

# ST5218\_\_Tut\_\_3

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Q1:

```
library(tseries)
library(moments)
library(MASS)
SP500 = get.hist.quote(instrument = "^gspc",
                        start="2014-01-01", end="2016-12-31",
                        quote=c("AdjClose"), provider="yahoo",
                        compress="d")
```

```
## time series starts 2014-01-02
## time series ends 2016-12-30
daily.log.return = diff(log(SP500), lag=1)
fit.t = fitdistr(daily.log.return, "t")
```

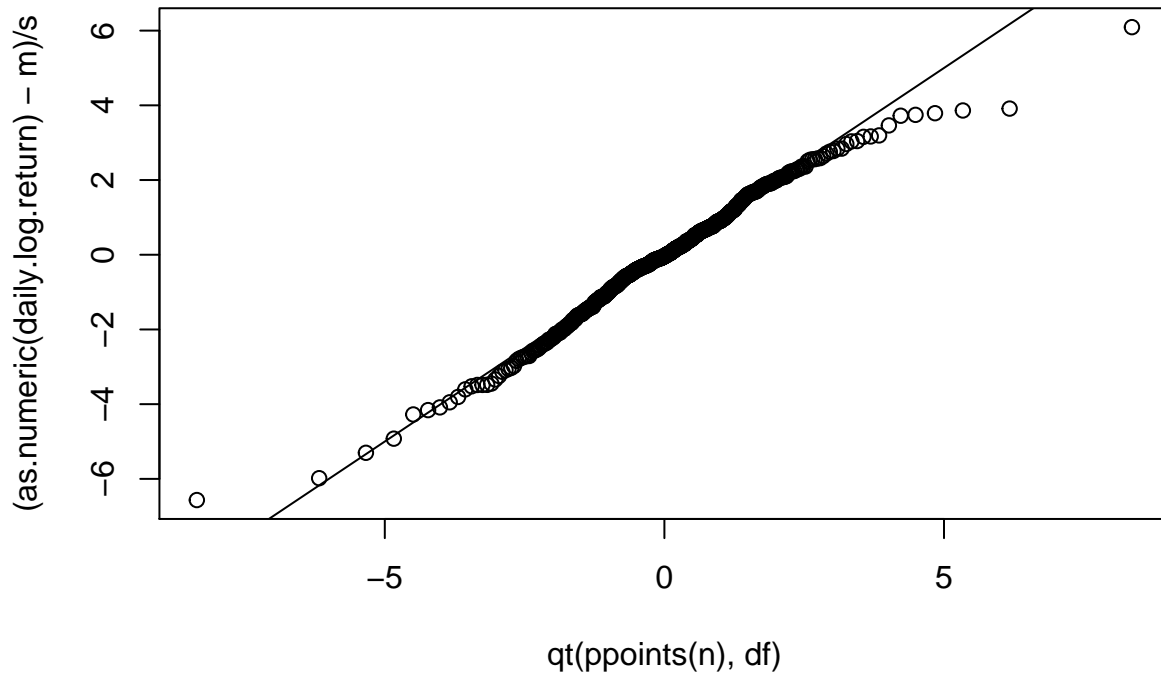
a)

```
fit.t
```

```
##           m           s           df
## 0.0005095120 0.0062012019 3.8465161900
## (0.0002662766) (0.0003014795) (0.6412577464)
```

b)

```
m = fit.t$estimate[1]
s = fit.t$estimate[2]
df = fit.t$estimate[3]
n = length(daily.log.return)
qqplot(qt(ppoints(n), df), (as.numeric(daily.log.return)-m)/s)
lines(qt(ppoints(n), df), qt(ppoints(n), df))
```



*# The distribution does not fit well for the data in the tails*

c)

$$Y = \frac{r - 0.0005095120}{0.0062012019} \sim t(3.8465161900)$$

With probability 0.001.

