Regression

Let us approach regression from a maximum likelihood perspective.

We assume that

$$y = f_{\mathsf{true}}(\mathbf{x}) + \epsilon,$$

where f_{true} is a function that is parameterized by $\boldsymbol{\theta}$ with ϵ being a random variable that follows some distribution \mathbb{P}_{ϵ} , with $\mathbb{E}[\epsilon]=0$. Further, we assume ϵ to be independent of \mathbf{x} .

It follows that

- $y \mid \mathbf{x}$ follows a distribution with mean $f_{\text{true}}(\mathbf{x})$ and variance $\text{Var}(\epsilon)$.
- We denote the corresponding density function by $p(y \mid \mathbf{x}, \theta)$.

Given data

$$\mathcal{D} = \left(\left(\boldsymbol{x}^{(1)}, y^{(1)} \right), \dots, \left(\boldsymbol{x}^{(n)}, y^{(n)} \right) \right)$$

the maximum-likelihood principle is to maximize the likelihood

$$\mathcal{L}(\theta) = \prod_{i=1}^{n} \rho\left(y^{(i)} \mid \mathbf{x}^{(i)}, \theta\right)$$

or to minimize the negative log-likelihood:

$$-\ell(\boldsymbol{\theta}) = -\sum_{i=1}^{n} \log \rho \left(y^{(i)} \mid \mathbf{x}^{(i)}, \boldsymbol{\theta} \right)$$

 Let us now simply define the negative log-likelihood as loss function

$$L(y, f(\mathbf{x} \mid \boldsymbol{\theta})) := -\log p(y \mid \mathbf{x}, \boldsymbol{\theta})$$

 Maximum-likelihood optimization can be formulated as an empirical risk minimization problem

$$\mathcal{R}_{\mathsf{emp}}(oldsymbol{ heta}) = \sum_{i=1}^n L\left(y^{(i)}, f\left(\mathbf{x}^{(i)} \mid oldsymbol{ heta}
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ight)$$

 We can even disregard multiplicative or additive constants in the loss as they do not change the minimizer.

- For every error distribution \mathbb{P}_{ϵ} we can derive an equivalent loss function, which leads to the same point estimator for the parameter vector $\boldsymbol{\theta}$ as maximum-likelihood.
- NB: The other way around does not always work: We cannot derive a probability density function or error distribution corresponding to every loss function – the Hinge loss is a prominent example.

GAUSSIAN ERRORS - L2-LOSS

Let us assume that errors are Gaussian, i.e. $\epsilon^{(i)} \sim \mathcal{N}(0, \sigma^2)$. Then

$$y = f_{\mathsf{true}}(\mathbf{x}) + \epsilon \sim N\left(f_{\mathsf{true}}(\mathbf{x}), \sigma^2\right).$$

The likelihood is then

$$\mathcal{L}(\boldsymbol{\theta}) = \prod_{i=1}^{n} \rho \left(y^{(i)} \mid f\left(\mathbf{x}^{(i)} \mid \boldsymbol{\theta}\right), \sigma^{2} \right)$$

$$\propto \exp \left(-\frac{1}{2\sigma^{2}} \sum_{i=1}^{n} \left(y^{(i)} - f\left(\mathbf{x}^{(i)} \mid \boldsymbol{\theta}\right) \right)^{2} \right).$$

GAUSSIAN ERRORS - L2-LOSS

It is easy to see that minimizing the negative log-likelihood is equivalent to the *L*2-loss minimization approach since

$$-\ell(\boldsymbol{\theta}) = -\log \left(\mathcal{L}(\boldsymbol{\theta})\right)$$

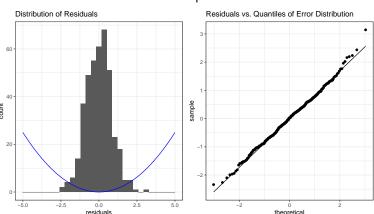
$$= -\log \left(\exp \left(-\frac{1}{2\sigma^2} \sum_{i=1}^n \left(y^{(i)} - f\left(\mathbf{x}^{(i)} \mid \boldsymbol{\theta}\right)\right)^2\right)\right)$$

$$\propto \sum_{i=1}^n \left(y^{(i)} - f\left(\mathbf{x}^{(i)} \mid \boldsymbol{\theta}\right)\right)^2.$$

Note: We use \propto as "proportional to ... up to multiplicative and additive constants".

