

MAP: for an exponential prior on the average of a gaussian distrib°:

$$x_i \sim \mathcal{N}(\mu, \sigma^2)$$

$\hookrightarrow \mu \sim \lambda e^{-\lambda \mu}$

$$\begin{aligned} \theta_{\text{MAP}}^* &= \underset{\theta}{\operatorname{argmax}} P(\theta | X) = \underset{\theta}{\operatorname{argmax}} \left(\frac{P(X|\theta) \cdot P(\theta)}{P(X)} \right) \\ &= \underset{\theta}{\operatorname{argmax}} \left(\underbrace{\log P(X|\theta)}_{\text{as MLE}} + \underbrace{\log P(\theta)}_{\text{prior}} \right) \end{aligned}$$

$$\theta_{\text{MAP}}^* = \underset{\mu}{\operatorname{argmax}} \left(\log \frac{1}{\sqrt{2\pi}\sigma} - \sum_{i=1}^N \frac{(x_i - \mu)^2}{2\sigma^2} + \log \lambda - \lambda \mu \right)$$

$$\mu: \frac{\partial J}{\partial \mu} = 0 \Leftrightarrow 0 = 0 + \frac{\lambda}{\lambda} \frac{\sum_{i=1}^N (x_i - \mu)}{\sigma^2} + 0 - \lambda$$

$$0 = \sum_{i=1}^N (x_i - \mu) - \lambda \sigma^2$$

$$\mu^* = \frac{1}{N} \sum_{i=1}^N x_i - \frac{\lambda \sigma^2}{N}$$

Remark: as always, when $N \rightarrow \infty$, the prior (λ) becomes irrelevant. Also, if $\sigma^2 \rightarrow 0$, the x_i 's are narrowly distributed, the prior is irrelevant.

Also, if $\lambda \rightarrow 0$ (average of expo. dist is $\rightarrow \infty$) then we have a (kind of) flat prior (although, on the half-space $\mu > 0$ only).

Remark: it would have made more sense to use this prior on a positive-definite variable, like σ^2 .