$$VaR_{0.001}(R) = 7.471026$$

$$VaR_{0.001}(R) = 7.471026 * 0.0062012019 - 0.0005095120 = 0.04581983$$

**Q2:**

$$\mu = \begin{pmatrix} \mu_1 \\ \mu_2 \end{pmatrix}$$

$$\Sigma = \begin{pmatrix} \sigma_1^2 & \rho\sigma_1\sigma_2 \\ \rho\sigma_1\sigma_2 & \sigma_2^2 \end{pmatrix}$$

$$f_x(x, y) = \frac{\Gamma[(\nu + 2/2)]}{\Gamma(\nu/2)(\nu\pi)} (\sigma_1^2\sigma_2^2(1 - \rho)^2)^{-\frac{1}{2}} \left[ 1 + \frac{1}{\nu(1 - \rho)^2} \left( \frac{(x - \mu_x^2)}{\sigma_x^2} + \frac{(y - \mu_y^2)}{\sigma_y^2} - \frac{2\rho(x - \mu_x)(y - \mu_y)}{\sigma_x\sigma_y} \right) \right]^{-\frac{(\nu+2)}{2}}$$

**Q3:**

a)

$$E(Y) = (0.2, 0.3, 0.2)^T$$

$$Cov(Y) = \begin{pmatrix} 0.8^2 + 1 & 0.8 * 0.7 & 0.8 * 0.9 \\ 0.7 * 0.8 & 0.7^2 + 1 & 0.7 * 0.9 \\ 0.9 * 0.8 & 0.9 * 0.7 & 0.9^2 + 1 \end{pmatrix}$$

b)

All the covariances are positive, because all that  $Y_i$  depend on common factor  $X$  with positive coefficients.

**Q4:**

a)

```
MMMdata = get.hist.quote(instrument="MMM",
                          start="2014-01-01", end="2014-12-31",
                          quote=("AdjClose"), provider="yahoo",
                          compression="d")
```

```
## time series starts 2014-01-02
```

```
## time series ends 2014-12-30
```

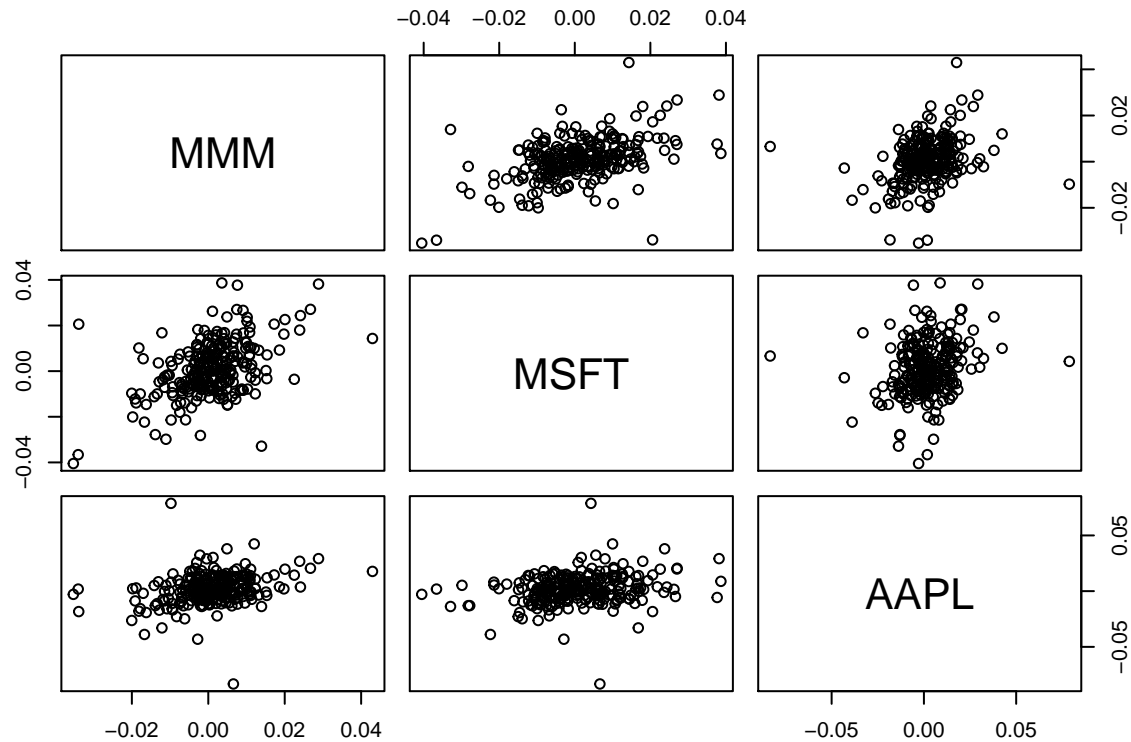
```
MSFTdata = get.hist.quote(instrument="MSFT",
                          start="2014-01-01", end="2014-12-31",
                          quote=("AdjClose"), provider="yahoo",
                          compression="d")
```

```
## time series starts 2014-01-02
```

```
## time series ends 2014-12-30
```

```
AAPLdata = get.hist.quote(instrument="AAPL",
                          start="2014-01-01", end="2014-12-31",
                          quote=("AdjClose"), provider="yahoo",
                          compression="d")
```

```
## time series starts 2014-01-02
## time series ends 2014-12-30
x = merge(MMMdata, MSFTdata, AAPLdata)
R = diff(log(x))
colnames(R)[1] = "MMM"
colnames(R)[2] = "MSFT"
colnames(R)[3] = "AAPL"
pairs(R)
```



```
cor(R)
```

```
##           MMM      MSFT      AAPL
## MMM  1.0000000  0.4742973  0.2955616
## MSFT  0.4742973  1.0000000  0.2338003
## AAPL  0.2955616  0.2338003  1.0000000
```