GAUSSIAN ERRORS - L2-LOSS

- We can plot the "empirical" error distribution, i.e. the distribution of the residuals after fitting a model w.r.t. L2-loss.
- With the help of a Q-Q-plot we can compare the empirical residuals vs. the theoretical quantiles of a Gaussian distribution.



LAPLACE ERRORS - L1-LOSS

Let us assume that errors are Laplacian, i.e. ϵ follows a Laplace distribution which has the density

$$\frac{1}{2\sigma}\exp\left(-\frac{|x|}{\sigma}\right), \sigma>0.$$

Then

$$y = f_{\mathsf{true}}(\mathbf{x}) + \epsilon$$

follows a Laplace distribution with mean $f(\mathbf{x}^{(i)} \mid \boldsymbol{\theta})$ and scale parameter σ .

LAPLACE ERRORS - L1-LOSS

The likelihood is then

$$\mathcal{L}(\boldsymbol{\theta}) = \prod_{i=1}^{n} \rho \left(y^{(i)} \mid f\left(\mathbf{x}^{(i)} \mid \boldsymbol{\theta}\right), \sigma \right)$$

$$\propto \exp \left(-\frac{1}{\sigma} \sum_{i=1}^{n} \left| y^{(i)} - f\left(\mathbf{x}^{(i)} \mid \boldsymbol{\theta}\right) \right| \right).$$

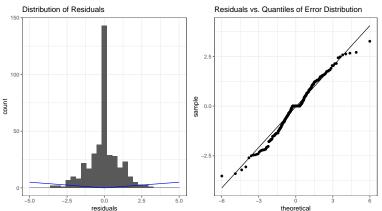
The negative log-likelihood is

$$-\ell(\boldsymbol{\theta}) \propto -\sum_{i=1}^{n} \left| y^{(i)} - f\left(\mathbf{x}^{(i)} \mid \boldsymbol{\theta}\right) \right|.$$

Minimizing the negative log-likelihood for Laplacian error terms corresponds to empirical risk minimization with L1-loss.

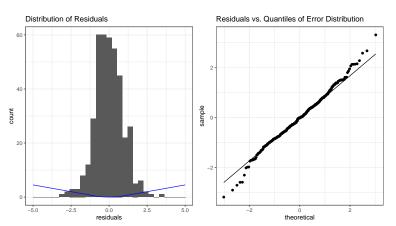
LAPLACE ERRORS - L1-LOSS

 Distribution of empirical residuals and their comparison to the theoretical quantiles of a Laplace-distribution.



OTHER ERROR DISTRIBUTIONS

 There are losses that do not correspond to "real" error densities, like the Huber loss. (In the QQ-plot below we show residuals against quantiles of a normal.)



OTHER ERROR DISTRIBUTIONS

However, intuitively, we see that a certain type of loss function corresponds to a certain error distribution.

Loss function	Error Distribution
L2-Loss	Gaussian Errors
L1-Loss	Laplace Errors
Huber Loss	"Huber Errors"

Classification

MAXIMUM LIKELIHOOD IN CLASSIFICATION

Let us assume the outputs $y^{(i)}$ to be Bernoulli-distributed, i.e.

$$y^{(i)} \sim \operatorname{Ber}(\pi(\mathbf{x}))$$

with probability $\pi(\mathbf{x})$ that depends on \mathbf{x} .

The maximization of the negative log-likelihood is based on

$$-\ell(\boldsymbol{\theta}) = -\sum_{i=1}^{n} \log \rho \left(y^{(i)} \mid \mathbf{x}^{(i)}, \boldsymbol{\theta} \right)$$
$$= \sum_{i=1}^{n} -y^{(i)} \log[\pi \left(\mathbf{x}^{(i)} \right)] - \left(1 - y^{(i)} \right) \log[1 - \pi \left(\mathbf{x}^{(i)} \right)].$$

MAXIMUM LIKELIHOOD IN CLASSIFICATION

This gives rise to the following loss function

$$L_{0,1}(y,\pi(\mathbf{x})) = -y \ln(\pi(\mathbf{x})) - (1-y) \ln(1-\pi(\mathbf{x}))$$

which we introduced as Bernoulli loss.

