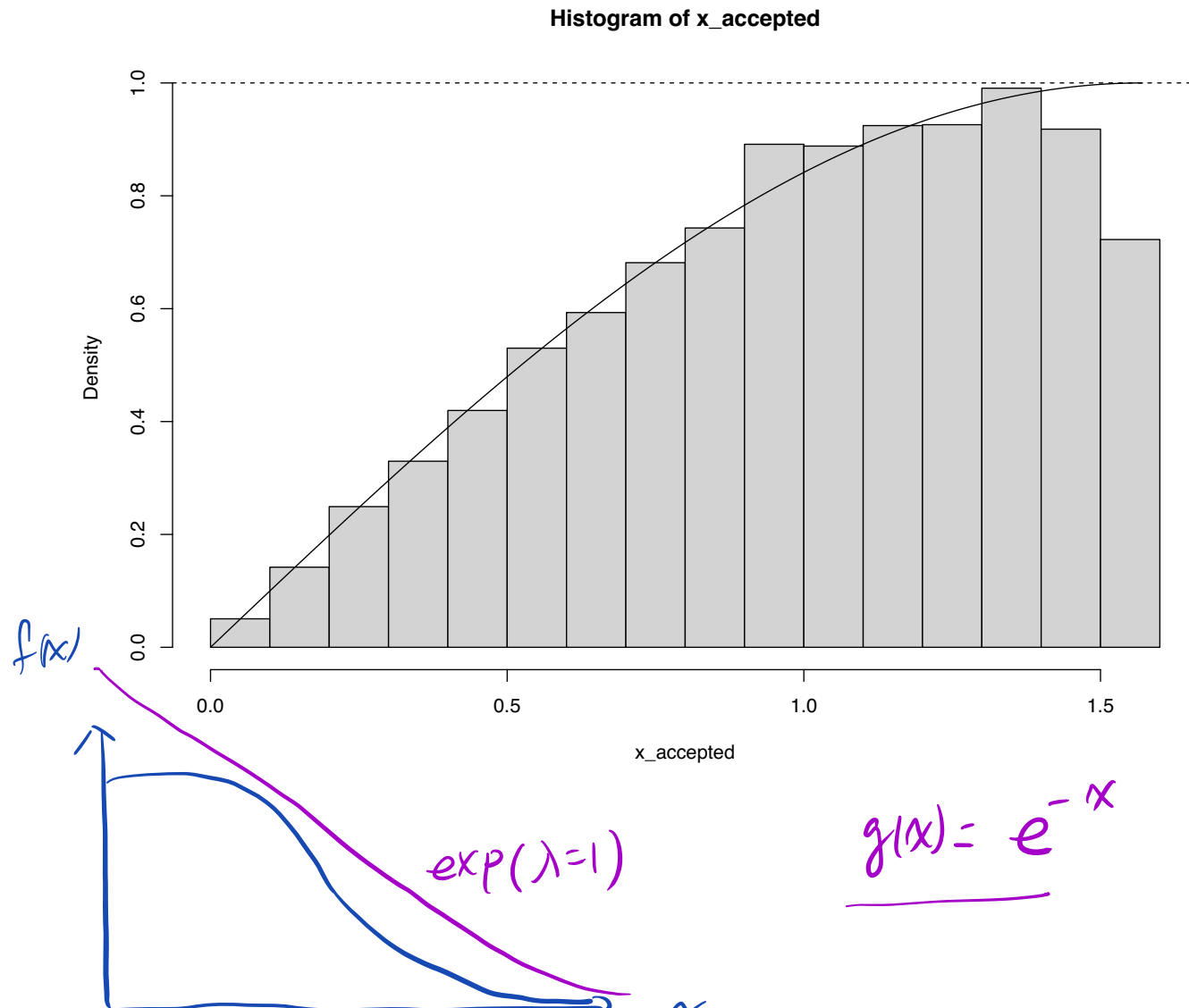
- ▶ Let $f(x) = \sin(x)$, for $x \in [0, \pi/2]$.
- ▶ $\int_0^{\pi/2} \sin(x) dx = 1$, so $f(x)$ is a PDF on $[0, \pi/2]$

