# **Regression model - Motor trend**

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## **Executive Summary**

Our work for Motor Trend, a magazine about the automobile industry. Looking at a data set of a collection of cars, they are interested in exploring the relationship between a set of variables and miles per gallon (MPG) (outcome). They are particularly interested in the following two questions:

- "Is an automatic or manual transmission better for MPG"
- "Quantify the MPG difference between automatic and manual transmissions"

## **Exploratory data analysis**

```
library("ggplot2")
library("GGally")
library("gridExtra")
library("dplyr")
# Load data
data(mtcars)
```

### **Compute summary statistics of data subsets:**

First, let's check the average.

The MT car seems to have a higher MPG.

# MPG Vs. Transmission Design

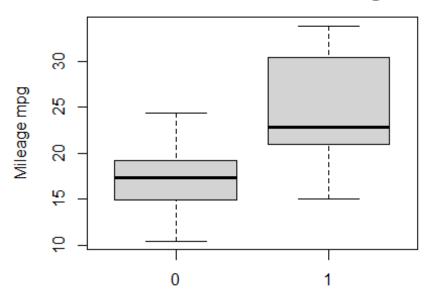


Fig 3. Transmission Type: 0-Auto, 1-Manual

Here we see the

boxplot that proves the average calculations for automatic and manual MPG.

#### **Calculate correlation:**

Calculate the correlation to see the relationship with other elements.

```
round(cor(mtcars), 2)[1, ]

## mpg cyl disp hp drat wt qsec vs am gear carb

## 1.00 -0.85 -0.85 -0.78 0.68 -0.87 0.42 0.66 0.60 0.48 -0.55
```

wt, disp, cyl and hp show high correlation.

#### **Fit Multiple Regression Models**

```
fit1 <- lm(mpg ~ am, mtcars)
fit2 <- lm(mpg ~ am + wt, mtcars)
fit3 <- lm(mpg ~ am + wt + disp, mtcars)
fit4 <- lm(mpg ~ am + wt + disp + cyl, mtcars)
fit5 <- lm(mpg ~ am + wt + disp + cyl + hp, mtcars)
anova(fit1, fit2, fit3, fit4, fit5)

## Analysis of Variance Table
##
## Model 1: mpg ~ am
## Model 2: mpg ~ am + wt
## Model 3: mpg ~ am + wt + disp
## Model 4: mpg ~ am + wt + disp</pre>
```

```
## Model 5: mpg \sim am + wt + disp + cyl + hp
##
    Res.Df
             RSS Df Sum of Sq F
                                       Pr(>F)
        30 720.90
## 1
## 2
        29 278.32 1
                       442.58 70.5432 7.017e-09 ***
## 3
        28 246.56 1
                        31.76 5.0628 0.033130 *
        27 188.43 1
## 4
                        58.13 9.2655 0.005289 **
## 5
        26 163.12 1
                        25.31 4.0336 0.055097 .
## ---
## Signif. codes: 0 '***' 0.001 '**' 0.05 '.' 0.1 ' ' 1
```

The **Model 4** p-value is near 0.005, so we will not reject the hypothesis. Model 4 (fit4) will fit better.

```
betterFit <- fit4
summary(betterFit)
##
## Call:
## lm(formula = mpg \sim am + wt + disp + cyl, data = mtcars)
## Residuals:
##
     Min
             1Q Median
                           3Q
                                 Max
## -4.318 -1.362 -0.479 1.354 6.059
## Coefficients:
               Estimate Std. Error t value Pr(>|t|)
##
## (Intercept) 40.898313 3.601540 11.356 8.68e-12 ***
## am
               0.129066 1.321512
                                   0.098
                                            0.92292
              -3.583425 1.186504 -3.020 0.00547 **
## wt
## disp
               0.007404
                          0.012081
                                     0.613
                                            0.54509
              -1.784173
                        0.618192 -2.886 0.00758 **
## cyl
## ---
## Signif. codes:
                  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
## Residual standard error: 2.642 on 27 degrees of freedom
## Multiple R-squared: 0.8327, Adjusted R-squared:
## F-statistic: 33.59 on 4 and 27 DF, p-value: 4.038e-10
```

This Multivariable Regression test now gives us an R-squared value of over .83, suggesting that 83% or more of variance can be explained by the multivariable model. P-values for cyl and wt are below 5%, suggesting that these are confounding variables in the relation between car Transmission and MPG.

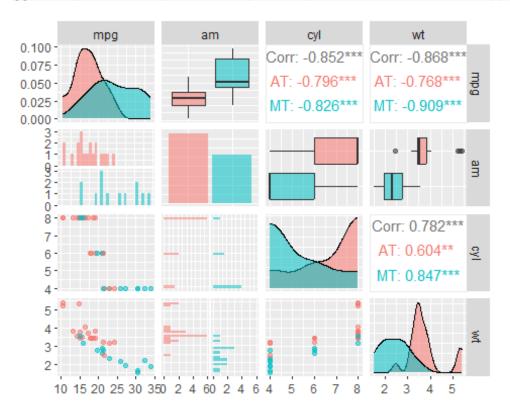
# **Residual and Diagnostics**

Now, we examine residual plots of our regression model and compute some of the regression diagnostics of our model to uncover outliers in the data set.

## **Appendix**

### Figure 1

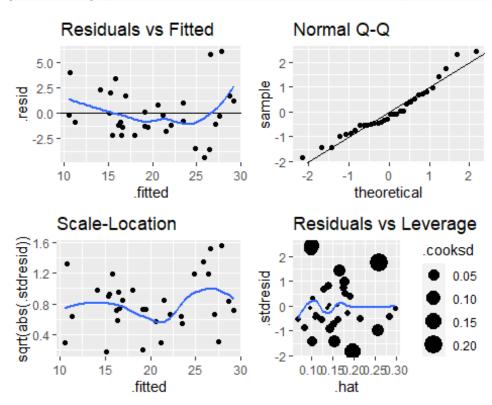
```
# Factorize
mtcars$am <- factor(mtcars$am, labels = c("AT", "MT"))
ggpairs(mtcars[, c(1, 9, 2, 6)], aes(color = am, alpha = .4))</pre>
```



## Figure 2

```
# Residuals vs Fitted
plot1 <- ggplot(betterFit, aes(.fitted, .resid)) +</pre>
        geom_point() +
        geom_hline(yintercept = 0) +
        geom_smooth(se = FALSE) +
        ggtitle("Residuals vs Fitted")
# Normal Q-Q
plot2 <- ggplot(betterFit) +</pre>
        stat_qq(aes(sample = .stdresid)) +
        geom_abline() +
        ggtitle("Normal Q-Q")
# Scale-Location
plot3 <- ggplot(betterFit, aes(.fitted, sqrt(abs(.stdresid)))) +</pre>
        geom_point() +
        geom_smooth(se = FALSE) +
        ggtitle("Scale-Location")
# Standardized Residuals vs Leverage
plot4 <- ggplot(betterFit, aes(.hat, .stdresid)) +</pre>
```

```
geom_point(aes(size = .cooksd)) +
    geom_smooth(se = FALSE) +
    ggtitle("Residuals vs Leverage")
grid.arrange(plot1, plot2, plot3, plot4, ncol = 2)
```



From Appendix Figure 2, we can make the following observations,

- The points in the Residuals vs Fitted plot seem to be randomly scattered on the plot and verify the independence condition.
- The Normal Q-Q plot consists of the points which mostly fall on the line indicating that the residuals are normally distributed.
- The Scale-Location plot consists of points scattered in a constant band pattern, indicating constant variance.