b)

```
n = length(MMMdata)
abs(cor(R)[1,2]) > qnorm(1-0.01/2)/sqrt(n-3)
```

```
## [1] TRUE
```

```
abs(cor(R)[1,3]) > qnorm(1-0.01/2)/sqrt(n-3)
```

```
## [1] TRUE
```

```
abs(cor(R)[2,3]) > qnorm(1-0.01/2)/sqrt(n-3)
```

```
## [1] TRUE
```

**Q5:**

$$\text{Var}(wR_1 + (1-w)R_2) = w^2\sigma_1^2 + 2w(1-w)\sigma_{12} + (1-w)^2\sigma_2^2$$

The first order condition is:

$$2w\sigma_1^2 + 2(1-2w)\sigma_{12} + 2(w-1)\sigma_2^2 = 0$$

The solution:

$$w = \frac{\sigma_2^2 - \sigma_{12}}{\sigma_1^2 + \sigma_2^2 - 2\sigma_{12}}$$

**Q6:**

Finding  $w_1 \cdots w_n$  to minimize  $\text{Var}(\sum_{i=1}^n w_i R_i)$  with  $w_1 + \cdots + w_n = 1$

Using Lagrange multipliers:

$$\min_{w_1, \dots, w_n, \lambda} \left\{ \text{Var}\left(\sum_{i=1}^n w_i R_i\right) + \lambda(w_1 + \cdots + w_n - 1) \right\}$$

Taking derivatives respect to  $w_k$  and  $\lambda$ :

$$2w_k\sigma_k^2 + \lambda = 0, \quad k = 1, \dots, n$$

The solution is:

$$w_k = -\frac{\lambda}{2\sigma_k^2}, \quad \lambda = -\frac{2}{\left(\frac{1}{\sigma_1^2}\right) + \cdots + \left(\frac{1}{\sigma_n^2}\right)}$$

$$\Rightarrow w_k = \frac{1/\sigma_k^2}{1/\sigma_1^2 + \cdots + 1/\sigma_n^2}$$

**Q7:**

**a)**

Let  $w_1, w_2$  be the weighting corresponding to A and B

$$w_1 + w_2 = 1$$

$$E(R_N) = w_1 E(R_1) + w_2 E(R_2) = w_1 r_1 + w_2 r_2$$

$$2.3\%w_1 + 4.5\%w_2 = 3\% \Rightarrow w_1 = 0.682, \quad w_2 = 0.318$$

Thus  $R_N = 0.682R_1 + 0.318R_2$

**b)**

$$\sigma_N^2 = w_1^2 \sigma_1^2 + w_2^2 \sigma_2^2 + 2\rho_{12}w_1w_2\sigma_1\sigma_2$$

$$5.5\% = 6\%w_1^2 + 11\%w_2^2 + 2 * 0.17 * 6\% * 11\%w_1w_2$$

$$w_1 + w_2 = 1$$

$$w_1 = 0.9404, \quad w_2 = 0.0596, \quad r_N = 2.43\%$$

or

$$w_1 = 0.411, \quad w_2 = 0.589, \quad r_N = 3.60\%$$

So the largest expected return is:

$$R_N = 0.411R_1 + 0.589R_2$$