The R code to generate samples from $f(x)$ using Rejection method.

```
set.seed(9999)
n <- 10000
u1 <- runif(n)
x <- u1 * (pi / 2)  这是个uniform dist的 inverse cdf
f_x <- sin(x)
M <- max(f_x) / (2 / pi)
u <- runif(n)
x_accepted <- x[which(u <= f_x / (M * (2 / pi)))]
```

Example 1: $f(x) = \sin x$ (Cont.)

```
hist(x_accepted, probability = T)
curve(sin, from = 0, to = pi/2, add = T)
abline(h = 1, lty = 2)
```



Example 2: Folded Normal Distribution

- ▶ Let $Z \sim N(0, 1)$ with pdf

$$f_Z(z) = \frac{1}{\sqrt{2\pi}} e^{-z^2/2}, \quad \text{for } z \in (-\infty, \infty).$$

- ▶ Suppose we want to sample from $X = |Z|$, which has PDF

$$f_X(x) = 2 \cdot \frac{1}{\sqrt{2\pi}} e^{-x^2/2} = \sqrt{\frac{2}{\pi}} e^{-x^2/2}, \quad \text{for } x \geq 0.$$

X has a **folded normal distribution**.

- ▶ How can we use rejection sampling to sample from X ?

$$M = \max \left(\frac{f(x)}{g(x)} \right) = \max_x \frac{\sqrt{\frac{2}{\pi}} e^{-x^2/2}}{e^{-x}} = \max_x \sqrt{\frac{2}{\pi}} e^{-\frac{(x^2}{2} - x)}$$

$$M = \frac{\sqrt{\frac{2}{\pi}} e^{-\frac{x^2}{2}}}{e^{-x}} = 1.315$$

$$\min_x \left(\frac{x^2}{2} - x \right) \\ \Rightarrow x^* = 1$$

R Code to sample from Folded Normal Distribution

写出 F, G

```
f <- function(x){  
  return(sqrt(2/pi) * exp(-x^2/2))  
}  
g <- function(x, lambda){  
  return(lambda * exp(-lambda * x))  
}
```

```
n <- 20000
```

```
v <- runif(n, 0, 1)
```

```
x <- -log(v) 用inverse cdf先找出gx的x
```

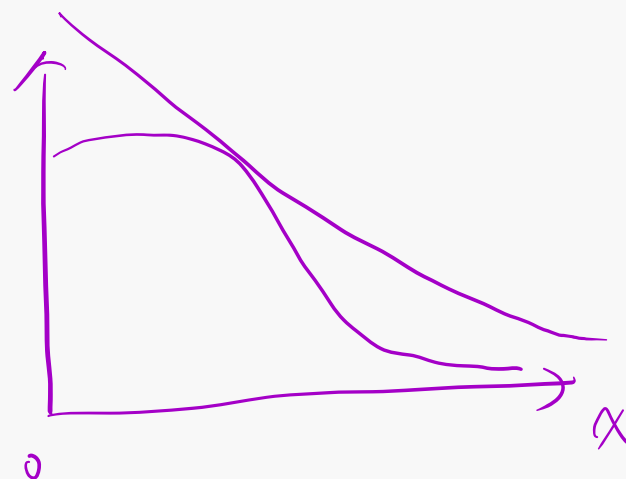
```
M <- 1.315
```

```
u <- runif(n)
```

判断是否小于 $f(x)/Mg(x)$

```
x_accepted <- x[which(u < f(x) / (M * g(x, 1)))]
```

$$\underline{g(x)} = \exp(\lambda=1) = e^{-x}$$



$$\min_x \left(\frac{x^2}{2} - x \right)$$

$$x^* = 1$$

$$M = \max \left(\frac{f(x)}{g(x)} \right)$$

$$\frac{\partial}{\partial x} \frac{f(x)}{g(x)} = 0$$

R Code to sample from Folded Normal Distribution (Cont.)

```
hist(x_accepted, prob = TRUE, las = 1,  
     main = "", ylim = c(0, M))  
curve(f(x), col = "blue", lwd = 2, add = T)  
curve(M * g(x, 1), lty = 2, lwd = 2, add = TRUE)  
legend("topright", c("f(x)", "M g(x)"),  
      inset = 0.1, lty = 1:2,  
      lwd = 2, col=c("blue", "black"))
```

